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Directional association test reveals high-quality putative cancer driver biomarkers including noncoding RNAs

Hua Zhong^{1†} and Mingzhou Song^{1,2*}

*Correspondence:

joemsong@cs.nmsu.edu

¹Department of Computer Science, New Mexico State University, University Ave, 88003, Las Cruces, NM, USA
Full list of author information is available at the end of the article
†huazhong@nmsu.edu

Abstract

Background: Most statistical methods used to identify cancer driver genes are either biased due to choice of assumed parametric models or insensitive to directional relationships important for causal inference. To overcome modeling biases and directional insensitivity, a recent statistical functional chi-squared test (FunChisq) detects directional association via model-free functional dependency. FunChisq examines patterns pointing from independent to dependent variables arising from linear, non-linear, or many-to-one functional relationships. Meanwhile, the Functional Annotation of Mammalian Genome 5 (FANTOM5) project surveyed gene expression at over 200,000 transcription start sites (TSSs) in nearly all human tissue types, primary cell types, and cancer cell lines. The data cover TSSs originated from both coding and noncoding genes. For the vast uncharacterized human TSSs that may exhibit complex patterns in cancer versus normal tissues, the model-free property of FunChisq provides us an unprecedented opportunity to assess the evidence for a gene's directional effect on human cancer.

Results: We first evaluated FunChisq and six other methods using 719 curated cancer genes on the FANTOM5 data. FunChisq performed best in detecting known cancer driver genes from non-cancer genes. We also show the capacity of FunChisq to reveal non-monotonic patterns of functional association, to which typical differential analysis methods such as *t*-test are insensitive. Further applying FunChisq to screen unannotated TSSs in FANTOM5, we predicted 1108 putative cancer driver noncoding RNAs, stronger than 90% of curated cancer driver genes. Next, we compared leukemia samples against other samples in FANTOM5 and FunChisq predicted 332/79 potential biomarkers for lymphoid/myeloid leukemia, stronger than the TSSs of all 87/100 known driver genes in lymphoid/myeloid leukemia.

Conclusions: This study demonstrated the advantage of FunChisq in revealing directional association, especially in detecting non-monotonic patterns. Here, we also provide the most comprehensive catalog of high-quality biomarkers that may play a causative role in human cancers, including putative cancer driver noncoding RNAs and lymphoid/myeloid leukemia specific biomarkers.

Keywords: FunChisq; non-monotonic directional association; human cancer; cancer driver gene; noncoding RNA; leukemia; biomarker

Background

Greatly outnumbering coding genes, noncoding RNA (ncRNA) genes remain elusive in our understanding of their function. Among various ncRNAs, microRNA,⁴⁵

long noncoding RNA, and enhancer RNA are the most heavily studied and some¹ are deregulated in cancer [1, 2, 3]. Due to technical challenges caused by their² typically low abundance, ncRNA profiles of cancer are yet widely available. For³ example, even in The Cancer Genome Atlas (TCGA) project [4], the expression of⁴ non-polyadenylated ncRNAs in tumor samples is not provided. Encouragingly, the⁵ Functional Annotation of Mammalian Genome 5 (FANTOM5) project [5] measured⁶ promoter-level transcriptome data at 209,911 transcription start sites (TSSs) in 752⁷ human samples covering all major human tissue types, primary cell types, and no-⁸ tably many cancer cell lines represented by 225 samples. Such a sampling diversity⁹ captured a wealth of system dynamics. Additionally, technical variations introduced¹⁰ in data acquisition are minimal because all samples in the project were sequenced¹¹ at the same facility housed in RIKEN, Japan. More than half (107,139) of the TSSs¹² are unannotated, pointing to most likely novel ncRNAs. Therefore, the FANTOM5¹³ data set opens up an enormous opportunity to study the role for ncRNAs in cancer.¹⁴

