



# The Bayesian Asteroseismology Data Modeling Pipeline and Its Application to K2 Data

Joel C. Zinn<sup>1</sup> , Dennis Stello<sup>1,2,3,4</sup> , Daniel Huber<sup>5,2,3,6</sup> , and Sanjib Sharma<sup>2,4</sup> <sup>1</sup> School of Physics, University of New South Wales, Barker Street, Sydney, NSW 2052, Australia; [j.zinn@unsw.edu.au](mailto:j.zinn@unsw.edu.au)<sup>2</sup> Sydney Institute for Astronomy (SfA), School of Physics, University of Sydney, NSW 2006, Australia<sup>3</sup> Stellar Astrophysics Centre, Department of Physics and Astronomy, Aarhus University, Ny Munkegade 120, DK-8000 Aarhus C, Denmark<sup>4</sup> ARC Centre of Excellence for All Sky Astrophysics in 3 Dimensions (ASTRO 3D), Australia<sup>5</sup> Institute for Astronomy, University of Hawai'i, 2680 Woodlawn Drive, Honolulu, HI 96822, USA<sup>6</sup> SETI Institute, 189 Bernardo Avenue, Mountain View, CA 94043, USA

Received 2018 June 11; revised 2019 September 2; accepted 2019 September 9; published 2019 October 16

## Abstract

We present the Bayesian Asteroseismology data Modeling (BAM) pipeline, an automated asteroseismology pipeline that returns global oscillation parameters and granulation parameters from the analysis of photometric time series. BAM also determines whether a star is likely to be a solar-like oscillator. We have designed BAM to specially process K2 light curves, which suffer from unique noise signatures that can confuse asteroseismic analysis, though it may be used on any photometric time series—including those from *Kepler* and *TESS*. We demonstrate that the BAM oscillation parameters are consistent within  $\sim 1.53\%$  (random)  $\pm 0.2\%$  (systematic) and  $1.51\%$  (random)  $\pm 0.6\%$  (systematic) for  $\nu_{\max}$  and  $\Delta\nu$  with benchmark results for typical K2 red giant stars in the K2 Galactic Archaeology Program's (GAP) Campaign 1 sample. Application of BAM to 13,016 K2 Campaign 1 targets not in the GAP sample yields 104 red giant solar-like oscillators. Based on the number of serendipitous giants we find, we estimate an upper limit on the average purity in dwarf selection among C1 proposals of  $\approx 99\%$ , which could be lower when considering incompleteness in BAM detection efficiency and proper-motion cuts specific to C1 Guest Observer proposals.

*Unified Astronomy Thesaurus concepts:* [Astronomy software \(1855\)](#); [Asteroseismology \(73\)](#); [Astronomy data analysis \(1858\)](#); [Giant stars \(655\)](#)

## 1. Introduction

Solar-like oscillators are stars that support standing acoustic waves excited by surface convection and whose global frequency characteristics are determined by the stellar density and surface gravity (e.g., Ulrich 1986; Brown et al. 1991; Kjeldsen & Bedding 1995). The frequencies may be measured in radial velocity variations or in photometric variability. Detecting mode frequencies in solar-like oscillators yields precise determinations of fundamental stellar parameters like mass and radius. However, only about a dozen stars had been observed to exhibit solar-like oscillations prior to the results from the space-based *CoRoT* (Baglin et al. 2006) and *Kepler* (Borucki et al. 2008) missions. With improved photometric precision compared to ground-based observations and continuous monitoring of many stars simultaneously for up to 4 yr with *Kepler*, solar-like oscillations have been photometrically detected in thousands of stars—mostly red giants (e.g., De Ridder et al. 2009; Hekker et al. 2009; Bedding et al. 2010; Mosser et al. 2010; Stello et al. 2013; Yu et al. 2018). In light of these large asteroseismic data sets, several pipelines have been developed in order to automatically extract asteroseismic parameters (e.g., OCT, Hekker et al. 2010; CAN, Kallinger et al. 2010, 2014, 2016; COR, Mosser & Appourchaux 2009; A2Z, Mathur et al. 2010).

Among these pipelines is SYD (Huber et al. 2009), much of whose success can be attributed to taking advantage of known scaling relations among stellar granulation, the frequency of maximum power ( $\nu_{\max}$ ), and the overtone frequency separation ( $\Delta\nu$ ; Kjeldsen & Bedding 2011) to provide accurate initial guesses for fitting parameters. A significant shortcoming of SYD (and other similar pipelines) is that it does not assess whether a given star shows excess power from oscillations in a

statistically robust way, hence requiring post-processing and often visual verification. This introduces significant unknown, and subjective, detection bias, which hampers population analyses of seismic samples. Ensuring reproducible selection functions is particularly important for applications aimed to perform Galactic archaeology studies (Stello et al. 2017).

In this paper we introduce a new pipeline, the Bayesian Asteroseismology data Modeling Pipeline (BAM), which builds on the SYD pipeline with an eye toward automatic, robust classification of light curves. BAM formalizes relations among granulation,  $\nu_{\max}$ , and  $\Delta\nu$  through a Bayesian framework in which these relations are implemented as priors. It is this Bayesian framework that then allows for a self-consistent, statistical separation of oscillators from nonoscillators.

BAM was also developed with the particular challenges involved in extracting asteroseismic parameters from the repurposed *Kepler* mission, K2 (Howell et al. 2014), in mind. Following the failure of two of its reaction wheels, the *Kepler* satellite was realigned to point in the ecliptic plane. As opposed to *Kepler*'s single field of view in Cygnus, the K2 pointing pattern covers the ecliptic plane with a footprint of about  $100^\circ$  along the ecliptic. However, periodic small-angle pointing corrections are performed every 6 hr by firing the spacecraft thrusters, which introduce instrumental signatures in K2 light curves. These features unfortunately correspond to typical frequencies of red giant oscillations and can mimic true asteroseismic oscillations near  $\sim 47 \mu\text{Hz}$  (the 6 hr thruster firing frequency period). Because this instrumental feature overlaps in frequency with where a typical red clump star shows maximum oscillation power, it can hinder recovering red clump stars, which compose the largest population of red giants in the Galaxy. BAM's Bayesian framework uses information like the

amplitude of the power excess and the shape of the rest of the power spectrum to distinguish between *K2* thruster firing noise and genuine oscillations. In addition to this instrumental feature, the *K2* white-noise level is typically larger than the white noise of the original *Kepler* mission by a factor of about two, depending on how the data are processed (however, several *K2* light curve processing pipelines have reported near-*Kepler* white-noise levels; Vanderburg & Johnson 2014; Armstrong et al. 2015; Lund et al. 2015; Aigrain et al. 2016; Luger et al. 2016).

In addition to describing how BAM works in this paper, we apply it to extract global oscillation parameters for red giants observed serendipitously by *K2* through Guest Observer (GO) programs targeting dwarf stars during Campaign 1. This new sample of giants therefore adds to the already-known red giant sample from Stello et al. (2017).

## 2. Data

In this paper, we work with two sets of *K2* light curves: (1) the Campaign 1 (C1) target sample from the *K2* Galactic Archaeology Program (GAP; Stello et al. 2015, 2017),<sup>7</sup> which comprises 8630 stars, and (2) all non-GAP C1 targets, 13,016 in total.<sup>8</sup> Results from BAM for the former sample have been published in Stello et al. (2017). We review some of those results here and extend the application of BAM to the latter sample in order to identify serendipitous red giants.

All our C1 light curves have been generated by Vanderburg & Johnson (2014) (VJ), who perform aperture photometry on *K2* images and remove trends associated with centroid errors caused by the spacecraft’s unstable pointing. We will show below that this pre-processing does not completely remove the thruster-induced instrumental features from the data and therefore requires additional processing in BAM.

We begin by first removing trends on timescales much longer than solar-like oscillation timescales for the stars we are interested in. For each light curve, we perform high-pass filtering by dividing the VJ light curve by a 4-day wide, boxcar-smoothed version of the light curve, thus imposing a high-pass cutoff frequency of  $\sim 3 \mu\text{Hz}$ ; frequencies below this limit are not considered in any of our analyses.<sup>9</sup> Next, we fill in small gaps in the light curve of up to three consecutive points with linear interpolation and remove  $4\sigma$  outliers. This procedure results in a smoother power spectrum and less contamination from the spectral window, without biasing global oscillation parameters (Stello et al. 2015). We will see, however, that for some stars additional measures are required to account for spectral window effects. We then calculate a power spectrum of the resulting light curve with a Lomb–Scargle periodogram (Scargle 1982).

Despite the efforts to remove systematic errors, the VJ light curves still exhibit non-negligible contamination at frequencies of 48.1 and 46.3  $\mu\text{Hz}$  owing to thruster firings. Generally, we do not find excess power at the nominal thruster firing frequency of 47.22  $\mu\text{Hz}$ . Figure 1 shows a median power spectrum across all GAP C1 spectra (8630 spectra in total) in a region around the thruster firing frequency. To calculate this

spectrum, we normalized each spectrum to the white-noise level, defined to be the median power density in a range from 250  $\mu\text{Hz}$  to the Nyquist frequency of 283  $\mu\text{Hz}$ .

In order to investigate whether the thruster firing noise features showed temporal variation over the course of the campaign, we computed a wavelet periodogram using the *astroML* library (Vanderplas et al. 2012). The chosen wavelet has the form

$$w(t, t_0, f_0, Q) \propto e^{-[f_0(t-t_0)/Q]^2} e^{2\pi i f_0(t-t_0)},$$

where  $t_0$  and  $f_0$  are the time and frequency of the 2D wavelet transform,  $t$  is the time coordinate for the entire baseline considered, and  $Q$  is a factor determining the time resolution of the wavelet transform:  $Q \rightarrow \infty$  recovers a Fourier transform and  $Q \rightarrow 0$  yields a wavelet periodogram with infinite temporal resolution. We set  $Q = 30$  for analyzing the noise feature of interest, which allows for resolving features in time of approximately 1/10 the baseline of C1, i.e., 8 days.

Two representative wavelet periodograms for C1 are shown in Figure 2. We find that there are definite temporal structures in the frequency domain of the *K2* thruster firing noise. We note that C1 light curves reduced by Angus et al. (2016) also exhibit qualitatively similar features.

Given that these noise features are present in most of the VJ light curves, we remove the affected regions of the power spectra in Fourier space by replacing each frequency bin in 0.2  $\mu\text{Hz}$  wide regions on either side of 47.2 and 48.1  $\mu\text{Hz}$ , as well as a 0.4  $\mu\text{Hz}$  wide region on either side of 46.3  $\mu\text{Hz}$ . We replace the power density in this region with power drawn from a chi-square distribution scaled to a linear interpolation between the median power in regions 5  $\mu\text{Hz}$  on either side of the affected regions.

## 3. Methods

After the pre-processing of the power spectrum with power,  $A_o(\nu_j)$ , at discrete frequencies,  $\nu_j$ , which constitutes our data,  $D$ , we then fit a smooth background component to the power spectrum, whose sets of parameters,  $\theta_{\text{meso}}$  and  $\theta_{\text{gran}}$ , are used as guesses for a subsequent stage of determining the global asteroseismic parameter  $\nu_{\text{max}}$  and the other parameters describing the oscillation excess,  $\theta_{\text{excess}}$ , which is finally used to guide fitting the global asteroseismic parameters related to  $\Delta\nu$ ,  $\theta_{\Delta\nu}$ .

We discuss each step in turn below.

### 3.1. Granulation Calculation

BAM first fits a two-component Harvey-like model that Kallinger et al. (2014) find best describes the smooth background component of *Kepler* red giant power spectra:

$$A(\nu_j) = \left[ W_N(\nu_j) \text{sinc}\left(\frac{\pi}{2} \frac{\nu_j}{\nu_{\text{Nyq}}}\right) \right]^2 \sum_{i=1,2} \frac{\sigma_i^2 \tau_i}{1 + (\pi \nu_j \tau_i)^4} + \text{WN} \quad (1)$$

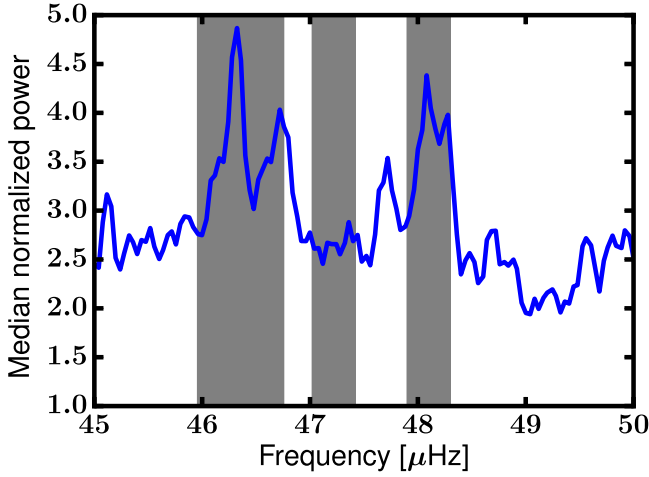
$$= A_{\text{meso}}(\nu_j) + A_{\text{gran}}(\nu_j) + \text{WN}, \quad (2)$$

where WN represents a white-noise term, which will dominate red giant power spectra at high frequencies;  $\sigma_i$  are amplitudes of each so-called Harvey component; and  $\tau_i$  are their characteristic timescales.  $A_{\text{meso}}(\nu_j)$  and  $A_{\text{gran}}(\nu_j)$  are defined here to be the two Harvey components of the granulation background. The sinc pre-factor with dependence on the

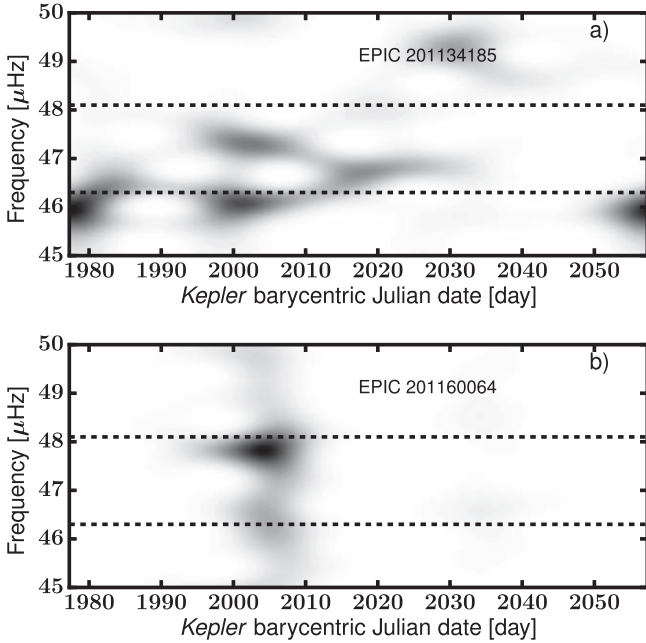
<sup>7</sup> <http://www.physics.usyd.edu.au/k2gap/> <https://archive.stsci.edu/prepds/k2gap/>

<sup>8</sup> We exclude the trans-Neptunian object, EPIC 200001049.

<sup>9</sup> We do, however, identify red giants with solar-like oscillations at frequencies  $\sim 3 \mu\text{Hz}$ , but the measured frequencies are upper limits and are not assigned errors.



**Figure 1.** Median spectrum for all C1 objects. We identify two regions particularly affected by *K2* noise in *VJ* light curves:  $46.3 \pm 0.4 \mu\text{Hz}$  (left) and  $48.1 \pm 0.2 \mu\text{Hz}$  (right). The middle gray shaded region ( $47.22 \pm 0.2 \mu\text{Hz}$ ) corresponds to the nominal thruster firing frequency of the spacecraft. These regions are treated specially in BAM, as described in the text.



**Figure 2.** Two examples of a wavelet analysis of the same frequency range around the nominal thruster firing frequency as shown in Figure 1, for EPIC 201134185 (top) and EPIC 201160064 (bottom). Clearly the 48.1 and 46.3  $\mu\text{Hz}$  instrumental features seen in the median spectrum (Figure 1) are not necessarily both present in every light curve at the same level and do not necessarily persist over the entire time baseline.

Nyquist frequency,  $\nu_{\text{Nyq}}$ , arises owing to *K2*'s finite exposure times, and  $W_N(\nu_j)$  is the spectral window function (see Kallinger et al. 2014 for more details).

