

# On climate change, water variability and conflicts

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1 **A growing empirical literature associates climate anomalies with in-  
2 creased risk of violent conflict. This association has been portrayed  
3 as a bellwether of future societal instability as the frequency and in-  
4 tensity of extreme weather events are predicted to increase. This  
5 paper investigates the theoretical foundation of this claim. A semi-  
6 nal microeconomic model of opportunity costs – a mechanism often  
7 thought to drive climate-conflict relationships – is extended by con-  
8 sidering realistic changes in the distribution of climate-dependent  
9 agricultural income. Results advise caution in using empirical as-  
10 sociations between short-run climate anomalies and conflicts to pre-  
11 dict the effect of sustained shifts in climate regimes: Although war  
12 occurs in bad years, conflict may decrease if agents expect more  
13 frequent bad years. Rather, theory suggests a non-monotonic rela-  
14 tion between climate variability and conflict that emerges as agents  
15 adapt and adjust their behavior to the new income distribution. We  
16 identify three measurable statistics of the income distribution that  
17 are each unambiguously associated with conflict likelihood. Jointly,  
18 these statistics offer a unique signature to distinguish opportunity  
19 costs from competing mechanisms that may relate climate anom-  
20 alies to conflict.**

civil conflict | climate change | water resources | agriculture

1 Climate change is commonly portrayed as one of the most  
2 important potential threats to human, ecosystem, and  
3 societal well-being (e.g., 1). Perhaps the most direct of these  
4 threats is the purported link between climate anomalies and  
5 violent conflicts, a notion that is presently shaping political,  
6 military, and popular discourse (2). This attention underscores  
7 the need for understanding the institutional, economic, and  
8 psychological factors that collectively drive individuals and  
9 groups to fight. While there is growing consensus among  
10 academics that the relation between climate anomalies and  
11 conflicts is robust (3), competing explanations and notable  
12 exceptions remain. Interpretation and projection of empirical  
13 findings in the context of climate change requires careful theo-  
14 retical consideration of underlying mechanisms. In this study,  
15 we relate hydrologic and microeconomic theory to mechanisti-  
16 cally describe how changes in water resource availability might  
17 alter the emergence of negative income shocks, a potential  
18 driver of conflict that is sensitive to climate change (3).

19 Why do violent conflicts emerge and persist if they are so  
20 destructive? This paradox has long attracted the interest of  
21 political scientists and economists. The high cost of violence  
22 implies that peace is typically a better (Pareto-improving)  
23 alternative, and most grievances are believed to be resolved  
24 through bargaining (4). Violence might emerge from a bargaining  
25 breakdown that prevents a peaceful redistribution of land  
26 or resources (5). Among the suspected causes of bargaining  
27 breakdown (see 6) are the absence of institutional or social  
28 checks, which creates a disconnect between decision makers  
29 and foot soldiers who pay the price for violence; incomplete  
30 information, including miscalculations of opponents' strength

31 or strategic withholding of private knowledge; and the inability  
32 to commit to a bargain, for example due to fluctuations in  
33 resource availability. Our analysis focuses on the last factor,  
34 because it is perhaps most directly affected by climate change  
35 (3, 7), rather than by historical, cultural, institutional and  
36 socioeconomic contexts. A growing empirical literature high-  
37 lights the link between climate variability and negative income  
38 shocks as an important determinant of violence (7, 8): fighting  
39 tends to happen during bad years, particularly for non-state  
40 level conflicts short of civil war that do not require the levels  
41 of funding and mobilization necessary for organized armed  
42 rebellion (9).

43 In a seminal paper, Chassang and Padro-i Miquel (10) use  
44 an opportunity cost argument to provide a theoretical under-  
45 pinning to the empirical relation between income shocks and  
46 conflict. The basic idea is that attacking diverts productive  
47 resources but yields an offensive advantage. There is little to  
48 lose in diverting resources to attack in bad years, but much  
49 to be gained from the expected future returns of captured  
50 resources. In bad years, the returns from attack outweigh  
51 the returns from peace. This prevents peaceful bargaining  
52 over resources, and parties go to war. This causal associa-  
53 tion between anomalously bad weather shocks and conflict  
54 occurrence has been robustly documented in the empirical  
55 literature. Motivated by both the theory and the empirical  
56 observations, many have argued that opportunity costs may  
57 be an important mechanism by which climate change can  
58 increase the propensity for conflict (see 3, 7). More extreme

## Significance Statement

There is growing consensus among academics that climate change may amplify the risk of violent conflicts. While underlying mechanisms are poorly understood, negative income shocks associated with climate variability have been long-hypothesized to play an important role. We relate recent hydrologic and microeconomic advances to investigate the theoretical foundation of this claim. Results prescribe caution in interpreting empirical relations between climate variability and conflict in the context of climate change. While fighting preferentially occurs during climate anomalies, more frequent anomalies may not yield more conflicts. By shifting the entire distribution of rainfall, climate change effectively redefines the very notion of climate anomaly. Adaptation to this new normal can have a dominant, and often counterintuitive, effect on conflict probability.

