

A Covert Queueing Channel in FCFS Schedulers

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Abstract—We study covert queueing channels (CQCs), which are a kind of covert timing channel that may be exploited in shared queues across supposedly isolated users. In our system model, a user sends messages to another user via his pattern of access to the shared resource, which serves the users according to a first come first served (FCFS) policy. One example of such a channel is the cross-virtual network covert channel in data center networks, resulting from the queueing effects of the shared resource. First, we study a system comprising a transmitter and a receiver that share a deterministic and work-conserving FCFS scheduler, and we compute the capacity of this channel. We also consider the effect of the presence of other users on the information transmission rate of this channel. The achievable information transmission rates obtained in this paper demonstrate the possibility of significant information leakage and great privacy threats brought by CQCs in FCFS schedulers.

Index Terms—Covert queueing channel, first-come-first-served scheduler, capacity limit.

I. INTRODUCTION

THE existence of side and covert channels due to the fragility of isolation mechanisms is an important privacy and security threat in computer networks. Such channels may be created across users, which were supposed to be isolated, resulting in information leakage. By definition, a covert channel is a hidden communication channel, which is not intended to exist in the system and is created furtively by users [2]. Covert channels may be exploited by a trusted user, or possibly a malware inside a system with access to secret information to leak it to a distrusted user. On the other hand, in a side channel a malicious user attempts to learn private information by observing information not intended for him. In this scenario, there is no collaboration between the source of information and the recipient [3].

Given that a lot of sensitive organizations such as CIA, and US Navy and Air Force, are abandoning in-house

infrastructure and migrating to clouds, privacy has emerged as a serious risk for clouds. In the context of covert channels, a disgruntled employee, a leaker, or a malware, can easily sneak out extremely confidential and private data without explicitly communicating with an external party (data exfiltration). The employee could pretend to be talking to another trusted entity but sending out covert signals via shared queues to someone implicitly. This would be a severe threat to privacy leakage as it would be unbeknown to the tenant being targeted that their data is being exfiltrated and would bypass most defenses deployed at the host or network layer. Typical scenarios here could be stealing cryptographic keys, bank records, medical records, service records of military personal, names and locations of secret offices and networks of CIA/NSA, etc., i.e., small pieces of information that can have detrimental consequences if leaked.

Timing channels are one of the main types of covert/side channels, in which information is conveyed through timing of occurrence of events (e.g., inter-arrival times of packets). A special case of timing channels is covert/side queueing channels, which can arise between users who share a packet scheduler in a network. Packet schedulers serve packets from multiple streams, which are queued in a single queue. This causes dependencies between delays observed by users. Particularly, the delay that one user experiences depends on the amount of traffic generated by other streams, as well as his own traffic. Hence, a user can gain information about other users' traffic by observing delays of his own stream. This dependency between the streams can breach private information as well as create hidden communication channels between the users.

One example of a covert/side queueing channel is the cross-virtual network covert channel in data center networks and cloud environments. As mentioned earlier, in recent years, migrating to commercial clouds and data centers is becoming increasingly popular among companies that deal with data. The multi-tenant nature of cloud and sharing infrastructure between several users has made data protection and avoiding information leakage a serious challenge in such environments [4]. In data center networks, software-defined-networks are frequently used for load balancing [5]. This generates logically isolated virtual networks and prevents direct data exchange. However, since packet flows belonging to different VNs inevitably share underlying network infrastructure (such as a router or a physical link), it is possible to transfer data across VNs through timing channels resulting from the queueing effects of the shared resource(s).

For data centers, the underlying premise is that resources should be multiplexed and oversubscribed as much as possible

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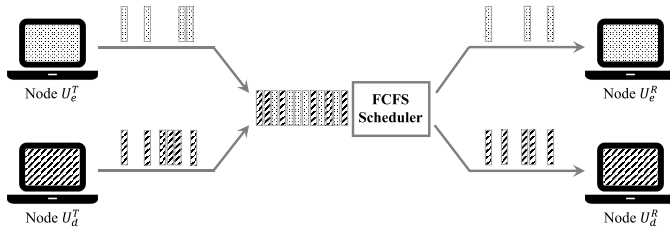


Fig. 1. Covert queueing channel in a system with 2 users.

to maximize utilization and consequently profits. Hence, most of the resources are shared (e.g., cache, system bus, memory, NICs, switches, links, etc.). Under such a scenario, it can be seen that there will be numerous opportunities for establishing covert channels. Hence, a queue-sharing based channel is certainly not the only option. However, it is the most convenient/practical and perhaps the only one that continues to exist even if tenants resort to dedicated infrastructure (fully dedicated servers are provided by cloud providers at a higher cost, but network is always shared and no option is provided for dedicated networking infrastructure). Two specific scenarios to argue this are as follows. In the first case, there are two mutually untrusting tenants that are placed on separate servers. However, the first hop router and the outbound links from then onwards are shared. In such a scenario, there are numerous shared queues which can be exploited to setup a covert channel. Since the two tenants are deployed on separate servers, none of the other media can be exploited to create covert channels, and the only option is to use a network-based queue-sharing channel. In the second scenario, consider a load balancing server. Two separate tenants, each own a VM on this load balancing server (such servers have typically very high-end NICs and cloud vendors try to maximize sharing on these expensive NICs). The VMs on this server simply load balance traffic between other VMs of the tenant. Again, there are numerous network and server-based queues that can be exploited to setup a covert channel.

In this paper, we study covert queueing channels (CQCs) in a shared deterministic and work-conserving first-come-first-served (FCFS) scheduler. We present an information-theoretic framework to describe and model the data transmission in this channel and calculate its capacity. First, we consider a two users setting depicted in Figure 1. In this model we have an encoder and a decoder user. Each user possesses a transmitter node and a receiver node. There is no direct communication channel between the users, but they share a packet scheduler. Hence, the delays observed by users are correlated. Therefore, the encoder user can encode a message in his traffic pattern and the decoder user can estimate the message by estimating the encoder's traffic pattern via the delays he experiences. Next, we extend the model to study the effect of the presence of a third user on the information transmission rate. The approach for analyzing the effect of the presence of the third user can be extended to calculate the capacity of the covert queueing channel serving any number of users.

The rest of the paper is organized as follows. We review related works in Section II. In Section III, we describe the system model. The capacities of the introduced channel for the two and three user cases are calculated in Sections IV and V, respectively. Our concluding remarks are presented in Section VI.

II. RELATED WORKS

The existing literature on covert/side timing channels has mainly concentrated on timing channels in which the receiver/adversary has direct access to the timing sequence produced by the transmitter/victim or a noisy version of it. However, in a covert/side queueing channel, the receiver/adversary does the inference based on the timing of his own packets which has been influenced by the original stream.

In a queuing side channel, where a malicious user, called an attacker, attempts to learn another user's private information, the main approach used by the attacker is traffic analysis. That is, the attacker tries to infer private information from the victim's traffic pattern. The attacker can have an estimation of the features of the other user's stream, such as packet size and timing by emitting frequent packets in his own sequence. Previous work shows that through traffic analysis, the attacker can obtain various private information including exact schedules of real-time systems [6], [7], visited web sites [8], sent keystrokes [9], and even inferring spoken phrases in a voice-over-IP connection [10].

In [11], Gong et al. proposed an attack where a remote attacker learns about a legitimate user's browser activity by sampling the queue sizes in the downstream buffer of the user's DSL link. The information leakage of a queueing side channel in an FCFS scheduler is analyzed in [12]. The analysis of more general work-conserving policies has been done in [13] and [14]. Gong and Kiyavash [14] presented an analytical framework for modeling information leakage in queueing side channels and quantify the leakage for several common scheduling policies.

Most of the work in covert timing channels is devoted to the case in which two users communicate by modulating the timings, and the receiver sees a noisy version of the transmitter's inputs [15]–[23]. Also, there are many works devoted to the detection of such channels [16], [24], [25]. The setup of CQC is new in the field of covert channels and as far as the authors are aware, there are very few works on this setup [1], [26], [27]. In [1] we studied a system comprised of only a transmitter and a receiver and computed the capacity of the covert channel. We have extended the results of [1] by studying the effect of the presence of other users on the information transmission rate of the CQC (Section V). Furthermore, the treatment of the proofs in the entire paper is now more rigorous and detailed. Specifically, we have provided a more in depth study of the function $\tilde{H}$ and its properties. This function indicates the highest amount of information transmittable for a given packet rate of the encoder user and a given inter-arrival time of packets of the decoder user, and plays a fundamental role in the calculation of the capacity.

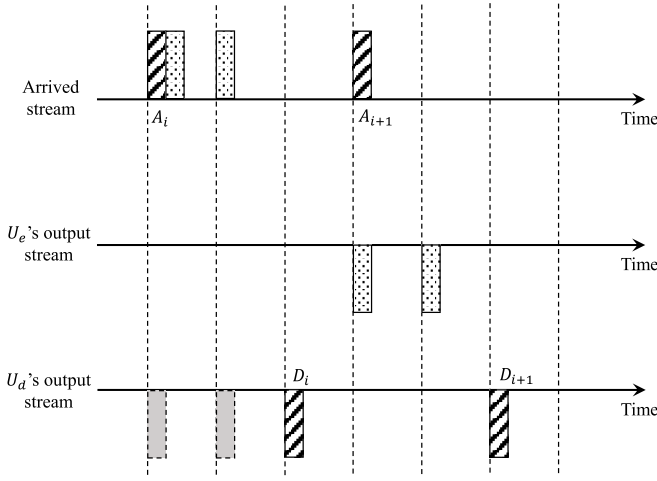


Fig. 2. An example of the input and output streams of the FCFS scheduler serving two users. Time slots are separated by dotted lines. Packets with dotted pattern belong to U_e and packets with diagonal pattern belong to U_d . Gray packets are existing packets in the buffer, which could belong to either of the users. We assume that one packet is buffered in the queue at time A_i .

III. SYSTEM DESCRIPTION

Consider the architecture depicted in Figure 1. In this model, a scheduler serves packets from 2 users, U_e and U_d . Each user U_i , $i \in \{e, d\}$, is modeled by a transmitter and a receiver node, denoted by U_i^T and U_i^R , respectively. U_i^R is the node which receives U_i^T 's packet stream. Note that U_i^T and U_i^R could correspond to the uplink and downlink of the same entity. U_e intends to send a message to U_d , but there is no direct channel between them. However, since U_e^T and U_d^T 's packets share the same queue, U_e^T can encode messages in the arrival times of its packets, which are passed onto U_d via queueing delays. Therefore, a timing channel is created between users via the delays experienced through the coupling of their traffic due to the shared scheduler.