Most statistical methods used to identify cancer marker genes [6, 7] are either¹⁵ biased due to parametric model choices, insensitive to directional causal relation-¹⁶ ships, or unable to reveal non-monotonic patterns. Table 1 summarizes advantages¹⁷ and disadvantages of several widely used biomarker detection methods. A symmet-¹⁸ ric association test reveals no directionality of a pattern, and thus cannot infer¹⁹ causality. Differential gene expression analysis methods are often unable to de-²⁰ tect non-monotonic patterns from gene to phenotype, commonly seen in biologi-²¹ cal systems. Logistic regression can fit a nonlinear function but requires a correct²² parametric model. To overcome these issues, the functional chi-squared test (Fun-²³ Chisq) [8, 9, 10] is a recently developed statistical test for directional association via²⁴ model-free functional dependency. The FunChisq test statistic is computed from a²⁵ contingency table, where the row variable represents independent variable X and²⁶ the column variable for dependent variable Y . When both X and Y are numeric²⁷ or ordinal, we can define the monotonicity of a pattern. X to Y is monotonically²⁸ increasing/decreasing if Y never decreases/increases as X increases. X to Y is²⁹ non-monotonic if Y can both increase at one point and decrease at another as X ³⁰ increases. The FunChisq test statistic is maximized by either one-to-one or many-to-³¹ one non-constant functions from X to Y given marginal sums of dependent variable³² Y . Thus, FunChisq is sensitive to both monotonic and non-monotonic functional³³ patterns. The original FunChisq test established an asymptotic chi-squared null³⁴ distribution for the test statistic [8]. An exact functional test using the same test³⁵ statistic has been developed to compute its statistical significance based on an ex-³⁶ act, instead of asymptotic, null distribution [9]. We also introduce function index ξ_f ,³⁷ derived from the FunChisq statistic, to measure the effect size of functional depen-³⁸ dency. The relationship of the index to the p -value of the FunChisq test statistic is³⁹ analogous to that of fold-change to p -value in differential gene expression analysis.⁴⁰ The pair of fold change and p -value is often visualized together in a volcano plot.⁴¹ Similarly, examining both the function index and the FunChisq p -value disfavors⁴² patterns either weak in functional dependency or statistically insignificant, leading⁴³ to increased confidence in causal inference.⁴⁴

The Heritage Provider Network (HPN)-Dialogue for Reverse Engineering Assess-⁴⁵ ments and Methods (DREAM) network inference challenges aimed to decipher⁴⁶

Table 1 Comparison of widely used biomarker detection methods.

Methods	Advantages	Disadvantages
Pearson's chi-squared test	Model free	No directionality
<i>t</i> -test	No discretization	No non-monotonicity
Wilcoxon test	Nonparametric	No non-monotonicity
Logistic regression	Nonlinear; No discretization	Requires a parametric model
DESeq2; edgeR	Generalized linear model	Requires a parametric model

causal gene networks connecting signaling proteins in human breast cancer [11].⁹ It evaluated network inference approaches employed or designed by about 80 participating teams for their effectiveness on revealing signaling networks. On the *in silico* data from a non-linear dynamical system model, FunChisq performed the best¹⁰ among all submissions. On the experimental phosphoprotein data measured from¹¹ cancer cell lines in response to stimuli, prior biological knowledge about molecular¹² interactions was allowed to be integrated. Notably, FunChisq, without incorporating¹³ any prior information, was ranked the 7th after six methods all using prior knowl-¹⁴ edge. In the post-challenge evaluation, combining prior knowledge with FunChisq¹⁵ led to substantial better performance over the best performer on the experimental¹⁶ data [11]. The outstanding performance of FunChisq supports its practicality in¹⁷ causal inference. Its advantage in distinguishing interaction directionality and sen-¹⁸ sitivity to non-monotonic patterns motivated us to study genes involved in cancer¹⁹ using FANTOM5 data.²⁰

On FANTOM5 data, we first evaluated FunChisq and six other methods using 719²¹ curated cancer genes. FunChisq performed best in detecting known cancer driver²² genes from non-cancer genes. We also show the capacity of FunChisq to reveal²³ non-monotonic patterns, to which typical differential analysis method such as *t*-²⁴ test are insensitive. We further applied FunChisq on unannotated human TSSs in²⁵ FANTOM5, and predicted 1108 ncRNAs as putative cancer drivers. They have²⁶ directional association to cancer stronger than 90% of the curated cancer driver²⁷ genes. Next, we compared leukemia samples against other samples in FANTOM5²⁸ and FunChisq predicted potential biomarkers for lymphoid leukemia and for myeloid²⁹ leukemia, stronger than all known driver genes of the two leukemia types.³⁰