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59 weather events and reduced crop productivity (e.g., 11) might  
 60 increase the frequency and intensity of income shocks during  
 61 which fighting tends to occur. This possibility is particularly  
 62 important to consider in institutionally weak and ethnically  
 63 fragmented regions where climate most directly impacts livelihoods  
 64 (12–14) – ironically, regions believed to be particularly  
 65 vulnerable to future climate change (15).

66 Two important knowledge gaps remain. First, existing  
 67 studies look at anomalous weather events, which affect the  
 68 cost but not the benefit of war. They find that parties go to  
 69 war in 'bad' years. A changed climate, however, alters the  
 70 *distribution* of annual rainfall. This affects the distribution  
 71 of income, which in turn affects both the costs and benefits  
 72 of fighting. Drought years will become more frequent as rain-  
 73 fall variability increases, which raises concerns about higher  
 74 conflict likelihood. However, captured resources will also be-  
 75 come less productive, which lowers the incentives for attack.  
 76 An internally consistent prediction on the conflict impact of  
 77 climate change has to account for both of these changes in  
 78 agents' cost-benefit analysis in a way that, to our knowledge,  
 79 existing projections do not. Second, competing mechanisms  
 80 (other than opportunity costs) can explain the observed link  
 81 between climate anomalies and conflict (see, e.g., 16) and cur-  
 82 rent studies do not conclusively speak to their relative salience.  
 83 Yet, effective policy design requires an accurate identification  
 84 of the underlying drivers for conflict.

85 We address these gaps by linking the opportunity cost  
 86 model proposed by Chassang and Padro-i Miquel (10) to  
 87 a parametric distribution of climate-related income that is  
 88 consistent with the current state of the art in hydrologic and  
 89 agronomic models (17–19). We perform a comparative statics  
 90 analysis (20) that accounts for agents' strategic adjustment to  
 91 a changed environment. Results yield important, and perhaps  
 92 counter-intuitive, insights on the two identified knowledge gaps.  
 93 First, one must be cautious in using empirical associations  
 94 between short-run climate anomalies and conflicts to predict  
 95 the effect of sustained shifts in climate regimes. If precipitation  
 96 becomes more variable, as climate models predict, conflicts  
 97 will not necessarily become more frequent. Rather, conflict  
 98 likelihood can go either up or down, as agents adapt and adjust  
 99 their response to the new income distribution. Even shifts in  
 100 climate *averages* will affect the income *variance*, and therefore  
 101 conflict, due to non-linear processes that link climate to income.  
 102 Second, we identify three measurable statistics of the income  
 103 distribution that individually have an unambiguous effect on  
 104 conflict and are jointly sufficient to predict the response of  
 105 conflict to a change in climate. These testable predictions may  
 106 help distinguish opportunity costs from competing mechanisms  
 107 relating climate anomalies to conflicts.

108 It is important to note that the model is not a tool for  
 109 making quantitative projections of climate-conflict trends in a  
 110 specific geopolitical context, particularly given the multiple  
 111 pathways by which societies can respond to climate or eco-  
 112 nomic shocks (see 6, 16). Rather, the primary objective of  
 113 the model is a careful theoretical treatment of opportunity  
 114 costs as a mechanism often thought to drive the relationship  
 115 between climate change and conflict. In doing so we elucidate  
 116 the rich dynamics, and often counterintuitive outcomes, that  
 117 emerge even under highly stylized theoretical representations  
 118 of human behavior and climate (21).

## Model overview

119 Consider two groups of farmers, whose annual income is sub-  
 120 ject to random rainfall variability, and who might fight for  
 121 control over limited land and labor resources (22). Each year,  
 122 the decision to attack is taken by weighing the immediate  
 123 opportunity costs of fighting against future expected returns  
 124 from the captured resources. The former is given by the cur-  
 125 rent year's rainfall draw and the latter is jointly determined by  
 126 the entire distribution of rainfall, by the probability of victory,  
 127 and by the endogenous risk of conflict occurring in future years  
 128 (see Materials and Method). Under these conditions, Chassang  
 129 and Padro-i Miquel (10) show that conflict emerges in 'bad'  
 130 years, when income falls below a threshold determined by its  
 131 underlying distribution. Insofar as income is influenced by  
 132 climate, their model offers a mechanism that can explain the  
 133 empirical findings that relate climate anomalies to conflicts  
 134 (16).

