

## Development of region-specific soil behavior type index correlations for evaluating liquefaction hazard in Christchurch, New Zealand



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### ABSTRACT

Utilizing an unprecedented database of field and laboratory test data from Christchurch, New Zealand, this study develops deterministic and probabilistic correlations relating the soil behavior type index ( $I_c$ ) to laboratory test-based liquefaction susceptibility and fines content. The proposed  $I_c$  correlations are in turn used to assess liquefaction hazard for 9623 case studies compiled from three earthquakes that impacted Christchurch, New Zealand, wherein the predicted liquefaction severity is compared to post-earthquake field observations. Furthermore, the accuracy of the predictions based on the region-specific  $I_c$  correlations are compared to those made using generic  $I_c$  correlations. The use of the Christchurch-specific  $I_c$  correlations, in general, only slightly improved the liquefaction severity predictions. These findings could be a result of shortcomings in the procedures used to predict liquefaction triggering and/or liquefaction manifestations. Nevertheless, the findings give credence to the use of generic correlations for typical projects. Finally, the overall framework of the study used herein can be applied to worldwide locals and is not just limited for use in the Canterbury region of New Zealand.

### 1. Introduction

Semi-empirical models based on post-earthquake field observations have become the standard of practice worldwide for predicting soil liquefaction. Since the inception of the Standard Penetration Test (SPT) based “simplified-procedure” by Whitman [62] and Seed and Idriss [55], variants based on other in-situ test metrics have been developed, including the Cone Penetration Test (CPT) (e.g., Stark and Olson [58]; Robertson and Wride [53]; Moss et al. [41]; Idriss and Boulanger [27]; and Boulanger and Idriss [8]; among others), small strain shear wave velocity ( $V_s$ ) (e.g., Andrus and Stokoe [1]; Kayen et al. [30]), flat plate dilatometer indices (e.g., Monaco et al. [40]), and others. Although the use of SPT-based variants of the simplified procedure (e.g., Youd et al. [63]; Cetin et al. [15]; Idriss and Boulanger [27]; and Boulanger and Idriss [8]) still predominates practice in much of the world, the CPT-based procedures offer significant advantages (NRC [43]): “*CPT soundings offer advantages over other methods of estimating liquefaction resistance in both the detection of thin layers that may affect liquefaction triggering and subsequent pore pressure redistribution and in the reproducibility of results. CPT results are less dependent on the equipment operator or setup than most other in situ test methods, and CPT can be performed with relative speed and economy.*” However, CPT-based variants of the simplified procedure have disadvantages too. Namely, because soil samples

are not recovered during CPT sounding, soils are often not characterized directly or tested further in the laboratory. While discrete, disturbed soil specimens may be obtained using a CPT-push sampler at sounding sites (Robertson [52]), the standard-of-practice is to infer from CPT indices certain soil properties and behaviors relevant to liquefaction evaluations, including liquefaction susceptibility and fines content ( $FC$ ).

The focus of this study is the use of the CPT soil behavior type index ( $I_c$ ) for assessing both liquefaction susceptibility and  $FC$  in conjunction with the Boulanger and Idriss (BI14) [8] CPT-based liquefaction evaluation procedure. Towards this end, the authors use a wealth of data from the 2010 – 2011 Canterbury, New Zealand, earthquake sequence (CES) and the 2016  $M_w 5.7$  Valentine's Day earthquake that impacted Christchurch, New Zealand (e.g., Cubrinovski and Green [19]; Cubrinovski et al. [20]) to develop region-specific deterministic and probabilistic  $I_c$  correlations for predicting liquefaction susceptibility and  $FC$ . These correlations are in turn used in conjunction with the BI14 CPT-based liquefaction evaluation procedure, operating within the Liquefaction Potential Index ( $LPI$ ) framework (Iwasaki et al. [28]), to predict surficial liquefaction manifestations during three Canterbury events. These predictions are compared to field observations and are evaluated against the results of parallel analyses that use generic  $I_c$  correlations. Furthermore, the optimal  $I_c$  values for defining

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liquefaction susceptibility based on correlations with laboratory test-based criteria and based on optimizing the accuracy of field predictions are compared. This leads to an assessment of the benefit of developing region-specific correlations. It should be noted that while this study focuses on BI14 and the Canterbury, New Zealand earthquakes, the overall framework of the study could be applied to other CPT-based procedures and to worldwide locals.

## 2. Background

### 2.1. Canterbury Earthquakes

The 2010–2011 CES and the 2016 Valentine's Day earthquake (henceforth, collectively referred to as the Canterbury earthquakes) resulted in liquefaction case-history data of unprecedented quantity and quality, presenting a unique opportunity to advance the science of liquefaction hazard assessment. A summary of the Canterbury earthquakes, to include tectonic and geologic settings, seismology, and environmental effects, is provided by Quigley et al. [50]. The present study uses data from three of the Canterbury earthquakes: the  $M_w7.1$ , 4 Sept. 2010 Darfield earthquake, the  $M_w6.2$ , 22 Feb. 2011 Christchurch earthquake, and the  $M_w5.7$ , 14 Feb. 2016 Valentine's Day earthquake. Ground motions from these events were recorded by a dense network of strong motion stations (e.g., Cousins and McVerry [18]; Bradley and Cubrinovski [12]; Bradley [9]), and due to the impact of liquefaction during the 2010–2011 CES, the New Zealand Earthquake Commission (EQC) funded an extensive geotechnical reconnaissance and characterization program (CERA [14]) that included ground water modeling and in-situ and laboratory soil testing. The resulting data are compiled in the New Zealand Geotechnical Database (NZGD [44]); the data used herein have been made available in this database.

### 2.2. Soil Behavior Index, $I_c$

First proposed by Jeffries and Davies [29] to define soil type boundaries,  $I_c$  was subsequently modified by Robertson and Wride [53] to better fit the Robertson [51]  $Q$ – $F$  classification scheme (Fig. 1), with the latter version becoming widely used in practice. Per Robertson and Wride [53]  $I_c$  is defined as:

$$I_c = \sqrt{(3.47 - \log_{10}Q)^2 + (1.22 + \log_{10}F)^2} \quad (1)$$

where  $Q$  and  $F$  are the normalized CPT penetration resistance and normalized CPT friction ratio, respectively. As shown in Fig. 1,  $I_c$  represents the radial distance between any point on this chart and the point defined by  $Q = 10^{3.47}$  and  $F = 10^{1.22}$  (i.e.,  $Q = 2951$  and  $F = 0.06026\%$ ). Circular arcs defined by constant  $I_c$  values approximate the boundaries between different soil behavior types, with  $I_c = 2.60$  shown in Fig. 1 as the approximate boundary between soil behavior types 4 and 5 (4: Silt Mixtures; 5: Sand Mixtures) (Robertson [52]).

Because the CPT is affected by many factors influencing soil behavior, such as soil plasticity, sensitivity, mineralogy, friction angle, age, cementation, and stress history (Robertson [51]), the relationship shown in Fig. 1 between  $I_c$  and global soil type is approximate and can vary regionally. Soil behavior type index correlations derived from global data can thus perform poorly on local scales. The applicability of such correlations to a region of interest should therefore be evaluated when possible using regional datasets, such as that compiled after the Canterbury earthquakes, and the correlations calibrated as necessary. Moreover, to assess liquefaction hazards in a fully probabilistic manner, the uncertainties of inferred soil properties must also be quantified.