Of the two Harvey-like components, the component at higher frequency is attributed to granulation, whereby the integrated light from the stellar disk varies owing to convective cell brightness variations. The lower-frequency component is attributed to mesogranulation, which is likely due to the variation in convective cell brightness for cells with sizes around 5–10 times that of granular cells (for a review of convection on the stellar surface, see Nordlund et al. 2009). For bookkeeping purposes, we require that the second component

**Table 1**

Priors Used for the Full Power Spectrum Fit, Equation (8), Adapted from Kallinger et al. (2010)

Parameter	Prior Distribution	Use
$\ln \sigma_{\text{gran}}$	$\mathcal{N}(-0.609 \ln \nu_{\text{max}} + 8.70, 0.165)$	Equations (4) and (8)
$\ln \tau_{\text{gran}}$	$\mathcal{N}(-0.992 \ln \nu_{\text{max}} - 1.09, 0.0870)$	Equations (4) and (8)
$\ln \sigma_{\text{meso}}$	$\mathcal{N}(-0.609 \ln \nu_{\text{max}} + 8.70, 0.165)$	Equations (4) and (8)
$\ln \tau_{\text{meso}}$	$\mathcal{N}(-0.970 \ln \nu_{\text{max}} + 0.00412, 0.970)$	Equations (4) and (8)
$\ln \frac{\tau_{\text{meso}}}{\tau_{\text{gran}}}$	$\mathcal{N}(1.386, 0.316)$	Equations (4) and (8)
$\ln b$	$\mathcal{N}(1.05 \ln \nu_{\text{max}} - 1.91, 0.198)$	Equation (8)
$\ln A_{\text{meso}} + \ln b$	$\mathcal{N}(-1.32 \ln \nu_{\text{max}} + 14.5, 1.22)$	Equation (8)

**Note.** The notation  $\mathcal{N}(a, b)$  indicates a Gaussian distribution with mean  $a$  and standard deviation  $b$ . Whether or not a given prior enters into Equation (4) or Equation (8) is indicated in the final column.

always be identified with the granulation background for which  $\tau_{\text{meso}} > \tau_{\text{gran}}$  and  $\sigma_{\text{gran}}^2 \tau_{\text{gran}} < \sigma_{\text{meso}}^2 \tau_{\text{meso}}$ .

We achieve a robust fit to the granulation background by taking advantage of scaling relations between  $\nu_{\text{max}}$  and the granulation parameters ( $\sigma$  and  $\tau$ ) noted by previous work (e.g., Kallinger et al. 2010; Kjeldsen & Bedding 2011). These relations naturally translate into priors in a Bayesian framework. We construct priors on the granulation parameters as detailed in Table 1. The final prior for a set of trial parameters is the product of the individual priors according to

$$\begin{aligned}
 P(\theta_{\text{meso}} = \{\sigma_{\text{meso}}, \tau_{\text{meso}}\}, \theta_{\text{gran}} = \{\sigma_{\text{gran}}, \tau_{\text{gran}}\}, \theta_{\text{excess}}) \\
 &= P(\sigma_{\text{meso}} | \tau_{\text{meso}}, \sigma_{\text{gran}}, \tau_{\text{gran}} | \theta_{\text{excess}}) \\
 &P(\tau_{\text{meso}}, \tau_{\text{gran}} | \sigma_{\text{gran}}, \theta_{\text{excess}}) P(\sigma_{\text{gran}} | \theta_{\text{excess}}) P(\theta_{\text{excess}}) \\
 &= P(\sigma_{\text{meso}} | \nu_{\text{max}}) P(\tau_{\text{meso}} | \nu_{\text{max}}) P(\tau_{\text{gran}} | \nu_{\text{max}}) P\left(\frac{\tau_{\text{meso}}}{\tau_{\text{gran}}}\right) \\
 &P(\sigma_{\text{gran}} | \nu_{\text{max}}),
 \end{aligned} \tag{3}$$

for which we introduce the notation  $\theta_{\text{excess}}$  to indicate parameters describing the solar-like oscillations (as distinguished from the granulation parameters), and whose parameters (other than  $\nu_{\text{max}}$ ) are defined later. The granulation priors are conditional upon  $\nu_{\text{max}}$ , and, in this sense,  $\nu_{\text{max}}$  is a latent variable that defines the relationships among all the granulation parameters.

Subsequently, we define a posterior probability given by

$$\begin{aligned}
 P(\theta_{\text{meso}}, \theta_{\text{gran}} | D = \{(\nu_j, A_o(\nu_j)), j = 0, 1, 2, \dots\}, \theta_{\text{excess}}) \\
 \propto P(\theta_{\text{meso}}, \theta_{\text{gran}} | \theta_{\text{excess}}) \prod_j \left[ \frac{1}{A(\nu_j)} \exp\left(-\frac{A_o(\nu_j)}{A(\nu_j)}\right) \right].
 \end{aligned} \tag{4}$$

Here,  $A_o(\nu_j)$  is the observed spectral density and  $A(\nu_j)$  is the model given by Equation (2). Note that the above expression assumes  $\chi^2$  statistics and not Gaussian statistics to describe  $A_o(\nu_j) / \langle A(\nu_j) \rangle \sim \chi^2(2)$ , where the observed spectrum is critically sampled and the observed spectrum is modeled by  $A(\nu_j)$ .

Given a Bayesian model for the data, we explore the parameter space with the Markov chain Monte Carlo (MCMC) method, as implemented in *emcee* (Foreman-Mackey et al. 2013), and report best-fitting parameters as the median of their marginalized posterior distributions and the uncertainty as the average of the range around the median encompassing 64% of the distribution. Of course, the prior factor,  $P(\theta_{\text{meso}}, \theta_{\text{gran}} | \theta_{\text{excess}})$ , depends on  $\nu_{\text{max}}$  (see Table 1). We simultaneously fit for  $\nu_{\text{max}}$  and the background parameters, with a guess for  $\nu_{\text{max}}$  calculated from a smoothed version of the spectrum, as in the SYD pipeline (Huber et al. 2009). Note that in this step the region of power excess is not explicitly modeled, and so  $\nu_{\text{max}}$  is implemented effectively as a dummy variable for this granulation model fitting stage of the process. The resulting best-fitting parameters are then used as initial guesses for a more complicated model that adds an additional component to describe the oscillation excess power, which we describe next.

Ultimately, BAM allows the user to choose which of the priors listed in Table 1 are to be used. The results presented in this paper do not use the first four priors of Table 1 for this granulation background fitting step, though they are used for the subsequent fitting step that determines  $\nu_{\text{max}}$  and  $A_{\text{max}}$ , as described in the next section. The extent to which the priors in Table 1 are applied does not significantly affect the resulting  $\nu_{\text{max}}$  value.

### 3.2. $\nu_{\text{max}}$ and $A_{\text{max}}$ Calculation

In the subsequent step, we add another component to the model such that

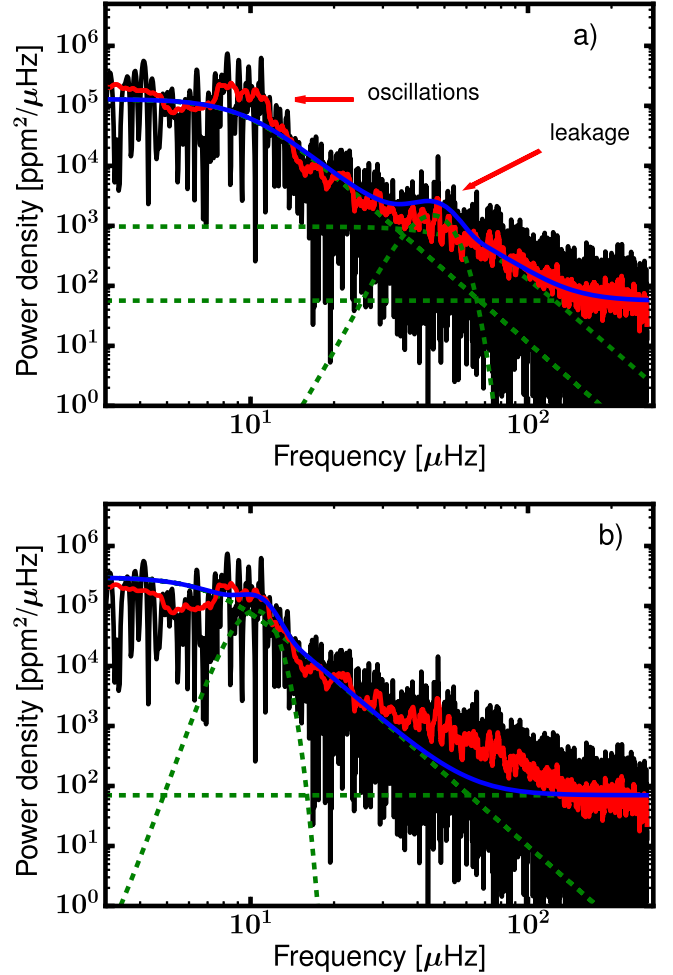
$$A_{\text{tot}}(\nu_j) = A_{\text{meso}}(\nu_j) + A_{\text{gran}}(\nu_j) + A_{\text{excess}}(\nu_j) + \text{WN}, \quad (5)$$

where  $A_{\text{excess}}$  represents the power excess from solar-like oscillations and  $A_{\text{meso}}(\nu_j)$ ,  $A_{\text{gran}}(\nu_j)$ , and WN are defined in Equation (2). We model  $A_{\text{excess}}$  as a Gaussian profile

$$A_{\text{excess}} = A_{\text{max}} \left[ W_N(\nu_j) \text{sinc}\left(\frac{\pi}{2} \frac{\nu_j}{\nu_{\text{Nyq}}}\right) \right]^2 \times \exp\left[-\frac{(\nu_j - \nu_{\text{max}})^2}{2b^2}\right]. \quad (6)$$

Our prior is now

$$\begin{aligned} P(\theta_{\text{meso}}, \theta_{\text{gran}}, \theta_{\text{excess}} = \{A_{\text{max}}, \nu_{\text{max}}, b\}) \\ &= P(\sigma_{\text{meso}} | \tau_{\text{meso}}, \tau_{\text{gran}}, \sigma_{\text{gran}}, \theta_{\text{excess}}) \\ &\quad P(\tau_{\text{meso}}, \tau_{\text{gran}} | \sigma_{\text{gran}}, \theta_{\text{excess}}) P(\sigma_{\text{gran}} | \theta_{\text{excess}}) P(\theta_{\text{excess}}) \\ &= P(\sigma_{\text{meso}} | \nu_{\text{max}}) P(\tau_{\text{meso}} | \nu_{\text{max}}) P(\tau_{\text{gran}} | \nu_{\text{max}}) P\left(\frac{\tau_{\text{meso}}}{\tau_{\text{gran}}}\right) \\ &\quad P(\sigma_{\text{gran}} | \nu_{\text{max}}) P(b, A_{\text{max}}, \nu_{\text{max}}) \\ &= P(\sigma_{\text{meso}} | \nu_{\text{max}}) P(\tau_{\text{meso}} | \nu_{\text{max}}) P(\tau_{\text{gran}} | \nu_{\text{max}}) P\left(\frac{\tau_{\text{meso}}}{\tau_{\text{gran}}}\right) \\ &\quad P(\sigma_{\text{gran}} | \nu_{\text{max}}) P(b | \nu_{\text{max}}) P(A_{\text{max}}, b | \nu_{\text{max}}). \end{aligned} \quad (7)$$



**Figure 3.** Raw (black) and smoothed (red) power spectrum of EPIC 201186616, and model fits (a) without and (b) with spectral window corrections (blue). Each component of the models is shown with green dashed curves (white noise, Gaussian excess, and Harvey components). The mesogranulation component does not contribute significantly to the fit upon spectral window correction, and so it is not shown in panel (b).

We construct a posterior probability given by

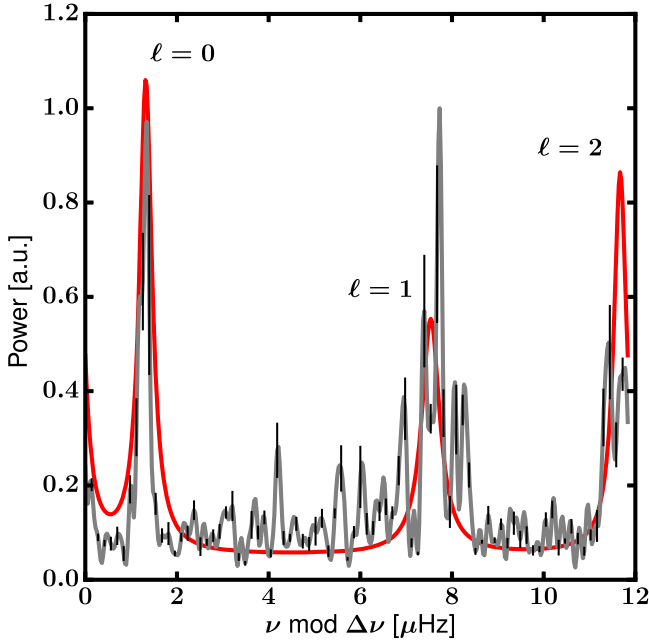
$$P(\theta_{\text{meso}}, \theta_{\text{gran}}, \theta_{\text{excess}} | D) \propto P(\theta_{\text{meso}}, \theta_{\text{gran}}, \theta_{\text{excess}}) \prod_j \left[ \frac{1}{A_{\text{tot}}(\nu_j)} \exp\left(-\frac{A_o(\nu_j)}{A_{\text{tot}}(\nu_j)}\right) \right]. \quad (8)$$

In this case, the total prior is a product over all priors listed in Table 1. By first fitting the parameters of the granulation as described in Section 3.1 and subsequently using these as priors for the fit involving both the granulation model and the Gaussian excess, we reduce the burn-in time and the chance of getting stuck at local maxima. It will also make more convenient our oscillator selection process, described in Section 3.6.

### 3.3. Low-frequency Oscillators

We find that objects oscillating at frequencies  $\nu_{\text{max}} \lesssim 15 \mu\text{Hz}$  exhibit significant spectral leakage at frequencies  $30 \mu\text{Hz} \lesssim \nu \lesssim 100 \mu\text{Hz}$ , often confusing the pipeline to fit a  $\nu_{\text{max}}$  at the location of the leakage, as shown in Figure 3(a). We





**Figure 4.** Example fit by BAM to the folded central power spectrum, with best-fitting model (red) and data (gray); black error bars are calculated as described in the text, of which every fifth is shown, for clarity.

correct for this leakage at each step in our MCMC chains: for each trial model granulation spectrum (Equation (2)), we compute an amplitude spectrum, with each frequency in the spectrum being assigned a random spectral phase. This amplitude spectrum is then convolved with the spectral window and squared to yield a power spectrum (see Murphy et al. 2013 for a worked example of how to contend with the spectral window in the context of asteroseismology, specifically). A lightly smoothed version of this convolved granulation power spectrum is added to the power excess term to create a model of the power spectrum that takes into account spectral leakage. This model is then fitted to the observed power spectrum within the Bayesian framework. Note that the trial power excess term is not convolved with the window function; as it turns out, it adds minimally to the spectral leakage compared to the granulation background, and it can lead to unstable fits in which the entire spectrum is modeled as a Gaussian excess plus its resulting spectral leakage. We find that this procedure results in correct  $\nu_{\max}$  identifications for  $\nu_{\max} \lesssim 15$ . Correcting for spectral leakage results in a statistically significant difference in fitted granulation parameters for low-frequency oscillators (Figure 3(b); note the difference in shape of the blue curve in regions dominated by granulation).

A caveat for these stars is that the lowest  $\nu_{\max}$  ( $\nu_{\max} \lesssim 4 \mu\text{Hz}$ ) values likely represent upper limits for  $\nu_{\max}$  because the *K2* resolution prevents an unambiguous determination of  $\nu_{\max}$ . Indeed, at frequencies near  $\sim 3 \mu\text{Hz}$ , there may only be three modes visible (e.g., Stello et al. 2014), which limits the precision with which a central  $\nu_{\max}$  may be defined using the Gaussian to model oscillation excess (Equation (6)).