This study demonstrates that FunChisq indeed detected many non-monotonic³¹ TSS-cancer association patterns, to which previous methods may be blind. As the³² TSS-cancer associations are predicted by directional functional dependency with-³³ out assuming a parametric model, we have provided the most comprehensive and³⁴ unbiased catalog of high-quality noncoding and coding RNA TSSs that may be³⁵ causative factors to human cancers.³⁶

Results

FunChisq is powerful in detecting known human cancer genes

We evaluated the performance of FunChisq and six other tests in distinguishing⁴² 719 curated cancer genes on FANTOM5 human data. The six other tests include⁴³ Pearson's chi-squared test [12], Wilcoxon test [13], *t*-test [14], logistic regression [15],⁴⁴ DESeq2 [16], and edgeR [17]. The curated cancer genes were obtained from Cancer⁴⁵ Gene Census [18] in COSMIC Release v83. The ground truth in the evaluation⁴⁶

was generated with true cancer driver genes and non-cancer-associated genes. For¹ each cancer driver gene, we extracted its representative TSS, which was the most² transcribed among all TSSs of the same gene. However, non-cancer-associated genes³ are not typically reported in the literature. Thus, excluding curated cancer genes, we⁴ randomly picked the same number of TSSs—most likely non-cancer TSSs. Then we⁵ evaluated all seven methods for their performance in revealing true cancer driver⁶ gene TSSs. DESeq2 and edgeR were tested on raw read count data, while the⁷ other methods on discrete data transformed from expression data in the unit of⁸ tags per million (TPM). Specifically, we used the R package *Ckmeans.1d.dp* [19,⁹ 20] to discretize the log-transformed TPM abundance from all samples for each¹⁰ TSS, before which numbers of discretization levels for each gene were automatically¹¹ determined by R package *mclust* [21] by fitting a finite Gaussian mixture model.¹²

The performance of the seven methods on detecting cancer TSSs from FANTOM5¹³ data is summarized in Figure 1. The receiver operating characteristic (ROC) curves¹⁴ in Figure 1a and precision-recall (PR) curves in Figure 1b indicate that FunChisq¹⁵ outperformed the other six methods. We repeated the same evaluation on 100 dif-¹⁶ ferent sets of randomly selected non-cancer TSSs. Figure 1c,d show that the areas¹⁷ under the ROC and PR curves of FunChisq are markedly better than all other¹⁸ six methods, demonstrating the advantage of FunChisq. The fact that directional¹⁹ FunChisq scored better than directionless Pearson's chi-squared test suggests the²⁰ importance of direction in detecting cancer genes. FunChisq also performed much²¹ better than the other five methods (Wilcoxon test, *t*-test, DESeq2, edgeR, and lo-²² gistic regression) not designed for detecting non-monotonic patterns, suggesting the²³ importance of detecting such patterns when analyzing cancer driver gene expression,²⁴ as demonstrated in the next subsection.²⁵

FunChisq is sensitive to non-monotonic patterns²⁷

On the whole-body FANTOM5 human transcriptome data, we showcase non-²⁸ monotonic interaction patterns between TSS abundance of two known cancer genes,²⁹ *KAT6A* (also known as *MYST3* and *MOZ*) [22] and *BRAF* [23], and their cancer³⁰ status of human samples in Figure 2. The non-monotonicity was detected only by³¹ FunChisq, while approaches based on comparison of means, such as *t*-test, would³² fail, because the means of non-monotonic patterns between cancer and non-cancer³³ samples may not differ significantly. *KAT6A* has been implicated to either promote³⁴ or inhibit senescence [24], important for tumor formation and growth [25]. *KAT6A*³⁵ is associated with oncogenesis [22] in both leukemia [26, 27, 28, 29] and breast³⁶ cancer [30], because of dysregulation of its histone acetyltransferase activity or its³⁷ aberrant expression. *KAT6A* was also hypothesized to suppress tumor when severe³⁸ DNA damage happened [31, 24]. Thus, *KAT6A* may both promote and suppress³⁹ cancer, playing competing roles depending on the cellular context. *BRAF* has long⁴⁰ been established as a proto-oncogene [32]. However, *BRAF* paradoxically inhibits⁴¹ stem cell renewal [33]; also in *BRAF*-driven mouse model of colon cancer, tumor⁴² formation is suppressed [33]. Therefore, *BRAF* may either promote or inhibit can-⁴³ cer depending on the context. Both examples illustrate the capacity of FunChisq⁴⁴ in recognizing non-monotonic patterns, which *t*-test and other statistical analysis⁴⁵ methods based on the comparison of group means may not manage to differentiate.⁴⁶

