

ALL-OR-NOTHING PHENOMENA: FROM SINGLE-LETTER TO HIGH DIMENSIONS

Galen Reeves* Jiaming Xu[†] Ilias Zadik[‡]

ABSTRACT

We consider the problem of estimating a p -dimensional vector β from n observations $Y = X\beta + W$, where $\beta_j \stackrel{\text{i.i.d.}}{\sim} \pi$ for a real-valued distribution π with zero mean and unit variance, $X_{ij} \stackrel{\text{i.i.d.}}{\sim} \mathcal{N}(0, 1)$, and $W_i \stackrel{\text{i.i.d.}}{\sim} \mathcal{N}(0, \sigma^2)$. In the asymptotic regime where $n/p \rightarrow \delta$ and $p/\sigma^2 \rightarrow \text{snr}$ for two fixed constants $\delta, \text{snr} \in (0, \infty)$ as $p \rightarrow \infty$, the limiting (normalized) minimum mean-squared error (MMSE) has been characterized by a single-letter (additive Gaussian scalar) channel.

In this paper, we show that if the MMSE function of the single-letter channel converges to a step function, then the limiting MMSE of estimating β converges to a step function which jumps from 1 to 0 at a critical threshold. Moreover, we establish that the limiting mean-squared error of the (MSE-optimal) approximate message passing algorithm also converges to a step function with a larger threshold, providing evidence for the presence of a computational-statistical gap between the two thresholds.

1. INTRODUCTION

Consider the classical linear regression model

$$Y = X\beta + W \quad (1)$$

where $X \in \mathbb{R}^{n \times p}$ with $X_{ij} \stackrel{\text{i.i.d.}}{\sim} \mathcal{N}(0, 1)$, $\beta \in \mathbb{R}^p$ with $\beta_j \stackrel{\text{i.i.d.}}{\sim} \pi$ for a distribution π with zero mean and unit variance, and $W \in \mathbb{R}^n$ with $W_i \stackrel{\text{i.i.d.}}{\sim} \mathcal{N}(0, \sigma^2)$. We are interested in estimating β from observation of (X, Y) . For a given estimator $\hat{\beta}(X, Y)$, the normalized mean squared-error of estimating β is given by

$$\text{MSE}(\hat{\beta}) := \frac{1}{p} \mathbb{E} \left[\left\| \beta - \hat{\beta} \right\|^2 \right].$$

Let MMSE denote the minimum of $\text{MSE}(\hat{\beta})$ among all possible estimators $\hat{\beta}$, or equivalently,

$$\text{MMSE} := \frac{1}{p} \mathbb{E} \left[\left\| \beta - \mathbb{E}[\beta | X, Y] \right\|^2 \right]. \quad (2)$$

*G. Reeves is with the Department of Electrical and Computer Engineering and the Department of Statistical Science, Duke University, Durham, NC 27708 USA; e-mail: galen.reeves@duke.edu.

[†]J. Xu is with the Fuqua School of Business, Duke University, Durham, NC 27708 USA; e-mail: jiamingxu.868@duke.edu.

[‡]I. Zadik is with the Operations Research Center, Massachusetts Institute of Technology, Cambridge, MA 02139 USA; e-mail: izadik@mit.edu.

In this paper, we focus on the asymptotic regime:

$$\frac{n}{p} \rightarrow \delta \quad \text{and} \quad \frac{p}{\sigma^2} \rightarrow \text{snr}, \quad \text{as } p \rightarrow \infty, \quad (3)$$

for two fixed constants $\delta, \text{snr} \in (0, \infty)$. Note that δ is the under-sampling ratio and snr is the signal-to-noise ratio in view of $\mathbb{E}[\|X\beta\|^2]/\mathbb{E}[\|W\|^2] = p/\sigma^2$.

Recent work [1–3] proves that under certain structural assumptions in terms of $(\pi, \delta, \text{snr})$, the limiting MMSE in the asymptotic regime (3) is characterized by the *replica-symmetric* (RS) formula through a single letter channel

$$y = \sqrt{s}\beta_0 + N, \quad (4)$$

where $s > 0$, $\beta_0 \sim \pi$ and $N \sim \mathcal{N}(0, 1)$ are independent. However, often the RS formula is too complicated to extract structural behavior of the limiting MMSE.