To receive the messages from U_e , user U_d sends a packet stream from the node U_d^T . He then uses the delays he experiences by receiving the packet stream at U_d^R to decode the message. Therefore, effectively, the nodes U_e^T and U_e^R are on the encoder side and the nodes U_d^T and U_d^R are on the decoder side of the channel of our interest. Throughout the paper, we call U_d 's sent stream the *probe stream*.

We consider an FCFS scheduler, which is commonly used in DSL routers. We assume this scheduler is deterministic and work-conserving. Time is discretized into slots, and the scheduler is capable of processing at most one packet per time slot. At each time slot, each user either issues one packet or remains idle. Furthermore, we assume that all packets are the same size. Throughout the paper, we assume that the priorities of the users are known. Particularly, without loss of generality, we assume that U_d has the highest priority among all users; i.e., in the case of simultaneous arrivals, U_d 's packet will be served first.

Figure 2 shows an example of the input and output streams of the system depicted in Figure 1 with an FCFS scheduler. In this figure, the first stream is the arrival stream i.e., arrivals

from both U_e^T and U_d^T , depicted by dotted and diagonal patterns, respectively. The second stream is the output stream of user U_e (received by U_e^R), and the third one is the output stream of user U_d (received by U_d^R). In this example, we assume that one packet is buffered in the queue at time A_i , where a packet arrives from both U_d^T and U_e^T . If user U_e had not sent the two packets (depicted by dotted pattern), the second packet of user U_d , which arrives at time A_{i+1} could have departed one time slot earlier. Therefore, U_d knows that U_e has issued two packets.

As mentioned earlier, at each time slot, each user is allowed to either send one packet or none; hence, the input and output packet sequences of each user could be viewed as a binary bitstream, where '1' and '0' indicates whether a packet was sent or not in the corresponding time slot.

Assume message W drawn uniformly from the message set $\{1, 2, \dots, M\}$ is transmitted by U_e^T , and $\hat{W}$ is U_d 's estimate of the sent message. Our performance metric is the average error probability, defined as follows.

$$P_e \triangleq P(W \neq \hat{W}) = \sum_{m=1}^M \frac{1}{M} P(\hat{W} \neq m | W = m).$$

U_e encodes each message into a binary sequence of length n , Δ^n , to create the codebook, which is known at the decoder, U_d .

In order to send a message, U_e^T emits a packet in the i -th time slot if $\Delta_i = 1$ and remains idle otherwise, i.e.,

$$\Delta_i = \begin{cases} 1 & \Rightarrow U_e^T \text{ issues a} \\ & \text{packet in time slot } i. \\ 0 & \Rightarrow U_e^T \text{ remains} \\ & \text{idle in time slot } i. \end{cases}$$

To decode this message, U_d^T sends a binary length n stream (the probe stream) to the scheduler during the same length n time period. User U_d will use this stream and the response stream received at node U_d^R to decode the sent message.

We define the code, rate of the code, and the channel capacity similar to the definitions in [15], [28], and [29], as follows.

Definition 1: An (n, M, ϵ) -code consists of a codebook of M equiprobable binary codewords, where messages take on average n time slots to be received, and the error probability satisfies $P_e \leq \epsilon$.

Definition 2: The information transmission rate, R , of a code is the amount of conveyed information (logarithm of the codebook size) normalized by the average number of used time slots for the message to be received, i.e., $R = (\log M)/n$. Rate may be interpreted as the average amount of information conveyed in a single time slot.

Definition 3 (Channel Capacity): The Shannon capacity C , for a channel is the maximum achievable rate at which one can communicate through the channel when the average probability of error goes to zero. In other words, C is the supremum of rates R , which satisfy the following property [29].

$$\forall \delta > 0, \exists (n, M, \epsilon_n)\text{-code} \\ \text{s.t. } \begin{cases} \frac{\log M}{n} > R - \delta \\ \epsilon_n \rightarrow 0 \end{cases} \text{ as } n \rightarrow \infty.$$

The following notations will be used throughout the paper.

- r_i : U_i^T 's packet rate.
- A_i : Arrival time of the i -th packet in the probe stream.
- D_i : Departure time of the i -th packet of the probe stream. We assume m packets are sent by U_d during n time slots and we have: $r_d = \lim_{n \rightarrow \infty} m/n$.
- X_i : Number of U_e 's packets sent in the interval $[A_i, A_{i+1})$. Note that $X_i = \sum_{j=A_i}^{A_{i+1}-1} \Delta_j$.
- $T_i = A_{i+1} - A_i$: Inter-arrival time between i -th and $(i+1)$ -th packet of the probe stream. We denote a realization of T by τ .
- $Y_i = D_{i+1} - D_i - 1$.
- $\hat{X}_i$: Estimation of X_i by decoder.
- $\hat{W}$: Decoded message.

In an FCFS scheduler, U_d can have an estimation of the number of the packets of other users between any of his own consecutive packets. The estimation of the number of packets in the interval $[A_i, A_{i+1})$ is accurate if the scheduler is deterministic and work-conserving and a sufficient number of packets is buffered in the queue at time A_i .¹ In that case, the number of other users' packets arriving in the interval $[A_i, A_{i+1})$ could be simply calculated by $Y_i = D_{i+1} - D_i - 1$. Note that U_d cannot pinpoint the location of the sent packets; that is, if the inter-arrival time is τ , U_d can distinguish between $\tau + 1$ different sets of bit streams sent during this time. Therefore, we look at any probe stream sent during n time slots as a combination of different inter-arrival times.

If the sum of the packet rates of the users used during sending a message of length n is on average larger than 1, then the message will be arrived on average during more than n time slots. Also, this will destabilize the input queue of the scheduler. For example, for a system with two users U_d and U_e , if U_d^T sends packets in every time slot, then sending a packet by U_e^T in any time slot would cause a delay in the serving of the next packet of U_d^T and hence could be detected. Therefore, in each time slot, U_e^T could simply idle to signal a bit '0' or send a packet to signal a bit '1', resulting in the information rate of 1/1.5 bit per time slot in the case that bits are equiprobable. But, this scheme is not feasible in practice as it would destabilize the queue and result in severe packet drops.

In our model, we do not have any restrictions on the number of time slots in which the data transmission is happening. Therefore, the arrival sequence can be arbitrarily long, and hence, queue stability is required. In order to have queue stability, it suffices that the total packet arrival rate does not exceed the service rate, which for a deterministic and work-conserving scheduler is equal to 1 (see Appendix VI for the proof of stability which is based on a Lyapunov stability

¹If the service rate of the scheduler is equal to 1, there should be at least $A_{i+1} - A_i - 1$ packets buffered in the queue at time A_i . Therefore, user U_d needs to know the queue length. This is feasible using the following formula.

$$q(A_i) = D_i - A_i - 1$$

where $q(A_i)$ denotes the queue length at the time that the i -th packet in the probe stream arrives at the queue. The extra 1 in the formula is the time needed for the i -th packet of the probe stream to be served. Therefore, user U_d should always be aware of the queue length and keep it sufficiently large by sending extra packets when needed.

argument for the general case that the serving rate is assumed to be $0 \leq \rho \leq 1$ and arbitrary number of users is considered). Specifically, for the case of two users we need

$$r_e + r_d < 1. \quad (1)$$

On the other hand, if the sum of the packet rates of the users used during sending a message of length n is on average less than 1, the length of the input queue may become less than the value required for accurate estimation of the number of the packets of other users between any of U_d 's consecutive packets. Consequently, this creates an extra source of error for U_d in estimating X_i 's. In other words, for any given scheme with sum of the packet rates less than 1, increasing r_d increases the resolution available for user U_d by removing the aforementioned extra source of error. Hence, U_d can have a better estimation of the number of other users' sent packets. To avoid this extra source of error, we focus on the coding schemes where the sum of the rates is 1. Therefore, in the case of two users, in order to achieve the highest information rate, the operation point should tend to the line $r_e + r_d = 1$.

IV. TWO-USER CASE

In this section, using achievability and converse arguments, the capacity of the introduced system is calculated for a system with a deterministic and work-conserving FCFS scheduler serving packets from two users.

As depicted in Figure 1, user U_e is attempting to send a message to U_d through the covert queueing channel between them. Note that since we have considered service rate of 1 for the FCFS scheduler and users can agree on the packet stream sent by U_d^T ahead of time, the feedback U_e^R is already available at the encoder. Therefore, the following Markov chain holds

$$W \rightarrow X^m \rightarrow Y^m \rightarrow \hat{W}. \quad (2)$$

Note that as mentioned earlier, if there is a sufficient number of packets buffered in the shared queue, $\hat{X}_i$ could be accurately estimated as Y_i .

The main result of this section is the following theorem, the proof of which is developed in the rest of the section.

Theorem 1: The capacity of the timing channel in a shared FCFS scheduler with service rate 1 depicted in Figure 1 is equal to 0.8114 bits per time slot, which can be obtained by solving the following optimization problem.

$$C = \max_{\alpha, \gamma_1, \gamma_2} \alpha \tilde{H}(\gamma_1, 1) + (1 - \alpha) \tilde{H}(\gamma_2, \frac{1}{2})$$

$$s.t. \quad \alpha(\gamma_1 + 1) + (1 - \alpha)(\gamma_2 + \frac{1}{2}) = 1, \quad (3)$$

where $0 \leq \alpha \leq 1$ and $0 \leq \gamma_1, \gamma_2 \leq \frac{1}{2}$ and the function $\tilde{H} : [0, 1] \times \{\frac{1}{k} : k \in \mathbb{N}\} \mapsto [0, 1]$ is defined as

$$\tilde{H}(\gamma, \frac{1}{k}) = \frac{1}{k} \max_{X \in \{0, 1, \dots, k\}} H(X), \quad k \in \mathbb{N}, 0 \leq \gamma \leq 1. \quad (4)$$

The function $\tilde{H}(\gamma, \frac{1}{k})$ indicates the ratio of highest amount of information that a random variable with support $\{0, \dots, k\}$ and mean $k\gamma$ can contain, to the number of bits used for

transmitting this information. We first investigate some of the properties of the function $\tilde{H}$.