135 We extend the existing model by specifying a rainfall dis-  
 136 tribution and an income-generating crop function that are  
 137 analytically tractable and consistent with governing meteorolog-  
 138 ical and hydrological processes (see Materials and Methods).  
 139 Doing so introduces a nonlinear relation between climate and  
 140 income, which implications for conflict we discuss in the fol-  
 141 lowing section. The parametric distribution of income also  
 142 allows us to compare predictions of conflict probabilities *across*  
 143 *distributions* by altering parameters to emulate the effect of  
 144 climate change (Figure 1). We initially focus on changes in  
 145 the relative variability of seasonal rainfall, quantified by its  
 146 coefficient of variation ( $CV_W$ ). The focus on  $CV_W$  places  
 147 our study at the intersection of empirical research exploring  
 148 historic associations between conflict, income, and short-run  
 149 anomalies of seasonal rainfall (see 7, 8) and climate modeling  
 150 research predicting an increase in rainfall variability (e.g., 11).  
 151 By performing a comparative statics exercise (20), we allow  
 152 agents to adapt to changed costs and benefits by adjusting  
 153 their fighting threshold. A changed climate affects both the  
 154 present opportunity cost and the future returns from conflict,  
 155 to which agents *adapt* by shifting the income threshold below  
 156 which they will decide to fight. For analytical tractability, we  
 157 favor this rather narrow definition of climate adaptation over  
 158 a broader interpretation that would allow agents to endoge-  
 159 nously optimize income distribution itself, e.g., through crop,  
 160 policy and infrastructure selection.

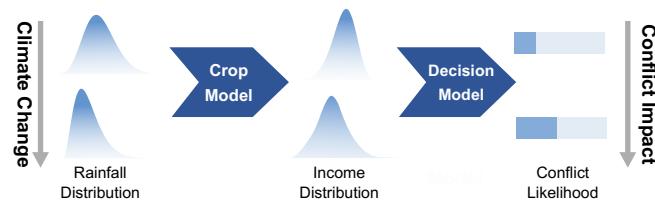
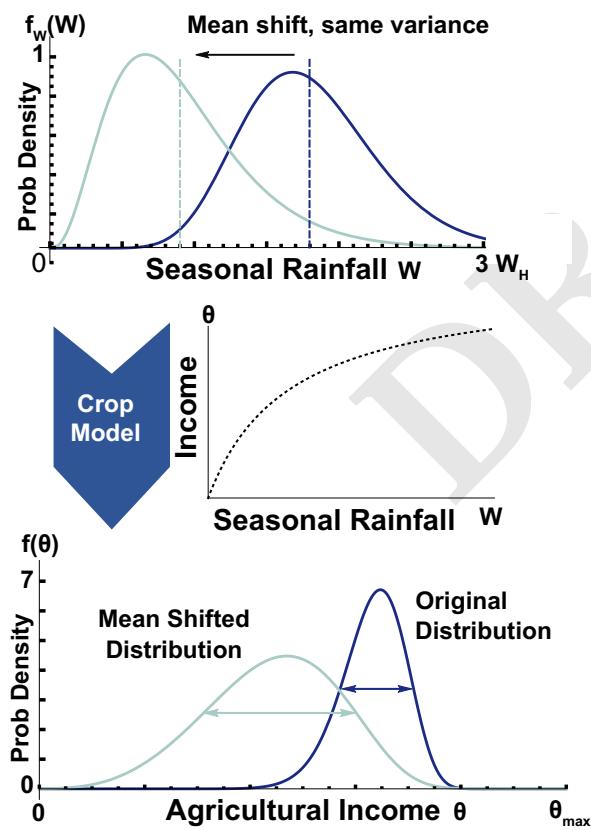


Fig. 1. Schematized relation between climate, crop and conflict models. Moving from left to right, rainfall distributions are related to income distributions via a deterministic model relating seasonal water availability to crop yield, taken as a proxy for income. Income distributions then inform a decision model for conflict. Changes in climate (top to bottom) alter the distribution of water availability, a change that propagates to an altered conflict likelihood.

## 162 Mean climate can affect income variance

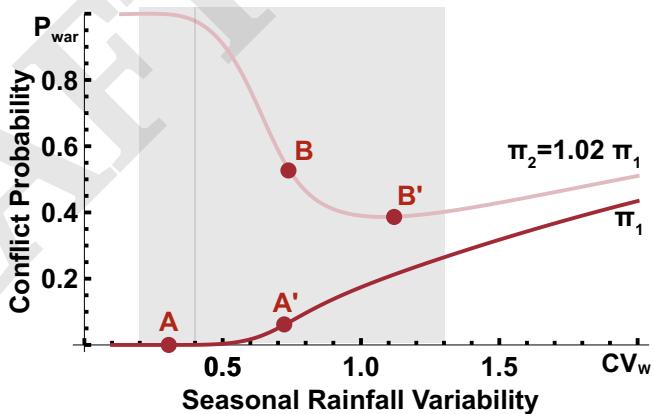
163 Crop yields do not generally scale linearly with water supply (e.g., 19), and so changes in *mean* water availability will  
 164 alter the *variance* of agricultural income. In particular, a  
 165 crop chosen to be robust to climate variations will have mean  
 166 water availability map to a flat region of its yield function  
 167 (dark blue line in Figure 2 top). For such a crop choice, the  
 168 effect of climate variability on income variability is minimal  
 169 under existing climate conditions (dark blue line in Figure 2  
 170 bottom). However, a systematic decrease in water availability  
 171 will enhance income variability due to the concave nature of  
 172 the crop yield curve. This effect is particularly pronounced  
 173 in the low water availability region (low  $W$ ) of the crop yield  
 174 curve, where curvature is maximal. There is a broad consensus  
 175 in climate predictions that points to an increase in rainfall  
 176 variability and an increase in mean temperatures (see, e.g.,  
 177 11). The discussion below focuses on changing drought char-  
 178 acteristics caused by an increase in rainfall variability. However,  
 179 the nonlinearity of the climate-income link implies similar  
 180 conclusions for sustained increases in mean temperature or for  
 181 excess precipitation (see *Supporting Information*).  
 182