#### 3.4. $\Delta\nu$ Calculation

We furthermore take advantage of the correlation between  $\nu_{\max}$  and  $\Delta\nu$  to place a prior on  $\Delta\nu$  in the same way we place priors on granulation parameters described in Sections 3.1 and 3.2. Because

of the short duration of *K2* light curves ( $\sim 80$  days), individual modes may not be well resolved, and therefore the large frequency separation can be difficult to measure. BAM measures  $\Delta\nu$  in two independent ways: one using the SYD autocorrelation method (see Huber et al. 2009), and the other using the  $\Delta\nu$ -folded power spectrum centered around  $\nu_{\max}$  and extending on  $3\Delta\nu$  on either side, as shown in Figure 4. The background contributions from the Harvey components of the model are divided out, and the folded power spectrum is computed by folding the spectrum on  $\Delta\nu$ , where each bin of the folded spectrum contains the sum over the power by folding the spectrum  $3\Delta\nu$  on either side of  $\nu_{\max}$  by  $\Delta\nu$ ; the bins are then normalized such that the highest peak of the folded power spectrum is unity. For the majority of red giants the folded spectrum shows three broad oscillation power excess regions corresponding to the radial, dipole, and quadrupole modes. We do not fit an octopole mode component because its low power usually makes it undetectable in *K2* data. We obtain  $\Delta\nu$  from this diagram by modeling it using three Lorentzian profiles, appropriate for solar-like oscillation modes, corresponding to the radial ( $\ell = 0$ ), dipole ( $\ell = 1$ ), and quadrupole ( $\ell = 2$ ) modes, as follows:

$$A_{\text{folded}}(\nu_j, (\nu_\ell, A_\ell, \text{FWHM}_\ell)_{\ell=0,1,2}, \Delta\nu, C) = \sum_{\ell=0}^{\ell=2} \frac{A_\ell}{1.0 + \frac{[(\nu_j \bmod \Delta\nu) - \nu_\ell]^2}{\text{FWHM}_\ell^2 / 4}} + C. \quad (9)$$

$C$  is a constant to model the imperfections when removing the background level in the vicinity of  $\nu_{\max}$ . The frequencies of the modes,  $\nu_\ell$ , in the folded central power spectrum are given by

$$\begin{aligned} \nu_0 &\equiv \epsilon \\ \nu_1 &= \nu_0 - \frac{1}{2}\Delta\nu + \delta\nu_{01} \\ \nu_2 &= \nu_0 - \delta\nu_{02}. \end{aligned}$$

The positions of the nonradial modes with respect to the radial mode,  $\epsilon$ , thus follow standard definitions (e.g., Bedding & Kjeldsen 2010), such that a given mode in the spectrum has a frequency,  $\nu$ , given by  $\nu \approx \Delta\nu(n + \ell/2 + \epsilon)$ , where  $n$  is the radial order of the mode.

Placing priors on the above parameters as detailed in Table 2 following the procedure in Sections 3.1 and 3.2 of the form

$$\begin{aligned} P(\theta_{\Delta\nu} &= \{(\delta_{01}, \delta_{02}), (A_0, A_1, A_2), \\ &(\text{FWHM}_0, \text{FWHM}_1, \text{FWHM}_2), \Delta\nu, C\} | \theta_{\text{excess}}) \\ &= P((\delta_{01}, \delta_{02}), (A_0, A_1, A_2), \\ &(\text{FWHM}_0, \text{FWHM}_1, \text{FWHM}_2), \Delta\nu | b, A_{\max}, \nu_{\max}) \\ &= P((\delta_{01}, \delta_{02}) | \nu_{\max}) P((A_0, A_1, A_2) | \nu_{\max}) \\ &P((\text{FWHM}_0, \text{FWHM}_1, \text{FWHM}_2) | \nu_{\max}) P(\Delta\nu | \nu_{\max}) \end{aligned}$$

yields a posterior probability

$$P(\theta_{\Delta\nu} | D, \theta_{\text{excess}}) \propto P(\theta_{\Delta\nu} | \theta_{\text{excess}}) \prod_j \left[ \frac{A_{o,\text{folded},j}(\nu_j, \Delta\nu)^{n_j-1}}{A_{\text{folded},j}(\nu_j, \theta_{\Delta\nu})^{n_j}} \exp\left(-n_j \frac{A_{o,\text{folded}}(\nu_j, \Delta\nu)}{A_{\text{folded}}(\nu_j, \theta_{\Delta\nu})}\right) \right], \quad (10)$$

where we use the statistics for an averaged spectrum derived in Appourchaux (2003).  $A_{o,\text{folded}}(\nu_j, \Delta\nu)$  is the power at frequency bin  $\nu_j$  in the observed folded spectrum for a given  $\Delta\nu$  and is a function of  $\Delta\nu$ : depending on  $\Delta\nu$ , the folding process will distribute the power in frequency bins,

**Table 2**  
Priors Used for the Fit to  $\Delta\nu$ , Equation (9)

Parameter	Prior Distribution
$\delta_{01}$	$\mathcal{N}(-0.025\Delta\nu^a, 0.1\Delta\nu^b)$
$\delta_{02}$	$\mathcal{N}(0.121\Delta\nu^a + 0.047^a, 0.1\Delta\nu^b)$
$A_0$	$\mathcal{N}(1.0^c, 0.15^b)$
$A_1$	$\mathcal{N}(0.5^c, 0.15^b)$
$A_2$	$\mathcal{N}(0.8^c, 0.15^b)$
$\text{FWHM}_0$	$\mathcal{U}(0.035\Delta\nu_{\text{guess}}^b, 0.45\Delta\nu_{\text{guess}}^b)$
$\text{FWHM}_1$	$\mathcal{U}(0.035\Delta\nu_{\text{guess}}^b, 0.9\Delta\nu_{\text{guess}}^b)$
$\text{FWHM}_2$	$\mathcal{U}(0.035\Delta\nu_{\text{guess}}^b, 0.45\Delta\nu_{\text{guess}}^b)$
$\Delta\nu$	$\mathcal{N}(\Delta\nu_{\text{guess}}^b, 0.15\Delta\nu_{\text{guess}}^b)$
$C$	$\mathcal{U}(0.00^b, 0.1^b)$

**Note.** The notation  $\mathcal{U}(a, b)$  indicates a uniform distribution between  $a$  and  $b$ . Priors adapted from Huber et al. (2010) (superscript <sup>a</sup>), this work (superscript <sup>b</sup>), and Stello et al. (2016) (superscript <sup>c</sup>).  $\Delta\nu_{\text{guess}}$  is the expected  $\Delta\nu$  given a  $\nu_{\text{max}}$ , from Stello et al. (2009):  $\Delta\nu_{\text{guess}} \equiv 0.263\nu_{\text{max}}^{0.772}$ .

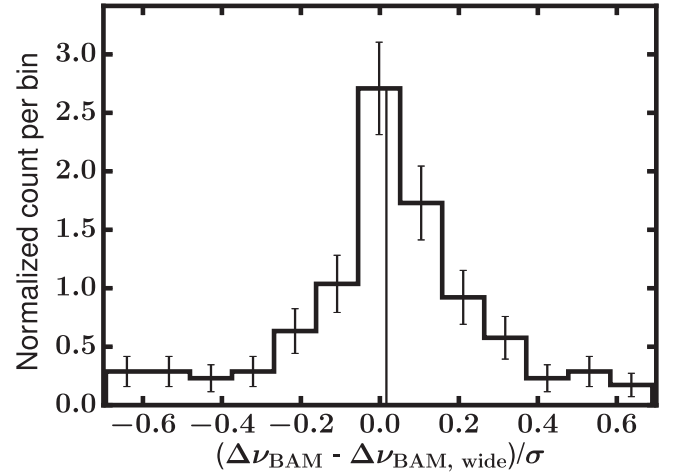
$A_{0, \text{folded}}(\nu_j, \Delta\nu)$ , differently. In practice, what this requires is recomputing the folded spectrum for each trial  $\Delta\nu$  in our MCMC.  $A_{\text{folded}}(\nu_j, \theta_{\Delta\nu})$  is the model for the folded spectrum (Equation (9)), and  $n_j$  is the number of points that went into the sum over power for that bin in the folded power spectrum.

Using the folded spectrum is particularly useful for determining  $\Delta\nu$  from K2 data because individual mode frequencies are not very well resolved. What complicates the recovery of  $\Delta\nu$  in the presence of degraded spectral resolution is that observed mode amplitudes and phases (and hence frequencies) are not stable with time and have intrinsic scatter. This is because the oscillations are stochastically driven and damped (e.g., Woodard 1984), which causes continuous variation in the centroid of mode frequencies and their amplitudes. The random behavior of the stochastic mode profile can only be mitigated by averaging spectra that are independent in frequency or in time (for a review of power spectrum statistics in the context of solar-like oscillations, see Anderson et al. 1990, and references therein). The folded spectrum approach therefore effectively averages out the random behavior of the modes and increases their signal-to-noise ratio, and it is what would be called an “ $m$ -averaged” spectrum (Anderson et al. 1990) in the context of solar modes.

To find the optimal  $\Delta\nu$ , we start with a guess value derived from the  $\Delta\nu$ – $\nu_{\text{max}}$  relation by Stello et al. (2009),

$$\Delta\nu_{\text{guess}} = 0.263\nu_{\text{max}}^{0.772}. \quad (11)$$

We determine best-fitting values by MCMC, in which  $\Delta\nu$  is constrained to be  $0.7\Delta\nu_{\text{guess}} < \Delta\nu < 1.3\Delta\nu_{\text{guess}}$ , and apply priors as described in Table 2. BAM returns  $\Delta\nu$  values for stars for which there is agreement to within  $2\sigma$  with  $\Delta\nu$  computed using the SYD autocorrelation method and for which the uncertainty on  $\Delta\nu$  is less than the spread in the  $\Delta\nu$  prior. The latter requirement captures information about how reliably the modes have been fit and serves as a means of determining which stars have more information about  $\Delta\nu$  than our prior choice. Note that BAM’s second, separate  $\Delta\nu$  value from an autocorrelation approach acts as a sort of second opinion. This autocorrelation  $\Delta\nu$  will not in general be the same  $\Delta\nu$  that a stand-alone application of the SYD pipeline to the same star would: the autocorrelation method requires a  $\nu_{\text{max}}$  to identify the region of the power spectrum that contains the power



**Figure 5.** Difference in best-fitting  $\Delta\nu$  when using a  $\Delta\nu$  prior of width  $0.9\Delta\nu_{\text{guess}}$  ( $\Delta\nu_{\text{BAM, wide}}$ ) vs. the nominal  $0.15\Delta\nu_{\text{guess}}$ , normalized by the error in the difference,  $\sigma$ ; error bars on the histogram bins correspond to Poisson uncertainties. The vertical line corresponds to the median of the distribution. This indicates that the differences between BAM runs with an expanded prior on  $\Delta\nu$  result in insignificant differences—10 times smaller than the error on  $\Delta\nu$ —in the resulting  $\Delta\nu$ .

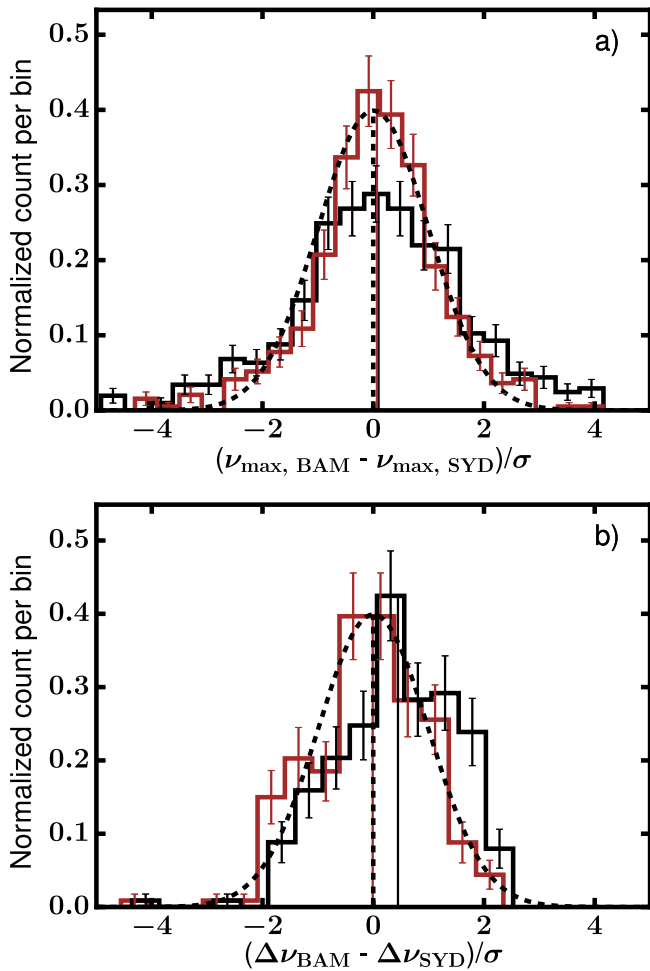
excess, and it also requires a removal of the smooth background of the power spectrum, both of which are independent of SYD in this case (for details of the autocorrelation approach to calculating  $\Delta\nu$ ; see Huber et al. 2009). We show an example of a model fit to the folded spectrum from this process in Figure 4.

Importantly, the priors that are placed on  $\Delta\nu$  are not too stringent. We tested the sensitivity of our  $\Delta\nu$  results on priors by increasing the spread in the  $\Delta\nu$  prior to  $0.9\Delta\nu_{\text{guess}}$  from  $0.15\Delta\nu_{\text{guess}}$  (see Table 2). For confirmed oscillators in the C1 K2GAP sample, our best-fitting  $\Delta\nu$  values are not significantly different when using our fiducial prior or a widened prior. We show the difference in best-fitting  $\Delta\nu$  using these two different priors in Figure 5. The spread is less than  $0.1\sigma$  for the majority of objects, indicating that the priors indeed do not significantly impact the determination of  $\Delta\nu$ .

### 3.5. Comparison to SYD

BAM parameters agree favorably with those computed by other techniques via different pipelines, as demonstrated in Stello et al. (2017). As a point of comparison to a well-established asteroseismic pipeline, Figure 6 shows BAM  $\nu_{\text{max}}$  and  $\Delta\nu$  values compared to those from SYD for the C1 GAP oscillator sample. The BAM parameters for this comparison exercise have been rederived using a slightly different methodology than described in the GAP Data Release 1 (GAP DR1) release paper (Stello et al. 2017) so as to be consistent with the methodology presented in this work. SYD values for  $\Delta\nu$  and  $\nu_{\text{max}}$  are taken directly from GAP DR1. Only giant candidates that were verified to be such by eye in Stello et al. (2017) and that BAM selects as giants according to Section 3.6 are considered in this comparison exercise.

The median in the normalized distribution of differences between BAM and SYD  $\Delta\nu$  values for this GAP comparison sample (solid black vertical line in Figure 6(b)) indicates a systematic offset of  $\sim 0.6\%$ . The red histogram in Figure 6(b) shows the  $\Delta\nu$  difference distribution if the BAM values are



**Figure 6.** Distributions of the differences between BAM and SYD (a)  $\nu_{\max}$  and (b)  $\Delta\nu$ , normalized by the sum in quadrature of their errors,  $\sigma \equiv \sqrt{\sigma_{\nu_{\max, \text{BAM}}}^2 + \sigma_{\nu_{\max, \text{SYD}}}^2}$  and  $\sigma \equiv \sqrt{\sigma_{\Delta\nu_{\text{BAM}}}^2 + \sigma_{\Delta\nu_{\text{SYD}}}^2}$ . The medians of both distributions are shown as vertical solid black lines; error bars on the histogram bins correspond to Poisson uncertainties. The red distributions in each panel indicate the distributions of differences in BAM and SYD values after systematic differences in central value and/or uncertainties are corrected, according to the text. The dotted curve is a Gaussian, to guide the eye; the vertical dashed line is centered at zero. Stars plotted here are drawn from the C1 GAP sample deemed from manual inspection to be definite oscillators (see Stello et al. 2017) and such that both SYD and BAM as implemented in this work returned  $\nu_{\max}$  or  $\Delta\nu$  values.

rescaled downward by 0.6%, which brings the distribution into better alignment with the expected Gaussian (black dashed curve). The median in the distribution of  $\nu_{\max}$  differences indicates a marginally significant ( $1\sigma$ ) systematic offset between the two  $\nu_{\max}$  scales (solid black vertical line in Figure 6(a)), which corresponds to a difference in BAM and SYD  $\nu_{\max}$  scales of  $\sim 0.2\%$ . There does appear to be an underestimation of either BAM or SYD  $\Delta\nu$  uncertainties (black histogram in Figure 6(a) is wider than the expected Gaussian; black dashed curve), which is ameliorated by rescaling the error on the difference upward by 30% (red histogram in Figure 6(a)).