### 2.3. Use of susceptibility criteria and $FC$ in liquefaction-potential evaluations

Antecedent to using any CPT-based liquefaction evaluation

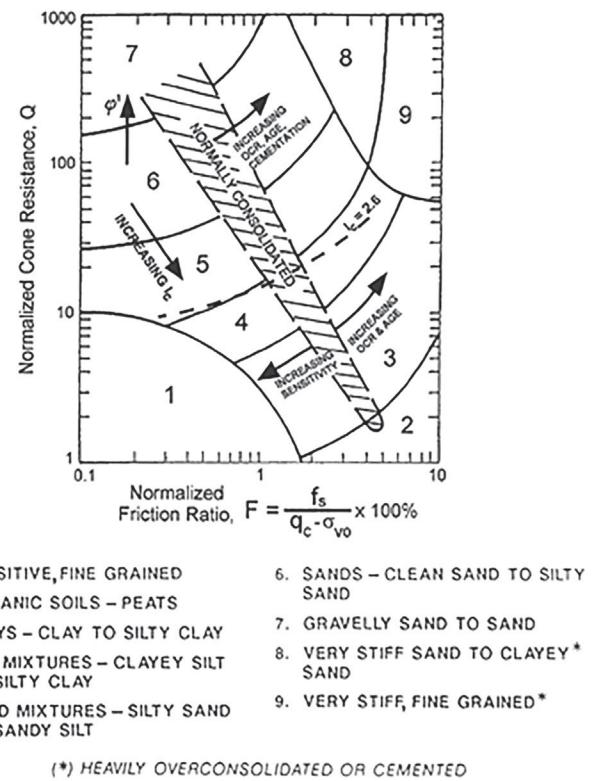


Fig. 1. CPT-based Soil Behavior Type chart (Robertson [51]; Robertson and Wride [53]).

procedure (e.g., BI14), liquefaction-susceptible soils must first be identified. CPT-based liquefaction procedures are intended to evaluate the potential for liquefaction triggering in soils susceptible to liquefaction. They should not be applied to high plasticity, fine grained, “non-liquefiable” soils, which could result in less accurate liquefaction hazard assessments (Maurer et al. [35]), and for which other more appropriate methods exist to predict cyclic behavior (e.g., Boulanger and Idriss [7]).

Various studies have proposed criteria for assessing liquefaction susceptibility based on laboratory test indices (e.g., Seed and Idriss [56]; Polito [47]; Seed et al. [57]; Bray and Sancio [13]; Boulanger and Idriss [6]). Overviews of commonly used susceptibility criteria are given in Armstrong and Malwick [2,3], Green and Ziotopoulou [26], and NRC [43], among others. However, this study focuses on the Boulanger and Idriss (BI06) [6] criterion because it was explicitly developed to determine the most appropriate analysis procedure for evaluating cyclic behavior, based on whether the soil's expected cyclic response is “sand-like” or “clay-like.” Soils that exhibit sand-like behavior experience a significant stiffness reduction at low to moderate strains and elevated excess pore pressures. In contrast, soils that exhibit clay-like behavior retain moderate stiffness at all levels of strain and excess pore pressure. Boulanger and Idriss [6] proposed that soils with Plasticity Index ( $PI$ )  $\leq 3$  exhibit sand-like behavior; soils with  $PI \geq 7$  exhibit clay-like behavior. However, if a soil classifies as CL-ML according to the Unified Soil Classification System (USCS) (ASTM D-2487-11, [4]), the criterion for clay-like behavior may reduce to  $PI \geq 5$ . Soils with  $3 < PI < 7$  (or 5 for CL-ML) may exhibit intermediate behavior and should be tested further. The liquefaction potential of soils that are determined to be susceptible to liquefaction per BI06 can be evaluated using the BI14 CPT-based simplified procedure, among others. However, inherent to implementing BI14 the  $FC$  of the soil needs to be determined/estimated. Accordingly,  $FC$  must be measured or estimated for every soil stratum expected to have “sand-like” cyclic response.

Thus, to assess liquefaction hazard per the BI14 liquefaction

evaluation procedure requires that samples be collected and tested to measure *PI* (to determine liquefaction susceptibility) and *FC* (to evaluate liquefaction potential of soils susceptible to liquefaction). However, continuous sampling and laboratory index-testing is prohibitively expensive for many projects, including hazard assessments with low misprediction-costs and those involving many sites. Alternative to sampling and laboratory testing,  $I_c$  is often used to assess liquefaction-susceptibility and *FC* (e.g., Robertson and Wride [53]). For susceptibility, an  $I_c = 2.6$  threshold is often used, such that soils with  $I_c < 2.6$  are inferred to be liquefiable (Robertson and Wride [53]). However, because  $I_c$  boundaries between soil types are approximate and may need regional refinement (e.g., Yi [61]), the  $I_c < 2.6$  criterion may in some cases be inappropriate (e.g., Zhang et al. [64]; Li et al. [31]; Pease [46]). For this reason, Youd et al. [63] recommended that soils with  $I_c > 2.4$  be sampled and tested to evaluate their susceptibility. Similarly, *FC* is commonly estimated using an  $I_c - FC$  correlation. However, owing to the uncertainties discussed above, these correlations often require regional calibration for optimal accuracy (e.g., Boulanger and Idriss [8]).

### 3. Christchurch-specific correlations for liquefaction susceptibility and *FC*

As mentioned above, following the damaging liquefaction in Christchurch, New Zealand, resulting from the 2010–2011 CES, an extensive geotechnical characterization program was initiated to assess regional liquefaction hazards and inform land-use planning. In this regard, acknowledgement is owed to the many constituents providing data to the Canterbury Geotechnical Database, and to Robinson et al. [54], Lees et al. [32], and Leeves et al. [33] for their seminal research on  $I_c$  correlations in Christchurch. A subset of the subsurface characterization data from the Canterbury Geotechnical Database is utilized herein to develop Christchurch-specific (ChCh)  $I_c$  correlations for liquefaction susceptibility and *FC*. The data used include: (a) SPT borings performed at 825 sites throughout Christchurch, from which 2620 samples were obtained in the split-spoon sampler; (b) laboratory measurements performed on each sample, to include *FC*, Liquid Limit (*LL*), *PI*, and natural moisture content ( $w_n$ ); and (c) CPT soundings performed within 5 m of each of the 825 SPTs. With respect to (b), Atterberg Limits were determined for 574 of the 2620 samples collected.

Several maximum offset distances between the borings and CPT locations were considered in assembling the above database. The advantage using a larger offset distance of 10 m is that 185 more borehole-CPT pairs could be included in the database, but the disadvantage is the increase in uncertainty due to the lateral spatial variability of the deposits that the  $I_c$  value corresponded to the soil sampled and tested in the laboratory. Conversely, the disadvantage of using a smaller maximum offset distance of 2.5 m is that 260 less borehole-CPT pairs could be included in the database, but the advantage is a decrease potential for the influence of heterogeneity of soil conditions (spatial variability) on the uncertainty that the  $I_c$  value corresponded to the soil sampled and tested in the laboratory. Ultimately, a maximum horizontal offset of 5 m was used to assemble the database used herein as a compromise of these advantages and disadvantages. Additionally, while this offset distance undoubtedly resulted in some incorrect pairings of  $I_c$  and soil properties, they are not thought to be pervasive and should have only increased the uncertainty in the developed  $I_c$  correlations for liquefaction susceptibility and *FC*, not influence the mean values.