**Fig. 2.** Changes in mean climate can cause changes in income variability. A decrease in mean rainfall  $W$  (with variance conserved, top), causes the distribution to map to a more concave region of the crop model (middle), resulting in distributions of income  $\theta$  with a larger variance  $\sigma_\theta^2$  (bottom). This increasing variance compounds the effect of a decreased mean income  $\mu_\theta$  and increases the coefficient of variation  $CV_\theta = \sigma_\theta / \mu_\theta$  of income. Model parameters (see Materials and Methods):  $W_H = 150$  mm,  $\theta_{max} = 3$  currency units,  $\sigma_W = 69.6$  mm,  $\mu_W = 270$  mm (dark lines) and 135 mm (light lines).

## 183 Non-monotonic effect of climate variability on conflict

184 Despite the stylized nature of the opportunity cost model,  
 185 changes in the coefficient of variability of water elicit complex  
 186 nonlinear, and at times non-monotone, effects on the proba-  
 187 bility of conflict. Figure 3 illustrates how conflict probability  
 188 increases monotonically with climate variability, captured by  
 189 the coefficient of variation  $CV_W$  of rainfall, for some param-  
 190 eter combinations (red line), but the relationship becomes  
 191 non-monotonic for others (pink line). Indeed, it is possible  
 192 that conflict prevalence decreases with climate variability for  
 193 small enough values of  $CV_W$  and a large enough offensive  
 194 advantage in the odds of victory (see *Supplementary Infor-  
 195 mation*). This behavior suggests that the opportunity cost  
 196 framework does *not* consistently predict that a more variable  
 197 climate will give rise to more prevalent conflicts. This insight  
 198 is important to consider when using the framework to interpret  
 199 empirical results. For instance, an empirical study finding an  
 200 insignificant (Figure 3 point A) or negative (Figure 3 point  
 201 B) relation between climate variability and conflict may not  
 202 be incompatible with the opportunity cost framework. It also  
 203 does not dismiss the possibility that a positive relation will  
 204 emerge as  $CV_W$  increases under the effect of climate change  
 205 (as seen in positive slopes at A' and B' on Figure 3).



**Fig. 3.** Subtle changes in economic parameters can substantially alter the qualitative relationship between conflict probability and climate variability. A slight (2%) increase in the probability of first strike victory  $\pi$  (see Materials and Methods) introduces a non-monotone relationship between the coefficient of variation of seasonal rainfall ( $CV_W$ ) and the predicted probability of conflict ( $P_{war}$ ): a higher climate variability successively *decreases* and then increases the probability of conflict for a higher value of  $\pi$  (pink line), whereas the relationship remains monotonically increasing for a lower first strike advantage (red line). The shaded area shows the mean (vertical line) and 99% confidence interval of  $CV_W$  observed for seasonal (3-monthly) rainfall, constructed from daily observations at 671 locations within the United States (23). Model Parameters (see Materials and Methods):  $\pi_1 = 0.5148$  (red),  $\pi_2 = 0.5252$  (pink),  $c = 0.9$ ,  $\delta = 0.9$ ,  $\mu_W = W_H = 150$  mm,  $\theta_{max} = 3$ .

## 206 Governing Statistics and Strategic Adaptation

207 Changes in rainfall variability ( $CV_W$ ) might cause farmers  
 208 to alter the income threshold below which they will engage  
 209 in conflict. This adaptation response can strongly influence  
 210 the probability of conflict ( $P_{war}$ ) as farmers weigh the current  
 211 opportunity costs of attack against expected future profits.  
 212 Opportunity costs are lower during a negative climate shock  
 213 due to decreased crop productivity. Attacking then increases  
 214 potential future profits for two reasons. First, the victor will  
 215 capture her opponent's resources and permanently increase

Table 1. Qualitative effect of the three governing statistics of income distribution on the predicted probability of conflict

Marginal Change in income Statistic	Direct Effect on $P_{\text{war}}$	Adaptation effect on $P_{\text{war}}$
↑ Income Shock Frequency $F(\tilde{\theta})$	↑	↑
↓ Mean Income $E[\theta]$	-	↓
↓ Conflict destructivity $cE[\theta   \theta < \tilde{\theta}]$ (or ↑ income shock intensity)	-	↓

her own agricultural profits. Second, triggering a conflict during an income shock hedges against the possibility of a conflict ever occurring in future periods, when opportunity costs are higher on average. In the stark language of our simple model, both incentives rely on the assumption that the defeated opponent exits the game forever. However, qualitatively similar incentives emerge if the defeated agent temporarily loses his land and ability to fight back, particularly when agents place a high emphasis on short-term profits.