Given that Kallinger et al. (2014) found systematic differences of up to  $\sim 5\%$  in  $\nu_{\max}$  depending on the model used for the mesogranulation and granulation background, any small systematic difference in  $\nu_{\max}$  could easily be due to the different treatment of the background between BAM and SYD. For example, the sinc term in Equation (2) is not included in

the SYD pipeline. This difference in methodology could plausibly explain the 0.6% systematic difference in  $\Delta\nu$ , as well: the positions of the modes used to measure  $\Delta\nu$  will be affected by the choice of the mesogranulation and granulation background, which are removed before calculating the folded spectrum.

Apart from these systematic differences, we find that BAM parameters are consistent with SYD to within  $\sim 1.53\%$  and  $1.51\%$  for  $\nu_{\max}$  and  $\Delta\nu$ , which correspond to the BAM GAP sample mean fractional errors on  $\nu_{\max}$  and  $\Delta\nu$ , respectively. There is some ambiguity as to the agreement in  $\nu_{\max}$ , where the errors on  $\nu_{\max}$  for either BAM or SYD may be underestimated by up to 30%, given the non-Gaussianity of the  $\nu_{\max}$  difference distribution (black histogram in Figure 6(a)). Non-Gaussianity in comparisons across pipelines was also found in Stello et al. (2017) and in part is caused by under- and overestimation of errors in *K2* asteroseismic parameters (Pinsonneault et al. 2018, J. C. Zinn et al. 2019, in preparation).

### 3.6. Bayesian Oscillator Selection

Because our approach for measuring the oscillation and granulation parameters will always provide a best-fitting model, even if there is no solar-like oscillation signal, we still need to determine whether a fit corresponds to a true detection. As mentioned in Section 1, BAM’s Bayesian approach means that we can use the parameter fits to determine which stars are, and are not, true oscillators.

This is essentially a problem in model comparison: does the model with a power excess term (Equation (5)) describe a star’s power spectrum better, or does one without power excess (Equation (2))? Jeffreys (1935) first formalized model comparison in a Bayesian approach using what is now called the Bayes factor, defined to be the ratio of the posterior odds in favor of a model to its prior odds. The Bayes factor derives simply from the Bayes theorem, by which the posterior odds of  $M_1$  can be written as

$$\frac{P(M_1|D)}{P(M_2|D)} = \frac{P(D|M_1) P(M_1)}{P(D|M_2) P(M_2)}. \quad (12)$$

In our case, the probability densities  $P(D|M_1)$  and  $P(D|M_2)$  correspond to integrals of Equations (8) and (4) over all of parameter space, and we assume that, a priori, a star is as likely to be a nonoscillator as an oscillator, in which case the prior odds of  $M_1$  are  $\frac{P(M_1)}{P(M_2)} = 1$ . The Bayes factor is defined as

$$B \equiv \frac{P(D|M_1)}{P(D|M_2)}.$$

To compute the Bayes factor, one needs to integrate the conditional probability densities of Equations (8) and (4) over all of parameter space. Though these conditional probability densities share the same priors on granulation parameters,  $P(\theta_{\text{meso}}, \theta_{\text{gran}}|\theta_{\text{excess}})$ , they do not neatly cancel out when computing the Bayes factor because  $P(D|M_1)$  and  $P(D|M_2)$  in Equation (12) are each separate integrals involving these priors. Such integrals are often computationally expensive to do and analytically intractable. Fortunately, there are various methods available to approximate the Bayes factor (e.g., Green 1995; Chib & Jeliazkov 2001; Skilling 2004). We use the widely applicable Bayesian information criterion (WBIC; Watanabe 2013) to compute the Bayes factor. This method generalizes the Bayesian information criterion (Schwarz 1978), such that the WBIC approximates the Bayes factor in the limit of weak priors and with the assumption that the posterior is

**Table 3**  
Campaign 1 Non-GAP BAM Asteroseismic Parameters for Giants and Giant Candidates

EPIC	2MASS	R.A.	Decl.	$\nu_{\max}$	$\sigma_{\nu_{\max}}$	$\Delta\nu$	$\sigma_{\Delta\nu}$	$A_{\max 2}$	$\sigma_{A_{\max 2}}$	$\sigma_{\text{meso}}$	$\sigma_{\sigma_{\text{meso}}}$	$\tau_{\text{meso}}$	$\sigma_{\tau_{\text{meso}}}$	$\sigma_{\text{gran}}$	$\sigma_{\sigma_{\text{gran}}}$	$\tau_{\text{gran}}$	$\sigma_{\tau_{\text{gran}}}$	WN	$\sigma_{\text{WN}}$	Giant?
		(deg)	(deg)	( $\mu\text{Hz}$ )	( $\mu\text{Hz}$ )	( $\mu\text{Hz}$ )	( $\mu\text{Hz}$ )	(ppm)	(ppm)	(ppm)	(ppm)	( $\mu\text{Hz}^{-1}$ )	( $\mu\text{Hz}^{-1}$ )	(ppm)	(ppm)	( $\mu\text{Hz}^{-1}$ )	( $\mu\text{Hz}^{-1}$ )	(ppm <sup>2</sup> )	(ppm <sup>2</sup> )	
								( $\mu\text{Hz}^{-1}$ )	( $\mu\text{Hz}^{-1}$ )									( $\mu\text{Hz}^{-1}$ )	( $\mu\text{Hz}^{-1}$ )	
201147434	11452628–0521209	176.359512	−5.355842	67.604	2.676	6.179	0.696	2.994e+03	9.890e+02	4.081e+02	4.279e+01	1.977e−02	1.790e−03	4.936e+02	3.934e+01	4.681e−03	4.733e−04	1.125e+03	1.211e+01	1
201151637	11360854–0515250	174.035575	−5.256981	57.160	4.415	4.894	0.021	2.166e+03	1.287e+03	4.491e+02	3.038e+01	2.230e−02	3.120e−03	5.273e+02	2.435e+01	5.841e−03	3.477e−04	5.719e+02	1.328e+01	1
201156121	11315343–0508579	172.972654	−5.149433	51.849	3.264	...	...	2.314e+03	1.125e+03	3.754e+02	3.378e+01	2.725e−02	3.432e−03	4.534e+02	2.530e+01	6.410e−03	5.875e−04	1.159e+03	2.382e+01	1
201163464	11353033–0458260	173.876408	−4.973911	52.509	3.362	5.149	0.018	4.768e+03	3.751e+03	5.603e+02	4.441e+01	2.319e−02	2.310e−03	6.859e+02	3.957e+01	5.976e−03	4.948e−04	2.425e+03	5.196e+01	1
201198517	11391931–0408141	174.830492	−4.137267	42.417	0.645	...	...	4.554e+03	6.148e+02	1.235e+03	5.464e+01	3.896e−02	2.187e−03	5.012e+02	4.072e+01	5.776e−03	3.189e−04	6.436e+01	5.346e+00	2
201214537	11461882–0345352	176.578525	−3.759797	51.098	4.421	...	...	5.040e+03	3.956e+03	5.377e+02	3.964e+01	2.843e−02	3.819e−03	7.327e+02	3.243e+01	5.561e−03	3.705e−04	2.272e+03	4.871e+01	1
201228269	11245189–0332138	171.216150	−3.537211	50.382	2.234	...	...	3.211e+03	2.136e+03	4.665e+02	3.551e+01	2.524e−02	3.378e−03	5.765e+02	2.944e+01	5.763e−03	4.952e−04	1.290e+03	3.113e+01	1
201244712	11233476–0317217	170.894883	−3.289417	49.686	0.704	...	...	3.406e+03	6.118e+02	8.720e+02	3.486e+01	2.863e−02	1.788e−03	6.132e+02	2.231e+01	6.004e−03	1.920e−04	8.177e+01	2.313e+00	2
201251246	11291517–0311199	172.313354	−3.188886	3.309	...	...	...	2.566e+06	8.814e+05	4.222e+03	5.350e+02	3.516e−01	3.131e−02	3.295e+03	2.028e+02	8.672e−02	6.170e−03	6.972e+01	4.575e+00	1
201253257	...	179.712362	−3.158483	55.077	4.786	...	...	4.172e+03	3.050e+03	4.662e+02	4.151e+01	2.424e−02	3.321e−03	7.106e+02	3.180e+01	5.625e−03	3.720e−04	1.646e+03	3.615e+01	1
201262747	11542756–0300552	178.614854	−3.015358	98.614	4.590	...	...	5.610e+02	2.737e+02	3.355e+02	1.901e+01	1.457e−02	1.457e−03	3.358e+02	1.925e+01	3.774e−03	3.295e−04	5.86e+02	8.709e+00	1
201269306	11235529–0254559	170.980417	−2.915544	114.376	0.673	...	...	1.065e+03	9.483e+01	4.929e+02	1.667e+01	1.176e−02	6.325e−04	3.587e+02	1.476e+01	3.408e−03	1.805e−04	8.608e+01	2.869e+00	2
201272934	11190372–0251388	169.765562	−2.860719	89.965	7.486	...	...	1.427e+03	9.458e+02	4.507e+02	3.229e+01	1.387e−02	1.517e−03	5.431e+02	3.221e+01	3.682e−03	3.046e−04	1.025e+03	2.614e+01	1
201305005	11374947–0222333	174.456163	−2.375928	30.647	2.360	4.379	0.058	1.281e+04	9.055e+03	1.019e+03	7.423e+01	4.373e−02	3.892e−03	8.133e+02	6.819e+01	9.908e−03	9.029e−04	6.525e+03	1.110e+02	1
201310650	11351161–0217353	173.798379	−2.293164	47.040	4.038	3.891	0.022	5.000e+03	3.903e+03	5.739e+02	4.526e+01	2.915e−02	3.419e−03	7.150e+02	3.875e+01	6.180e−03	5.059e−04	2.635e+03	5.559e+01	1
201348966	11305698–0143174	172.737462	−1.721542	38.852	1.057	...	...	4.974e+03	1.102e+03	5.837e+02	3.973e+01	3.047e−02	2.635e−03	7.666e+02	2.502e+01	7.802e−03	2.314e−04	1.159e+02	2.557e+00	2
201361246	12005920–0132144	180.246675	−1.537347	47.547	5.332	...	...	3.093e+03	2.187e+03	5.869e+02	4.346e+01	2.613e−02	4.348e−03	6.068e+02	3.927e+01	7.203e−03	7.921e−04	1.866e+03	4.179e+01	1
201361404	11112565–0132069	167.856942	−1.535272	100.372	0.903	...	...	1.482e+03	3.635e+02	4.316e+02	2.491e+01	1.144e−02	1.093e−03	4.287e+02	2.480e+01	3.316e−03	2.431e−04	1.349e+02	1.135e+01	2
201370145	11130872–0124345	168.286287	−1.409614	82.959	1.019	...	...	1.266e+03	1.318e+02	5.517e+02	2.180e+01	1.510e−02	8.499e−04	4.355e+02	1.811e+01	4.215e−03	1.461e−04	4.157e+01	1.464e+00	2
201371239	11111931–0124305	167.830442	−1.394597	57.161	0.813	5.421	0.018	2.287e+03	3.498e+02	5.523e+02	2.636e+01	1.894e−02	1.235e−03	5.211e+02	2.139e+01	5.413e−03	1.759e−04	4.157e+01	1.183e+00	2
201374034	11062714–0121136	166.613092	−1.353819	98.965	6.816	9.528	0.028	4.333e+02	2.267e+02	2.891e+02	1.616e+01	1.252e−02	1.068e−03	7.279e+02	1.612e+01	3.694e−03	3.425e−04	2.290e+02	5.629e+00	1
201375598	11194625–0119486	169.942750	−1.330181	4.559	0.941	0.731	0.016	3.717e+05	3.404e+05	2.476e+03	5.092e+02	2.255e−01	5.046e−02	2.637e+03	4.187e+02	7.317e−02	1.506e−02	8.340e+04	1.384e+03	1
201379262	11234562–0116277	170.940087	−1.274369	40.242	5.022	4.699	0.051	6.862e+03	5.470e+03	6.450e+02	5.438e+01	3.529e−02	5.425e−03	8.063e+02	4.876e+01	7.269e−03	8.057e−04	4.611e+03	9.516e+01	1
201379481	...	168.657175	−1.270828	37.570	0.758	...	...	8.026e+03	1.018e+03	1.091e+03	4.372e+01	3.813e−02	2.345e−03	5.955e+02	3.462e+01	5.662e−03	2.493e−04	1.954e+01	3.093e+00	2
201389394	11082147–0107156	167.089500	−1.121086	137.060	1.990	13.075	0.049	4.438e+02	5.241e+01	3.329e+02	1.265e+01	1.006e−02	5.729e−04	2.704e+02	1.439e+01	2.838e−03	2.119e−04	1.439e+02	4.161e+00	2
201400095	11210314–0057399	170.263029	−0.961072	50.276	4.850	...	...	6.613e+03	5.340e+03	6.171e+02	4.821e+01	2.535e−02	2.918e−03	7.880e+02	4.464e+01	6.060e−03	5.457e−04	3.689e+03	7.561e+01	1
201411377	11232145–0047048	170.839412	−0.784719	38.573	3.427	3.782	0.032	5.361e+03	3.500e+03	7.352e+02	5.184e+01	3.525e−02	3.285e−03	7.006e+02	4.463e+01	7.732e−03	6.768e−04	3.371e+03	6.706e+01	1
201414784	11252575–0044016	171.357304	−0.733783	4.869	0.363	...	...	7.325e+05	3.186e+05	2.949e+03	5.548e+02	2.350e−01	2.549e−02	3.273e+03	2.135e+02	6.460e−02	4.725e−03	2.207e+04	4.225e+02	1
201415775	11333341–0043060	173.389225	−0.718339	27.298	0.294	...	...	1.405e+04	2.691e+03	9.926e+02	7.152e+01	4.837e−02	2.932e−03	5.915e+02	3.877e+01	9.928e−03	4.971e−04	3.234e+01	2.713e+00	2
201420175	11334091–0039085	173.420575	−0.652361	33.244	0.805	...	...	6.643e+03	1.246e+03	6.136e+02	4.458e+01	3.751e−02	3.055e−03	7.234e+02	2.309e+01	7.648e−03	2.576e−04	1.427e+02	3.323e+00	2
201438887	11361104–0022409	174.046021	−0.378047	30.272	2.411	...	...	1.274e+04	9.403e+03	9.159e+02	7.223e+01	4.535e−02	3.834e−03	8.394e+02	5.464e+01	9.200e−03	8.658e−04	5.250e+03	1.055e+02	1
201454791	11453665–0009037	176.402708	−0.151056	130.534	1.069	...	...	4.305e+02	3.249e+01	3.025e+02	1.218e+01	9.235e−03	5.605e−04	2.150e+02	1.463e+01	2.935e−03	2.704e−04	3.790e+01	1.434e+00	2
201459357	11255698–0005024	171.487429	−0.084022	3.972	...	...	...	2.560e+05	2.209e+05	2.544e+03	4.206e+02	2.651e−01	3.377e−02	2.398e+03	1.666e+02	8.400e−02	8.039e−03	1.014e+04	1.748e+02	1
201467358	11422234+0002007	175.593208	0.033297	113.094	0.658	9.942	0.032	2.495e+02	2.104e+01	2.694e+02	8.544e+00	1.167e−02	4.811e−04	1.895e+02	7.096e+00	3.102e−03	1.254e−04	8.390e+00	4.052e−01	2
201472519	11441126+0006347	176.046942	0.109667	29.995	0.735	...	...	1.290e+04	2.561e+03	9.782e+02	7.404e+01	3.403e−02	2.911e−03	9.752e+02	5.070e+01	1.020e−02	4.305e−04	3.790e+01	1.323e+00	2
201503634	11302841+0034387	172.618392	0.577436	5.408	0.378	...	...	3.108e+05	1.452e+05	2.163e+03	3.068e+02	2.004e−01	2.316e−02	2.238e+03	1.779e+02	6.368e−02	6.244e−03	1.052e+04	1.785e+02	1
201508025	11110893+0038392	167.787258	0.644258	88.027	7.028	...	...	7.613e+02	5.025e+02	3.710e+02	2.034e+01	1.372e−02	1.205e−03	3.613e+02	1.763e+01	4.171e−03	3.231e−04	2.669e+02	6.714e+00	1
201508509	11260798+0039130	171.533283	0.653625	34.168	4.443	2.927	0.029	9.747e+03	9.512e+03	7.774e+02	5.589e+01	3.514e−02	3.750e−03	7.343e+02	4.661e+01	8.281e−03	8.927e−04	3.624e+03	5.932e+01	1
201512825	11432842+0043107	175.868425	0.719658	79.011	0.804	...	...	1.109e+03	9.890e+01	3.580e+02	1.692e+01	1.495e−02	1.033e−03	3.385e+02	1.569e+01	4.117e−03	1.831e−04	3.418e+01	1.180e+00	2
201515047	11504834+0045163	177.701492	0.754572	249.915	5.820	17.348	0.006	1.538e+02	1.885e+01	2.526e+02	8.249e+00	7.260e−03	4.611e−04	2.237e+02	8.993e+00	1.578e−03	1.442e−04	6.159e+01	4.456e−01	1
201525065	11450722+0054340	176.280079	0.909461	23.455	0.507	...	...	1.784e+04	2.212e+03	1.727e+03	4.306e+01	6.146e−02	2.692e−02	5.076e+02	5.357e+01	1.706e−02	1.278e−03	1.775e+01	2.482e+00	2
201526688	11461610+0056057	176.567104	0.934933	36.565	4.112	4.703	0.047	7.871e+03	6.151e+03	7.356e+02	6.506e+01	3.705e−02	4.139e−03							