The near-continuous nature of CPT sounding data (i.e., 1 or 2 cm measurement intervals) can result in significant changes in the measured  $I_c$  over short depth intervals. Accordingly, to study relationships between  $I_c$  and laboratory test indices, the measured  $I_c$  was averaged over a 300 mm depth interval (e.g., Boulanger et al. [5]), centered on the midpoint of the split-spoon sample. In addition, samples with large variation in  $I_c$  were identified by computing the standard deviation ( $\sigma$ ) of  $I_c$  within the sample interval. The locations of all test sites are shown

in Fig. 2, along with an example profile from a site in Christchurch showing CPT data and *FC* measurements.

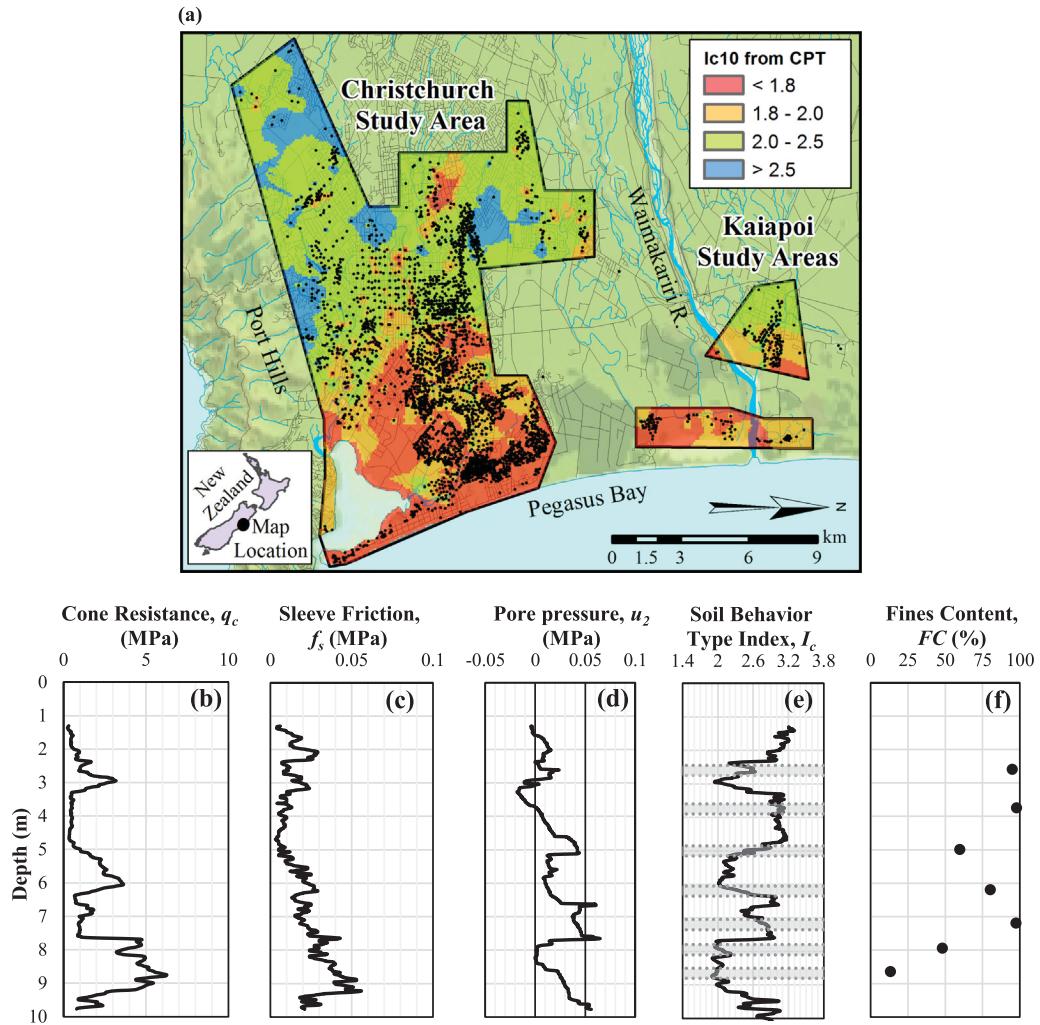
#### 3.1. Liquefaction susceptibility correlation

##### 3.1.1. Deterministic correlation

This study seeks to determine (a) how well the  $I_c$  index segregates soils susceptible to liquefaction from soils not susceptible to liquefaction, as defined by the BI06 criterion and (b) the  $I_c$  thresholds, or cut-offs, at which the efficiencies of these segregations are optimized. To make these determinations, a standard analysis is needed to assess the performance of diagnostic tests. Receiver operating characteristic (ROC) analyses are herein adopted for this purpose. ROC analyses have been widely used to study the performance of classifier systems, including extensive use in medical diagnostics (e.g., Zou [66]), but by comparison, their use in geotechnical engineering is limited (Chen et al. [16]; Oommen et al. [45]; Mens et al. [39]; Maurer et al. [35]; Zhu et al. [65]). In any ROC application, the distributions of “positives” (e.g., soil is susceptible to liquefaction per BI06) and “negatives” (e.g., soil is not susceptible to liquefaction per BI06) overlap when the frequency of the distributions are expressed as a function of index test results (e.g.,  $I_c$  values). In such cases, optimal decision thresholds for the index test are selected considering the rates of true positives ( $R_{TP}$ ) (e.g., soil is predicted to be susceptible to liquefaction based on  $I_c$  and is predicted to exhibit sand-like behavior per BI06) and false positives ( $R_{FP}$ ) (e.g., soil is predicted to be susceptible to liquefaction based on  $I_c$  but is predicted to exhibit clay-like behavior per BI06). Setting the  $I_c$  threshold too high will result in a higher  $R_{FP}$ , the cost of which could be excessive spending on site remediation. Conversely, setting the threshold too low results in a higher rate of false negatives (e.g., soil is predicted to be not susceptible to liquefaction based on  $I_c$  but is predicted to exhibit sand-like behavior per BI06), the cost of which is liquefaction-induced damage during a future event. Thresholds should thus be selected to minimize the rates of false-positive and false-negative predictions.

Receiver operating characteristic curves plot  $R_{TP}$  versus  $R_{FP}$  for varying threshold values. Fig. 3 illustrates the relationship among the positive and negative distributions, the threshold value, and the ROC curve. Fig. 3b also illustrates how a ROC curve is used to assess the efficiency of a diagnostic test and to select an optimum threshold. In ROC space, random guessing is indicated by a 1:1 line through the origin (i.e., equivalent number of correct and incorrect predictions), while a perfect model plots as a bi-linear function that passes through the points (0,0; 0,1; 1,1), indicating the existence of a threshold value which perfectly segregates the dataset (e.g., all soils susceptible to liquefaction have  $I_c$  below the threshold; all soils not susceptible to liquefaction have  $I_c$  above the threshold). While no single parameter can fully characterize model performance, the area under a ROC curve (AUC) is commonly used for this purpose, where AUC is statistically equivalent to the probability that “positives” have lower index test values than “negatives” (e.g., Fawcett [22]). As such, a higher AUC indicates better model performance. The optimum decision threshold is defined herein as the threshold which minimizes the rate of misprediction [i.e.,  $R_{FP} + (1 - R_{TP})$ ]. As such, contours of the quantity [ $R_{FP} + (1 - R_{TP})$ ] map points of equivalent performance in ROC space, as shown in Fig. 3b. Notably, this definition implicitly treats the costs of false positives and false negatives to be equal. For further overview of ROC analyses, and for demonstration of how project-specific misprediction consequences can be incorporated into ROC analyses, the reader is referred to Fawcett [22] and Maurer et al. [37].