FunChisq is empirically efficient in runtime¹
 We measured the total runtime of the seven methods evaluating the relationship² of all TSSs to cancer, as summarized in Table 2. The input to each method is the³ FANTOM5 data covering 209,911 TSSs across 752 samples, including 527 cancer⁴ cell lines and 225 normal primary/tissue cells. The program ran on a single thread⁵ of a server with $12 \times 2.40\text{GHz}$ Intel(R) Xeon(R) CPU E5645 and 192GB RAM under⁶ openSUSE Leap 15.0 OS. FunChisq, Pearson's chi-squared test, Wilcoxon test and⁷ *t*-test took the least time of less than 10 minutes. Logistic regression and edgeR took⁸ much longer time fitting default models. DESeq2 costed most time due to raw read⁹ count normalization, dispersion estimation, and generalized linear model fitting. In¹⁰ summary, the empirical runtime comparison suggests that FunChisq is practically¹¹ efficient.¹²

Table 2 Empirical runtime of seven methods in evaluating association of 209,911 transcription start sites with cancer. The methods are sorted in the increasing order of runtime.

Methods	Runtime
<i>t</i> -test	2m 26s
Pearson's chi-squared test	8m 32s
FunChisq	8m 40s
Wilcoxon test	8m 41s
edgeR	43m 44s
Logistic regression	44m 01s
DESeq2	54h 08m

FunChisq reveals putative cancer driver noncoding RNAs²³

The latest FANTOM5 annotation has identified most coding genes in the human²⁴ genome. Thus, we hypothesize that the majority of the 107,139 unannotated TSSs²⁵ may belong to potential novel ncRNAs. To identify the directional effect from TSS²⁶ to cancer, we applied FunChisq on the expression of each TSS in cancer versus²⁷ non-cancer samples to report function indices and *p*-values. Figure 3 shows the²⁸ distribution of function index of representative TSSs from the 719 known cancer²⁹ genes, versus that of all other TSSs. The two distributions demonstrate that known³⁰ cancer TSSs have a greater average function index than other TSSs, indicating that³¹ the cancer status has stronger dependency on known cancer TSSs than other TSSs.³²

Rather than picking a fixed function index cutoff, we selected the threshold at 90³³ percentile of known cancer TSS function index values (Figure 3). The criterion is³⁴ stringent to select the most relevant candidates. At the 90 percentile function index³⁵ cutoff of 0.40 and an adjusted *p*-value threshold of 0.05, we selected 1108 unanno-³⁶ tated TSSs with a directional effect on cancer status. Thus they are stronger than³⁷ 90% of representative TSSs of all known cancer driver genes, constituting putative³⁸ cancer driver ncRNAs. Figure 4 shows two such predicted ncRNAs, one with a³⁹ monotonic interaction pattern with cancer status and the other a non-monotonic⁴⁰ pattern. All 1108 predicted noncoding cancer TSSs are listed in **Additional file 1**.⁴¹ We expect cancer biologists to find these ncRNA biomarkers interesting and to⁴² apply either RNA silencing or gene editing to study their functions in cancer.⁴³

Putative cancer-type specific biomarkers for lymphoid and myeloid leukemias⁴⁴

Both lymphoid and myeloid leukemia samples have the largest sample size among⁴⁵ all cancer types sequenced by the FANTOM5 project. We contrast samples of a⁴⁶

cancer type and all remaining samples which also include other cancer types, such¹ that the markers identified are only specific to the cancer type of interest. This² strategy is only possible with FANTOM5 data in that they cover all major tissue,³ cell, and cancer types in human.⁴