Consequently, the opportunity cost model points to three fundamental statistics of the income distribution that govern the relationship between climate change, adaptation response and conflict: (i) The frequency of income shocks, defined as the probability  $F(\tilde{\theta})$  that income falls below agents' conflict threshold  $\tilde{\theta}$ , has a direct effect on the probability of conflict. Any change to the income distribution that increases shock frequency will directly increase  $P_{\text{war}}$ . However, this direct effect also causes war to occur sooner in expectation, thereby reducing the future profits from (current) peace. This gives rise to a second, indirect, effect of an increase in income shock frequency: Because peace becomes less advantageous, farmers respond by increasing the income thresholds below which they attack, an adaptation that further exacerbates  $P_{\text{war}}$ . (ii) A decrease in mean income  $E[\theta]$ , for instance associated with a permanent decrease in mean rainfall, makes victory less profitable. Agents adapt to this by adjusting their income threshold for conflict. This causes a decrease in agents' threshold and, all other statistics being held constant, a decrease in  $P_{\text{war}}$ . (iii) Expected income during conflict-inducing shocks  $E[\theta | \theta < \tilde{\theta}]$ , as a measure of the intensity of income shocks, is proportional to the destructivity of conflicts and affects agents' incentive to fight through the hedging motive described above. More intense income shocks cause smaller expected losses in income during conflict years. Future conflicts are then less costly on average, which incentivizes agents to postpone fighting. Consequently (and perhaps surprisingly), the anticipation of more intense income shocks has a negative effect on  $P_{\text{war}}$ .

Changes in the three income statistics discussed above have independent and consistent (either positive or negative) effects on conflict prevalence, as summarized in Table 1. However, the influence of changes in *climate* on all three *income* statistics is complex and driven by the specific shape of the distribution that governs inter-seasonal climate variability. For a realistic distribution of water availability (see *Materials and Methods*), these relations are displayed in Figure 4 (top) and show that changes in water variability ( $CV_W$ ) elicit different changes in each of the three income statistics, in terms of sign and magnitude. In particular, the income shock frequency response can either be positive or negative, depending on the value of  $CV_W$  and the equilibrium threshold for fighting (Figure 4 top, dark

blue). In the bottom panel, we decompose the overall changes in conflict probability caused by increased climate variability ( $\frac{\partial P_{\text{war}}}{\partial CV_W}$ , in red) into its previously described fundamental components (Table 1). Depending on the relative magnitude of the responses, the overall relation between climate variability and conflict may itself be non-monotone (Figure 4 and Figure 3, pink). In particular, the figure shows that the relation can be dominated by agents' response to changes (dashed) in both mean income (gray dashed) and in the intensity of income shocks (i.e. conflict destructivity, light blue dashed). This insight is relevant in the context of recent literature focusing almost exclusively on the effect of changes in the *frequency* of income shocks on conflicts (e.g., 7). Our theoretical results suggest that farmer adaptation to other climate-driven income statistics, such as the *intensity* of income shocks, may be equally important to consider.