Receiver operating characteristic analyses were performed to determine how well the  $I_c$  index correlates to liquefaction susceptibility per the BI06 criterion. Shown in Fig. 4 are frequency distributions of the 574 classified samples, plotted as a function of the corresponding  $I_c$ . As may be observed in this figure, there is almost a complete overlap of the distributions for the soils classified as “not susceptible” and “intermediate.” Accordingly, the ROC analysis is performed using the

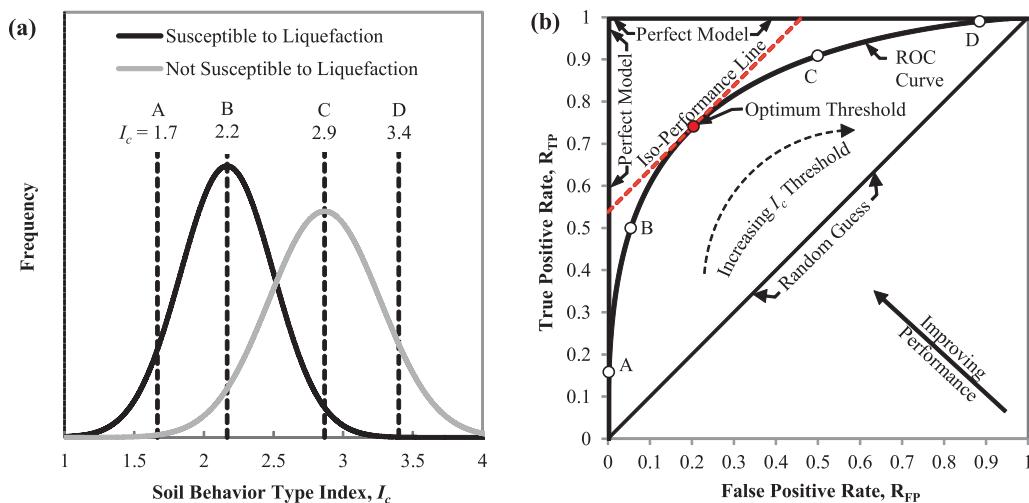


**Fig. 2.** (a) Map of the soil behavior type index computed from CPT data and averaged over the uppermost 10 m,  $I_{c10}$ , using inverse-distance-weighting interpolation between CPT locations. Black dots represent all soil sampling locations and all CPT sounding locations. (b) Example profiles of cone resistance,  $q_c$ , (c) sleeve friction,  $f_s$ , (d) pore pressure,  $u_2$ , (e) soil behavior type index,  $I_c$ , and (f) fines content,  $FC$ , from a site in Christchurch; 300 mm intervals over which  $I_c$  statistics were recorded, each corresponding to a soil sampling location, are shown in (e).

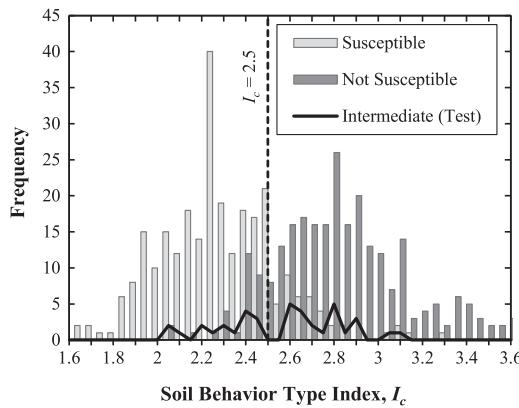
“susceptible” and “not susceptible” distributions shown in Fig. 4, with the results shown in Fig. 5. The performance of  $I_c$  in assessing liquefaction susceptibility is indicated by  $AUC = 0.92$  (i.e., there is ~92% probability that the measured  $I_c$  of a soil “not susceptible” is greater

than that of a “susceptible” soil). In other words,  $I_c$  is well-correlated to the BI06 PI criterion.

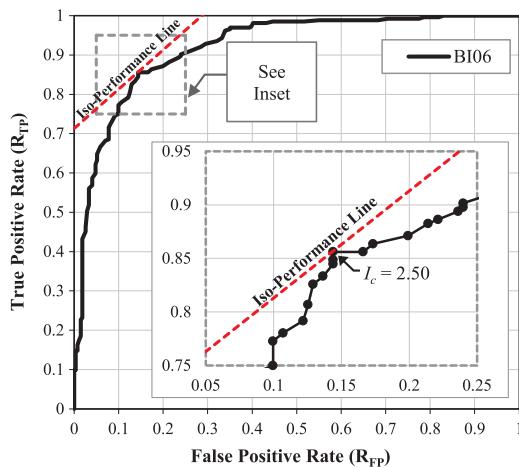
As highlighted in Figs. 4 and 5, the optimal  $I_c$  threshold for the BI06 criterion is 2.50 (Table 1), which is in general accord with the  $I_c$  criteria



**Fig. 3.** ROC analyses: (a) frequency distributions of soils susceptible and not susceptible to liquefaction per BI06 as a function of measured  $I_c$ , with four different threshold  $I_c$  values shown; (b) corresponding ROC curve, and illustration of how a ROC curve is used to assess the efficiency of a diagnostic test. The optimum decision threshold is that for which the rate of misprediction is minimized.



**Fig. 4.** Frequency distributions of samples classified by the BI06 liquefaction susceptibility criterion plotted as a function of measured  $I_c$ . The optimal  $I_c$  threshold for identifying liquefaction-susceptible soils is identified by a vertical dotted line.



**Fig. 5.** ROC analysis of  $I_c$  index performance in segregating soils susceptible to liquefaction from soils not susceptible to liquefaction per BI06. The optimum threshold is identified as  $I_c = 2.50$ .

**Table 1**

Optimal deterministic  $I_c$  thresholds and coefficients for probabilistic  $I_c$  relationship (Eq. (2)) for different liquefaction susceptibility criteria.

| Criterion  | Optimal Deterministic $I_c$ Threshold | Coef. for Probabilistic $I_c$ Relationship (Eq. 2) |        |
|------------|---------------------------------------|--|--------|
|            |                                       | $\beta$  | $x_m$  |
| BI06 [6]   | 2.50                                  | 0.0851   | 2.5031 |
| P01 [47]   | 2.55                                  | 0.0988   | 2.5474 |
| Sea03 [57] | 2.60                                  | 0.1348   | 2.6214 |
| BS06 [13]  | 2.75                                  | 0.1275   | 2.7315 |

used in practice (e.g., Robertson and Wride [53]; Youd et al. [63]). However, it should be recognized that this threshold is that which minimized the misprediction rate and is only optimal if false positives and false negatives have equal cost. For example, if the cost of false negatives using the BI06 criterion was instead greater than the cost of false positives, the optimum  $I_c$  threshold would increase in response. Operating at the optimal threshold of  $I_c = 2.50$ , the overall accuracy of using  $I_c$  as a proxy for the BI06 criterion for liquefaction susceptibility is 86%. This should not be interpreted to mean that using a threshold of  $I_c = 2.50$  will correctly determine the liquefaction susceptibility of a soil 86% of the time, but rather, that using  $I_c = 2.50$  will correctly determine whether the soil is susceptible, as defined by BI06, 86% of the

time. It should be emphasized that while the BI06 criterion offers a pragmatic approach to evaluating liquefaction susceptibility, it is not the definitive test of susceptibility. Ideally, cyclic laboratory tests on undisturbed samples would also be performed to corroborate or re-calibrate susceptibility thresholds; this will be done in Christchurch as results become available. Nonetheless, these analyses suggest that  $I_c$  can be an efficient and cost-effective index of liquefaction susceptibility.