In this work, we discover that the limiting MMSE exhibits an all-or-nothing phenomena. More precisely, consider a family $(\pi_\epsilon, \delta_\epsilon, \text{snr}_\epsilon)$ indexed by a positive parameter ϵ where π_ϵ has finite entropy $H_\epsilon := H(\pi_\epsilon)$. We show that if the MMSE of the single letter channel (4) as a function of s converges to a step function as $\epsilon \rightarrow 0$, then the limiting MMSE of the linear regression model (1) also converges to a step function, which jumps from 1 to 0 at a critical threshold $\delta_\epsilon = \delta_{\epsilon, \text{MMSE}}$, where

$$\delta_{\epsilon, \text{MMSE}} := \frac{2H_\epsilon}{\log(1 + \text{snr}_\epsilon)}. \quad (5)$$

In other words, an all-or-nothing phenomena in the single letter channel implies an all-or-nothing phenomena in the high-dimensional linear regression model. Moreover, we establish that the limiting MSE of the (MSE-optimal) approximate message passing (AMP) algorithm also converges to a step function, which jumps from 1 to 0 at a larger threshold $\delta_\epsilon = \delta_{\epsilon, \text{AMP}}$, where

$$\delta_{\epsilon, \text{AMP}} := \frac{2H_\epsilon(1 + \text{snr}_\epsilon)}{\text{snr}_\epsilon}. \quad (6)$$

An important application of our general result is the binary linear regression model where $\beta_j \stackrel{\text{i.i.d.}}{\sim} \text{Bern}(\epsilon)$. In this case, we show that the MMSE function of the single letter channel converges to a step function as the sparsity $\epsilon \rightarrow 0$. Then we obtain from our general result that the limiting MMSE of the binary linear regression model converges to a step function which

jumps from 1 to 0 at the critical threshold $\delta_\epsilon = \frac{2\epsilon \log(1/\epsilon)}{\log(1+\text{snr}_\epsilon)}$. This coincides with the all-or-nothing phenomena established in [4] for the binary linear regression model where β is chosen uniformly at random from the set of binary k -sparse vectors, in the highly sparse and high signal-to-noise ratio regime where $k/\sqrt{p} \rightarrow 0$ and k/σ^2 is above a sufficiently large constant. Furthermore, we deduce from our general result that the limiting MSE of the (MSE-optimal) AMP converges to a step function which jumps from 1 to 0 at the critical threshold $\delta_\epsilon = \frac{2\epsilon \log(1/\epsilon)(1+\text{snr}_\epsilon)}{\text{snr}_\epsilon}$. This coincides with the computational threshold for a number of computationally efficient methods in the literature such as LASSO or Orthogonal Matching Pursuit. In particular, our result adds to the existing evidence for the presence of a computational-statistical gap (see [5, 6] for an extended discussion and literature review on the presence of this computational-statistical gap).

2. PRELIMINARIES

2.1. The Replica Symmetric Formulas

To describe the RS formulas, we first define the mutual information and MMSE functions for the single letter channel (4):

$$I(s) := I(\beta_0; \sqrt{s}\beta_0 + N), \quad s > 0 \quad (7)$$

$$M(s) := \text{mmse}(\beta_0 | \sqrt{s}\beta_0 + N), \quad s > 0 \quad (8)$$

where $\beta_0 \sim \pi$ and $N \sim \mathcal{N}(0, 1)$ are independent. Both of these functions are non-negative and the unit variance assumption on π means that for any $s > 0$, (see [7] for details)

$$I(s) \leq \frac{1}{2} \log(1 + s) \leq \frac{s}{2}, \quad (9)$$

$$M(s) \leq \frac{1}{1 + s} \leq 1. \quad (10)$$

Next, we define the potential function $\mathcal{F} : [0, \infty) \rightarrow [0, \infty)$ according to

$$\mathcal{F}(s) := I(s) + \frac{\delta}{2} \phi\left(\frac{s}{\delta \text{snr}}\right), \quad (11)$$

where $\phi(x) = x - \log x - 1$, and δ, snr are respectively the undersampling ratio and the signal-to-noise ratio of our original model. Note that $\phi(x)$ is convex and non-negative on $(0, \infty)$.

