

Comment on “Bayesian Regression Tree Models for Causal Inference” by Hahn, Murray and Carvalho

Stefan Wager
Stanford University

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The literature on heterogeneous treatment effect estimation has been extremely active over the past few years, and the paper by [Hahn, Murray, and Carvalho \[2020\]](#) is a major addition to it. [Hahn et al. \[2020\]](#) show how to design priors for heterogeneous treatment effects that are robust to what the authors call “regularization-induced confounding” and “targeted selection” and, as such, open the door to considerably more robust and reliable Bayesian inference of treatment heterogeneity under unconfoundedness. The authors convincingly show that their innovations add considerable value over a simpler Bayesian forest approach following, e.g., early work from [Hill \[2011\]](#).

The publication of this paper provides a nice opportunity to reflect on just how fast this area has developed over the past years. When [Hahn et al. \[2020\]](#) released the first draft of their manuscript on arXiv in 2017, there were still major questions about how best to approach the problem of heterogeneous treatment effect estimation as evidenced by, e.g., discussions at that year’s Atlantic Causal Inference Conference. By now, in contrast, there appears to be fairly widespread consensus on conceptual ideas that underpin good estimators of treatment heterogeneity. This comment offers one take on 3 ideas which, I believe, have played a major role in pushing the field forward. Of course, these ideas manifest themselves differently depending on methodological context (e.g., frequentist vs. Bayesian methods, trees vs. lasso), but overall they seem to have broad applicability.

Dedicated regularization for treatment effects Following notation from [Hahn et al. \[2020\]](#), we want to estimate the effect of a binary treatment $Z_i \in \{0, 1\}$ on an outcome $Y_i \in \mathbb{R}$ as a function of covariates $X_i \in \mathcal{X}$. Following the Neyman–Rubin causal model [[Imbens and Rubin, 2015](#)], we posit potential outcomes $\{Y_i(0), Y_i(1)\}$ such that $Y_i = Y_i(Z_i)$, and seek to estimate the conditional average treatment effect function $\tau(x) = \mathbb{E}[Y_i(1) - Y_i(0) \mid X_i = x]$. For purposes of identification, we assume unconfoundedness [[Rosenbaum and Rubin, 1983](#)], i.e., that treatment is as good as random conditionally on X_i : $\{Y_i(0), Y_i(1)\} \perp\!\!\!\perp Z_i \mid X_i$.

One can readily check that, under unconfoundedness, we have $\tau(x) = \mu(x; 1) - \mu(x; 0)$ with $\mu(x; z) = \mathbb{E}[Y_i \mid X_i = x, Z_i = z]$. This suggests a simple strategy for non-parametric estimation of the conditional average treatment effect function $\tau(x)$: First learn $\hat{\mu}(x; 0)$ and $\hat{\mu}(x; 1)$ by fitting separate predictive models to the control and treated samples respectively, and then set $\hat{\tau}(x) = \hat{\mu}(x; 1) - \hat{\mu}(x; 0)$. This may, however, lead to problems. In general, modern machine learning methods operate via some type of representation learning; and, if

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we use different representations to express $\hat{\mu}(x; 0)$ and $\hat{\mu}(x; 1)$, then $\hat{\tau}(x)$ may be excessively noisy or biased. As a simple example, consider the case where $\hat{\mu}(x; 0)$ and $\hat{\mu}(x; 1)$ are both decision trees—but with different splits. In this case, $\hat{\tau}(x) = \hat{\mu}(x; 1) - \hat{\mu}(x; 0)$ would be quite unstable, and in particular would have a more complicated shape than either $\hat{\mu}(x; 0)$ or $\hat{\mu}(x; 1)$ on its own. [Künzel, Sekhon, Bickel, and Yu \[2019\]](#) provide further examples of this issue.

A first major advance in the literature on treatment heterogeneity was the realization that, in a high-dimensional or non-parametric setting, it’s important to use dedicated regularizers that directly push $\hat{\tau}(x)$ to have a simple form. A simple application of this idea arises in the case of the lasso [[Hastie, Tibshirani, and Wainwright, 2015](#)]. Assume that $\mu(x; z) = x \cdot \beta_{(z)}$ for some high-dimensional vector z , so that $\tau(x) = x \cdot (\beta_{(1)} - \beta_{(0)})$. A naïve analysis might fit separate lasso regressions on the treated and control units, but such an approach may learn different sparsity patterns for $\beta_{(0)}$ and $\beta_{(1)}$, thus resulting in an unstable $\tau(x)$ estimate. A better approach is to reparametrize $\mu(x; z) = x \cdot b + (2z - 1)x \cdot \delta$, where $b = (\beta_{(0)} + \beta_{(1)})/2$ and $\delta = \beta_{(1)} - \beta_{(0)}$; then, we can apply separate sparsity penalties on b and δ . This is a simple idea but, by directly pushing the treatment effect parameter δ towards sparsity, it often improves performance considerably.

[Hahn et al. \[2020\]](#) show how to adapt it to Bayesian prior design, while [Athey and Imbens \[2016\]](#) discuss a modification of regression trees that directly target $\tau(x)$. More subtle ideas for algorithmically regularizing the treatment effect function include the X-learner of [Künzel, Sekhon, Bickel, and Yu \[2019\]](#), and refitting predictions from a first step analysis as in the Virtual Twins method of [Foster, Taylor, and Ruberg \[2011\]](#).