Lemma 1: Let $U_k \sim \text{Unif}(\{0, 1, \dots, k\})$. The distribution which achieves the optimum value in (4) is the tilted version of $\text{Unif}(\{0, 1, \dots, k\})$ with parameter λ ; that is

$$P_X(i) = \frac{e^{i\lambda}}{\sum_{i=0}^k e^{i\lambda}}, \quad i \in \{0, 1, \dots, k\},$$

where $\lambda = (\psi'_{U_k})^{-1}(k\gamma)$, where the function $\psi'_{U_k}(\cdot)$ is the derivative of the log-moment generating function of U_k .

See Appendix VI for the proof of Lemma 1.

Lemma 2: The function $\tilde{H}$ could be computed using the following expression.

$$\tilde{H}(\gamma, \frac{1}{k}) = \frac{1}{k} [\log_2(k+1) - \psi_{U_k}^*(k\gamma) \log_2 e], \quad (5)$$

where $U_k \sim \text{Unif}(\{0, 1, \dots, k\})$, and the function $\psi_{U_k}^*(\cdot)$ is the rate function given by the Legendre-Fenchel transform of the log-moment generating function $\psi_{U_k}(\cdot)$,

$$\psi_{U_k}^*(\gamma) = \sup_{\lambda \in \mathbb{R}} \{\lambda\gamma - \psi_{U_k}(\lambda)\}. \quad (6)$$

In order to prove this lemma, first we note that for any random variable X defined over the set $\{0, 1, \dots, k\}$,

$$\begin{aligned} H(X) &= \sum_{i=0}^k P_X(i) \log \frac{1}{P_X(i)} \\ &= \sum_{i=0}^k P_X(i) \log(k+1) - \sum_{i=0}^k P_X(i) \log \frac{P_X(i)}{\frac{1}{k+1}} \\ &= \log(k+1) - D(P_X || U_k), \end{aligned}$$

where $D(P_X || U_k)$ denotes the KL-divergence between P_X and U_k . Therefore, in order to maximize $H(X)$, we need to minimize $D(P_X || U_k)$. Using the following well-known fact [30], concludes the lemma.

$$\min_{\mathbb{E}[X]=k\gamma} D(P_X || U_k) = \psi_{U_k}^*(k\gamma) \log_2 e. \quad (7)$$

Figure 3 shows the function $\tilde{H}(\gamma, \frac{1}{k})$ for different values of γ and $k \in \{1, 2, 3\}$.

Lemma 3: The function $\tilde{H}(\cdot, \cdot)$ is concave in pair $(\gamma, \frac{1}{k})$ in the sense that for integers k_1, k_2, k_3 , and for values $0 \leq \gamma_1, \gamma_2, \gamma_3 \leq 1$, and for $\alpha \in [0, 1]$, such that $\alpha(\gamma_1, \frac{1}{k_1}) + (1 - \alpha)(\gamma_3, \frac{1}{k_3}) = (\gamma_2, \frac{1}{k_2})$, we have

$$\alpha \tilde{H}(\gamma_1, \frac{1}{k_1}) + (1 - \alpha) \tilde{H}(\gamma_3, \frac{1}{k_3}) \leq \tilde{H}(\gamma_2, \frac{1}{k_2}). \quad (8)$$

See Appendix VI for the proof of Lemma 3.

Substituting (5) in (3) and solving it, the capacity of the timing channel in the shared FCFS scheduler with service rate 1 depicted in Figure 1 is equal to 0.8114 bits per time slot, achieved by $\alpha = 0.177$, $\gamma_1 = 0.43$ and $\gamma_2 = 0.407$.

Lemma 4: For all $\gamma \in [0, 1]$ and $k \in \mathbb{N}$, we have

$$\tilde{H}(\gamma, \frac{1}{k}) = \tilde{H}(1 - \gamma, \frac{1}{k}).$$

See Appendix VI for the proof of Lemma 4.

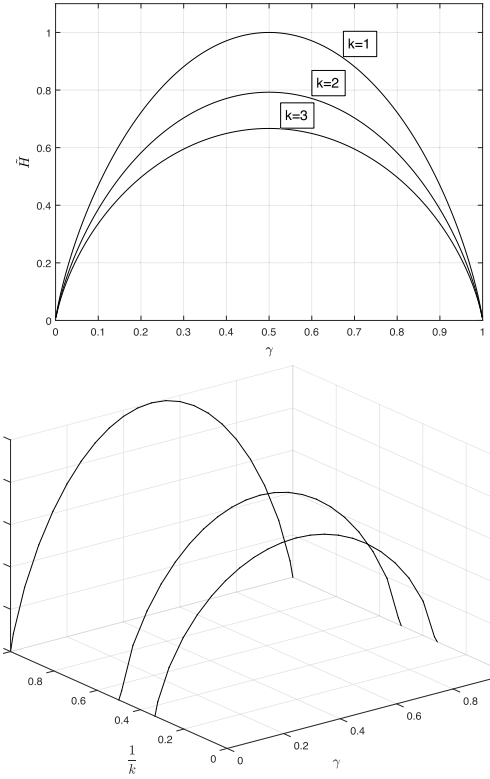


Fig. 3. $\tilde{H}(\gamma, \frac{1}{k})$ for different values of γ and $k \in \{1, 2, 3\}$.

In the following, the proof of Theorem 1 is given. The proof is based on converse and achievability arguments.

A. Converse

In the converse side, the ultimate goal is to prove that

$$\begin{aligned} C &\leq \max_{\alpha, \gamma_1, \gamma_2} \alpha \tilde{H}(\gamma_1, 1) + (1 - \alpha) \tilde{H}(\gamma_2, \frac{1}{2}) \\ \text{s.t. } &\alpha(\gamma_1 + 1) + (1 - \alpha)(\gamma_2 + \frac{1}{2}) = 1, \end{aligned}$$

where $0 \leq \alpha \leq 1$ and $0 \leq \gamma_1, \gamma_2 \leq \frac{1}{2}$. We break the proof into two lemmas. First in Lemma 5, we find an upper bound on the information rate, which consists of a weighted summation of possible maximum information rates for different inter-arrival times of the packets in the probe stream, which satisfies the stability constraint. Then in Lemma 6, we upper bound the summation with one which only corresponds to inter-arrival times of 1 and 2.

Lemma 5: For the timing channel in a shared FCFS scheduler with service rate 1 depicted in Figure 1, any code consisting of a codebook of M equiprobable binary codewords, where messages take on average n time slots to be received, satisfies

$$\frac{1}{n} \log M \leq \sum_{\tau=1}^n [\pi_{\tau} \tilde{H}(\mu_{\tau}, \frac{1}{\tau})] + \epsilon_n, \quad (9)$$

where $\sum_{\tau=1}^n \pi_{\tau}(\mu_{\tau} + \frac{1}{\tau}) = 1$, and for all τ , $0 \leq \mu_{\tau} \leq \frac{1}{2}$. In this expression, $\epsilon_n = \frac{1}{n}(H(P_e) + P_e \log_2(M - 1))$, π_{τ} is the

portion of time that user U_d sends packets with inter-arrival time equal to τ in the probe stream, and μ_τ is U_e^T 's average packet rate when the inter-arrival time is equal to τ .

Proof: We first we note that

$$\begin{aligned} \frac{1}{n} \log M &\stackrel{(a)}{=} \frac{1}{n} H(W) \\ &\stackrel{(b)}{=} \frac{1}{n} H(W|\tau^m) \\ &= \frac{1}{n} I(W; \hat{W}|\tau^m) + \frac{1}{n} H(W|\hat{W}, \tau^m) \\ &\stackrel{(c)}{\leq} \frac{1}{n} I(W; \hat{W}|\tau^m) + \epsilon_n \\ &\stackrel{(d)}{\leq} \frac{1}{n} I(X^m; Y^m|\tau^m) + \epsilon_n, \end{aligned}$$

where (a) holds because W is a uniform random variable over the set of messages $\{1, \dots, M\}$, (b) follows from the fact that the chosen message is independent of the inter-arrival time of decoder's packets, (c) follows from Fano's inequality with $\epsilon_n = \frac{1}{n}(H(P_e) + P_e \log_2(M-1))$, and (d) follows from data processing inequality in Markov chain in (2). Therefore,

$$\begin{aligned} \frac{1}{n} \log M &\leq \frac{1}{n} [H(X^m|\tau^m) - H(X^m|Y^m, \tau^m)] + \epsilon_n \\ &\leq \frac{1}{n} H(X^m|\tau^m) + \epsilon_n \\ &\leq \frac{1}{n} \sum_{j=1}^m H(X_j|\tau^m) + \epsilon_n \\ &\leq \frac{1}{n} \sum_{j=1}^m \max_{P_{X_j|\tau^m}} H(X_j|\tau^m) + \epsilon_n. \end{aligned}$$

In the maximization above, the mean of the distribution $P_{X_j|\tau^m}$ is $\mathbb{E}[X_j|\tau^m]$. As mentioned in Section III, in order to find the maximum information rate while having stability, we are interested in the asymptotic regime in which the operating point is converging to the line $r_e + r_d = 1$. Therefore, the information rate is upper bounded by having the set of means, $\{\mathbb{E}[X_1|\tau^m], \mathbb{E}[X_2|\tau^m], \dots, \mathbb{E}[X_m|\tau^m]\}$ satisfying the constraint $\frac{1}{n} \sum_{j=1}^m \mathbb{E}[X_j|\tau^m] + r_d = 1$. Let $\xi_j = \frac{\mathbb{E}[X_j|\tau^m]}{\tau_j}$. Using (4), we have

$$\max_{\substack{P_{X_j|\tau^m} \\ \mathbb{E}[X_j|\tau^m] = \tau_j \xi_j}} H(X_j|\tau^m) = \tau_j \tilde{H}(\xi_j, \frac{1}{\tau_j}), \quad (10)$$

where as mentioned in Lemma 1, the distribution for each X_j which achieves the maximum value in (10) is the tilted distribution of U_{τ_j} with parameter λ , such that $\lambda = (\psi'_{U_{\tau_j}})^{-1}(\tau_j \xi_j)$. Therefore, we will have:

$$\frac{1}{n} \log M \leq \frac{1}{n} \sum_{j=1}^m \tau_j \tilde{H}(\xi_j, \frac{1}{\tau_j}) + \epsilon_n,$$

such that the set $\{\xi_1, \xi_2, \dots, \xi_m\}$ satisfies the constraint $\frac{1}{n} \sum_{j=1}^m \tau_j \xi_j + r_d = 1$. The inter-arrival times take values in the set $\{1, 2, \dots, n\}$. Therefore, in the summation above we can fix the value of inter-arrival time on the value τ and count the number of times that τ_j has that value. Defining m_τ as the number of times that the inter-arrival time is equal to

τ (note that $n = \sum_{\tau=1}^n \tau \cdot m_\tau$), we can break the summation above as follows.

$$\begin{aligned} \frac{1}{n} \log M &\leq \frac{1}{n} \sum_{\tau=1}^n [\sum_{k=1}^{m_\tau} \tau \tilde{H}(\mu_{\tau,k}, \frac{1}{\tau})] + \epsilon_n \\ &= \frac{1}{n} \sum_{\tau=1}^n [\tau \sum_{k=1}^{m_\tau} \tilde{H}(\mu_{\tau,k}, \frac{1}{\tau})] + \epsilon_n \\ &= \frac{1}{n} \sum_{\tau=1}^n [\tau \cdot m_\tau \sum_{k=1}^{m_\tau} \frac{1}{m_\tau} \tilde{H}(\mu_{\tau,k}, \frac{1}{\tau})] + \epsilon_n, \end{aligned}$$

where $\mu_{\tau,k}$ is equal to the k -th ξ_j which has $\tau_j = \tau$.