Lemma 1. *All stationary points of $\mathcal{F}(s)$ lie on the open interval between $\delta \text{snr} / (1 + \text{snr})$ and δsnr .*

Proof. By differentiation with respect to s and the I-MMSE relation for the single-letter channel [7], we have that for any $s > 0$

$$\mathcal{F}'(s) \propto M(s) + 1/\text{snr} - \delta/s.$$

The fact that $M(s) > 0$ for all s implies that $\mathcal{F}'(s)$ is strictly positive for all $s \geq \delta \text{snr}$, and thus $\mathcal{F}(s)$ is strictly increasing

on $[\delta \text{snr}, \infty)$. Alternatively, the fact that $M(s) < 1$ for all $s > 0$ implies that $\mathcal{F}'(s)$ is strictly negative for all $s \leq \delta \text{snr} / (1 + \text{snr})$, and thus $\mathcal{F}(s)$ is strictly decreasing on $(0, \delta \text{snr} / (1 + \text{snr})]$. \square

In view of Lemma 1, the minimum of the potential function and the smallest and largest minimizers can be defined as follows:

$$\mathcal{F}^* := \min_s \mathcal{F}(s), \quad (12)$$

$$\underline{s}^* := \min\{s : \mathcal{F}(s) = \mathcal{F}^*\}, \quad (13)$$

$$\bar{s}^* := \max\{s : \mathcal{F}(s) = \mathcal{F}^*\}. \quad (14)$$

Note that $\underline{s}^* = \bar{s}^*$ if and only if the minimum is attained at a unique point.

Proposition 2 (RS MMSE [2, 3, 8]). *For any $(\delta, \text{snr}, \pi)$ for which (snr, π) satisfies the single-crossing property [2] and π has finite fourth moment¹, the mutual information and MMSE satisfy*

$$\lim_{p \rightarrow \infty} \frac{1}{p} I(\beta; X, Y) = \mathcal{F}^*, \quad (15)$$

$$\limsup_{p \rightarrow \infty} \frac{1}{p} \mathbb{E} \left[\|\beta - \mathbb{E}[\beta | y, X]\|^2 \right] \leq M(\bar{s}^*), \quad (16)$$

$$\liminf_{p \rightarrow \infty} \frac{1}{p} \mathbb{E} \left[\|\beta - \mathbb{E}[\beta | y, X]\|^2 \right] \geq M(\underline{s}^*), \quad (17)$$

where the limits are taken as $(n = n_p, p, \sigma^2 = \sigma_p^2)$ scale to infinity with $p \rightarrow +\infty$, $n/p \rightarrow \delta$ and $p/\sigma^2 \rightarrow \text{snr}$.

Next, we turn to the family of *approximate message passing* (AMP) [9, 10] algorithms and specifically to the case of MMSE-AMP which is proven in to be optimal among AMP algorithms in minimizing the MSE of the recovery problem of interest [11]. For simplicity from now on when we say AMP we refer to the AMP-MMSE algorithm. We show how a related formula to the one described in Proposition 2 describes the asymptotic MSE associated with AMP.

The smallest stationary point is defined as

$$s^{\text{AMP}} := \inf\{s : \mathcal{F}'(s) = 0\}. \quad (18)$$

It is rather straightforward to check that s^{AMP} is attained by some positive value s and therefore its a stationary point of $\mathcal{F}(s)$. In particular, by Lemma 1 we have $s^{\text{AMP}} \in (\delta \text{snr} / (1 + \text{snr}), \delta \text{snr})$.

For the next result, for $T \in \mathbb{N}$ let $\hat{\beta}_{\text{AMP}, T}(Y, X)$ the output of the AMP estimator [11, Section II.C] with input data (Y, X) after T iterations.