The propensity score as a covariate Once we’ve dealt with egregious instability of $\hat{\tau}(x)$ by using appropriate regularizers, a next concern is whether confounding effects may bleed into treatment effect estimates due to finite sample effects. As is well explained in the section on “targeted selection” in [Hahn et al. \[2020\]](#), this concern arises whenever the propensity score $\pi(x) = \mathbb{P}[Z_i = 1 \mid X_i = x]$ is associated with the baseline effect $\mu(x; 0)$. If $\mu(x; 0)$ takes on a complicated non-parametric specification that cannot be perfectly captured in finite samples and we use a method that underfits the baseline function such that $\mu(x; 0) - \hat{\mu}(x; 0)$ is positively (or negatively) correlated with $\pi(x)$, then we may easily have this baseline error push us to over- (or under-) estimate $\tau(x)$.

Considerations of this type have attracted considerable attention in the causal inference community for several decades [[Robins and Ritov, 1997](#)], and play a key role in discussions of how best to do variable selection when estimating global causal parameters [e.g., [Belloni, Chernozhukov, and Hansen, 2014](#)]. The general message is that, in order to be robust to associations between $\pi(x)$ and $\mu(x; 0)$, one needs to fit the propensity score $\hat{\pi}(x)$ and adjust for it in the final modeling step. [Hahn et al. \[2020\]](#) propose the simple idea of using a propensity estimate as a feature when estimating the baseline effect, i.e., they fit the baseline as $\hat{\mu}(x, \hat{\pi}(x); 0)$, and find this to work well empirically.

Treatment-focused loss functions A final idea that unlocks a general suite of tools for heterogeneous treatment effect estimation is the use of loss functions that directly isolate the treatment effect function $\tau(x)$. The idea of using target-specific loss functions has a long history in causal inference, going back at least to [van der Laan and Dudoit \[2003\]](#). In the context heterogeneous treatment effect estimation, a simple instance of this idea is the “transformed outcome” method, which starts from the following observation. Under unconfoundedness, we can check that $\mathbb{E}[\Delta_i \mid X_i = x] = \tau(x)$, where $\Delta_i = Z_i Y_i / \pi(X_i) -$

$(1 - Z_i)Y_i/(1 - \pi(X_i))$. Thus, in a randomized trial where propensity scores $\pi(x)$ are known a-priori, we can estimate $\tau(x)$ by first forming the modified outcomes Δ_i and then running a non-parametric regression of Δ_i on X_i [Tian, Alizadeh, Gentles, and Tibshirani, 2014]. Baseline effects $\mu(x; 0)$ never even appear in this specification, which largely obviates concerns related to regularizing or under-fitting of this term.

One limitation of the transformed outcome method is that it is not robust to errors in estimating the propensity score when $\pi(x)$ is not known a-priori, and several “robust” loss functions for treatment effect estimation that remedy this issue have recently been proposed. The R-learner [Nie and Wager, 2020] builds on the partially linear model estimator of Robinson [1988] to develop a loss function that is first-order robust to errors in estimating $\pi(x)$ and baseline effects; see also Athey, Tibshirani, and Wager [2019] and Zhao, Small, and Ertefaie [2017] for variants of this idea applied to random forests and the lasso specifically. Meanwhile, the DR-learner [Fan, Hsu, Lieli, and Zhang, 2019, Kennedy, 2020, Zimmert and Lechner, 2019] estimates $\tau(x)$ by regressing the augmented inverse-propensity weighted scores of Robins, Rotnitzky, and Zhao [1994] against X_i . Overall, the promise of such robust loss functions is that they enable accurate estimation of $\tau(x)$ even when $\pi(x)$ and $\mu(x; 0)$ may be difficult to estimate, thus generalizing well known results on semiparametric inference for “global” targets like the average treatment effect; see Chernozhukov, Chetverikov, Demirer, Duflo, Hansen, Newey, and Robins [2018], and references therein.

Closing thoughts Recent advances in methods for treatment heterogeneity have enabled a large toolkit of practical methods available to applied researchers. As a community, we now appear to be at a point where we can quickly remix these ideas to develop new methods for treatment heterogeneity that can address different application-specific challenges, and have a formal understanding of how dedicated methods are able to accurately target $\tau(x)$. Several open questions remain, however. In particular, while reading the paper of Hahn, Murray, and Carvalho [2020], I was left wondering about the following two:

- When using treatment-focused loss functions, it’s possible to show that (under appropriate conditions), the accuracy with which we can estimate $\tau(x)$ is insensitive to our rate of convergence on the nuisance components $\pi(x)$ and $\mu(x; z)$, even when $\hat{\pi}(x)$ and $\hat{\mu}(x; z)$ may converge an order of magnitude slower than $\hat{\tau}(x)$ [Kennedy, 2020, Nie and Wager, 2020]. Do analogous results hold for the approach of Hahn et al. [2020], where $\hat{\pi}(x)$ is used as a covariate when fitting $\hat{\mu}(\cdot)$? In the context of estimating an average treatment effect, Hirano, Imbens, and Ridder [2003] showed that using an estimated propensity score for weighting could sometimes be enough to achieve efficiency. Does any intuition of this type carry over to the problem heterogeneous treatment effect estimation?
- In Hahn et al. [2020], the propensity score is only used as a covariate to make us robust to potential confounding effects. In some applications, however, there may also be interest in the propensity score as an effect modifier. For example, when studying the returns to college education, Brand and Xie [2010] argue that students who are least likely to attend college a-priori may be the ones who benefit the most from attendance; and, in applications like these, it may be of interest to allow $\tau(x)$ explicitly depend on $\pi(x)$. Of particular interest here would be to understand how $\tau(x)$ varies with the true propensity score $\pi(x)$, and not just the estimate $\hat{\pi}(x)$.

Finally, I want to thank Hahn, Murray, and Carvalho [2020] for preparing this very nice paper, and look forward to seeing how the community builds on their results in the future.

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