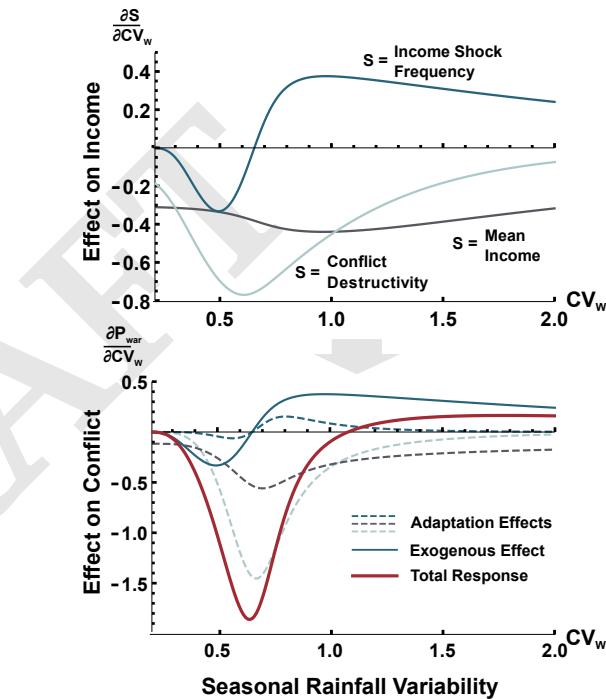


Fig. 4. Components of the climate-income-conflict relationship. *Top*: A marginal increase in rainfall variability affects each of the three governing statistics  $S$  of the income distribution: shock frequency  $F(\tilde{\theta})$ , mean income  $E[\theta]$  and conflict destructivity  $cE[\theta | \theta < \tilde{\theta}]$  (as a measure of shock intensity). The magnitude of each effect is expressed as a partial derivative with respect to  $CV_W$ . Variables  $\theta$ ,  $\tilde{\theta}$ ,  $c$  and  $CV_W$  respectively indicate annual income (a random variable with cumulative density function  $F$ ), the income threshold for conflict, the opportunity cost parameter and the coefficient of variation of rainfall (see Materials and Method). *Bottom*: Income shock frequency has a direct and axiomatic impact on the probability of conflict  $P_{\text{war}} = F(\tilde{\theta})$  (solid blue). However, changes in all three income statistics affect  $P_{\text{war}}$  because agents adapt by changing their income threshold  $\tilde{\theta}$  (dashed lines; see Table 1). The total contribution of these effects determines the non-monotonic response of  $P_{\text{war}}$  (red), which is also expressed as a partial derivative with respect to  $CV_W$ . Parameters (see Materials and Methods):  $\pi = 0.523$ ,  $c = 0.9$ ,  $\delta = 0.9$ ,  $\mu_W = W_H = 150 \text{ mm}$ ,  $\theta_{\text{max}} = 3$ .

## Relation to Empirical Regularities

Chassang and Padro-i Miquel (10) point to two stylized facts that persistently emerge from the empirical literature on income and conflict: (i) conflicts tend to happen during bad income shocks and (ii) conflicts are more prevalent in low-

287 income countries. At first sight, these regularities may appear  
288 at odds with our theoretical predictions suggesting that more  
289 intense income shocks and lower average income both *decrease*  
290 the propensity for conflict (Table 1).

291 At a closer look however, stylized fact (i) is a statement  
292 about low individual draws from a *given* distribution (horizontal  
293 direction in Figure 1), whereas Table 1 concerns sustained  
294 shifts in the *distribution* of income (vertical direction in Figure  
295 1). In line with (10), conflict occurs in our model when income  
296 falls below a certain threshold. Table 1 is saying that a sus-  
297 tained shift in the distribution towards more extreme droughts  
298 causes agents to lower that threshold. In other words, agents  
299 fight in *anomalously* dry years for a given distribution, but  
300 they think twice about fighting for a given draw if dry years  
301 become the 'new normal'.

302 Regarding the second stylized fact, it is important to point  
303 out that the theoretical results in Table 1 concern *marginal*  
304 changes in each income statistic, with the two other statistics  
305 held constant. Any non-marginal change in distribution will  
306 also change the other two statistics because they are them-  
307 selves determined by the threshold  $\tilde{\theta}$ . For instance, scaling  
308 annual income by a constant factor affects all three statistics  
309 in a way that they exactly cancel out (see *Supplementary In-*  
310 *formation*). This gives rise to the invariance of  $P_{\text{war}}$  to income  
311 scaling noted by Chassang and Padró-i Miquel (10). Similarly,  
312 a constant upward *shift* in income results in a *decrease* in  $P_{\text{war}}$   
313 under reasonable assumptions, as shown in *Supplementary In-*  
314 *formation*. The reality may be best captured by a combination  
315 of the two: Rich countries have more income, and also less  
316 volatile income. The model would then indeed predict a lower  
317 probability of war.

## 318 Practical Implications

319 The theoretical arguments in this paper are a strong simplifi-  
320 cation of reality. The economic incentives we discuss represent  
321 a small subset of the social, political and historical processes  
322 that together give rise to violent conflicts. Nonetheless, they  
323 capture important dynamics through which climate-related  
324 income shocks may cause rational agents to be amenable to  
325 conflict. Theoretical insights from the model have three impor-  
326 tant implications that can guide policy and empirical research.

327 First, it is important to distinguish *climate* from *income*  
328 variability when examining their implications for conflicts. The  
329 non-linear and highly local effect of climate on agricultural  
330 income has been highlighted in several studies (e.g., 24, 25)  
331 and has a strong qualitative impact on conflict incentives. It  
332 emerges from a combination of natural (timing of rain events  
333 (26)), technical (crop choice (27)), economic (agricultural  
334 prices (28, 29)) and institutional (insurance and regulation  
335 (30)) processes that were often put in place precisely to  
336 decouple income from climate variability (31). However, as  
337 climate variability begins to exceed historical ranges, these  
338 hedging mechanisms may become less effective. For instance,  
339 a crop that is adapted to a certain precipitation range will  
340 be more susceptible to variation at lower rainfall levels due  
341 to the increased curvature of the crop function (see Figure 2).  
342 This curvature causes a change in *mean* climate to affect the  
343 *variability* of income, which propagates to conflict incentives.  
344 This stylized example highlights the necessity of a careful  
345 empirical characterization of the climate-income relationship  
346 to understand implications for conflicts.