**Table 3**  
(Continued)

EPIC	2MASS	R.A.	Decl.	$\nu_{\max}$	$\sigma_{\nu_{\max}}$	$\Delta\nu$	$\sigma_{\Delta\nu}$	$A_{\max}$ (ppm <sup>2</sup> )	$\sigma_{A_{\max}}$ (ppm <sup>2</sup> )	$\sigma_{\text{meso}}$	$\sigma_{\sigma_{\text{meso}}}$	$\tau_{\text{meso}}$	$\sigma_{\tau_{\text{meso}}}$	$\sigma_{\text{gran}}$	$\sigma_{\sigma_{\text{gran}}}$	$\tau_{\text{gran}}$	$\sigma_{\tau_{\text{gran}}}$	WN (ppm <sup>2</sup> )	$\sigma_{\text{WN}}$ (ppm <sup>2</sup> )	Giant? <sup>a</sup>
		(deg)	(deg)	( $\mu\text{Hz}$ )	( $\mu\text{Hz}$ )	( $\mu\text{Hz}$ )	( $\mu\text{Hz}$ )	(ppm <sup>2</sup> )	(ppm <sup>2</sup> )	(ppm)	(ppm)	( $\mu\text{Hz}^{-1}$ )	( $\mu\text{Hz}^{-1}$ )	(ppm)	(ppm)	( $\mu\text{Hz}^{-1}$ )	( $\mu\text{Hz}^{-1}$ )	(ppm <sup>2</sup> )	(ppm <sup>2</sup> )	
201680913	12014373+0319161	180.432225	3.321144	55.107	2.315	4.207	0.013	2.271e+03	1.026e+03	3.333e+02	2.600e+01	2.490e-02	3.454e-03	4.304e+02	2.612e+01	5.496e-03	4.999e-04	1.001e+03	2.259e+01	1
201696051	12044299+0334230	181.179146	3.573111	88.357	3.333	...	...	8.036e+02	1.932e+02	4.894e+02	2.670e+01	1.251e-02	8.702e-04	4.415e+02	2.593e+01	4.323e-03	2.850e-04	1.950e+02	5.450e+00	2
201697539	11482814+0335534	177.117221	3.598100	15.789	0.769	...	...	1.080e+04	5.597e+03	1.201e+03	1.332e+02	8.025e-02	6.970e-03	6.971e+02	6.035e+01	1.973e-02	1.773e-03	2.106e+03	3.720e+01	1
201700607	11282997+0338594	172.124917	3.649897	81.719	5.318	7.558	0.052	9.614e+02	4.903e+02	3.906e+02	2.278e+01	1.605e-02	1.690e-03	3.830e+02	2.185e+01	4.127e-03	3.552e-04	5.028e+02	1.114e+01	1
201702907	...	178.026996	3.688792	74.988	7.398	...	...	1.618e+03	1.135e+03	4.119e+02	3.300e+01	1.480e-02	1.662e-03	5.439e+02	2.490e+01	4.621e-03	2.693e-04	5.542e+02	1.268e+01	1
201704568	11261591+0342583	171.566321	3.716222	97.876	2.453	...	...	6.111e+02	1.180e+02	2.721e+02	1.385e+01	1.457e-02	1.240e-03	2.750e+02	1.497e+01	3.569e-03	2.790e-04	2.393e+02	5.845e+00	1
201713224	11264919+0351517	171.704958	3.864342	31.338	0.406	3.038	0.002	5.379e+03	9.801e+02	4.816e+02	3.372e+01	3.413e-02	3.187e-03	5.307e+02	2.088e+01	9.077e-03	3.144e-04	1.661e+01	4.746e-01	2
201720476	11143920+0359154	168.663354	3.987589	83.725	7.192	6.103	0.722	1.801e+03	1.140e+03	5.019e+02	4.172e+01	1.395e-02	1.486e-03	5.545e+02	3.616e+01	4.061e-03	3.630e-04	1.023e+03	2.439e+01	1
201722766	11240739+0401380	171.030762	4.027258	88.538	4.371	...	...	6.930e+02	2.708e+02	1.907e+01	1.441e-02	1.370e-03	3.437e+02	1.855e+01	3.847e-03	3.201e-04	3.768e+02	9.177e+00	1	
201723568	11360352+0402289	174.014671	4.041367	5.372	0.363	0.743	0.024	1.565e+06	9.312e+05	2.169e+03	3.964e+02	1.931e-01	2.499e-02	2.648e+03	3.212e+02	6.169e-02	6.429e-03	1.907e+05	2.827e+03	1
201724514	11582220+0403262	179.592525	4.057289	29.111	0.611	...	...	1.355e+04	2.703e+03	1.070e+03	7.104e+01	3.660e-02	2.477e-03	9.358e+02	4.733e+01	9.489e-03	4.001e-04	4.201e+01	1.246e+00	2
201724852	11122791+0403471	168.116350	4.063117	94.344	6.219	10.897	0.046	5.750e+02	2.997e+02	3.066e+02	1.903e+01	1.342e-02	1.369e-03	3.039e+02	1.823e+01	3.961e-03	3.720e-04	3.552e+02	8.529e+00	1
201726163	11150267+0405078	168.761167	4.085508	93.244	2.575	8.602	0.038	1.338e+03	3.435e+02	3.942e+02	3.385e+01	1.153e-02	1.159e-03	5.179e+02	2.848e+01	4.104e-03	2.688e-04	5.624e+02	1.442e+01	1
201729267	11230257+0408173	170.760737	4.138147	92.008	2.000	...	...	5.634e+02	1.083e+02	2.784e+02	1.703e+01	1.185e-02	8.909e-04	3.272e+02	1.350e+01	3.366e-03	1.919e-04	1.166e+02	3.677e+00	1
201733406	11213386+0412299	170.391112	4.208306	90.679	1.445	9.181	1.372	5.055e+02	7.600e+01	2.350e+02	1.280e+01	1.409e-02	1.177e-03	2.298e+02	1.339e+01	3.770e-03	3.028e-04	1.147e+02	3.052e+00	2
201743103	11170064+0421565	169.252658	4.365717	84.644	5.676	...	...	1.348e+03	6.447e+02	4.056e+02	2.541e+01	1.562e-02	1.488e-03	4.229e+02	2.519e+01	4.144e-03	3.356e-04	6.171e+02	1.551e+01	1
201747404	11390558+0426188	174.773221	4.438600	156.608	1.034	12.324	0.674	3.123e+02	2.315e+01	2.616e+02	8.140e+00	9.993e-03	6.149e-04	2.087e+02	8.171e+00	2.669e-03	1.708e-04	3.179e+01	1.310e+00	2
201749662	11153895+0428400	168.912304	4.477769	90.575	1.455	...	...	4.238e+02	5.398e+01	2.296e+02	1.100e+01	1.732e-02	1.492e-03	1.801e+02	1.138e+01	3.903e-03	3.090e-04	9.589e+01	2.272e+00	2
201750985	11292465+0429584	172.352729	4.499581	27.465	0.976	...	...	1.390e+04	4.681e+03	9.214e+02	7.938e+01	4.705e-02	4.105e-03	7.151e+02	6.105e+01	1.146e-02	1.011e-03	5.044e+03	9.507e+01	1
201751998	11160874+0431029	169.036404	4.517472	54.093	1.772	...	...	2.776e+03	1.131e+03	3.731e+02	2.817e+01	2.365e-02	2.788e-03	5.354e+02	2.422e+01	5.059e-03	3.211e-04	8.587e+02	1.943e+01	1
201752633	11121723+0431442	168.071796	4.528894	106.785	6.842	...	...	3.175e+02	1.511e+02	3.065e+02	1.540e+01	1.126e-02	9.700e-04	3.141e+02	1.355e+01	3.527e-03	2.658e-04	1.077e+02	3.486e+00	1
201758449	11175773+0437487	169.490575	4.630186	99.409	5.084	...	...	6.210e+02	2.615e+02	3.469e+02	1.888e+01	1.279e-02	1.011e-03	3.648e+02	1.952e+01	3.452e-03	3.073e-04	3.053e+02	8.457e+00	1
201761560	11152022+0441023	168.834287	4.683992	88.898	3.367	...	...	1.347e+03	5.248e+02	4.625e+02	2.959e+01	1.286e-02	1.096e-03	4.554e+02	2.774e+01	4.162e-03	3.229e-04	7.687e+02	6.284e+00	1
201763504	11483335+0443022	177.138908	4.717314	8.609	0.139	1.177	0.158	8.146e+05	1.573e+05	5.359e+03	3.332e+02	1.075e-01	6.684e-03	1.519e+03	2.472e+02	4.300e-02	3.608e-03	3.880e-01	9.678e-01	2
201765667	11145908+0445177	168.746204	4.754925	85.684	4.221	7.375	0.050	1.266e+03	4.148e+02	3.232e+02	2.353e+01	1.742e-02	1.930e-03	4.094e+02	2.394e+01	4.006e-03	3.893e-04	6.068e+02	1.375e+01	1
201766812	11150311+0446301	168.762996	4.775025	100.462	4.872	...	...	1.175e+03	4.130e+02	4.083e+02	2.828e+01	1.300e-02	1.230e-03	4.734e+02	2.518e+01	3.561e-03	3.001e-04	5.436e+02	1.521e+01	1
201772439	1112479+0452234	167.853371	4.873231	98.648	2.035	...	...	1.127e+03	2.578e+02	4.209e+02	2.540e+01	1.434e-02	1.340e-03	4.923e+02	1.984e+01	3.732e-03	2.733e-04	3.643e+02	1.059e+01	1
201774359	11162408+0454236	169.100333	4.906556	72.644	5.287	...	...	1.120e+03	5.598e+02	3.210e+02	1.958e+01	1.865e-02	1.844e-03	3.651e+02	1.782e+01	4.662e-03	3.498e-04	4.071e+02	9.163e+00	1
201774883	11163452+0454529	169.143833	4.914697	83.505	5.950	...	...	1.197e+03	6.494e+02	3.223e+02	2.548e+01	1.453e-02	1.803e-03	3.942e+02	2.029e+01	4.235e-03	3.786e-04	6.295e+02	1.350e+01	1
201781960	11103146+0502268	167.631062	5.040825	81.685	6.281	9.842	0.731	1.279e+03	8.737e+02	4.916e+02	3.903e+01	1.377e-02	1.557e-03	4.964e+02	3.245e+01	4.500e-03	3.862e-04	8.033e+02	1.852e+01	1
201786083	11125055+0506500	168.210683	5.113931	80.724	5.662	5.613	0.018	1.103e+03	6.570e+02	3.848e+02	2.605e+01	1.420e-02	1.407e-03	4.465e+02	2.663e+01	4.407e-03	3.542e-04	7.038e+02	1.520e+01	1
201788284	11101971+0509085	167.582175	5.152367	90.274	4.934	...	...	9.987e+02	4.206e+02	2.776e+02	2.451e+01	1.322e-02	1.486e-03	3.942e+02	2.029e+01	3.674e-03	2.790e-04	5.087e+02	4.786e+00	1
201797512	11235883+0519178	170.995146	5.321614	113.952	0.674	...	...	7.773e+02	7.031e+01	3.354e+02	1.110e+01	1.249e-02	6.951e-04	2.692e+02	1.141e+01	3.120e-03	1.892e-04	5.433e+01	1.841e+00	2
201797810	11110894+0519382	167.787300	5.327325	82.795	4.936	8.705	0.027	1.988e+03	8.873e+02	3.949e+02	3.293e+01	1.766e-02	2.020e-03	5.200e+02	3.276e+01	3.878e-03	3.651e-04	9.518e+02	2.400e+01	1
201825690	...	167.827858	5.836272	61.013	2.837	...	...	3.625e+03	1.530e+03	5.475e+02	3.983e+01	2.117e-02	2.681e-03	6.541e+02	4.011e+01	5.117e-03	4.068e-04	1.739e+03	4.163e+01	1
201830769	11134232+0555598	168.426371	5.933281	83.880	4.531	6.318	0.702	1.087e+03	4.195e+02	4.978e+02	2.741e+01	1.443e-02	1.171e-03	4.417e+02	2.863e+01	4.251e-03	3.460e-04	6.077e+02	1.508e+01	1
201839927	...	174.142337	6.106439	27.190	0.344	...	...	2.258e+04	4.173e+03	1.576e+03	2.150e+02	5.011e-02	4.196e-03	8.107e+02	6.529e+01	9.456e-03	8.908e-04	1.334e+02	1.323e+01	2
201843056	11392070+0609576	174.836300	6.166011	93.862	2.045	6.147	0.013	4.633e+02	9.604e+01	6.335e+02	1.424e+01	1.622e-02	1.213e-03	2.968e+02	1.352e+01	3.880e-03	2.572e-04	1.565e+02	4.012e+00	1
201843394	11105997+0610194	167.749917	6.172050	84.029	5.168	5.683	0.012	9.430e+02	4.859e+02	4.234e+02	2.726e+01	1.553e-02	1.771e-03	5.008e+02	2.251e+01	4.261e-03	3.092e-04	4.734e+02	1.274e+01	1
201843809	11113052+0610473	167.877196	6.179844	93.392	3.923	...	...	8.797e+02	3.696e+02	4.266e+02	2.785e+01	1.140e-02	1.027e-03	4.573e+02	2.556e+01	3.769e-03	2.820e-04	3.918e+02	1.103e+01	2
201846331	11535849+0613495	178.493754	6.230472	74.569	4.161	5.347	0.005	7.916e+02	3.287e+02	4.056e+02	1.744e+01	1.933e-02	1.399e-03	3.334e+02	1.620e+01	4.535e-03	2.940e-04	1.820e+02	4.964e+00	1
201852681	11423531+0621280	175.647167	6.357800	34.347	0.918	3.053	0.016	6.068e+03	1.227e+03	6.915e+02	4.496e+01	3.609e-02	2.687e-03	7.504e+02	2.751e+01	8.389e-03	2.798e-04	9.884e+01	2.325e+00	2
201877455	11554329+0651459	178.930404																		

asymptotically normal:

$$\ln B \approx \Delta_{\text{WBIC}} \equiv \langle \ln \mathcal{L}_1 \rangle_{P(\theta|D)} - \langle \ln \mathcal{L}_2 \rangle_{P(\theta|D)}, \quad (13)$$

where  $\langle \cdot \rangle_{P(\theta|D)}$  indicates a mean taken over the modified posteriors of Equations (14) and (15) (see below), and the likelihoods are from

$$\text{Equations (8) and (4)} \quad \left( \mathcal{L}_1 \equiv \prod_j \left[ \frac{1}{A_{\text{tot}}(\nu_j)} \exp\left(-\frac{A_o(\nu_j)}{A_{\text{tot}}(\nu_j)}\right) \right] \text{ and } \mathcal{L}_2 \equiv \prod_j \left[ \frac{1}{A(\nu_j)} \exp\left(-\frac{A_o(\nu_j)}{A(\nu_j)}\right) \right] \right).$$

Crucially, the WBIC approach means that the Bayes factor can be computed trivially in an MCMC setting. We compute the means  $\langle \ln \mathcal{L}_1 \rangle_{P(\theta|D)}$  and  $\langle \ln \mathcal{L}_2 \rangle_{P(\theta|D)}$  using our two-step MCMC method, recalling that we perform fits to the data both with and without a power excess term (Equations (5) and (2)). For the purposes of approximating the Bayes factor, then, we run each MCMC an additional time, except using modified conditional posteriors so that instead of Equations (8) and (4), we have

$$P(\theta_{\text{meso}}, \theta_{\text{gran}}, \theta_{\text{excess}}|D) \propto P(\theta_{\text{meso}}, \theta_{\text{gran}}, \theta_{\text{excess}}) \prod_j \left[ \frac{1}{A_{\text{tot}}(\nu_j)} \exp\left(-\frac{A_o(\nu_j)}{A_{\text{tot}}(\nu_j)}\right) \right]^\beta \quad (14)$$

and

$$P(\theta_{\text{meso}}, \theta_{\text{gran}}|D = \{(\nu_j, A_o(\nu_j)), j = 0, 1, 2, \dots\}, \theta_{\text{excess}}) \propto P(\theta_{\text{meso}}, \theta_{\text{gran}}|\theta_{\text{excess}}) \prod_j \left[ \frac{1}{A(\nu_j)} \exp\left(-\frac{A_o(\nu_j)}{A(\nu_j)}\right) \right]^\beta, \quad (15)$$

where  $\beta \equiv 1/\ln N$ , with  $N$  being the number of points in the power spectrum being fit. While performing an MCMC fit using posteriors from Equations (14) and (15) in place of Equations (8) and (4), we save the original likelihoods from Equations (8) and (4) at each link in our MCMC chains. In the end, we take an average of those likelihoods, insert into Equation (13), and in this way compute the Bayes factor.