¹Different set of assumptions on $(\delta, \text{snr}, \pi)$ for which the Proposition holds can be found in [3, 8]

Proposition 3 (AMP, [10, 11]). *For any $(\delta, \text{snr}, \pi)$ where π has a finite fourth moment, AMP satisfies*

$$\lim_{T \rightarrow +\infty} \lim_{p \rightarrow +\infty} \frac{1}{p} \mathbb{E} \left[\left\| \beta - \widehat{\beta}_{\text{AMP}, T}(Y, X) \right\|^2 \right] = M(s^{\text{AMP}}) \quad (19)$$

where the outer limit is taken as $(n = n_p, p, \sigma^2 = \sigma_p^2)$ scale to infinity with $p \rightarrow +\infty$, $n/p \rightarrow \delta$ and $p/\sigma^2 \rightarrow \text{snr}$.

Remark 1. The results stated above imply that AMP is optimal whenever $\underline{s}^* = \bar{s}^* = s^{\text{AMP}}$.

Remark 2. For a proof of Proposition 3 we refer the reader to the statement and proof of [11, Theorem 6].

3. MAIN RESULTS

Let us consider now a family of coefficient distributions $(\pi_\epsilon)_{\epsilon > 0}$ indexed by a positive-valued parameter $\epsilon > 0$. We assume throughout the section that for each $\epsilon > 0$ the distributions π_ϵ has zero mean, unit variance and finite entropy H_ϵ . Our results are all based on the following assumption on the family π_ϵ .

Assumption 1. Let $(\pi_\epsilon)_{\epsilon > 0}$ be a family of distributions with unit variance and finite entropy H_ϵ . The MMSE function $M_\epsilon(s)$ of the single letter channel, as defined in (8), for π_ϵ coefficient distribution is assumed to converge pointwise to a step function as $\epsilon \rightarrow 0$ in the following sense

$$\lim_{\epsilon \rightarrow 0} M_\epsilon(2H_\epsilon t) = \begin{cases} 1, & t \in [0, 1) \\ 0, & t \in (1, \infty). \end{cases} \quad (20)$$

Remark 3. It can be straightforwardly checked using the I-MMSE relation for the single-letter channel [7] that the rescaling in the argument of M_ϵ by twice the entropy term, i.e. by $2H_\epsilon$, is necessary for the convergence of M_ϵ to the step function.

Remark 4. As we establish later in the section, Assumption 1 is satisfied for the family of (normalized) Bernoulli distributions with probability ϵ . We expect the assumption to hold under greater generality.

We now present our two main results assuming the family of distributions $(\pi_\epsilon)_{\epsilon > 0}$ satisfies Assumption 1.

Theorem 4. *Let $(\pi_\epsilon)_{\epsilon > 0}$ satisfying Assumption 1. Given a number $r \in (0, 1) \cup (1, \infty)$, let $(\delta_\epsilon, \text{snr}_\epsilon, \pi_\epsilon)_{\epsilon > 0}$ be a family of triplets such that*

$$\lim_{\epsilon \rightarrow 0} \frac{\delta_\epsilon}{\delta_{\epsilon, \text{MMSE}}} = r. \quad (21)$$

Then, the minimizers of the RS potential function exhibit the all-or-nothing behavior in the small- ϵ limit depending on whether r is greater than or less than one:

$$r \in (0, 1) \implies M_\epsilon(\bar{s}_\epsilon^*) \rightarrow 0 \quad (22)$$

$$r \in (1, \infty) \implies M_\epsilon(\underline{s}_\epsilon^*) \rightarrow 1. \quad (23)$$

Combining Theorem 4 with Proposition 2 we obtain the following Corollary.