By Lemma 3, the function $\tilde{H}(\cdot, \cdot)$ is a concave function of its first argument. Therefore, by Jensen's inequality,

$$\sum_{k=1}^{m_\tau} \frac{1}{m_\tau} \tilde{H}(\mu_{\tau,k}, \frac{1}{\tau}) \leq \tilde{H}(\mu_\tau, \frac{1}{\tau}), \quad (11)$$

where $\mu_\tau = \frac{1}{m_\tau} \sum_{k=1}^{m_\tau} \mu_{\tau,k}$. Using (11) and the equation $n = \sum_{\tau=1}^n \tau \cdot m_\tau$, we have

$$\begin{aligned} \frac{1}{n} \log M &\leq \sum_{\tau=1}^n [\frac{\tau \cdot m_\tau}{\sum_{\tau=1}^n \tau \cdot m_\tau} \tilde{H}(\mu_\tau, \frac{1}{\tau})] + \epsilon_n \\ &= \sum_{\tau=1}^n [\pi_\tau \tilde{H}(\mu_\tau, \frac{1}{\tau})] + \epsilon_n, \end{aligned} \quad (12)$$

where $\pi_\tau = (\tau \cdot m_\tau) / (\sum_{\tau=1}^n \tau \cdot m_\tau)$.

The packet rates of the users could be written as follows.

$$\begin{aligned} r_e &= \frac{1}{n} \sum_{j=1}^m \tau_j \xi_j = \frac{1}{n} \sum_{\tau=1}^n \sum_{k=1}^{m_\tau} \tau \mu_{\tau,k} \\ &= \frac{1}{n} \sum_{\tau=1}^n \tau m_\tau \frac{1}{m_\tau} \sum_{k=1}^{m_\tau} \mu_{\tau,k} = \sum_{\tau=1}^n \frac{1}{n} \tau m_\tau \mu_\tau \\ &= \sum_{\tau=1}^n \frac{\tau \cdot m_\tau}{\sum_{\tau=1}^n \tau \cdot m_\tau} \mu_\tau = \sum_{\tau=1}^n \pi_\tau \mu_\tau, \end{aligned}$$

and

$$r_d = \sum_{\tau=1}^n \pi_\tau \frac{1}{\tau}.$$

Therefore, the constraint could be written as follows.

$$\sum_{\tau=1}^n \pi_\tau (\mu_\tau + \frac{1}{\tau}) = 1. \quad (13)$$

Suppose the set of pairs $\{(\mu_\tau, \frac{1}{\tau})\}_{\tau=1}^n$ satisfies (13) and maximizes the right hand side of (12). By Lemma 4, there exists another set of pairs $\{(\hat{\mu}_\tau, \frac{1}{\tau})\}_{\tau=1}^n$ with $\hat{\mu}_\tau$ defined as

$$\hat{\mu}_\tau = \begin{cases} \mu_\tau & \text{if } 0 \leq \mu_\tau \leq \frac{1}{2}, \\ 1 - \mu_\tau & \text{if } \frac{1}{2} \leq \mu_\tau \leq 1, \end{cases}$$

that gives the same value for the right-hand side of (12), but it has $\sum_{\tau=1}^n \pi_\tau (\hat{\mu}_\tau + \frac{1}{\tau}) \leq \sum_{\tau=1}^n \pi_\tau (\mu_\tau + \frac{1}{\tau})$. Therefore, U_d can increase his packet rate and increase the information rate using the values $\hat{\mu}_\tau$. Hence, in the maximizing set, for all τ , we have $0 \leq \mu_\tau \leq \frac{1}{2}$. Therefore, the optimal operating

point will be on the line $r_e + r_d = 1$, with $0 \leq r_e \leq \frac{1}{2}$ and $\frac{1}{2} \leq r_d \leq 1$. $\square$

Applying Lemma 3, we can replace all pairs of form $(\mu_\tau, \frac{1}{\tau})$, $\tau \geq 2$, with a single pair of form $(\mu, \frac{1}{2})$.

Lemma 6: For any set of pairs $\mathcal{S} = \{(\mu_\tau, \frac{1}{\tau}), \tau \in [n]\}$, where for all τ , $0 \leq \mu_\tau \leq \frac{1}{2}$, with weights $\{\pi_\tau, \tau \in [n]\}$ with operating point on the line $0 \leq r_e \leq \frac{1}{2}$ and $\frac{1}{2} \leq r_d \leq 1$, there exists $0 \leq \alpha \leq 1$ and $0 \leq \gamma_1, \gamma_2 \leq \frac{1}{2}$ such that

$$\alpha(\gamma_1 + 1) + (1 - \alpha)(\gamma_2 + \frac{1}{2}) = 1,$$

and

$$\sum_{\tau=1}^n [\pi_\tau \tilde{H}(\mu_\tau, \frac{1}{\tau})] \leq \alpha \tilde{H}(\gamma_1, 1) + (1 - \alpha) \tilde{H}(\gamma_2, \frac{1}{2}).$$

Proof: For all $\tau \in \{3, \dots, n\}$, there exists $\beta_\tau \in [0, 1]$, such that

$$\beta_\tau(\mu_1, 1) + (1 - \beta_\tau)(\mu_\tau, \frac{1}{\tau}) = (\mu_2^\tau, \frac{1}{2}).$$

Clearly, the set $\{(\mu_1, 1), (\mu_2, \frac{1}{2}), (\mu_2^3, \frac{1}{2}), \dots, (\mu_2^n, \frac{1}{2})\}$ can also give the same operating point as $\mathcal{S}$ does. By Lemma 3,

$$\beta_\tau \tilde{H}(\mu_1, 1) + (1 - \beta_\tau) \tilde{H}(\mu_\tau, \frac{1}{\tau}) \leq \tilde{H}(\mu_2^\tau, \frac{1}{2}), \quad \forall \tau \in \{3, \dots, n\}.$$

Therefore,

$$\begin{aligned} & \sum_{\tau=1}^n \pi_\tau \tilde{H}(\mu_\tau, \frac{1}{\tau}) \\ &= \zeta_1 \tilde{H}(\mu_1, 1) + \zeta_2 \tilde{H}(\mu_2, \frac{1}{2}) + \sum_{\tau=3}^n \zeta_\tau (\beta_\tau \tilde{H}(\mu_1, 1) \\ & \quad + (1 - \beta_\tau) \tilde{H}(\mu_\tau, \frac{1}{\tau})) \\ &\leq \zeta_1 \tilde{H}(\mu_1, 1) + \zeta_2 \tilde{H}(\mu_2, \frac{1}{2}) + \sum_{\tau=3}^n \zeta_\tau \tilde{H}(\mu_2^\tau, \frac{1}{2}) \\ &\leq \zeta_1 \tilde{H}(\mu_1, 1) + (1 - \zeta_1) \tilde{H}(\frac{\zeta_2 \mu_2 + \sum_{\tau=3}^n \zeta_\tau \mu_2^\tau}{1 - \zeta_1}, \frac{1}{2}), \quad (14) \end{aligned}$$

where $\pi_1 = \zeta_1 + \sum_{\tau=3}^n \zeta_\tau \beta_\tau$, $\pi_2 = \zeta_2$ and $\pi_\tau = \zeta_\tau (1 - \beta_\tau)$ for $3 \leq \tau \leq n$, and we have used Lemma 3 again in the last inequality. $\square$

From Lemmas 5 and 6 we have

$$\begin{aligned} \frac{1}{n} \log M &\leq \alpha \tilde{H}(\gamma_1, 1) + (1 - \alpha) \tilde{H}(\gamma_2, \frac{1}{2}) + \epsilon_n \\ &\leq \max_{\alpha, \gamma_1, \gamma_2} \alpha \tilde{H}(\gamma_1, 1) + (1 - \alpha) \tilde{H}(\gamma_2, \frac{1}{2}) + \epsilon_n. \end{aligned}$$

Letting $n \rightarrow \infty$, ϵ_n goes to zero and we have

$$\begin{aligned} C &\leq \max_{\alpha, \gamma_1, \gamma_2} \alpha \tilde{H}(\gamma_1, 1) + (1 - \alpha) \tilde{H}(\gamma_2, \frac{1}{2}) \\ \text{s.t. } &\alpha(\gamma_1 + 1) + (1 - \alpha)(\gamma_2 + \frac{1}{2}) = 1, \end{aligned}$$

where $0 \leq \alpha \leq 1$ and $0 \leq \gamma_1, \gamma_2 \leq \frac{1}{2}$. This completes the proof of the converse part.

B. Achievability

The sequence of steps in our achievability scheme is as follows:

Generating the Codebook: The codebook $\mathcal{C}$ is generated by combining two codebooks $\mathcal{C}_1$ and $\mathcal{C}_2$, which are designed for the cases that the inter-arrival times in the probe stream are 1 and 2, respectively.