We interpret the strength of evidence for the Gaussian excess model following Kass (1995), who recommend that  $\ln B > 1$  would indicate positive evidence for the Gaussian excess model. We also require that the granulation component be resolved by imposing that the white noise be lower than the granulation component power (i.e., that the white noise should not dominate the power spectrum). Note that these selection criteria do not include information about  $\Delta\nu$ : identifying excess power corresponding to  $\nu_{\text{max}}$  is easier than identifying  $\Delta\nu$ , especially in the presence of mixed modes exhibited in red clump stars. The sample of non-GAP red giants that we will discuss in Section 4 are these candidates that had evidence according to the Bayes factor of exhibiting solar-like oscillations ( $\ln B > 1$ ): 316 giant candidates are chosen in this way from the non-GAP sample of 13,016 objects.

For every star in this sample of oscillating red giant candidates, we confirmed BAM’s selections as bona fide giants or not by visual inspection of the power spectra. We categorized each of BAM’s giant candidates into one of three categories: as having (1) a spectrum with oscillation modes that are discernible individually by eye or with excess power that is conspicuous by eye (“yes” oscillator), (2) a spectrum with

marginal evidence of excess power at a frequency consistent with the shape of the granulation and mesogranulation components (“maybe” oscillator), or (3) a spectrum that shows at best very weak evidence of excess power or whose model power spectrum is in clear disagreement with the observed one (“no” oscillator). The  $\nu_{\text{max}}$  inferred by eye in the “yes” and “maybe” cases must be within 3–283  $\mu\text{Hz}$ , such that giants that show evidence of a granulation spectrum at low frequencies are not selected as oscillators if the power excess is not visible above 3  $\mu\text{Hz}$ . In this discernment process, the amplitude of the power spectrum, which has a relation to  $\nu_{\text{max}}$  (as formalized, e.g., in Kallinger et al. 2014 and in Table 1), is allowed to be 10–50 times smaller than might be expected of a giant, to allow for cases where light from a nonoscillator contaminates the light curve, hence reducing the fractional brightness variation from granulation and oscillations. This effect can be significant. For instance, if a foreground dwarf of the same brightness as a background giant falls on the giant’s aperture mask, it would dilute the signal of the giant’s power spectrum by a factor of four.

Upon this visual verification, 31 of BAM’s non-GAP giant candidates were certain oscillators; 73, possible oscillators; and 212, not oscillating giants.

## 4. Results and Discussion

We apply the BAM pipeline to 13,016 C1 targets with *VJ* light curves not in the GAP sample, which have been selected for a wide range of science programs—mostly detection of planets around dwarfs. We identify 31 red giants that have detectable oscillation excesses that satisfy the BAM selection criteria of Section 3.6 and that have been validated by individual inspection—21 of these are from GO proposal target lists that did not intentionally target giants. An additional 73 objects are potential giants, though they cannot be definitely confirmed as such; 70 of these “maybe” cases are from programs that did not intentionally target giants. Combined, these 104 red giants and red giant candidates represent an 8% increase in the number of giants identified from C1 compared to those from the GAP sample (Stello et al. 2017), which expressly targeted giants. The global oscillation parameters and granulation parameters for the red giants and red giant candidates are given in Table 3.

### 4.1. Completeness and Purity of Observed Non-GAP Giants

The magnitude distribution of the stars we find in this serendipitous sample, shown in Figure 7, demonstrates that BAM can recover red giant oscillations in *K2* down to  $Kp \sim 14$  ( $H \sim 12$ ). All the adopted magnitudes and colors we use in the following are taken from the Ecliptic Input Catalog (EPIC; Huber et al. 2016).<sup>10</sup> Note that even though the majority of the non-GAP C1 targets have  $Kp \gtrsim 15$  (dashed green), the non-GAP giant sample from this work mostly has  $Kp \lesssim 15$  (solid green). This is due to white noise dominating the spectra of giants at fainter magnitudes and is the reason why the number of GAP giants also drops beyond  $Kp \gtrsim 13$  (solid blue). We adopt a conservative  $Kp = 13$  as our fiducial completeness

<sup>10</sup> A few objects had photometry in the EPIC that did not correspond to the giant in question, and these mismatches were corrected by searching for the nearest, brighter neighbor in the EPIC. The EPIC IDs affected were 201269306, 201472519, and 201724514.

limit, which we test in the next section by comparing to a model of the C1 non-GAP oscillators.

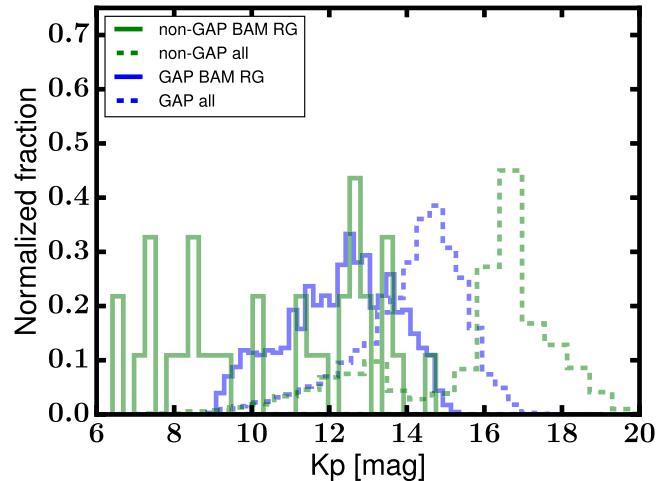
The purity of the non-GAP giant sample from BAM can be thought of as how many giants are verified visually as giants out of all the candidates that BAM believes are giants (i.e., 31 out of 316). Given that the majority of the non-GAP targets were selected by GO programs to be dwarfs, it is unsurprising that there are giant impostors that BAM mistakenly selected as giant candidates. Encouragingly, we find that BAM does not mistake the power in the frequency spectra from *K2*'s regular thruster firing for genuine oscillator excess. Instead, the objects mistakenly flagged as oscillators are due to one of a handful of failure modes. A full half of the false positives are objects exhibiting sharp, periodic signals overlaid on smooth, power-law spectra. Unlike genuine solar-like oscillators, however, objects falling into the latter failure mode generally exhibit multiple peaks (e.g., in Figure 8(a)). In future work, power spectra of periodic signals could be separated from those of giants by adding a second power excess component in Equation (5). If the best-fitting model preferred two regions of power excess instead of one, the spectrum would be rejected as a possible periodic case and not a giant. The other half of the false positives are either borderline “maybe”/“no” cases where the power excess is seemingly absent, but a granulation signal is present; cases in which BAM has converged on an incorrect  $\nu_{\max}$  (in which case, even if the giant is oscillating, it is assigned a “no” category); or dwarfs that have enough low-frequency activity to mimic a noisy giant granulation spectrum. Examples of these false positives are shown in Figures 8(a) and (b), in addition to an example of a potential giant oscillator (Figure 8(c)) and examples of bona fide oscillators (Figures 8(d)–(f)).

To get a better idea of the completeness of the sample, and to better understand the distributions of the observed properties of the non-GAP giant sample, we compare to a simulation that we describe in the next section.

#### 4.2. *Galaxia* Simulation of Non-GAP Giants

We model the non-GAP giant population using a *Galaxia* synthetic population of all stars in the field of Campaign 1 (see Sharma et al. 2011 for a description of *Galaxia* and Stello et al. 2017 for a comparison of this synthetic population to observed asteroseismic red giants from the GAP targets). Non-GAP *Galaxia* giants are defined to have  $3 \mu\text{Hz} < \nu_{\max} < 290 \mu\text{Hz}$ ,  $Kp < 13$ , and a probability of detection greater than 95% according to the same procedure used in Chaplin et al. (2011). However, here we assume  $\sqrt{A_{\max}} = 2.5(L/L_{\odot})^{0.9}(M/M_{\odot})^{-1.7}(T_{\text{eff}}/T_{\text{eff},\odot})^{-2.0}$  (Stello et al. 2011) and noise equal to that of *K2*. The use of a stellar population model of C1 like this is to make population-level statements about the concordance between the observed non-GAP giant population and a simulated one, and ideally to come to conclusions regarding the completeness and purity of the BAM non-GAP giant sample. In what follows, we will argue that there are likely inadequacies in both the recovered observed distribution due to selection effects, as well as inadequacies on the modeling side due to a difficult selection function and a probable metallicity offset in *Galaxia*'s underlying stellar models.

In order to make a fair comparison between the observed non-GAP targets and the non-GAP *Galaxia* stars, we resampled the *Galaxia* simulation such that it reproduced



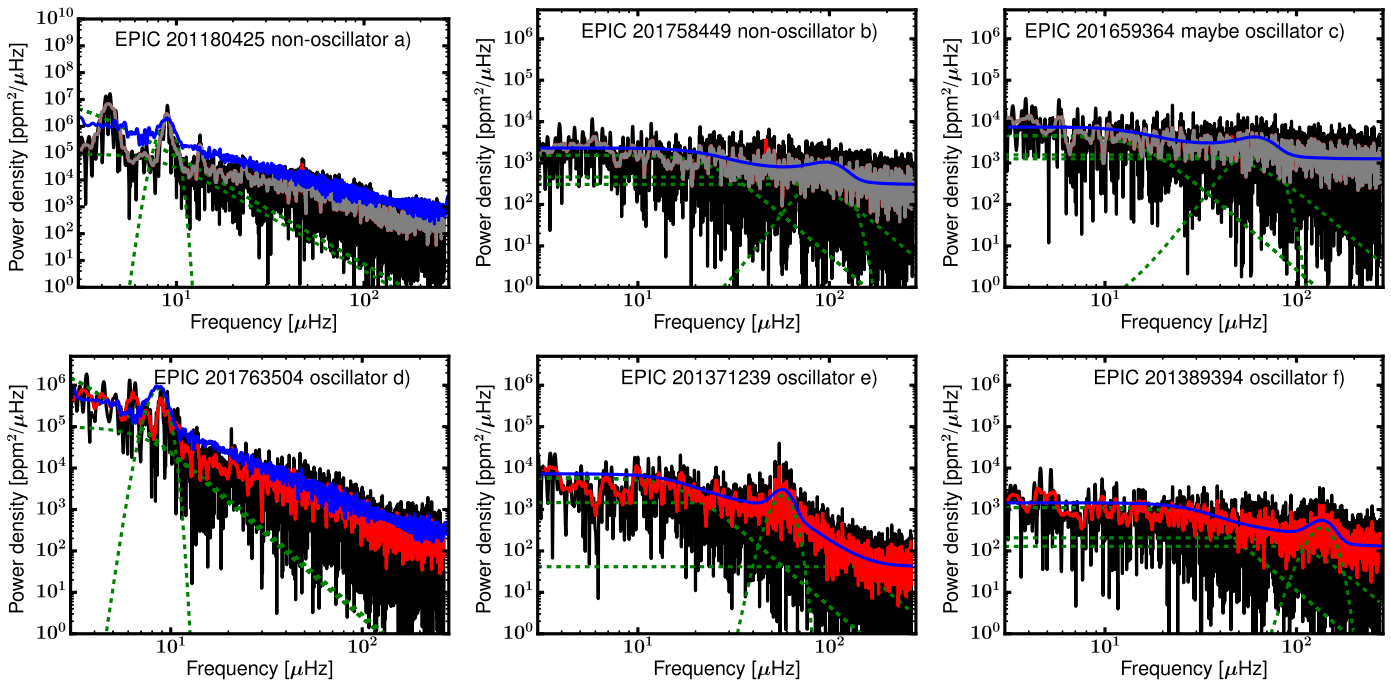
**Figure 7.** Magnitude distribution of the BAM non-GAP giant sample of this work (solid green line), compared to all observed non-GAP C1 targets (dashed green line), all GAP targets (dashed blue line), and GAP oscillators from Stello et al. (2017) (solid blue line).

the observed non-GAP distribution in  $(J - K_s, H)$  space. We first binned the observed non-GAP stars in  $(J - K_s, H)$  space and assigned each bin a probability of sample membership proportional to the number of stars in that bin. We then binned the *Galaxia* non-GAP stars using the same bins and resampled the stars by drawing a star one by one with a probability equal to the aforementioned sample membership probability of the bin in which it falls. The bins were chosen to optimize agreement with the simulated and observed distributions in  $(J - K_s, H)$  space and were approximately (0.05 mag, 1 mag) in width. The resampling stopped when the number of stars with  $Kp < 13$  equaled the number of stars in the observed non-GAP sample with  $Kp < 13$  (2080 stars in total).<sup>11</sup> This process results in some stars having the same properties because there are not enough unique *Galaxia* stars to match the number of observed stars. For this reason, we added a spread of 3% on the simulated giants'  $\nu_{\max}$ ,  $\Delta\nu$ , and 2% on photometry to avoid a sample with identical stars. The resampled *Galaxia* distribution is shown in the gray contours in Figure 9. The blue contours show the observed non-GAP population that we wanted to simulate, which shows that the simulation is consistent with the observations. The simulated giants within this sample, defined as mentioned above to have  $3 \mu\text{Hz} < \nu_{\max} < 290 \mu\text{Hz}$ ,  $Kp < 13$ , and a probability of detection greater than 95%, are shown by the gray circles.

#### 4.3. Comparison to *Galaxia*

With the *Galaxia* model for the non-GAP giants in hand, we can proceed to evaluate the agreement between simulation and observation, with implications for both the purity/completeness of the BAM sample and the fidelity of the *Galaxia* simulation in its description of the data. Figure 9 shows that the recovered giants (magenta and green circles) occupy two primary magnitude–color loci: (1) bright, red objects ( $H < 7$  and  $J - K_s > 0.5$ ), which were not targeted in GAP because of the the brightness cut in GAP of  $H > 7$ , and (2) giants at a typical magnitude, but bluer than typical giants

<sup>11</sup> A total of 12,839 out of the 13,016 non-GAP stars had valid  $Kp$  values in the EPIC, 11,579 of those had valid  $J - K_s$  colors, and 2080 of those also had  $Kp < 13$ .