### 3.1.2. Probabilistic correlation

To assess liquefaction hazards in a fully probabilistic manner, the uncertainty of liquefaction susceptibility should be adequately accounted for. A probabilistic correlation is developed herein using an approach similar to that described by Porter et al. [48] and Porter [49] to create fragility functions compatible with performance-based methods in earthquake engineering. While the adopted approach is outlined below, the reader is referred to the aforementioned publications for complete details.

The probability that a soil is “not susceptible” to liquefaction, given a measured  $I_c$  value, is herein denoted  $F_{\text{not susceptible}}(I_c)$  and idealized by a log-normal distribution:

$$F_{\text{not susceptible}}(I_c) = \Phi \left[ \frac{\ln \left( \frac{I_c}{x_m} \right)}{\beta} \right] \quad (2)$$

where  $\Phi$  denotes the Gaussian cumulative distribution function;  $x_m$  is the median value of the distribution (equal to the mean of the natural logarithm of  $x$ ); and  $\beta$  is the logarithmic standard deviation. The probability that a soil is “susceptible” to liquefaction, denoted as  $F_{\text{susceptible}}(I_c)$ , is then given by:

$$F_{\text{susceptible}}(I_c) = 1 - F_{\text{not susceptible}}(I_c) \quad (3)$$

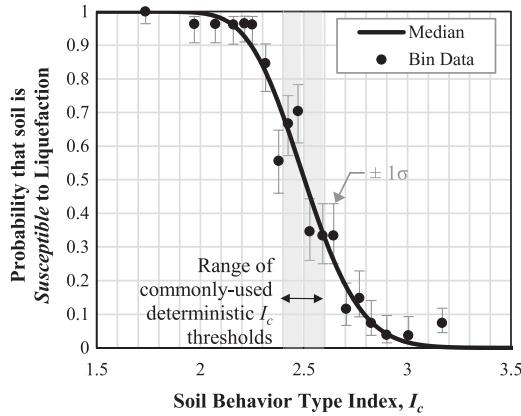
While many approaches exist for fitting functions to data, this study utilizes the maximum likelihood method described in Porter [49], which identifies the model parameters with the highest likelihood of producing the observed data. Specifically, the data are first grouped into  $j$  bins of similar  $I_c$ , where bins have index  $i$ , average value  $I_{ci}$ , and contain  $n_i$  samples, of which  $f_i$  are samples not susceptible to liquefaction. Assuming quantity  $f_i$  can be estimated from a binomially-distributed random variable,  $F_i$ , Eq. (4) gives the probability of observing quantity  $f_i$  among  $n_i$  samples, if the probability of an individual sample being not susceptible is given by Eq. (2).

$$P[F_i = f_i] = \frac{n_i!}{f_i!(n_i - f_i)!} \cdot p_i^{f_i} \cdot (1 - p_i)^{n_i - f_i} \quad (4)$$

In Eq. (4),  $p_i$  is defined by Eq. (2), evaluated at  $I_{ci}$ . Lastly, the values of parameters  $x_m$  and  $\beta$  that maximize the likelihood of producing the observed data are determined. This likelihood is given by the product of the probabilities in Eq. (4), multiplied over all bins:

$$L(X_m, \beta) = \prod_{i=1}^j P[F_i = f_i] \quad (5)$$

Using the approach outlined above, a probabilistic correlation was developed from the “susceptible” and “not susceptible” distributions shown in Fig. 4. The resulting function is plotted in Fig. 6 and defined by Eqs. (2) and (3) in conjunction with the coefficients in Table 1. The proposed function allows for the uncertainty of whether a soil is susceptible to liquefaction to be incorporated into fully probabilistic hazard assessments. It can be seen from Fig. 6 that the  $I_c$  value corresponding to a 50% probability of susceptibility is equal to the deterministic threshold developed from ROC analysis. For  $I_c$  thresholds traditionally used in practice (i.e.,  $2.4 \leq I_c \leq 2.6$ ), the probability of susceptibility ranges from 0.69 to 0.33 for  $I_c = 2.4$  and 2.6, respectively. As with the deterministic approach, it should be emphasized that the BI06 criterion is not a definitive test of susceptibility. Ideally, cyclic laboratory tests would also be used to classify susceptibility and



**Fig. 6.** The probability of liquefaction susceptibility per the BI06 criterion as a function of measured  $I_c$ . The range of deterministic  $I_c$  thresholds commonly used in practice is also highlighted.

develop  $I_c$  functions.

### 3.1.3. Deterministic and probabilistic correlations for other liquefaction susceptibility criteria

Maurer et al. [38] used the same approaches as those outlined above to develop deterministic and probabilistic correlations between  $I_c$  and the laboratory test-based liquefaction susceptibility criteria proposed by Polito (P01) [47], Seed et al. (Sea03) [57], and Bray and Sancio (BS06) [13], as well as by Boulanger and Idriss (BI06) [6]. For the deterministic correlations, the AUC for P01, Sea03, and BS06 are very similar to that for BI06. The optimal  $I_c$  thresholds for P01, Sea03, and BS06 were determined to be 2.55, 2.60, and 2.75, respectively (Table 1), which except for BS06 are in accord with the range generally used in practice (i.e.,  $2.4 \leq I_c \leq 2.6$ ) (e.g., Robertson and Wride [53]; Youd et al. [63]). Table 1 also lists the coefficients for Eq. (2) for the probabilistic  $I_c$ -liquefaction susceptibility correlations for P01, Sea03, and BS06.

### 3.2. Deterministic and probabilistic fines content correlations

Plotted in Fig. 7 are  $I_c$  vs. FC data from 2620 soil samples, from which a Christchurch-specific correlation is developed using the classical (Eisenhart [21]) approach to calibration and the least squares estimator of a linear model having the form:

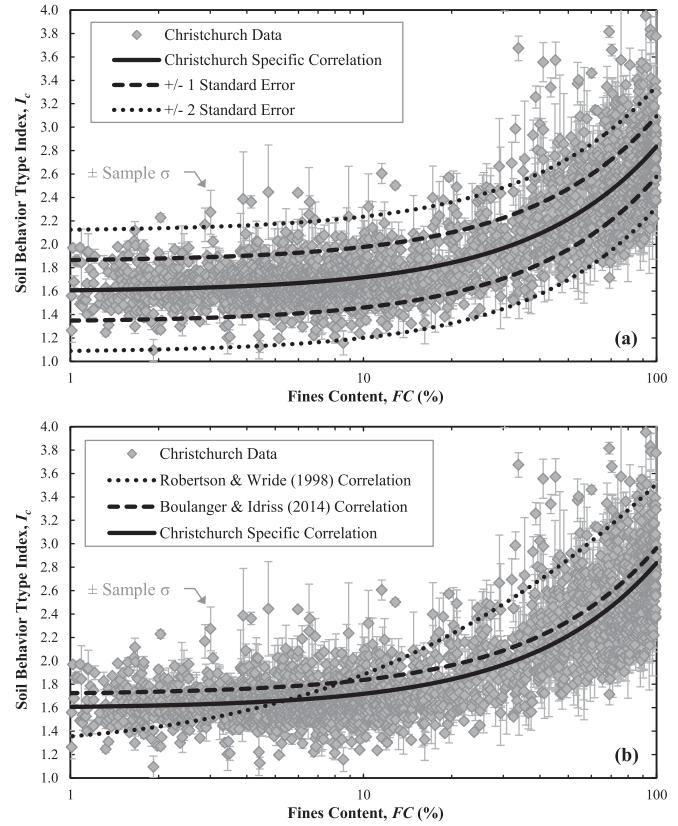
$$FC = \beta_0 + \beta_1 I_c + \varepsilon \quad (6)$$