Corollary 5 (All-or-nothing MMSE behavior). *Let $r \in (0, 1) \cup (1, \infty)$. For any family of triplets $(\delta_\epsilon, \text{snr}_\epsilon, \pi_\epsilon)_{\epsilon > 0}$, suppose that for any $\epsilon > 0$, $(\text{snr}_\epsilon, \pi_\epsilon)$ satisfies the single-crossing property, π_ϵ has finite fourth moment, $(\pi_\epsilon)_{\epsilon > 0}$ satisfies Assumption 1, and (21) holds. Then it holds that*

$$\lim_{\epsilon \rightarrow 0} \lim_{p \rightarrow \infty} \frac{1}{p} \mathbb{E} \left[\left\| \beta - \mathbb{E}[\beta | Y, X] \right\|^2 \right] = \begin{cases} 1, & r \in [0, 1) \\ 0, & r \in (1, \infty). \end{cases} \quad (24)$$

where the inner limits are taken as $(n = n_p, p, \sigma^2 = \sigma_p^2)$ scale to infinity with $p \rightarrow +\infty$, $n/p \rightarrow \delta_\epsilon$ and $p/\sigma^2 \rightarrow \text{snr}_\epsilon$.

We next present our second main result.

Theorem 6. *Let $(\pi_\epsilon)_{\epsilon > 0}$ satisfying Assumption 1. Given a number $r \in (0, 1) \cup (1, \infty)$, let $(\delta_\epsilon, \text{snr}_\epsilon, \pi_\epsilon)$ be such that*

$$\lim_{\epsilon \rightarrow 0} \frac{\delta_\epsilon}{\delta_{\epsilon, \text{AMP}}} = r \quad (25)$$

Then, the smallest stationary point s^{AMP} exhibits the all-or-nothing behavior in the small- ϵ limit depending on whether r is greater than or less than one:

$$r \in (0, 1) \implies M_\epsilon(s_\epsilon^{\text{AMP}}) \rightarrow 1 \quad (26)$$

$$r \in (1, \infty) \implies M_\epsilon(s_\epsilon^{\text{AMP}}) \rightarrow 0. \quad (27)$$

For the result, for $T \in \mathbb{N}$ let $\widehat{\beta}_{\text{AMP}, T}(Y, X)$ the output of the AMP estimator [11, Section II.C] with input data (Y, X) after T iterations. Combining Theorem 6 with Proposition 3 we obtain the following Corollary on the performance of AMP.

Corollary 7 (All-or-nothing AMP behavior). *Let $r \in (0, 1) \cup (1, \infty)$. For any family of triplets $(\delta_\epsilon, \text{snr}_\epsilon, \pi_\epsilon)_{\epsilon > 0}$, suppose that each π_ϵ has a finite fourth moment, the family $(\pi_\epsilon)_{\epsilon > 0}$ satisfies Assumption 1, and (25) holds. Then it holds that*

$$\lim_{\epsilon \rightarrow 0} \lim_{T \rightarrow +\infty} \lim_{p \rightarrow \infty} \frac{1}{p} \mathbb{E} \left[\left\| \beta - \widehat{\beta}_{\text{AMP}, T}(Y, X) \right\|^2 \right] = \begin{cases} 1, & r \in [0, 1) \\ 0, & r \in (1, \infty). \end{cases} \quad (28)$$

where the inner limits are taken as $(n = n_p, p, \sigma^2 = \sigma_p^2)$ scale to infinity with $p \rightarrow +\infty$, $n/p \rightarrow \delta_\epsilon$ and $p/\sigma^2 \rightarrow \text{snr}_\epsilon$.

4. APPLICATION: SPARSE BINARY REGRESSION

We now present our main application of our two results to sparse binary regression, where $\beta_i \stackrel{\text{i.i.d.}}{\sim} \text{Bern}(\epsilon)$. To this end, we first consider the case where β_i is i.i.d. drawn from the following two-point distribution:

$$\pi_\epsilon = (1 - \epsilon) \delta_{\mu_1} + \epsilon \delta_{\mu_2}, \quad (29)$$

where δ_x denotes a Dirac distribution with mass at $x \in \mathbb{R}$, and $\mu_1 = -\sqrt{\epsilon/(1-\epsilon)}$ and $\mu_2 = \sqrt{(1-\epsilon)/\epsilon}$ are chosen such that π_ϵ has zero mean and unit variance. The following Lemma holds for the family of MMSE functions $(M_\epsilon(s))_{\epsilon>0}$:

Lemma 8. *The distribution π_ϵ in (29) has entropy $H_\epsilon = -\epsilon \log \epsilon - (1-\epsilon) \log(1-\epsilon)$ and MMSE function*

$$M_\epsilon(s) = \mathbb{E} \left[\frac{1}{1 - \epsilon + \epsilon \exp\left(\frac{s}{2\epsilon(1-\epsilon)} + \sqrt{\frac{s}{\epsilon(1-\epsilon)}} N\right)} \right], \quad (30)$$

where $N \sim \mathcal{N}(0, 1)$. Furthermore, the distribution π_ϵ satisfies the single-crossing condition [2] for all $\text{snr} > 0$ and

$$\lim_{\epsilon \rightarrow 0} H_\epsilon / (\epsilon \log 1/\epsilon) = 1 \quad (31)$$

$$\lim_{\epsilon \rightarrow 0} \sup_{s>0} \left| M_\epsilon(s) - Q\left(\frac{s - 2\epsilon \log(1/\epsilon)}{2\sqrt{s\epsilon}}\right) \right| = 0, \quad (32)$$

where $Q(z) = \int_z^\infty (2\pi)^{-1/2} \exp(-t^2/2) dt$.

An immediate implication of the result is that the family of distributions $(\pi_\epsilon)_{\epsilon>0}$ satisfies Assumption 1 as well as the conditions of Corollaries 5 and 7. Hence, all-or-nothing phase transitions hold for the limiting MMSE around $\delta_{\epsilon, \text{MMSE}}$ and for the MSE of the AMP around $\delta_{\epsilon, \text{AMP}}$. Using that $\lim_{\epsilon \rightarrow 0} H_\epsilon / (\epsilon \log 1/\epsilon) = 1$, we can further simplify the phase transition points given in (5), (6) by observing

$$\lim_{\epsilon \rightarrow 0} \delta_{\epsilon, \text{MMSE}} / \left(\frac{2\epsilon \log(1/\epsilon)}{\log(1 + \text{snr}_\epsilon)} \right) = 1 \quad (33)$$

and

$$\lim_{\epsilon \rightarrow 0} \delta_{\epsilon, \text{AMP}} / \left(\frac{2(1 + \text{snr}_\epsilon)\epsilon \log(1/\epsilon)}{\text{snr}_\epsilon} \right) = 1. \quad (34)$$

Next, we extend the above results to the sparse binary regression problem of interest, where $\beta_i \stackrel{\text{i.i.d.}}{\sim} \text{Bern}(\epsilon)$. We denote by $k = \epsilon p$ the (expected) number of non-zero coordinates of β . Define

$$\tilde{\beta} = \frac{\beta - \mathbb{E}[\beta]}{\sqrt{\epsilon(1-\epsilon)}}.$$

Then $\tilde{\beta}_i \stackrel{\text{i.i.d.}}{\sim} \pi_\epsilon$ as given in (29). Moreover, define

$$\tilde{Y} = \frac{Y - X\mathbb{E}[\beta]}{\sqrt{\epsilon(1-\epsilon)}}, \quad \tilde{W} = \frac{W}{\sqrt{\epsilon(1-\epsilon)}}.$$

Then it follows that $\tilde{Y} = X\tilde{\beta} + \tilde{W}$. Since $\tilde{W}_i \stackrel{\text{i.i.d.}}{\sim} \mathcal{N}(0, \tilde{\sigma}^2)$ with $\tilde{\sigma} = \sigma/\sqrt{\epsilon(1-\epsilon)}$, it follows that

$$\text{snr}_\epsilon = \frac{p}{\tilde{\sigma}^2} = \frac{p\epsilon(1-\epsilon)}{\sigma^2}. \quad (35)$$

Hence, according to Corollary 5, (33), (35), we obtain that the limiting MMSE exhibits an all-or-nothing behavior at

$$\delta_{\epsilon, \text{MMSE}} = 2\epsilon \log(1/\epsilon) / \log(1 + \epsilon(1-\epsilon)p/\sigma^2),$$

which using $k = \epsilon p$ as $\epsilon \rightarrow 0$ simplifies with negligible multiplicative error to

$$\delta_{\epsilon, \text{MMSE}} = 2(k/p) \log(p/k) / \log(1 + k/\sigma^2).$$

Note that this is the exact information-theoretic threshold for which an all-or-nothing phenomenon has been proven to hold when $\limsup_p \log k / \log p < 0.5$ in [4].