**Figure 8.** Examples of the raw (black) and smoothed (red) power spectra of giant candidates selected by BAM, by requiring that the WBIC favor Equation (5) over Equation (2) (see Section 3.6). Each component of the models is shown in green dashed curves (white noise, Gaussian excess, and Harvey components), with the total model in blue. The top row shows BAM giant candidates determined to be false positives by visual inspection: EPIC 201180425 (panel (a)) shows a periodic signal, an alias of which BAM has mistaken for solar-like oscillations; EPIC 201758449 (panel (b)) shows a dwarf-like power spectrum that is at best a borderline no/maybe case—BAM has converged on a suboptimal model in this case, in addition; and EPIC 201659364 (panel (c)) shows what may be a giant spectrum with no discernible oscillation modes. In all panels in this row, shown in gray is a smoothed VJ spectrum when the thruster firing has been removed according to the procedure described in Section 2. The bottom row shows BAM giant candidates confirmed by visual inspection. The model of EPIC 201763504 (panel (d)) has been convolved with the spectral window, which allows BAM to fit the correct  $\nu_{\max}$  at  $\sim 8 \mu\text{Hz}$  rather than the spectral noise at  $\sim 50 \mu\text{Hz}$  (see text).

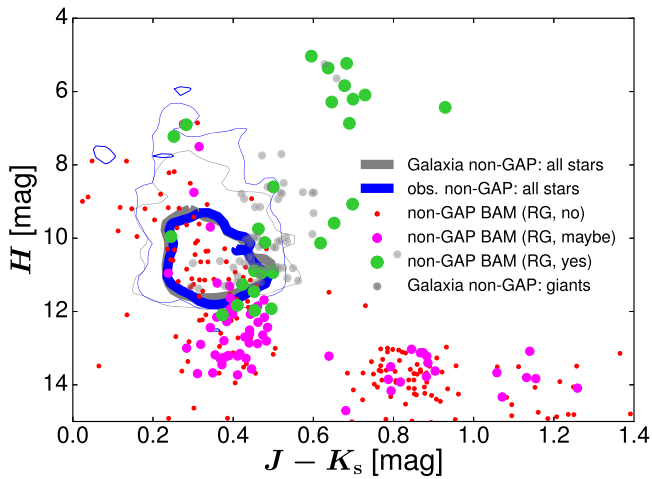
( $7 < H < 13$  and  $J - K_s < 0.5$ ), which were not in GAP because they have  $J - K_s < 0.5$ . First, let us consider the blue ( $J - K_s < 0.5$ ) giants, which are the more numerous population. That *Galaxia* predicts the presence of this population (gray circles) is the best indicator of agreement between our simulations and observations. Indeed, we expect that the blue population of non-GAP giants is a result of at least two factors: (1) the GAP  $J - K_s > 0.5$  selection is arbitrary and there are genuine oscillators with  $J - K_s < 0.5$ , and (2) due to photometric errors (taken to be  $\sim 0.02$  in the *Galaxia* C1 simulation), some oscillating giants with  $J - K_s > 0.5$  will be scattered to  $J - K_s < 0.5$ . The *Galaxia* simulation also successfully predicts that the bright ( $H < 7$ ) giants should exist. Note that our simulations only extend to our completeness cut of  $Kp = 13$ , and so we do not comment on *Galaxia* agreement in the regime of  $H > 12$ .

If the non-GAP sample were drawn from a similar distribution to our *Galaxia* simulation, we would expect the ratio of red ( $J - K_s > 0.5$ ) to blue ( $J - K_s < 0.5$ ) giants in *Galaxia* to agree with that of recovered BAM giants. We take the ratio of the observed number of published “yes” and “maybe” oscillators from K2GAP DR1 (Stello et al. 2017; with  $Kp < 13$  and  $J - K_s > 0.5$  cuts applied) to those with  $J - K_s < 0.5$  from the new, non-GAP giant sample presented here and compare it to the expected ratio from *Galaxia*. For this test, the  $(J - K_s, H)$  distribution of the GAP population was simulated in *Galaxia* following the sample membership probability procedure described above, only using the GAP targets instead of the non-GAP targets. Giants were then chosen to have  $3 \mu\text{Hz} < \nu_{\max} < 290 \mu\text{Hz}$ , a probability of

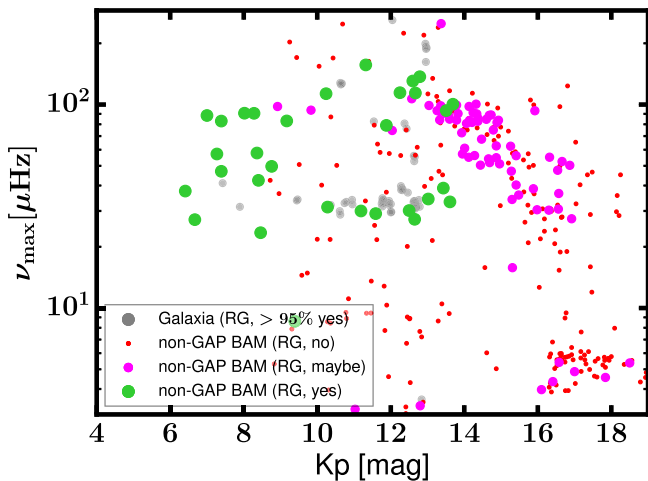
detection greater than 95%, and  $Kp < 13$ . The resulting ratio for *Galaxia* of  $13 \pm 2$  is significantly less than the same ratio for the BAM distribution of “yes” and “maybe” GAP giants of  $38 \pm 9.0$ , accounting for Poisson errors. Either the number of GAP giants is at odds with predictions, the number of non-GAP giants is, or both. Looking at the absolute numbers of giants in this ratio, 651/17 for observed BAM giants and 821/64 for *Galaxia*, the GAP giants agree better in number with what is expected from *Galaxia* than do the non-GAP giants. The 70% deficit in observed giants compared to *Galaxia* for the blue, non-GAP giants indicates that *Galaxia* predicts too many blue giants and/or BAM recovers too few blue giants. We consider both effects, in turn.

One of the primary effects that might result in an overprediction in our *Galaxia* model’s number of non-GAP giants is an incorrect selection function. The *Galaxia* non-GAP sample as we have constructed it only reproduces the color–magnitude distribution of the many GO proposal targets that compose the non-GAP sample. We expect this approach to globally describe the complex selection function of the sample, given that the GO proposals select objects based on color and magnitude cuts. Indeed, the non-GAP sample does describe well the observed sample (Figure 9). However, the majority of the GO proposals that compose the non-GAP sample also use proper-motion or reduced proper-motion cuts to choose dwarfs. Although these cuts will be functions of color and magnitude, we cannot precisely reproduce them in color and magnitude space. Therefore, we tested how many *Galaxia* non-GAP giants remained after applying a rather conservative (i.e., preserving more giants than dwarfs) reduced proper-motion cut





**Figure 9.** Color-magnitude diagram for Galaxia stars not passing GAP selection criteria (gray contours); Galaxia giants not passing GAP selection criteria, with  $>95\%$  probability of detection (gray circles); and observed non-GAP stars (blue contours). Contours enclose 68% (thick lines) and 95% (thin lines) of stars in the plotted region. Contours have been smoothed for illustrative purposes. Overlaid are stars from the non-GAP C1 target sample returned by BAM that visual inspection classified as definitely oscillators (green circles; 31 stars), maybe oscillators (magenta circles; 73 stars), and not oscillators (red circles; 212 stars).



**Figure 10.**  $\nu_{\max}$ - $Kp$  distribution of Galaxia-predicted detections of non-GAP oscillating giants (gray). Overlaid are stars from the non-GAP C1 target sample returned by BAM that visual inspection classified as definitely oscillators (green circles; 31 stars), maybe oscillators (magenta circles; 73 stars), and not oscillators (red circles; 212 stars).

of  $V + 5\log_{10} \mu > 20(V - J) - 25$ . (These cuts use the kinematic information that is stored as part of a Galaxia simulation). Only 11 non-GAP stars remained after this reduced proper-motion cut, which indicates that the GO reduced proper-motion cuts could explain the difference between the observed number of non-GAP giants (17) and that otherwise predicted by Galaxia (64). Another selection function could still be at work within the Galaxia model itself: an incorrect metallicity distribution of disk stars could result in too many blue giants, whose colors naturally depend on metallicity. A metallicity effect could also explain the offset in red clump position with respect to the observed red clump in K2 data, which is discussed in the next section.

With the reduced proper-motion cut’s role in mind, we still anticipate that some of the deficit in observed numbers of non-

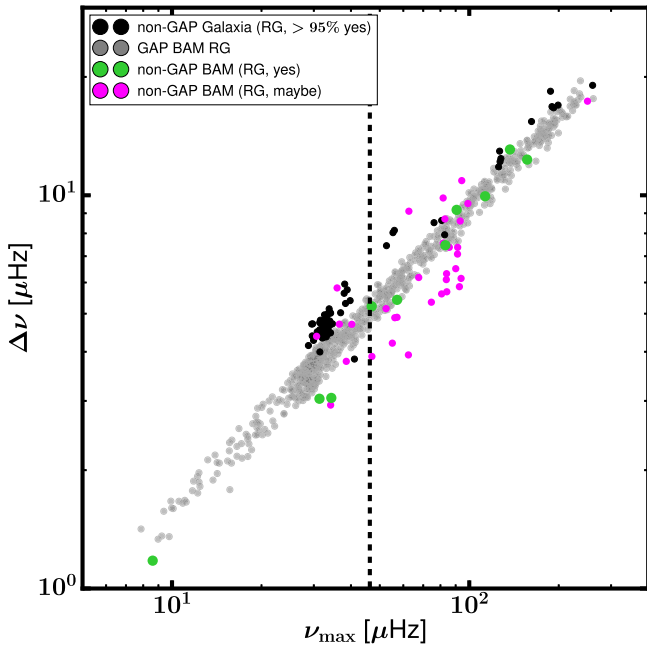
GAP giants is likely to reflect genuine incompleteness in the BAM giant sample. For example, in a handful of cases in the false-positive (“no”) sample, BAM performed a poor fit to the data, which will mean that its Bayesian model comparison will not be valid. Also, blended light from dwarfs would also strongly select against recovery with BAM because of a dilution of the oscillation signal resulting in significant departures from the amplitudes imposed by BAM’s priors in Table 1. We note also that asteroseismic giant detection with K2 will miss giants with  $\nu_{\max} \lesssim 3 \mu\text{Hz}$  and  $\nu_{\max} > 283 \mu\text{Hz}$ —the most evolved giants, and those closest to the base of the red giant branch. Establishing robust completeness and efficiency estimates is not the purpose of this paper, however, and we will explore these concerns more thoroughly in the next K2GAP data release (J. C. Zinn et al. 2019, in preparation).

We can also compare the Galaxia non-GAP red giant sample and the observed BAM non-GAP red giant sample in magnitude- $\nu_{\max}$  space, as shown in Figure 10. Kolmogorov-Smirnov tests indicate that both the  $\nu_{\max}$  distribution and  $Kp$  distribution for the definite BAM red giants are in  $\sim 3.2\sigma$  and  $\sim 4.0\sigma$  tension with the Galaxia  $\nu_{\max}$  and  $Kp$  distributions, assuming our adopted detection limit of  $Kp < 13$ . We note at this point that the procedure to match observed and Galaxia magnitude and color distributions (Section 4.2) is stochastic because the distributions are matched by drawing from probability distributions. This results in the Galaxia giants having  $\nu_{\max}$  and  $Kp$  distributions that vary in their agreement with the observed non-GAP giant distributions, fluctuating at the  $0.3\sigma$  and  $0.4\sigma$  level, respectively. Keeping this caveat in mind, there is still a tension in the simulated and observed  $\nu_{\max}$  distributions when marginalizing over realizations of the Galaxia  $\nu_{\max}$  distribution. That the tension in  $\nu_{\max}$  space decreases by  $\sim 1\sigma$  with a reduced proper-motion cut (see Section 4.2) indicates that this difference might be due to the unmodeled non-GAP selection function effects of individual GO proposals. There could also certainly be a  $\nu_{\max}$ -dependent efficiency in BAM identifying giants. Indeed, the latter effect is seen across various pipelines when comparing to a ground truth set of giants in K2 fields identified by eye, even while Galaxia giant predictions as a function of  $\nu_{\max}$  agree very well with the ground truth (K2GAP DR2; J. C. Zinn et al. 2019, in preparation).

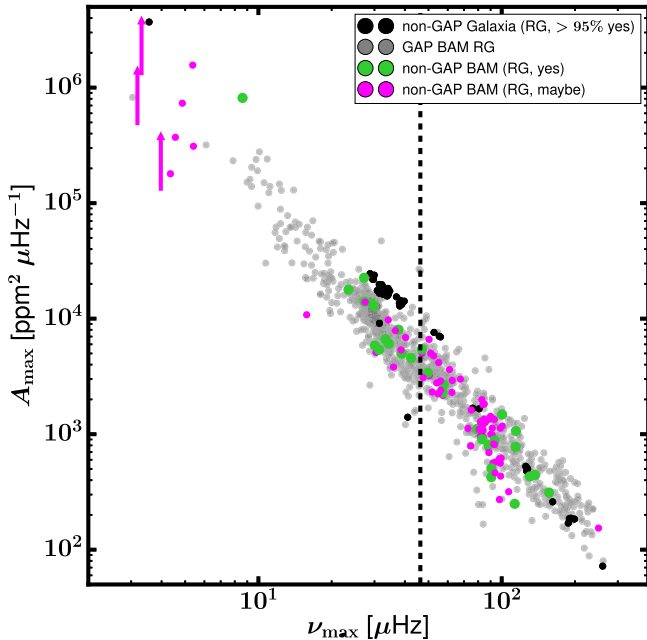
#### 4.4. Properties of Galaxia and Observed Non-GAP Giants

We show in Figures 11 and 12 the  $\Delta\nu$ - $\nu_{\max}$  and  $A_{\max}$ - $\nu_{\max}$  relations for this sample (colored points), as well as for the Galaxia model (black points). We have also included BAM GAP giants published in Stello et al. (2017), for reference (gray points). The agreement between model and observed properties in these spaces is good, except for the clump, for which Galaxia predicts a too-high  $\Delta\nu$  and  $A_{\max}$ . We can see that Galaxia overpredicts  $\Delta\nu$  and  $A_{\max}$  (and does not underpredict  $\nu_{\max}$ ) because the  $\nu_{\max}$  location of the overdensity in GAP BAM stars at  $\nu_{\max} \sim 30 \mu\text{Hz}$  agrees with the location of the overdensity in the non-GAP Galaxia stars. Figure 13 shows a modified Kiel diagram, in which  $J - K_s$  color is used instead of temperature and  $\nu_{\max}$  instead of gravity<sup>12</sup>. In this space, we can see that nearly all of the observed non-GAP sample is found at or below the clump (at  $\nu_{\max} \sim 30 \mu\text{Hz}$ ) and

<sup>12</sup> Note the reversed y-axis: a smaller  $\nu_{\max}$  means a smaller gravity and so is in the sense of a normal Kiel diagram.

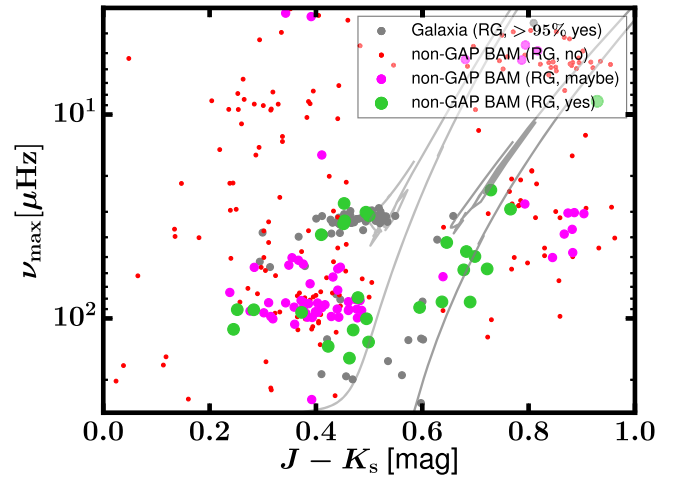


**Figure 11.** The  $\Delta\nu$ – $\nu_{\max}$  relation, with the non-GAP giant sample shown as colored circles as in Figure 10, comprising stars that have both  $\Delta\nu$  and  $\nu_{\max}$  measured by BAM. Gray points are BAM results from K2GAP DR1 (Stello et al. 2017), and black points are from our *Galaxia* simulation of the non-GAP giant sample. The dashed line corresponds to the nominal *K2* thruster firing frequency.