347 Second, theoretical results may inform empirical research  
348 that seeks to disentangle opportunity cost motives from other  
349 mechanisms that predict conflict during bad years. Alternative  
350 hypotheses (see 16) include weakened government structures  
351 (caused by a drop in tax revenue), increased (perceived) in-  
352 equality, climate-induced migration, as well as cognitive and  
353 physiological factors that contribute to aggression. All of these  
354 competing mechanisms also predict that current conflict is  
355 negatively correlated with current income. However, since  
356 none of the alternative explanations are forward-looking, they  
357 would predict either none, or perhaps a negative, correlation  
358 between current conflict and the income in prior years (see  
359 discussion in *Supplementary Information*). Opportunity costs  
360 are different: If agents update their belief about future incomes  
361 in a Bayesian way (some evidence of it is given in Deryugina  
362 (32)), a sequence of good years leads agents to expect greater  
363 gains from attack, and thus render them *more*, not less, ag-  
364 gressive in subsequent years. This is a testable implication  
365 that is unique to the opportunity cost argument and can thus  
366 serve to empirically assess its explanatory power.

367 Finally, caution must be exercised in using micro-economic  
368 income shock arguments to interpret empirical analyses of his-  
369 toric data and draw extrapolations for climate change. While  
370 the model does suggest a positive correlation between weather  
371 anomalies and conflict, it does not support the argument that  
372 conflicts will always be more prevalent if these anomalies occur  
373 more frequently due to climate change. Rather, the theory  
374 suggests a complex, and potentially non-monotonic, relation  
375 between climate variability and conflict. This complexity  
376 emerges both from non-linear climate to income relationships,  
377 and from strategic adaptation by agents to a changing income  
378 distribution. By affecting the entire distribution of climate,  
379 climate change will effectively define a "new normal". Agents  
380 strategically adapt to multiple facets of climate change by  
381 adjusting their response to income variability. In doing so,  
382 they redefine the very notion of climate anomalies and associ-  
383 ated negative income shocks as they pertain to climate-related  
384 conflicts.

## 385 Materials and Methods

386 **Conflict.** Two groups of farmers occupy a common territory over  
387 an infinite number of periods (growing seasons). Three productive  
388 inputs determine crop yields and agricultural income: land, labor  
389 and water availability. Land and labor are equitably distributed  
390 between the two players (unequal distribution can be resolved  
391 through peaceful bargaining (see 10)) and constant across periods.  
392 However, rainfall varies randomly across periods, following a known  
393 probability distribution and affecting both groups identically. In  
394 each period, both groups observe rainfall and either group can  
395 unilaterally launch an attack to seize permanent control of the  
396 entire territory. If neither group attacks, peace prevails, all labor  
397 is put to productive use, and both groups keep control of their  
398 own land and labor. If either side attacks, violence prevails, and  
399 both groups divert a fixed share  $c$  of labor to armed conflict. In a  
400 one-sided attack, the attacker has an offensive advantage and wins  
401 with probability  $\pi > 0.5$ . In a simultaneous attack, both groups  
402 win with equal probability. The winner controls the entire territory  
403 forever, and the loser exits the game.

404 The decision to attack in each season  $t$  relies on weighing the  
405 expected future benefits of victory against the current opportunity  
406 cost of conflict. Peace will prevail if the expected returns of peace,  
407  $E[\mathcal{P}]$ , are larger than the expected returns of launching a surprise

408 attack,  $E[\mathcal{W}]$ :

$$\underbrace{\theta_t}_{\text{current season}} + \underbrace{\delta V^{\mathcal{P}}}_{\text{future seasons}} > \underbrace{\pi 2\theta_t(1-c)}_{E[\mathcal{P}]} + \underbrace{\pi \delta V^V}_{E[\mathcal{W}]} \quad [1]$$

409 where  $\theta_t$  is an income sampled from the PDF  $f(\theta)$ ;  $V^{\mathcal{P}}$  are the future  
410 expected returns of peacefully farming one's own land (discounted  
411 by a constant factor,  $\delta$ );  $\pi$  is the probability of victory in a surprise  
412 attack;  $c$  is the fractional cost of the present season's production  
413 devoted to war; and  $V^V$  represents the expected returns of victory  
414 (discounted by  $\delta$ ). The factor 2 appears because the victorious  
415 farmer obtains both plots of land.