In Eq. (6),  $FC$  and  $I_c$  are as previously defined;  $\beta_0$  and  $\beta_1$  are regression coefficients; and  $\varepsilon$  is a random error term. With the goal of developing an unbiased model (i.e., one for which prediction errors sum to zero),  $\varepsilon$  is assumed to have an approximately normal distribution about the mean prediction. The standard deviation of  $\varepsilon$  ( $\sigma_\varepsilon$ ) is unknown, however, and is thus estimated by the residual standard error. The resulting Christchurch-specific correlation is given by Eq. (7):

$$\mu_{FC} = 80.645I_c - 128.5967 \quad (7)$$

where  $\mu_{FC}$  is the mean estimate of  $FC$  (%), limited to  $0\% \leq FC \leq 100\%$ , and  $I_c$  is as previously defined. The uncertainty in this prediction, represented by the standard deviation of  $\varepsilon$ , is  $\sigma_\varepsilon = 16.56\%$ . This indicates that approximately 68% of the samples have  $FC$  within  $\pm 16.56\%$  of the mean prediction (Eq. (7)); approximately 95% have  $FC$  within  $\pm 33.1\%$  of this prediction. The proposed mean,  $\pm 1\sigma$ , and  $\pm 2\sigma$  correlations are plotted in Fig. 7a.

To demonstrate potential uses of the proposed correlation, two example cases are given below. For each case, a hypothetical  $I_c$  measurement of 2.05 has been obtained from CPT sounding data in a soil stratum. From Eq. (7), the mean (i.e., deterministic) estimate of  $FC$  is 36.73%.



**Fig. 7.** (a) Christchurch-specific  $I_c$  – FC data and proposed correlation (Eq. (7)); (b) Comparison with the Robertson and Wride [53] and Boulanger and Idriss [8] generic  $I_c$  – FC correlations.

- Case I: When performing fully probabilistic liquefaction hazard assessments, it may be desirable to account for the uncertainty in the estimated  $FC$ . Using such an approach in conjunction with the total probability theorem, the probabilities of various  $FC$  values can be computed from a Gaussian probability density function. As an example, the probability of the soil stratum having  $49.5\% \leq FC \leq 50.5\%$  is estimated to be 0.0175, as illustrated by Eq. (8), for which the corresponding Microsoft Excel command is also given.

$$\begin{aligned} P(49.5\% \leq FC \leq 50.5\%) &= \Phi_{FC} \left[ \frac{50.5\% - \mu_{FC}}{\sigma_\varepsilon} \right] - \Phi_{FC} \left[ \frac{49.5\% - \mu_{FC}}{\sigma_\varepsilon} \right] \\ &\approx \Phi_{FC} \left[ \frac{50\% - \mu_{FC}}{\sigma_\varepsilon} \right] \cdot \Delta FC = \Phi_{FC} \left[ \frac{50\% - 36.73\%}{16.56\%} \right] \cdot (50.5\% - 49.5\%) \\ &= \text{NORM. DIST}(50, 36.73, 16.56, \text{FALSE}) \cdot (50.5\% - 49.5\%) = 0.0175 \end{aligned} \quad (8)$$

- Case II: For soil classification or other probabilistic applications, it may be desirable to compute the probability of the  $FC$  being less than or greater than a particular value. In this case, the desired probabilities are computed from a Gaussian cumulative distribution function. As an example, the probability of the soil stratum having  $FC \leq 50\%$  (i.e., a coarse-grained soil) is estimated to be 0.7886, as illustrated by Eq. (9), for which the corresponding Microsoft Excel command is also given.

$$\begin{aligned} P(FC \leq 50\%) &= \Phi_{FC} \left[ \frac{50\% - \mu_{FC}}{\sigma_\varepsilon} \right] = \Phi_{FC} \left[ \frac{50\% - 36.73\%}{16.56\%} \right] \\ &= \text{NORM. DIST}(50, 36.73, 16.56, \text{TRUE}) = 0.7886 \end{aligned} \quad (9)$$

In Fig. 7b, the mean Christchurch-specific (ChCh)  $I_c$  – FC correlation is compared with the Robertson and Wride [53] and Boulanger and Idriss [8] correlations developed from global data. The Robertson and Wride (RW98) [53] correlation is defined by Eq. (10) and has been

widely used since its inception:

$$FC = 1.75I_c^{3.25} - 3.7 \quad (10)$$

The Boulanger and Idriss (BI14) [8]  $I_c$  –  $FC$  correlation is defined by Eq. (11) and was developed largely from liquefaction case history data:

$$FC = 80 (I_c + C_{FC}) - 137 \quad (11)$$

In Eq. (11),  $C_{FC}$  is a calibration parameter, such that the general correlation (i.e.,  $C_{FC} = 0$ ) may be calibrated to site- or region-specific conditions. The BI14 correlation thus recognizes the uncertainty of global correlations and the need for site-specific  $I_c$  –  $FC$  investigations. On this note, the ChCh  $I_c$  –  $FC$  correlation (i.e., Fig. 7a and Eq. (7)) may be closely approximated using  $C_{FC} = 0.13$  in Eq. (11). To compare the performance of the three correlations shown in Fig. 7b for the Christchurch data, the Nash-Sutcliffe model efficiency coefficient,  $E$ , is adopted. Commonly used in hydrology, the Nash-Sutcliffe efficiency is defined as (Nash and Sutcliffe [42]):

$$E = 1 - \frac{\sum_{i=1}^n (O_i - P_i)^2}{\sum_{i=1}^n (O_i - \bar{O})^2} \quad (12)$$

where  $O_i$  and  $P_i$  are the measured and predicted values of  $FC$ , respectively;  $\bar{O}$  is the mean of the measured values of  $FC$ ; and  $n$  is the number of data points (2620). An efficiency coefficient of 1.0 indicates a perfect model, while an efficiency coefficient less than zero indicates that the mean  $FC$  of the dataset (i.e., 30.2%) would have been a better predictor than the correlation. For the compiled dataset, the Christchurch-specific, BI14, and RW98 correlations have efficiencies of 0.69, 0.66, and 0.37, respectively. Thus, the estimate of  $FC$  for the Christchurch soils using the Christchurch-specific correlation is slightly more accurate than if the BI14 correlation is used and significantly more accurate than if the RW98 correlation is used.

#### 4. Liquefaction hazard assessment

The severity of surficial liquefaction manifestations was predicted for 9623 case studies from the Canterbury earthquakes. In these predictions, the BI14 simplified liquefaction evaluation procedure was used in conjunction with the LPI framework. The predictions were repeated multiple times assuming  $I_c$  susceptibility thresholds ranging from 1.8 to 2.8 and using the RW98, BI14, and ChCh  $I_c$ - $FC$  correlations. The predicted surface-manifestations were compared with post-earthquake field observations to determine the accuracy of the predictions using each of the  $I_c$  correlations. Specifics of the liquefaction hazard assessment follow.