- Set $\alpha = 0.177 - \delta$, for a small and positive value of δ .
- To generate $\mathcal{C}_1$, fix a binary distribution P_1 such that $P_1(1) = 0.43$ and $P_1(0) = 0.57$. Generate a binary codebook $\mathcal{C}_1$ containing 2^{anR_1} sequences of length an of i.i.d. entries according to P_1 .
- To generate $\mathcal{C}_2$, fix a ternary distribution P_2 over set of symbols $\{a_0, a_1, a_2\}$ such that $P_2(a_0) = 0.43$, $P_2(a_1) = 0.325$ and $P_2(a_2) = 0.245$. Generate a ternary codebook $\mathcal{C}_2$ containing $2^{(1-\alpha)nR_2}$ sequences of length $\frac{1}{2}(1-\alpha)n$ of i.i.d. entries according to P_2 . Substituting a_0 with 00, a_1 with 10, and a_2 with 11, we will have $2^{(1-\alpha)nR_2}$ binary sequences of length $(1-\alpha)n$.
- Combine $\mathcal{C}_1$ and $\mathcal{C}_2$ to get $\mathcal{C}$, such that $\mathcal{C}$ has $2^{n(\alpha R_1 + (1-\alpha)R_2)}$ binary sequences of length n where we concatenate i -th row of $\mathcal{C}_1$ with j -th row of $\mathcal{C}_2$ to make the $((i-1)(2^{(1-\alpha)nR_2}) + j)$ -th row of $\mathcal{C}$ (note that $2^{(1-\alpha)nR_2}$ is the number of rows in $\mathcal{C}_2$). Rows of $\mathcal{C}$ are our codewords.

In above, n should be chosen such that anR_1 , an , $(1-\alpha)nR_2$ and $\frac{1}{2}(1-\alpha)n$ are all integers.

Encoding: To send message m , U_e^T sends the corresponding row of $\mathcal{C}$, that is, he sends the corresponding part of m from $\mathcal{C}_1$ in the first an time slots and the corresponding part of m from $\mathcal{C}_2$ in the rest of $(1-\alpha)n$ time slots.

Decoding: Recall that U_d^T sends a binary length n stream to the scheduler during the same length n time period. Here, U_d^T sends the stream of all ones (one packet in each time slot) in the first an time slots and sends bit stream of concatenated 10's for the rest of $(1-\alpha)n$ time slots. Assuming the queue is not empty,² since there is no noise in the system, the decoder can always learn the exact sequence sent by U_e .

Consequently, we will have:

$$\begin{aligned} C &\geq \frac{\log_2 2^{n(\alpha R_1 + (1-\alpha)R_2)}}{n} \\ &= \alpha R_1 + (1 - \alpha)R_2. \end{aligned}$$

In infinite block-length regime, where $n \rightarrow \infty$, we can choose $R_1 = H(P_1)$, $R_2 = \frac{1}{2}H(P_2)$ and find codebooks $\mathcal{C}_1$ and $\mathcal{C}_2$ such that this scheme satisfies the rate constraint. Therefore,

$$C \geq \alpha H(P_1) + (1 - \alpha) \frac{1}{2} H(P_2).$$

Substituting the values in the expression above, and letting δ go to zero, we see that the rate 0.8114 bits per time slot is achievable.

V. THREE-USER CASE

As an extension to the basic problem, in this section we consider the case that a third user is also using the shared

²Since in our achievable scheme U_d^T 's packets are spaced by either one or two time slots, it is enough to have one packet buffered in the queue, where since we are working in the heavy traffic regime, it will not be a problem.

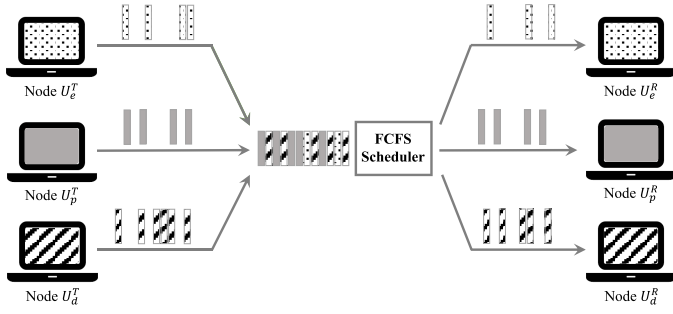


Fig. 4. Covert queueing channel in a system with three users.

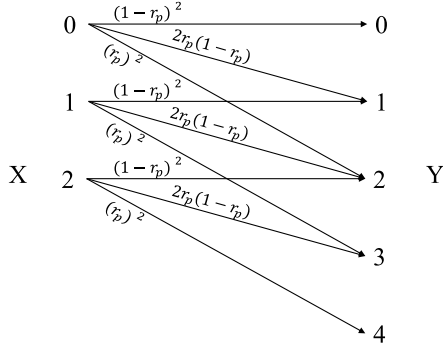


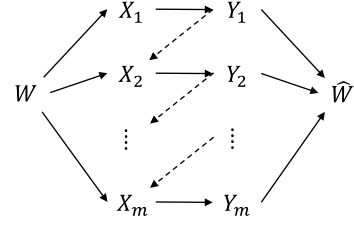
Fig. 5. Channel between the encoder and the decoder of the system for the case that the inter-arrival time of two packets of the probe stream is 2.

scheduler. We add user U_p to our basic system model. This user has nodes U_p^T and U_p^R as his transmitter and receiver nodes, respectively (Figure 4). We assume that node U_p^T sends packets according to a Bernoulli process with rate r_p to the shared scheduler. The shared scheduler is again assumed to be FCFS with service rate 1 and we analyze the capacity for coding schemes satisfying queueing stability condition in the asymptotic regime where the operating point is converging to the line $r_e + r_p + r_d = 1$. Also, in this section we consider the extra assumption that the inter-arrival time of the packets in the probe stream is upper bounded by the value τ_{max} .

Assuming that a sufficient number of packets are buffered in the shared queue, user U_d can still count the number of packets sent by the other two users between any of his own consecutive packets, yet he cannot distinguish between packets sent by user U_e and the packets sent by user U_p . Hence, user U_d has uncertainty in estimating the values of X . We model this uncertainty as a noise in receiving X . Suppose U_d^T sends two packets with $\tau_i = 2$. Each of the other users can possibly send at most 2 packets in the interval $[A_i, A_{i+1})$ and hence, $Y \in \{0, 1, 2, 3, 4\}$. Therefore, we have the channel shown in Figure 5 for this example. In the general case, for the inter-arrival time τ , given $X = x$, we have $Y \in \{x + 0, \dots, x + \tau\}$ such that

$$Pr(Y = i + x | X = x) = \binom{\tau}{i} (r_p)^i (1 - r_p)^{\tau-i}, \quad i \in \{0, \dots, \tau\}, \quad (15)$$

which is a binomial distribution $Bin(\tau, r_p)$. Therefore, the support of the random variable Y is $\{0, 1, \dots, 2\tau\}$. For the

Fig. 6. The graphical model representing the statistical relation between W , X^m , Y^m and $\hat{W}$.

mean of Y , we have

$$\mathbb{E}[Y|\tau] = \mathbb{E}[\mathbb{E}[Y|X, \tau]|\tau] = \mathbb{E}[X + \tau r_p|\tau] = \tau(r_e + r_p). \quad (16)$$

Because of user U_p 's stream, the encoder is not aware of the stream received at node U_e^R beforehand and this output can provide information to the encoder about U_p^T 's stream. The more packets node U_e^T sends to the scheduler, the more information this stream contains about U_p^T 's stream. Using this information, the encoder can have an estimation of the output of the channel at the decoder's side and hence, it could be considered as a noisy feedback to the encoder. Figure 6 shows the graphical model for random variables in our system.

The main result of this section is evaluation of the capacity of the introduced channel, presented in Theorem 2. In the following, the subscript r_p denotes that the calculation is done when the rate of U_p is r_p .

Theorem 2: If the rate of U_p is r_p , the capacity of the timing channel in a shared FCFS scheduler with service rate 1 depicted in Figure 4 is given by

$$C(r_p) = \max_{\alpha, \gamma_1, \gamma_2, \tau} \alpha \tilde{I}_{r_p}(\gamma_1, \frac{1}{\tau}) + (1 - \alpha) \tilde{I}_{r_p}(\gamma_2, \frac{1}{\tau + 1}) \quad (17)$$

$$s.t. \quad \alpha(\gamma_1 + \frac{1}{\tau}) + (1 - \alpha)(\gamma_2 + \frac{1}{\tau + 1}) = 1 - r_p,$$

where $0 \leq \alpha \leq 1$ and $0 \leq \gamma_1, \gamma_2 \leq 1$ and $1 \leq \tau \leq \tau_{max} - 1$. The function $\tilde{I}_{r_p} : [0, 1] \times \{\frac{1}{k} : k \in \mathbb{N}\} \mapsto [0, 1]$ is defined as

$$\tilde{I}_{r_p}(\gamma, \frac{1}{k}) = \frac{1}{k} \max_{\substack{X \in \{0, 1, \dots, k\} \\ \mathbb{E}[X] = k\gamma}} I_{r_p}(X; Y), \quad k \in \mathbb{N}, 0 \leq \gamma \leq 1. \quad (18)$$

The proof is based on converse and achievability arguments. Before giving the proof, we first investigate some of the properties of the function $\tilde{I}$.

Lemma 7: The function $\tilde{I}_{r_p}$ could be computed using the following expression.

$$\tilde{I}_{r_p}(\gamma, \frac{1}{k}) = \frac{1}{k} \check{H}_{r_p}(\gamma, \frac{1}{k}) - \frac{1}{k} H(Bin(k, r_p)), \quad (19)$$

where $\check{H}_{r_p}(\gamma, \frac{1}{k}) = \max_{\substack{X \in \{0, 1, \dots, k\} \\ \mathbb{E}[X] = k\gamma}} H_{r_p}(Y)$ and the second term is the entropy of the binomial distribution with parameters k and r_p .