416 A key characteristic of the model is that the current opportunity  
417 cost is driven by an individual draw  $\theta$ , while the future benefits are  
418 affected by the entire probability distribution  $F$  of income. Groups  
419 go to war when current income falls below a threshold  $\tilde{\theta}$ , which  
420 depends on economic parameters and the distribution  $F$ . Chassang  
421 and Padro-i Miquel (10) show that  $V^V = 2E[\theta]/(1 - \delta)$ , where  
422  $E[\cdot]$  is the expectation operator. In contrast,  $V^{\mathcal{P}}$  is an implicit  
423 equation that depends on the attack threshold,  $\tilde{\theta}$ , defined as the  
424  $\theta$  at which  $E[\mathcal{P}] = E[\mathcal{W}]$  (see *Supplementary Information*). An  
425 implicit expression for  $\tilde{\theta}$  is found by substituting  $V^V$  and  $V^{\mathcal{P}}$  into  
426 1, setting  $E[\mathcal{W}] = E[\mathcal{P}]$ , and rearranging:

$$428 \tilde{\theta} = \frac{\delta}{1 - 2P(1 - c)} \left[ (2P - 1) \frac{E[\theta]}{1 - \delta} + \frac{F(\tilde{\theta}) \cdot cE[\theta | \theta < \tilde{\theta}]}{1 - \delta(1 - F(\tilde{\theta}))} \right] \quad [2]$$

429 where  $F(\tilde{\theta}) = \int_0^{\tilde{\theta}} f(x)dx$ , and  $E[| \cdot |]$  is the conditional expectation  
430 operator. The probability of war in any season is simply  $P_{\text{war}} = F(\tilde{\theta})$ .

431 **Climate, water availability and crop productivity.** We assume that  
432 both farmer groups are subject to the same crop productivity ( $\theta_t$ )  
433 governed by seasonal water volume,  $W$  [L], normalized by catchment  
434 area. We use a model for lumped crop yield potential [ $M L^{-2}$ ] as  
435 a proxy for agricultural income,  $\theta$ . Water supply is assumed to  
436 be the yield-limiting factor (19), allowing us to map  $f(\theta)$  directly  
437 to the distribution of water supply,  $f_W(W)$ . Although additional  
438 factors such as intraseasonal dry spells are known to affect crop  
439 yields, we do not include them in our model since our principal aim  
440 is to maintain emphasis on the human decision model, and yields  
441 have been shown to be primarily determined by total precipitation.  
442 Based on observations reported in (19), we specify a parsimonious  
443 boundary function relation for yield,  $B(W)$ :

$$444 \theta = \theta_{\max} \cdot \frac{W}{W + W_H}, \quad [3]$$

445 where  $W_H$  is a half-saturation constant, and  $\theta_{\max}$  is the maximum  
446 productivity. We assume land to be spatially homogeneous and  
447 situated in a watershed sufficiently flat for hydrologic conditions to  
448 be driven by vertical rainfall infiltration into the soil layer (33). We  
449 assume that water is derived from rainfall, allowing  $f_W(W)$  to be  
450 approximated using a Gamma distribution (see 18, and *Supplementary Information*). Under these assumptions, an exact expression  
451 for  $f(\theta)$  is:

$$452 f(\theta) = \frac{\exp \left( -\frac{\theta}{(\theta_{\max} - \theta)(\mu_W W_H^{-1})CV_W^2} \right) \left( \frac{\theta}{(\theta_{\max} - \theta)(\mu_W W_H^{-1})CV_W^2} \right)^{\frac{1}{CV_W^2}}}{\frac{\theta}{\theta_{\max}} (\theta_{\max} - \theta) \Gamma \left( \frac{1}{CV_W^2} \right)} \quad [4]$$

453 where  $\mu_W$  [L] and  $CV_W$  [-] are the mean and coefficient of variation  
454 of  $f_W(W)$ , respectively, and  $\Gamma(-)$  is the Gamma function.

455 **Response of  $P_{\text{war}}$  to Changing Water Resources.** We determine the  
456 response of  $P_{\text{war}} = F(\tilde{\theta})$  to water variability by numerically differen-

tiating  $F(\tilde{\theta})$  with respect to  $CV_W$ :

$$\frac{dP_{\text{war}}}{dCV_W} = \underbrace{\frac{\partial P_{\text{war}}}{\partial CV_W}}_{\text{mechanistic effect}} + \underbrace{\sum_{n=1}^3 \frac{\partial P_{\text{war}}}{\partial \tilde{\theta}} \cdot \frac{\partial \tilde{\theta}}{\partial S_n} \cdot \frac{\partial S_n}{\partial CV_W}}_{\text{farmer adaptation}} \quad [5]$$

457 where  $S \in \{E[\theta], F(\tilde{\theta}), E[\theta | \theta < \tilde{\theta}]\}$  are the three fundamental  
458 statistics that govern  $\tilde{\theta}$  (Equation 2). Total sensitivity of  $P_{\text{war}}$  is  
459 partitioned into direct and adaptation effects (following Burke  
460 et al. (7), Eq. 5). Changes in  $f_W(W)$  alter the probability of an  
461 income shock in a given period (direct effect), thereby changing the  
462 probability that farmers will attack,  $F(\tilde{\theta})$ . The direct change to  $f(\theta)$   
463 also alters the expected returns from peace,  $V^{\mathcal{P}}$  (see *Supplementary Information*). Farmers therefore adapt  $\tilde{\theta}$  to a value that again  
464 satisfies 1 with equality (adaptation effect).

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