**Figure 12.**  $A_{\max}$ – $\nu_{\max}$  relation, with the non-GAP giant sample shown as colored circles as in Figure 10. Giants for which  $\nu_{\max} \lesssim 4 \mu\text{Hz}$  are considered upper limits. Gray points are BAM results from K2GAP DR1 (Stello et al. 2017), and black points are from our *Galaxia* simulation of the non-GAP giant sample. BAM K2GAP DR1 amplitudes were not published in Stello et al. (2017), though they are reproduced here. The dashed line corresponds to the nominal *K2* thruster firing frequency.

that the location of the *Galaxia* clump overlaps with several of the presumable observed red clump stars, confirming that the *Galaxia* clump  $\nu_{\max}$  locus is not discrepant with the observed locus. That the modeled clump  $\Delta\nu$  locus is offset from the



**Figure 13.** Modified Kiel diagram, with the non-GAP giant sample shown as colored circles as in Figure 10. The gray points are predictions from a simulation of the non-GAP stellar population in Campaign 1 using *Galaxia* (Sharma et al. 2011). See text for details. Evolutionary tracks for a  $1.3 M_{\odot}$  star with  $[\text{Fe}/\text{H}] = 0$  (dark gray) and  $[\text{Fe}/\text{H}] = -1$  (light gray) from MIST (Choi et al. 2016; Dotter 2016) are shown for visualization purposes.

observed clump  $\Delta\nu$  locus is another indication that the *Galaxia* models could be relying on a Galactic metallicity distribution at odds with the actual one—a conclusion that one arrives at when comparing *Galaxia* stellar parameters to those from asteroseismology in other *K2* campaigns (Sharma et al. 2019).

#### 4.5. Implications for Dwarf Selection Purity

A summary of the number of “yes” and “maybe” giants broken down by the GO target list from which they arise is shown in Table 4. Of the sample of non-GAP giants, 21 are serendipitous: they are only targets from GO proposals that do not intentionally select giants. This, in turn, allows us to say that the purity of giant exclusion across *K2* C1 GO proposals is  $\sim 99\%$ , based on the observed confirmed number of serendipitous giants found among the GO target lists that do not purport to select giants (those that intentionally target giants are not included in our calculation of dwarf purity and are noted in Table 4). The purity decreases a negligible amount if also including the BAM non-GAP “maybe” giants. This estimated dwarf selection purity is an upper bound because we have certainly not recovered all the giants owing to reasons discussed in Section 4.2. In this estimate, we have only counted targets that are within our completeness limit of  $Kp < 13$ . In this sense, we confirm that the *K2* dwarf samples chosen with color and proper-motion cuts are generally free from giants for  $Kp < 13$ .

## 5. Conclusion

In this paper, we have presented the BAM pipeline, which calculates global oscillation parameters in a Bayesian framework. A major advantage of the Bayesian fitting method we have employed is its natural basis for probabilistic selection of likely true oscillators among a collection of light curves. In the process of developing this pipeline and applying it to *K2* Campaign 1 (C1) stars, including both GAP (Stello et al. 2015, 2017) giant targets and non-GAP dwarf targets, we have found the following:

**Table 4**

The Number of Confirmed and Marginal Giants Discussed in This Paper Found in the Observed Targets of Various Guest Observer Proposals Gives an Indication of the Success at Rejecting Giants Using Color and Proper-motion Cuts

Guest Observer ID	Giant Fraction (yes)	Giant Fraction (maybe)	Notes
GO1001	0/3	0/3	
GO1002	1/30	0/30	
GO1003	0/2	0/2	Targeted extremely red stars, many likely to be AGB and long-period variables, which would not have been selected by BAM because their frequencies would be below our cutoff of 3 $\mu\text{Hz}$
GO1005	0/16	0/16	
GO1006	0/20	0/20	
GO1014	0/1	0/1	
GO1021	0/1	0/1	
GO1023	0/4	0/4	
GO1026	0/2	0/2	Targeted eclipsing binaries, some of which may be giants
GO1027	2/50	1/50	Targeted AF-type stars, the coolest of which might be oscillating giants
GO1029	0/1	0/1	
GO1030	0/1	0/1	
GO1036	0/38	0/38	
GO1038	0/12	1/12	Targeted potential oscillators
GO1040	5/6	0/6	Targeted bright giants
GO1043	0/25	0/25	
GO1046	0/3	0/3	Targeted bright stars, among them three subgiants, which likely will not oscillate below the long-cadence Nyquist frequency of $\sim 283 \mu\text{Hz}$
GO1052	0/1	0/1	
GO1053	0/1	0/1	
GO1054	9/2092	5/2092	
GO1055	0/39	0/39	
GO1057	0/1	0/1	Targeted giant oscillators, and this object was missed by BAM
GO1061	2/7	0/7	
GO1062	3/4	0/4	
GO1066	3/3	0/3	Targeted subgiants
GO1068	0/3	0/3	Targeted eclipsing binaries, some of which may be giants
GO1069	0/6	0/6	
GO1072	0/4	0/4	
GO1073	0/10	0/10	
GO1074	1/3	0/3	Targeted extragalactic objects

**Note.** Note that the tabulated numbers only include targets that had long-cadence data. Unless otherwise noted above, the GO proposals did not, to our knowledge, target giants. We have not listed GO1059, because that is the GAP.

1. We have identified an as-of-yet-unidentified noise pattern present in Vanderburg & Johnson (2014) light curves of C1 stars that causes a splitting of the nominal thruster firing frequency artifact at 47.22  $\mu\text{Hz}$  in a time-dependent manner.
2. We have additionally shown that it is necessary to account for the spectral window in fitting the spectra of solar-like oscillators in order to model the unphysical spectral leakage in the power spectrum of oscillators with  $\nu_{\text{max}} \lesssim 15 \mu\text{Hz}$ . In this work, we have done so by convolving models of the granulation with the observed window function.
3. We have benchmarked our asteroseismic parameters against the existing SYD asteroseismic pipeline and quantified statistical and systematic errors for BAM parameters accordingly. We find typical errors for K2 BAM giants in  $\nu_{\text{max}}$  and  $\Delta\nu$  of  $\sim 1.53\%$  (random)  $\pm 0.2\%$  (systematic) and  $1.51\%$  (random)  $\pm 0.6\%$  (systematic).
4. As an example application of BAM, we have also presented a sample of 104 non-GAP BAM red giants and

red giant candidates from C1 identified by their solar-like oscillations, 91 of which were not selected by GO proposals to be giants and hence are serendipitous discoveries.




5. The size of the non-GAP BAM red giant sample suggests that K2 C1 dwarf samples chosen with color and proper-motion cuts are generally free from giants for  $Kp < 13$  to a high degree (upper bound of  $\sim 99\%$  pure).
6. Simulated *Galaxia* C1 non-GAP giant populations are in tension with the  $Kp$  and  $\nu_{\text{max}}$  distributions of observed non-GAP giants with  $Kp < 13$  found by BAM. When considering also the higher-than-observed number of blue ( $J - K_s < 0.5$ ) giants in the *Galaxia* model, the disagreement between model and observation can be explained by the proper-motion cuts used to select the non-GAP targets. There is also likely incompleteness in the BAM giant detection process, which will be addressed in future work. Finally, the *Galaxia* metallicity distribution is likely different from the distribution of the non-GAP stars (Sharma et al. 2019).

BAM promises to robustly identify and characterize solar-like oscillators in *K2* and the *TESS* mission (Ricker et al. 2014), which is observing hundreds of thousands of red giants with at least 30-minute cadence. Though it will perform at least as well as *K2* in resolving oscillations on the lower giant branch, the majority of *TESS*'s red giant data will have roughly half the temporal baseline of *K2* and therefore will be a factor of two worse in spectral resolution. Spectral resolution is particularly important in identifying the low-frequency oscillators like those presented here. In this sense, BAM's Bayesian fitting techniques will take advantage of the information in ("global") features of the power spectrum that are less sensitive to degraded frequency resolution, in order to robustly identify  $\nu_{\max}$  for *TESS* giants.

We thank the referee for their comments, which improved both BAM and the contents of this paper. J.Z. acknowledges support from NASA grants 80NSSC18K0391 and NNX17AJ40G. D.H. acknowledges support by the National Science Foundation (AST-1717000) and the National Aeronautics and Space Administration under grants NNX14AB92G and NNX16AH45G issued through the *Kepler* Participating Scientist Program and the *K2* Guest Observer Program. D.S. is the recipient of an Australian Research Council Future Fellowship (project No. FT1400147). Parts of this research was conducted by the Australian Research Council Centre of Excellence for All Sky Astrophysics in 3 Dimensions (ASTRO 3D), through project No. CE170100013. This research was supported in part by the National Science Foundation under grant No. NSF PHY-1748958 and by the Heising-Simons Foundation.

This paper includes data collected by the *K2* mission. We thank all NASA employees that have contributed to the remarkable success of the *K2* mission. Funding for the *K2* mission is provided by the NASA Science Mission directorate. Some of the data presented in this paper were obtained from the Mikulski Archive for Space Telescopes (MAST).

### ORCID iDs

Joel C. Zinn  <https://orcid.org/0000-0002-7550-7151>  
 Dennis Stello  <https://orcid.org/0000-0002-4879-3519>  
 Daniel Huber  <https://orcid.org/0000-0001-8832-4488>  
 Sanjib Sharma  <https://orcid.org/0000-0002-0920-809X>

### References

- Aigrain, S., Parviainen, H., & Pope, B. J. S. 2016, *MNRAS*, **459**, 2408  
 Anderson, E. R., Duvall, T. L., Jr., & Jefferies, S. M. 1990, *ApJ*, **364**, 699  
 Angus, R., Foreman-Mackey, D., & Johnson, J. A. 2016, *ApJ*, **818**, 109  
 Appourchaux, T. 2003, *A&A*, **412**, 903  
 Armstrong, D. J., Kirk, J., Lam, K. W. F., et al. 2015, *A&A*, **579**, A19  
 Baglin, A., Michel, E., Auvergne, M., & COROT Team 2006, in ESA Special Publication 624, Proc. SOHO 18/GONG 2006/HELAS I, Beyond the Spherical Sun, ed. K. Fletcher (Paris: ESA), 34  
 Bedding, T. R., Huber, D., Stello, D., et al. 2010, *ApJL*, **713**, L176  
 Bedding, T. R., & Kjeldsen, H. 2010, *CoAst*, **161**, 3  
 Borucki, W., Koch, D., Basri, G., et al. 2008, in IAU Symp. 249, Exoplanets: Detection, Formation and Dynamics, ed. Y.-S. Sun, S. Ferraz-Mello, & J.-L. Zhou (Cambridge: Cambridge Univ. Press), 17  
 Brown, T. M., Gilliland, R. L., Noyes, R. W., & Ramsey, L. W. 1991, *ApJ*, **368**, 599  
 Chaplin, W. J., Kjeldsen, H., Bedding, T. R., et al. 2011, *ApJ*, **732**, 54  
 Chib, S., & Jeliazkov, I. 2001, *J. Am. Stat. Assoc.*, **96**, 270  
 Choi, J., Dotter, A., Conroy, C., et al. 2016, *ApJ*, **823**, 102  
 De Ridder, J., Barban, C., Baudin, F., et al. 2009, *Natur*, **459**, 398  
 Dotter, A. 2016, *ApJS*, **222**, 8  
 Foreman-Mackey, D., Hogg, D. W., Lang, D., & Goodman, J. 2013, *PASP*, **125**, 306  
 Green, P. J. 1995, *Biometrika*, **82**, 711  
 Hekker, S., Broomhall, A.-M., Chaplin, W. J., et al. 2010, *MNRAS*, **402**, 2049  
 Hekker, S., Kallinger, T., Baudin, F., et al. 2009, *A&A*, **506**, 465  
 Howell, S. B., Sobeck, C., Haas, M., et al. 2014, *PASP*, **126**, 398  
 Huber, D., Bedding, T. R., Stello, D., et al. 2010, *ApJ*, **723**, L607  
 Huber, D., Bryson, S. T., Haas, M. R., et al. 2016, *ApJS*, **224**, 2  
 Huber, D., Stello, D., Bedding, T. R., et al. 2009, *CoAst*, **160**, 74  
 Jeffreys, H. 1935, *PCPS*, **31**, 203  
 Kallinger, T., De Ridder, J., Hekker, S., et al. 2014, *A&A*, **570**, A41  
 Kallinger, T., Hekker, S., Garcia, R. A., Huber, D., & Matthews, J. M. 2016, *SciA*, **2**, 1500654  
 Kallinger, T., Mosser, B., Hekker, S., et al. 2010, *A&A*, **522**, A1  
 Kass, R. E., & Raftery, A. E. 1995, *J. Am. Stat. Assoc.*, **90**, 773  
 Kjeldsen, H., & Bedding, T. R. 1995, *A&A*, **293**, 87  
 Kjeldsen, H., & Bedding, T. R. 2011, *A&A*, **529**, L8  
 Luger, R., Agol, E., Kruse, E., et al. 2016, *AJ*, **152**, 100  
 Lund, M. N., Handberg, R., Davies, G. R., Chaplin, W. J., & Jones, C. D. 2015, *ApJ*, **806**, 30  
 Mathur, S., García, R. A., Régulo, C., et al. 2010, *A&A*, **511**, A46  
 Mosser, B., & Appourchaux, T. 2009, *A&A*, **508**, 877  
 Mosser, B., Belkacem, K., Goupil, M.-J., et al. 2010, *A&A*, **517**, A22  
 Murphy, S. J., Shibahashi, H., & Kurtz, D. W. 2013, *MNRAS*, **430**, 2986  
 Nordlund, Å., Stein, R. F., & Asplund, M. 2009, *LRSP*, **6**, 2  
 Pinsonneault, M. H., Elsworth, Y. P., Tayar, J., et al. 2018, *ApJS*, **239**, 32  
 Ricker, G. R., Winn, J. N., Vanderspek, R., et al. 2014, *Proc. SPIE*, **9143**, 914320  
 Scargle, J. D. 1982, *ApJ*, **263**, 835  
 Schwarz, G. 1978, *AnSta*, **6**, 461  
 Sharma, S., Bland-Hawthorn, J., Johnston, K. V., & Binney, J. 2011, *ApJ*, **730**, 3  
 Sharma, S., Stello, D., Bland-Hawthorn, J., et al. 2019, arXiv:1904.12444  
 Skilling, J. 2004, AIP Conference Proceedings, **735**, 395  
 Stello, D., Cantiello, M., Fuller, J., Garcia, R. A., & Huber, D. 2016, *PASA*, **33**, e011  
 Stello, D., Chaplin, W. J., Basu, S., Elsworth, Y., & Bedding, T. R. 2009, *MNRAS*, **400**, L80  
 Stello, D., Compton, D. L., Bedding, T. R., et al. 2014, *ApJL*, **788**, L10  
 Stello, D., Huber, D., Bedding, T. R., et al. 2013, *ApJL*, **765**, L41  
 Stello, D., Huber, D., Kallinger, T., et al. 2011, *ApJL*, **737**, L10  
 Stello, D., Huber, D., Sharma, S., et al. 2015, *ApJL*, **809**, L3  
 Stello, D., Zinn, J., Elsworth, Y., et al. 2017, *ApJ*, **835**, 83  
 Ulrich, R. K. 1986, *ApJL*, **306**, L37  
 Vanderburg, A., & Johnson, J. A. 2014, *PASP*, **126**, 948  
 Vanderplas, J., Connolly, A., Ivezić, Ž., & Gray, A. 2012, in Conf. on Intelligent Data Understanding (CIDU) (Piscataway, NJ: IEEE), 47  
 Watanabe, S. 2013, *Journal of Machine Learning Research*, **14**, 867  
 Woodard, M. F. 1984, PhD thesis, Univ. California, San Diego  
 Yu, J., Huber, D., Bedding, T. R., et al. 2018, *ApJS*, **236**, 42