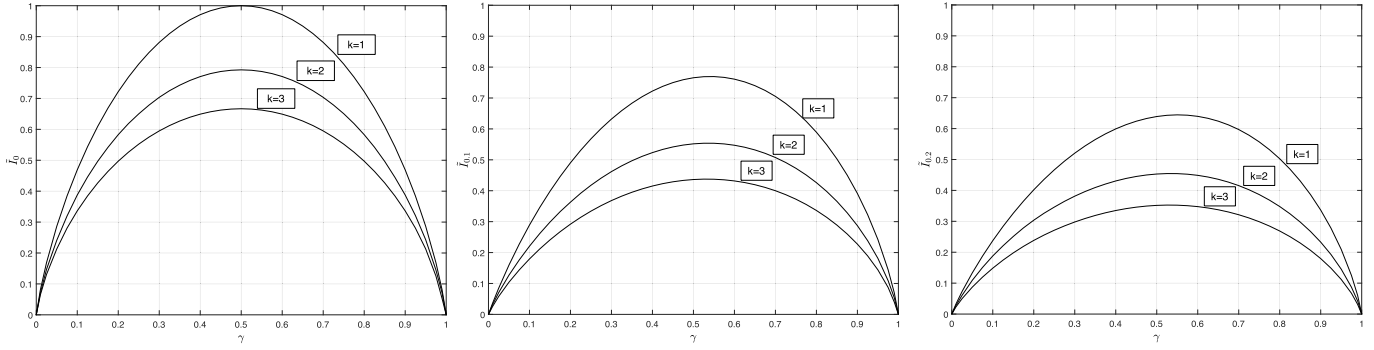


Fig. 7. $\tilde{I}_{r_p}(\gamma, \frac{1}{k})$ for different values of γ and $k \in \{1, 2, 3\}$ and $r_p \in \{0, 0.1, 0.2\}$.

See Appendix VI for the proof of Lemma 7.

In order to calculate $\tilde{H}_{r_p}(\gamma, \frac{1}{k})$, the following optimization problem should be solved.

$$\begin{aligned} \max_{P_X \geq 0} \quad & \log_2 e \sum_{i=0}^{2k} P_Y(i) \ln\left(\frac{1}{P_Y(i)}\right) \\ \text{s.t.} \quad & \begin{cases} \sum_{i=0}^k i P_X(i) = \mathbb{E}[X] = k\gamma, \\ \sum_{i=0}^k P_X(i) = 1, \end{cases} \end{aligned} \quad (20)$$

where $P_Y = P_X * P_{\text{Bin}(k, r_p)}$, that is,

$$P_Y(i) = \sum_{j=0}^k P_X(j) P_{\text{Bin}(k, r_p)}(i - j), \quad i \in \{0, 1, \dots, 2k\}. \quad (21)$$

Figure 7 shows the functions $\tilde{I}_0(\gamma, \frac{1}{k})$, $\tilde{I}_{0.1}(\gamma, \frac{1}{k})$ and $\tilde{I}_{0.2}(\gamma, \frac{1}{k})$ for different values of γ and $k \in \{1, 2, 3\}$.

Lemma 8: For all $0 \leq r_p \leq 1$, integers $1 \leq k_1, k_2, k_3 \leq \tau_{\max}$, values $0 \leq \gamma_1, \gamma_2, \gamma_3 \leq 1$, and $\alpha \in [0, 1]$, such that $\alpha(\gamma_1, \frac{1}{k_1}) + (1 - \alpha)(\gamma_3, \frac{1}{k_3}) = (\gamma_2, \frac{1}{k_2})$, we have

$$\alpha \tilde{I}_{r_p}(\gamma_1, \frac{1}{k_1}) + (1 - \alpha) \tilde{I}_{r_p}(\gamma_3, \frac{1}{k_3}) \leq \tilde{I}_{r_p}(\gamma_2, \frac{1}{k_2}). \quad (22)$$

See Appendix VI for the proof of Lemma 8.

Using the mentioned properties, the capacity of the timing channel in the shared FCFS scheduler of Figure 4 for different values of r_p can be calculated. Figure 8 shows the value of the capacity with respect to r_p .

The following proof of Theorem 2 is based on converse and achievability arguments.

A. Converse

For the converse part, similar to the approach in Section IV, first we find an upper bound on the information rate, which consists of a weighted summation of possible maximum information rates for different inter-arrival times of the packets in the probe stream, which satisfies the stability constraint; and then we further bound it with a summation which only corresponds to two inter-arrival times.

Suppose the rate of U_p is r_p . Similar to the proof of Lemma 5, for the timing channel in a shared FCFS scheduler with service rate 1 depicted in Figure 4, any code consisting

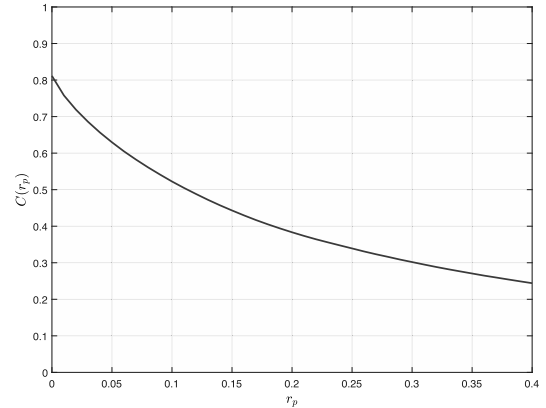


Fig. 8. Capacity of the timing channel in the shared FCFS scheduler of Figure 4 for different values of r_p .

of a codebook of M equiprobable binary codewords, where messages take on average n time slots to be received, satisfies

$$\begin{aligned} \frac{1}{n} \log M &\leq \frac{1}{n} I_{r_p}(W; \hat{W} | \tau^m) + \epsilon_n \\ &\stackrel{(a)}{\leq} \frac{1}{n} I_{r_p}(W; Y^m | \tau^m) + \epsilon_n, \end{aligned}$$

where $\epsilon_n = \frac{1}{n}(H(P_e) + P_e \log_2(M - 1))$ and (a) follows from data processing inequality in the model in Figure 6. Therefore,

$$\begin{aligned} \frac{1}{n} \log M &\leq \frac{1}{n} \sum_{j=1}^m I_{r_p}(W; Y_j | Y^{j-1} \tau^m) + \epsilon_n \\ &\leq \frac{1}{n} \sum_{j=1}^m I_{r_p}(W, Y^{j-1}; Y_j | \tau^m) + \epsilon_n \\ &\stackrel{(a)}{\leq} \frac{1}{n} \sum_{j=1}^m I_{r_p}(X_j; Y_j | \tau^m) + \epsilon_n \\ &\leq \frac{1}{n} \sum_{j=1}^m \max_{P_{X_j | \tau^m}} I_{r_p}(X_j; Y_j | \tau^m) + \epsilon_n, \end{aligned}$$

where (a) again follows from data processing inequality in the model in Figure 6. In the maximization above, the mean of the distribution $P_{X_j | \tau^m}$ is $\mathbb{E}[X_j | \tau^m]$ and in order to find the maximum information rate, the set of means, $\{\mathbb{E}[X_1 | \tau^m], \mathbb{E}[X_2 | \tau^m], \dots, \mathbb{E}[X_m | \tau^m]\}$, should satisfy

the constraint $r_e + r_p + r_d = 1$, that is, $\frac{1}{n} \sum_{j=1}^m \mathbb{E}[X_j | \tau^m] + r_p + r_d = 1$. Let $\zeta_j = \frac{\mathbb{E}[X_j | \tau^m]}{\tau_j}$. Using (18), we have

$$\max_{\substack{P_{X_j | \tau^m} \\ \mathbb{E}[X_j | \tau^m] = \tau_j \zeta_j}} I_{r_p}(X_j; Y_j | \tau^m) = \tau_j \tilde{I}_{r_p}(\zeta_j, \frac{1}{\tau_j}). \quad (23)$$

This implies that

$$\frac{1}{n} \log M \leq \frac{1}{n} \sum_{j=1}^m \tau_j \tilde{I}(\zeta_j, \frac{1}{\tau_j}) + \epsilon_n.$$

Next, using Lemma 8 and similar to the proof of Lemma 5, by breaking the summation, using Jensen's inequality and the equation $n = \sum_{\tau=1}^{\tau_{\max}} \tau \cdot m_\tau$, we will have

$$\frac{1}{n} \log M \leq \sum_{\tau=1}^{\tau_{\max}} \pi_\tau \tilde{I}_{r_p}(\mu_\tau, \frac{1}{\tau}) + \epsilon_n, \quad (24)$$

where μ_τ is the average of ζ_j 's which have $\tau_j = \tau$ and $\pi_\tau = (\tau \cdot m_\tau) / (\sum_{\tau=1}^{\tau_{\max}} \tau \cdot m_\tau)$. In this expression, π_τ could be interpreted as the portion of time that user U_d sends packets with inter-arrival time equal to τ .

Also, using the same approach as the one used in the proof of Lemma 5, the constraint of the problem could be written as follows.

$$\sum_{\tau=1}^{\tau_{\max}} \pi_\tau (\mu_\tau + \frac{1}{\tau}) = 1 - r_p. \quad (25)$$

Suppose the set of pairs $\mathcal{S} = \{(\mu_\tau, \frac{1}{\tau}), \tau \in [\tau_{\max}]\}$ with weights $\{\pi_\tau, \tau \in [\tau_{\max}]\}$ gives $\sum_{\tau=1}^{\tau_{\max}} \pi_\tau \tilde{I}_{r_p}(\mu_\tau, \frac{1}{\tau})$ and has its operating point on the line $r_e + r_d = 1 - r_p$, and we have $\frac{1}{\tau^*} \leq \sum_{\tau=1}^{\tau_{\max}} \pi_\tau \frac{1}{\tau} \leq \frac{1}{\tau^*+1}$, for some $1 \leq \tau^* \leq \tau_{\max} - 1$. We have

$$\begin{aligned} \beta_\tau (\mu_{\tau^*+1}, \frac{1}{\tau^*+1}) + (1 - \beta_\tau) (\mu_\tau, \frac{1}{\tau}) &= (\mu_{\tau^*}, \frac{1}{\tau^*}), \\ \tau &\leq \tau^* - 1, \\ \beta_\tau (\mu_{\tau^*}, \frac{1}{\tau^*}) + (1 - \beta_\tau) (\mu_\tau, \frac{1}{\tau}) &= (\mu_{\tau^*+1}, \frac{1}{\tau^*+1}), \\ \tau &\geq \tau^* + 2, \end{aligned}$$

for some $\beta_\tau \in [0, 1]$. Clearly, set $\{(\mu_{\tau^*}, \frac{1}{\tau^*}), \dots, (\mu_{\tau^*}^{\tau^*-1}, \frac{1}{\tau^*}), (\mu_{\tau^*}, \frac{1}{\tau^*}), (\mu_{\tau^*+1}, \frac{1}{\tau^*+1}), (\mu_{\tau^*+2}, \frac{1}{\tau^*+2}), \dots, (\mu_{\tau_{\max}}, \frac{1}{\tau_{\max}})\}$ can give the same operating point as $\mathcal{S}$ does. Therefore, using the technique presented in (14) and twice use of Lemma 8 we have

$$\begin{aligned} \frac{1}{n} \log M &\leq \alpha \tilde{I}_{r_p}(\gamma_1, \frac{1}{\tau^*}) + (1 - \alpha) \tilde{I}_{r_p}(\gamma_2, \frac{1}{\tau^*+1}) + \epsilon_n \\ &\leq \max_{\alpha, \gamma_1, \gamma_2, \tau} \alpha \tilde{I}_{r_p}(\gamma_1, \frac{1}{\tau}) + (1 - \alpha) \tilde{I}_{r_p}(\gamma_2, \frac{1}{\tau+1}) + \epsilon_n. \end{aligned}$$

Letting $n \rightarrow \infty$, ϵ_n goes to zero and we get the desired result.