##### 4.1. CPT Soundings

Drawing from a subset of data available in the New Zealand Geotechnical Database (NZGD [44]), this study utilizes 3834 CPT soundings performed at sites where the severity of liquefaction manifestation was well-documented following the 4 Sept 2010 Darfield, 22 Feb 2011 Christchurch, and 14 Feb 2016 Valentine's Day earthquakes, resulting in a combined 9623 case studies. Extended coverage of CPT data is provided in Maurer et al. [34,35].

##### 4.2. Liquefaction Severity

Observations of liquefaction and the severity of surface manifestation were made by the authors for each CPT location following the three aforementioned earthquakes. The severity of manifestation was classified using the criteria of Green et al. [25] (Table 2), such that manifestations are ranked as “none,” “marginal,” “moderate,” or “severe.” These rankings characterize the quantity of vented sand and do not apply to cases of ground oscillation or lateral spreading, which are not considered in this study. Of the 9623 cases compiled from the

Canterbury earthquakes, 58% are cases of “none” and 42% are cases where manifestations were observed and classified.

#### 4.3. Estimation of peak ground acceleration (PGA)

To evaluate the factor of safety against liquefaction triggering ( $FS_{liq}$ ) for use in computing  $LPI$  values, the Peak Ground Accelerations (PGAs) at the ground surface were computed with the Bradley [10] procedure, which has been used by many Canterbury earthquake studies (e.g., Green et al. [23,24]; Green et al. [25]; Maurer et al. [36]; van Ballegooij et al. [60]; among others). The Bradley [10] procedure combines unconditional PGA distributions estimated by the Bradley [11] ground motion prediction equation, recorded PGAs from strong motion stations, and the spatial correlation of intra-event residuals to compute PGAs at sites of interest.

#### 4.4. Estimation of ground-water table (GWT) depth

Given the sensitivity of computed liquefaction hazard to ground-water table (GWT) depth (e.g., Chung and Rogers [17]; Maurer et al. [34]), accurate measurement of the GWT is critical. For this study, GWT depths were sourced from the robust, event-specific regional ground water models by van Ballegooij et al. [59]. These models, which reflect seasonal and localized fluctuations across the region, were derived in part using monitoring data from a network of ~1000 piezometers and provide a best-estimate of GWT depths immediately prior to each earthquake.

#### 4.5. Liquefaction evaluation

##### 4.5.1. Liquefaction potential index, LPI

While the simplified liquefaction evaluation procedure (e.g., BI14) is central to most liquefaction hazard assessments, its output is not a direct quantification of liquefaction damage potential, but rather is the  $FS_{liq}$  in a soil stratum at depth. To link  $FS_{liq}$  at depth to damage potential at the ground surface, Iwasaki et al. [28] proposed the  $LPI$ , which is widely used in current practice.  $LPI$  is computed as:

$$LPI = \int_0^{z_{max}} F(FS_{liq}) \cdot w(z) dz \quad (2)$$

where  $F(FS_{liq})$  and  $w(z)$  are functions that weight the respective influences of  $FS_{liq}$  and depth on surface manifestation; and  $z$  is depth below the ground surface. Specifically,  $F(FS_{liq}) = 1 - FS_{liq}$  for  $FS_{liq} \leq 1$  and  $F(FS_{liq}) = 0$  otherwise;  $w(z) = 10 - 0.5z$ ; and  $z_{max}$  = is the maximum depth of liquefiable soil in the soil profile, down to a maximum depth of 20 m. Thus,  $LPI$  assumes the severity of liquefaction manifestation depends on the cumulative thickness of liquefied strata in the upper 20 m of the profile, the proximity of those strata to the ground surface, and the amount by which  $FS_{liq}$  in each stratum is less than 1.0. Given this definition,  $LPI$  can range from zero to a maximum of 100. Using SPT data from 45 liquefaction sites in Japan, Iwasaki et al. [28] found that 80% of the sites had  $LPI > 5$ , while 50% had  $LPI > 15$ . Based on this data, Iwasaki et al. [28] proposed that severe liquefaction should be expected for sites where  $LPI > 15$  but should not be expected for sites where  $LPI < 5$ .

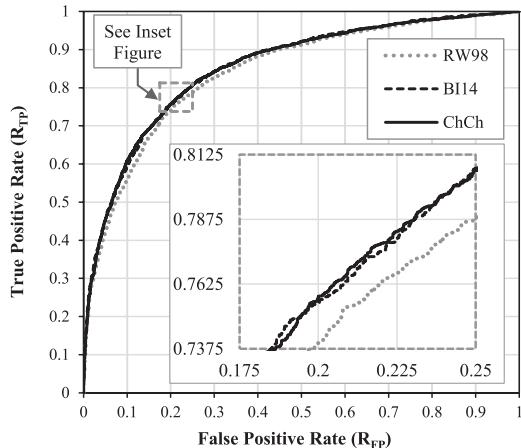
##### 4.5.2. Results

As stated previously, the severity of surficial liquefaction manifestation was predicted for 9623 case studies from the Canterbury earthquakes using the BI14 CPT-based simplified liquefaction evaluation procedure in conjunction with the  $LPI$  framework. The assessment was repeated multiple times assuming  $I_c$  susceptibility thresholds ranging from 1.8 to 2.8 and using the RW98, BI14, and ChCh  $I_c$ - $FC$  correlations. The predictions were compared with post-earthquake field observations to determine the accuracy of the predictions using each of the  $I_c$  correlations.

**Table 2**

Liquefaction severity classification criteria (Green et al. [25]).

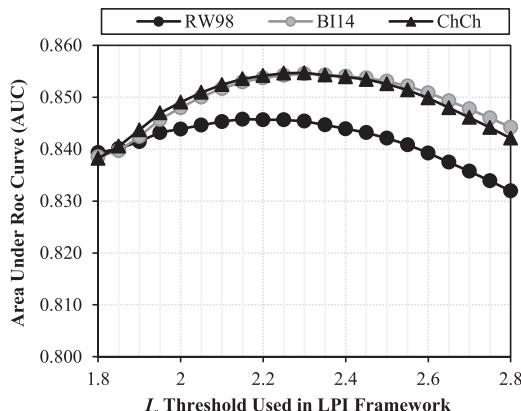
| Classification        | Criteria  |
|-----------------------|---|
| No Liquefaction       | No surficial liquefaction manifestation or lateral spread cracking  |
| Minor Liquefaction    | Small, isolated liquefaction features; streets had traces of ejecta or wet patches less than a vehicle width; < 5% of ground surface covered by ejecta      |
| Moderate Liquefaction | Groups of liquefaction features; streets had ejecta patches greater than a vehicle width but were still passable; 5–40% of ground surface covered by ejecta |
| Severe Liquefaction   | Large masses of adjoining liquefaction features, streets impassable due to liquefaction, > 40% of ground surface covered by ejecta                          |



**Fig. 8.** ROC analyses of LPI performance, utilizing the BI14 liquefaction evaluation procedure and an  $I_c$  threshold value of 2.50, in segregating sites with and without manifestations of liquefaction. The performance of three  $I_c$ -FC correlations (i.e., global correlations proposed by RW98 and BI14 and the Christchurch-specific correlation developed herein) are compared.