Similarly, according to Corollary 7, (34), and (35), the limiting MSE of the AMP exhibits an all-or-nothing behavior at:

$$\delta_{\epsilon, \text{AMP}} = 2 \left(1 + \frac{\sigma^2}{p\epsilon(1-\epsilon)} \right) \epsilon \log(1/\epsilon),$$

which using $k = \epsilon p$ as $\epsilon \rightarrow 0$ simplifies with negligible multiplicative error to

$$\delta_{\epsilon, \text{AMP}} = 2(k + \sigma^2) \log(p/k) / p.$$

Note that this is the exact computational threshold for a number of computationally efficient methods in the literature such as LASSO or Orthogonal Matching Pursuit (see [5, 6] for references). Our result suggest that the threshold corresponds to a barrier also for AMP in a strong sense.

5. REFERENCES

- [1] Dongning Guo and S. Verdú, "Randomly spread cdma: asymptotics via statistical physics," *IEEE Transactions on Information Theory*, vol. 51, no. 6, pp. 1983–2010, June 2005.
- [2] Galen Reeves and Henry D. Pfister, "The replica-symmetric prediction for random linear estimation with Gaussian matrices is exact," *IEEE Trans. Inform. Theory*, vol. 65, no. 4, pp. 2252–2283, Apr. 2019.
- [3] Jean Barbier, Mohamad Dia, Nicolas Macris, and Florent Krzakala, "The mutual information in random linear estimation," in *Proc. Annual Allerton Conf. on Commun., Control, and Comp.*, Monticello, IL, 2016.
- [4] Galen Reeves, Jiaming Xu, and Ilias Zadik, "The all-or-nothing phenomenon in sparse linear regression," in *Conference On Learning Theory (COLT)*, 2019, [Online]. Available <https://arxiv.org/abs/1903.05046>.
- [5] David Gamarnik and Ilias Zadik, "High dimensional regression with binary coefficients. estimating squared error and a phase transition," in *COLT*, 2017.

- [6] David Gamarnik and Ilias Zadik, “Sparse high-dimensional linear regression. algorithmic barriers and a local search algorithm,” 2017, [Online]. Available <https://arxiv.org/abs/1711.04952>.
- [7] Dongning Guo, S. Shamai, and S. Verdu, “Mutual information and mmse in gaussian channels,” in *International Symposium on Information Theory, 2004. ISIT 2004. Proceedings.*, June 2004, pp. 349–349.
- [8] Jean Barbier, Florent Krzakala, Nicolas Macris, Léo Miolane, and Lenka Zdeborová, “Optimal errors and phase transitions in high-dimensional generalized linear models,” *Proceedings of the National Academy of Sciences*, vol. 116, no. 12, pp. 5451–5460, 2019.
- [9] David L. Donoho, Arian Maleki, and Andrea Montanari, “Message-passing algorithms for compressed sensing,” *Proceedings of the National Academy of Sciences*, vol. 106, no. 45, pp. 18914–18919, Nov. 2009.
- [10] M. Bayati and A. Montanari, “The dynamics of message passing on dense graphs, with applications to compressed sensing,” *IEEE Trans. Inform. Theory*, vol. 57, no. 2, pp. 764–785, Feb. 2011.
- [11] Galen Reeves and Michael Gastpar, “The sampling rate-distortion tradeoff for sparsity pattern recovery in compressed sensing,” *IEEE Trans. Inf. Theor.*, vol. 58, no. 5, pp. 3065–3092, May 2012.