B. Achievability

For a given value of r_p , we first use the expression obtained in the converse to obtain optimal choice for parameters α , γ_1 , γ_2 and τ . Because of Lemma 8, in order to find the optimal τ , we can start with $\tau = 1$ and optimize other parameters and then calculate $\alpha \tilde{I}_{r_p}(\gamma_1, \frac{1}{\tau}) + (1 - \alpha) \tilde{I}_{r_p}(\gamma_2, \frac{1}{\tau+1})$ and in each step, increase the value of τ by 1, stopping whenever the

obtained value is decreased compared to the previous step. For instance, for $r_p \leq 0.1$, the optimal τ is 1, and hence the procedure stops after checking two steps. After calculating optimal parameters α^* , γ_1^* , γ_2^* and τ^* , the optimal input distributions P_1^* and P_2^* could be obtained using optimization problem (20).

Achieving the proposed upper bound could be done by a method exactly similar to the one used in Subsection IV-B. The sequence of steps in our achievability scheme is as follows.

Generating the Codebook: The codebook $\mathcal{C}$ is generated by combining two codebooks $\mathcal{C}_1$ and $\mathcal{C}_2$, which are designed for the cases that the inter-arrival times in the probe stream are τ^* and $\tau^* + 1$, respectively.

- Set $\alpha = \alpha^* - \delta$, for a small and positive value of δ .
- To generate $\mathcal{C}_1$, consider the $(\tau^* + 1)$ -ary distribution P_1^* over set of symbols $\{a_0, \dots, a_{\tau^*}\}$. Generate a $(\tau^* + 1)$ -ary codebook $\mathcal{C}_1$ containing 2^{anR_1} sequences of length $\frac{1}{\tau^*+1}an$ of i.i.d. entries according to P_1^* . Substitute a_i with a binary sequence of i 1's followed by $\tau^* - i$ 0's. Therefore, we will have 2^{anR_1} binary sequences of length an .
- To generate $\mathcal{C}_2$, consider the $(\tau^* + 2)$ -ary distribution P_2^* over set of symbols $\{a_0, \dots, a_{\tau^*+1}\}$. Generate a $(\tau^* + 2)$ -ary codebook $\mathcal{C}_2$ containing $2^{(1-\alpha)nR_2}$ sequences of length $\frac{1}{\tau^*+1}(1 - \alpha)n$ of i.i.d. entries according to P_2^* . Substitute a_i with a binary sequence of i 1's followed by $\tau^* + 1 - i$ 0's. Therefore, we will have $2^{(1-\alpha)nR_2}$ binary sequences of length $(1 - \alpha)n$.
- Similar to the approach in Subsection IV-B, combine $\mathcal{C}_1$ and $\mathcal{C}_2$ to get $\mathcal{C}$, such that $\mathcal{C}$ has $2^{n(\alpha R_1 + (1-\alpha)R_2)}$ binary sequences of length n . Rows of $\mathcal{C}$ are our codewords.

The encoding and decoding parts are exactly similar to the approach in Subsection IV-B, and we avoid repeating them, except that here U_d^T sends a sequence with inter-arrival times of τ^* in the first an time slots and $\tau^* + 1$ for the rest of $(1 - \alpha)n$ time slots. Note that unlike the case with two users, here due to the presence of user U_p , we have noise in the channel. A standard random coding approach [28, Ch. 7] shows that in infinite block-length regime, where $n \rightarrow \infty$, we can choose $R_1 = \frac{1}{\tau^*} I_{r_p}(X; Y)$, with X having distribution P_1^* , with channel described in (15), i.e., $R_1 = \tilde{I}_{r_p}(\gamma_1^*, \frac{1}{\tau^*})$, and $R_2 = \frac{1}{\tau^*+1} I_{r_p}(X; Y)$, with X having distribution P_2^* , with channel described in (15), i.e., $R_2 = \tilde{I}_{r_p}(\gamma_2^*, \frac{1}{\tau^*+1})$. Therefore, we have

$$C \geq \alpha \tilde{I}_{r_p}(\gamma_1^*, \frac{1}{\tau^*}) + (1 - \alpha) \tilde{I}_{r_p}(\gamma_2^*, \frac{1}{\tau^*+1}).$$

Letting δ go to zero, completes the achievability proof.

VI. CONCLUSION

We studied convert queueing channels (CQCs) that can occur through delays experienced by users who are sharing a scheduler. As the scheduling policy plays a crucial role in the possible information transmission rate in this type of channel, we focused on deterministic and work-conserving FCFS scheduling policy. An information-theoretic framework was proposed to derive the capacity of the CQC under this

scheduling policy. First, we studied a system comprising a transmitter and a receiver that share a deterministic and work-conserving FCFS scheduler. We obtained the maximum information transmission rate in this CQC and showed that an information leakage rate as high as 0.8114 bits per time slot is possible. We also considered the effect of the presence of other users on the information transmission rate of this channel. We extended the model to include a third user who also uses the shared resource and studied the effect of the presence of this user on the information transmission rate. The solution approach presented in this extension can be applied to calculate the capacity of the covert queueing channel among any number of users. The achievable information transmission rates obtained from this study demonstrate the possibility of significant information leakage and great privacy threats brought by CQCs in FCFS schedulers. Based on this result, special attention must be paid to CQCs in high security systems. Finding the capacity of CQCs under other scheduling policies, especially non-deterministic policies, remains to be done in the research area of covert communications and is considered as the main direction for future work. Furthermore, a comprehensive study is required to design suitable scheduling policies that can simultaneously guarantee adequate levels of both security and throughput.

APPENDIX A PROOF OF STABILITY

Consider the system model with M users and service rate ρ , and let $\mu = \sum_{i=1}^M r_i$. We denote arrival, service and queue length at time k , with $a(k)$, $s(k)$ and $q(k)$, respectively, and we have

$$q(k+1) = (q(k) + a(k) - s(k))^+.$$

Using Foster-Lyapunov theorem with Lyapunov function $V(q(k)) = (q(k))^2$ and calculating the drift, we have

$$\begin{aligned} \mathbb{E}[q^2(k+1) - q^2(k) | q(k) = q] &\leq \mathbb{E}[(q + a - s)^2 - q^2] \\ &= \mathbb{E}[2q(a - s)] + \mathbb{E}[(a - s)^2], \end{aligned}$$

where $\mathbb{E}[(a - s)^2]$ is a constant and we denote it by K . Therefore, for some $\epsilon > 0$, if $\mu < \rho$, for large enough value of q , we have

$$\mathbb{E}[q^2(k+1) - q^2(k) | q(k) = q] \leq 2q(\mu - \rho) + K \leq -\epsilon,$$

which implies the stability.

APPENDIX B PROOF OF LEMMA 1

In order to find the optimum distribution, P_X , the optimization problem could be written as follows.

$$\begin{aligned} \max_{P_X \geq 0} \quad & \log_2 e \sum_{i=0}^k P_X(i) \ln\left(\frac{1}{P_X(i)}\right) \\ \text{s.t.} \quad & \begin{cases} \sum_{i=0}^k i P_X(i) = \mathbb{E}[X] = k\gamma, \\ \sum_{i=0}^k P_X(i) = 1, \end{cases} \end{aligned} \quad (26)$$

which could be solved using the Lagrange multipliers method. The Lagrangian function would be as follows.

$$\begin{aligned} \sum_{i=0}^k P_X(i) \ln\left(\frac{1}{P_X(i)}\right) + \lambda \left(\sum_{i=0}^k i P_X(i) - k\gamma \right) \\ + \rho \left(\sum_{i=0}^k P_X(i) - 1 \right). \end{aligned}$$

Setting the derivative with respect to $P_X(i)$ equal to zero, we get $\ln\left(\frac{1}{P_X(i)}\right) - 1 + i\lambda + \rho = 0$, which implies that

$$P_X(i) = e^{\rho-1} \cdot e^{i\lambda}. \quad (27)$$

Also, from the second constraint we have

$$\sum_{i=0}^k e^{\rho-1} \cdot e^{i\lambda} = 1 \Rightarrow e^{\rho-1} = \frac{1}{\sum_{i=0}^k e^{i\lambda}}. \quad (28)$$

Combining (27) and (28), we have

$$P_X(i) = \frac{e^{i\lambda}}{\sum_{i=0}^k e^{i\lambda}},$$

which is the tilted distribution of U_k with parameter λ . In order to calculate λ , from the first constraint,

$$\begin{aligned} k\gamma &= \sum_{i=0}^k i P_X(i) = \sum_{i=0}^k i \frac{e^{i\lambda}}{\sum_{i=0}^k e^{i\lambda}} = \frac{\sum_{i=0}^k i e^{i\lambda}}{\sum_{i=0}^k e^{i\lambda}} \\ &= \frac{\mathbb{E}[U_k e^{U_k \lambda}]}{\mathbb{E}[e^{U_k \lambda}]} = \frac{d}{d\lambda} (\ln \mathbb{E}[e^{U_k \lambda}]) = \psi'_{U_k}(\lambda). \end{aligned}$$

APPENDIX C PROOF OF LEMMA 3

First we note that

$$\begin{aligned} \tilde{H}\left(\gamma, \frac{1}{k}\right) &= \frac{1}{k} [\log_2(k+1) - \psi_{U_k}^*(k\gamma) \log_2 e] \\ &= \left[\frac{1}{k} \log_2(k+1) - \frac{1}{k} \sup_{\lambda} \{k\gamma \lambda - \log\left(\frac{\sum_{i=0}^k e^{i\lambda}}{k+1}\right)\} \log_2 e \right] \\ &= -\sup_{\lambda} \{ \gamma \lambda \log_2 e - \frac{1}{k} \log_2 \left(\sum_{i=0}^k e^{i\lambda} \right) \}. \end{aligned}$$

Therefore, if we can show that for any given λ the function $h(\gamma, \frac{1}{k}) = \gamma \lambda \log_2 e - \frac{1}{k} \log_2 \left(\sum_{i=0}^k e^{i\lambda} \right)$ is convex, then since the supremum of convex functions is convex, we can conclude the desired concavity of the function $\tilde{H}(\cdot, \cdot)$.