Similar to the ROC analysis used to evaluate the efficacy of using  $I_c$  as a proxy for the BI06 liquefaction susceptibility criterion, ROC analyses are used to assess the accuracy of the manifestation predictions using various combinations of  $I_c$  threshold and  $I_c$ -FC correlation. As an illustration, the ROC curves for an assumed deterministic  $I_c$  susceptibility threshold of 2.50 (i.e., soils with  $I_c \geq 2.50$  are assumed to be not susceptible to liquefaction) in combination with the RW98, BI14, and ChCh  $I_c$ -FC correlations are shown in Fig. 8. As may be observed from this figure, the AUCs for the ROC curves using the BI14 and ChCh  $I_c$ -FC correlations are approximately equal ( $AUC \approx 0.853$ ) and slightly higher than that of the curve using the RW98  $I_c$ -FC correlation ( $AUC \approx 0.842$ ).

Fig. 9 shows a plot of AUC as function of the assumed deterministic  $I_c$  susceptibility threshold, ranging from 1.8 to 2.8, in combination with the RW98, BI14, and ChCh  $I_c$ -FC correlations. Similar to the scenario



**Fig. 9.** Area under ROC curve (AUC) vs. soil-behavior-type index ( $I_c$ ) threshold used in the LPI framework, wherein the BI14 liquefaction evaluation procedure was used.

presented in Fig. 8, the AUCs for the ROC curves using the BI14 and ChCh  $I_c$ -FC correlations are approximately equal for the full range of deterministic  $I_c$  susceptibility thresholds and slightly higher than those for the curves using the RW98  $I_c$ -FC correlation. However, as may be observed in Fig. 9, the AUCs for the BI14 curves are slightly higher than the curves for the ChCh  $I_c$ -FC correlation, despite the ChCh  $I_c$ -FC correlation yielding more accurate estimations of FC for Christchurch soils. Possible reasons for this are: (1) less than optimal accounting for the influence of FC on liquefaction triggering in the BI14 simplified liquefaction evaluation procedure; and/or (2) shortcomings of the LPI framework in accounting for FC on severity of surficial liquefaction manifestations, either of which could fortuitously result in improved manifestation predictions despite the less accurate FC estimation. Regardless, the practical significance of the difference in the liquefaction manifestation predictions for the Canterbury earthquakes using the ChCh versus BI14  $I_c$ -FC correlation is negligible.

The shapes of the AUC curves shown in Fig. 9 are of interest and merit discussion. For low values of  $I_c$ -threshold, liquefaction-susceptible soils are treated as non-liquefiable and the computed LPI approaches zero for all sites. As a result, the computed AUC approaches 0.5, indicating that predictions are no better than random guessing, independent of the  $I_c$ -FC correlation used to estimate FC. However, as the  $I_c$ -threshold increases,  $I_c$ -FC correlations become increasing more important because the predicted FC influences the magnitude of the fines-correction factors applied to the measured cone tip resistance. This in turn influences the computed cyclic resistance of the soil and  $FS_{liq}/LPI$ . As a result, AUC curves increasingly diverge as the  $I_c$ -threshold increases. As the  $I_c$  threshold continues to increase soils that are likely too plastic to liquefy are treated as being susceptible to liquefaction and the  $FS_{liq}$  of such soils are computed by the BI14 simplified procedure. However, because the BI14 simplified procedure is not intended for use with soils that classify as either clayey silt, silty clay, or clay, the accuracy of the predictions decrease, resulting in lower AUC values.

Each of the AUC curves shown in Fig. 9 reaches a maximum in the intermediate  $I_c$ -threshold range:  $I_c \approx 2.2$  for the RW98  $I_c$ -FC correlation and  $I_c \approx 2.3$  for the BI14 and ChCh  $I_c$ -FC correlations. If we compare these  $I_c$  values to the optimal values for the laboratory test-based criteria (i.e., BI06:  $I_c = 2.50$ ; P01:  $I_c = 2.55$ ; Sea03:  $I_c = 2.60$ ; and BS06:  $I_c = 2.75$ ) it can be seen that the BI06 and BS06 are the most and least efficacious, respectively. However, the BI06 criterion is still considerably higher than the optimal threshold determined from the field data (i.e., 2.5 versus 2.2–2.3). This could be a result of shortcomings in the susceptibility criteria and/or in the procedures used to predict liquefaction triggering and liquefaction manifestation. Additionally, the  $I_c$  providing an optimal correlation to soil properties of one stratum in a profile may not be the same as that for another stratum in the same profile. Regardless, the practical significance of using deterministic  $I_c$  susceptibility thresholds ranging from 1.8 to 2.8 is relatively negligible for the assessed database.

## 5. Conclusions

Utilizing an unprecedented database of field and laboratory test data, this study investigated CPT-based soil characterization in Christchurch, New Zealand, with emphasis on obtaining accurate inputs for liquefaction hazard assessment. In particular, region-specific

deterministic and probabilistic  $I_c$  correlations were developed for predicting liquefaction susceptibility and  $FC$ . To predict liquefaction susceptibility, the BI06 susceptibility criteria was found to have an optimal deterministic  $I_c$  threshold of 2.50 (i.e., soils with an  $I_c > 2.50$  are predicted not to be susceptible to liquefaction). A probabilistic  $I_c$  - susceptibility correlation was proposed in Eqs. (2) and (3) and plotted in Fig. 6. As may be observed from this figure, the proposed deterministic threshold (i.e.,  $I_c = 2.50$ ) corresponds to a 50% probability that the soil is liquefaction-susceptible per the BI06 criterion. To predict  $FC$ , an  $I_c$  -  $FC$  correlation was developed in Eq. (7) and can be used in both deterministic and probabilistic analyses. The proposed correlation was plotted in Fig. 7 and scenarios demonstrating its use were presented.

The proposed  $I_c$  correlations were used to assess liquefaction hazard for 9623 case studies compiled from the 2010–2011 Canterbury earthquake sequence and the 2016 Valentine's Day earthquake (i.e., Canterbury earthquakes), wherein predictions of liquefaction manifestation were compared to field observations. In this assessment, the BI14 CPT-based simplified liquefaction evaluation procedure was used in conjunction with the LPI framework. The assessment was repeated multiple times assuming  $I_c$  susceptibility thresholds ranging from 1.8 to 2.8 and using the RW98, BI14, and Christchurch-specific  $I_c$ - $FC$  correlations, so as to assess the benefit of developing regional relationships. The efficacy of predictions using the Christchurch-specific  $I_c$ - $FC$  correlation was essentially equivalent to that when the BI14 generic correlation was used and slightly better than when the RW98 generic correlation was used. Also, although the BI06 ( $I_c = 2.50$ ) and BS06 ( $I_c = 2.75$ ) susceptibility thresholds were the most and least efficacious, respectively, of the criteria evaluated herein, none of the susceptibility criteria evaluated correspond to the deterministic  $I_c$  threshold value that yields the optimal manifestation predictions (i.e.,  $I_c = 2.2$ –2.3).

The above findings could be a result of shortcomings in the BI14 CPT-based simplified procedure and/or the LPI framework. However, in either case the choice of the region-specific versus generic  $I_c$  correlations for liquefaction susceptibility and  $FC$  had somewhat of a negligible impact on the accuracy of the predicted liquefaction severity for the assessed database. This gives credence to the use of generic correlations for typical projects involving artificial fills or fluvial deposits. However, the development and use of region-specific correlations may be more important when less common types of liquefiable soils are present. Finally, it is noted that while this study focused on BI14 and the Canterbury earthquakes, the overall framework of the study could be applied to other CPT-based procedures and to worldwide locals.

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