Noting that $\frac{1}{k} \log_2 \left(\sum_{i=0}^k e^{i\lambda} \right) = \frac{1}{k} \log_2 \left(\frac{1-e^{(k+1)\lambda}}{1-e^\lambda} \right)$, to prove the convexity of $h(\cdot, \cdot)$, it suffices to prove that the function $g(x) = x \log\left(\frac{1-e^{(\frac{1}{x}+1)\lambda}}{1-e^\lambda}\right)$, $0 < x \leq 1$, is concave. This is true from the concavity of the function $\hat{g}(x) = \log\left(\frac{1-e^{(x+1)\lambda}}{1-e^\lambda}\right)$, and the fact that for any function f , $xf(\frac{1}{x})$ is concave if $f(x)$ is concave. The concavity of the function $\hat{g}(\cdot)$ can be easily seen by taking its second derivative.

APPENDIX D PROOF OF LEMMA 4

For a given γ and support set $\{0, 1, \dots, k\}$, suppose the distribution P_X^* is defined over $\{0, 1, \dots, k\}$ and has mean $\mathbb{E}_{P^*}[X] = k\gamma$ and

$$\max_{\substack{X \in \{0, 1, \dots, k\} \\ \mathbb{E}[X] = k\gamma}} \frac{1}{k} H(X) = \frac{1}{k} H(P_X^*).$$

Define distribution Q_X as $Q_X(i) = P_X^*(k-i)$, $0 \leq i \leq k$. Therefore the entropy of Q_X will be the same as the entropy of P_X^* and we have

$$\begin{aligned} \mathbb{E}_{Q^*}[X] &= \sum_{i=0}^k i Q_X(i) = \sum_{i=0}^k i P_X^*(k-i) \\ &= - \sum_{i=0}^k (-k + (k-i)) P_X^*(k-i) \\ &= k - \sum_{i=0}^k (k-i) P_X^*(k-i) = k - k\gamma = k(1-\gamma). \end{aligned}$$

Hence, we have

$$\begin{aligned} \tilde{H}(\gamma, \frac{1}{k}) &= \max_{\substack{X \in \{0, 1, \dots, k\} \\ \mathbb{E}[X] = k\gamma}} \frac{1}{k} H(X) = \frac{1}{k} H(P_X^*) = \frac{1}{k} H(Q_X) \\ &\leq \max_{\substack{X \in \{0, 1, \dots, k\} \\ \mathbb{E}[X] = k(1-\gamma)}} \frac{1}{k} H(X) = \tilde{H}(1-\gamma, \frac{1}{k}). \end{aligned} \quad (29)$$

Similarly, suppose for the distribution Q_X^* , defined over $\{0, 1, \dots, k\}$ and with mean $\mathbb{E}_{Q^*}[X] = k(1-\gamma)$,

$$\max_{\substack{X \in \{0, 1, \dots, k\} \\ \mathbb{E}[X] = k(1-\gamma)}} \frac{1}{k} H(X) = \frac{1}{k} H(Q_X^*).$$

Define distribution P_X as $P_X(i) = Q_X^*(k-i)$, $0 \leq i \leq k$. Therefore the entropy of P_X will be the same as the entropy of Q_X^* and we have

$$\begin{aligned} \mathbb{E}_P[X] &= \sum_{i=0}^k i P_X(i) = \sum_{i=0}^k i Q_X^*(k-i) \\ &= - \sum_{i=0}^k (-k + (k-i)) Q_X^*(k-i) \\ &= k - \sum_{i=0}^k (k-i) Q_X^*(k-i) = k - k(1-\gamma) = k\gamma. \end{aligned}$$

Hence, we have

$$\begin{aligned} \tilde{H}(1-\gamma, \frac{1}{k}) &= \max_{\substack{X \in \{0, 1, \dots, k\} \\ \mathbb{E}[X] = k(1-\gamma)}} \frac{1}{k} H(X) = \frac{1}{k} H(Q_X^*) = \frac{1}{k} H(P_X) \\ &\leq \max_{\substack{X \in \{0, 1, \dots, k\} \\ \mathbb{E}[X] = k\gamma}} \frac{1}{k} H(X) = \tilde{H}(\gamma, \frac{1}{k}). \end{aligned} \quad (30)$$

Comparing (29) and (30) gives the desired result.

APPENDIX E PROOF OF LEMMA 7

$$\begin{aligned} \tilde{I}_{r_p}(\gamma, \frac{1}{k}) &= \max_{\substack{X \in \{0, 1, \dots, k\} \\ \mathbb{E}[X] = k\gamma}} \frac{1}{k} I_{r_p}(X; Y) \\ &= \max_{\substack{X \in \{0, 1, \dots, k\} \\ \mathbb{E}[X] = k\gamma}} \frac{1}{k} [H_{r_p}(Y) - H_{r_p}(Y|X)] \\ &= \max_{\substack{X \in \{0, 1, \dots, k\} \\ \mathbb{E}[X] = k\gamma}} \frac{1}{k} [H_{r_p}(Y) - \sum_{x=0}^k P_X(x) H_{r_p}(Y|X=x)] \\ &\stackrel{(a)}{=} \max_{\substack{X \in \{0, 1, \dots, k\} \\ \mathbb{E}[X] = k\gamma}} \frac{1}{k} [H_{r_p}(Y) - \sum_{x=0}^k P_X(x) H(\text{Bin}(k, r_p))] \\ &= \max_{\substack{X \in \{0, 1, \dots, k\} \\ \mathbb{E}[X] = k\gamma}} \frac{1}{k} [H_{r_p}(Y) - H(\text{Bin}(k, r_p))] \\ &= \max_{\substack{X \in \{0, 1, \dots, k\} \\ \mathbb{E}[X] = k\gamma}} \frac{1}{k} H_{r_p}(Y) - \frac{1}{k} H(\text{Bin}(k, r_p)) \\ &= \frac{1}{k} \tilde{H}_{r_p}(\gamma, \frac{1}{k}) - \frac{1}{k} H(\text{Bin}(k, r_p)), \end{aligned}$$

where (a) follows from (15).

APPENDIX F PROOF OF LEMMA 8

We first prove that the function $\tilde{I}(\cdot, \cdot)$ is concave in its first argument. Let $P_{X_1}^*$ and $P_{X_3}^*$ be the optimum distributions resulted from optimization problem (20) for parameters $(\gamma_1, \frac{1}{k})$ and $(\gamma_3, \frac{1}{k})$, respectively. Therefore for any $0 \leq \alpha \leq 1$,

$$\begin{aligned} \alpha \tilde{I}_{r_p}(\gamma_1, \frac{1}{k}) + (1-\alpha) \tilde{I}_{r_p}(\gamma_3, \frac{1}{k}) &\stackrel{(a)}{=} \frac{1}{k} \alpha H(P_{X_1}^* * \text{Bin}(k, r_p)) + \frac{1}{k} (1-\alpha) H(P_{X_3}^* * \text{Bin}(k, r_p)) \\ &\quad - \frac{1}{k} H(\text{Bin}(k, r_p)) \\ &\stackrel{(b)}{\leq} \frac{1}{k} H(\alpha(P_{X_1}^* * \text{Bin}(k, r_p)) + (1-\alpha)(P_{X_3}^* * \text{Bin}(k, r_p))) \\ &\quad - \frac{1}{k} H(\text{Bin}(k, r_p)) \\ &\leq \frac{1}{k} \max_{\substack{X \in \{0, 1, \dots, k\} \\ \mathbb{E}[X] = k(\alpha\gamma_1 + (1-\alpha)\gamma_3)}} H(P_X * \text{Bin}(k, r_p)) \\ &\quad - \frac{1}{k} H(\text{Bin}(k, r_p)) \\ &= \tilde{I}_{r_p}(\alpha\gamma_1 + (1-\alpha)\gamma_3, \frac{1}{k}), \end{aligned}$$

where (a) follows from Lemma 7 and (b) follows from the concavity of the entropy function.

Because of the complexity and lack of symmetry or structure in the function $\tilde{I}$, there is no straightforward analytic method for proving its concavity. But we notice that it suffices to show that for all $2 \leq k \leq \tau_{\max} - 1$, and α such that

$$\alpha \frac{1}{k-1} + (1-\alpha) \frac{1}{k+1} = \frac{1}{k}, \quad (31)$$

we have

$$\begin{aligned} \alpha \tilde{I}_{r_p}(\gamma_1, \frac{1}{k-1}) + (1-\alpha) \tilde{I}_{r_p}(\gamma_3, \frac{1}{k+1}) \\ \leq \tilde{I}_{r_p}(\alpha\gamma_1 + (1-\alpha)\gamma_3, \frac{1}{k}). \end{aligned} \quad (32)$$

From (31) we have $\alpha = \frac{k-1}{2k}$, hence using Lemma 7, (32) reduces to

$$\begin{aligned} 2\check{H}_{r_p}(\gamma_2, \frac{1}{k}) - \check{H}_{r_p}(\gamma_1, \frac{1}{k-1}) \\ - \check{H}_{r_p}(\gamma_3, \frac{1}{k+1}) + f(k, r_p) \geq 0, \end{aligned} \quad (33)$$

where $f(k, r_p) = H(\text{Bin}(k-1, r_p)) + H(\text{Bin}(k+1, r_p)) - 2H(\text{Bin}(k, r_p))$. Noting that the left-hand side is a Lipschitz continuous function of γ_1, γ_3 (away from zero), and r_p and the fact that k takes finitely many values, the validation of inequality (33) can be done numerically.

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