We first searched for potential biomarkers of lymphoid leukemia by testing the⁵ directional effect of each TSS on lymphoid leukemia status. Among all 752 sam-⁶ ples from FANTOM5, there are 23 lymphoid leukemia and 48 related normal lym-⁷ phoid samples. We divided the samples into two groups: the first group contains⁸ 23 lymphoid leukemia samples and the second group has all other 729 samples (in-⁹ cluding the 48 normal lymphoid samples and all cancer types other than lymphoid¹⁰ leukemia). We then performed the FunChisq test on each TSS to hunt for ones on¹¹ which lymphoid leukemia status functionally depend. By requiring a *p*-value under¹² 0.05 and a function index greater than all 87 known lymphoid leukemia driver gene¹³ TSSs, we identified 332 putative lymphoid leukemia biomarkers.¹⁴

Next we performed the same procedure to search for biomarkers for myeloid¹⁵ leukemia by contrasting the 28 myeloid leukemia samples with the remaining¹⁶ 724 samples (including 26 normal myeloid samples and all cancer types other¹⁷ than myeloid leukemia). We detected 79 statistically significant putative myeloid¹⁸ leukemia biomarkers, with a *p*-value no more than 0.05 and function index greater¹⁹ than the TSSs of all 100 known myeloid leukemia driver genes.²⁰

Figure 5 illustrates the expression patterns of four biomarker candidates that²¹ are distinct between the specific leukemia and other samples. Only in lym-²² phoid leukemia, *p1@SNX9* is under-expressed but not in any other samples (Fig-²³ ure 5a); *hg_153880.1* is mostly highly expressed only in lymphoid leukemia (Fig-²⁴ ure 5b). *p4@LMO2* is exclusively highly expressed in myeloid leukemia (Figure 5c);²⁵ *hg_35610.1* also exhibited the highest expression in myeloid leukemia (Figure 5d).²⁶

Distributions of detected biomarkers along each chromosome for lymphoid and²⁷ myeloid leukemias are shown in Figure 6. In lymphoid leukemia samples, chro-²⁸ mosomes 12 contain the highest number of biomarkers, while in myeloid leukemia²⁹ samples, chromosome 6 and 19 has much more biomarkers than others. In chronic³⁰ lymphocytic leukemia (CLL), trisomy 12 has been reported to be the third most³¹ frequent chromosomal aberration and is often present as a unique cytogenetic al-³² teration [34]. In acute myeloid leukemia (AML), trisomy chromosome 6 has been³³ reported as a sole cytogenetic abnormality in AML-M5 [35], and chromosome 19³⁴ abnormalities are commonly seen in AML-M7 [36]. Our findings of the biomarker³⁵ genomic locations are consistent with these known chromosomal abnormalities in³⁶ subtypes of leukemia, which supports potential cancer-related functions of the pu-³⁷ tative biomarkers detected.³⁸

The predicted biomarkers of both lymphoid and myeloid leukemias are reported³⁹ in **Additional file 2** (see section Additional files).⁴⁰

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Discussion⁴²

FunChisq measures the functional strength from row variable *X* to column variable⁴³ *Y* in a contingency table via a model-free approach. Given the column sums, a con-⁴⁴ tingency table maximizes the FunChisq statistic if and only if column variable *Y* is a⁴⁵ non-constant mathematical function of row variable *X*. This theoretical optimality⁴⁶

makes FunChisq model-free in promoting all forms of functional patterns regardless¹ of parametric family, linearity, or monotonicity. This flexibility unconstrained by² functional forms offers one a greater capacity in inferring causality with reduced³ biases than other methods.⁴

The model-free property of FunChisq aligns well to the need of unbiased knowl-⁵ edge discovery in the analysis of vast uncharacterized human noncoding genes as⁶ uncovered by the FANTOM5 project, providing us a powerful instrument to assess⁷ objectively the evidence for a gene's directional effect on human cancer.⁸

Conclusions

We have shown that the FunChisq statistical method is powerful in detecting direc-¹¹ tional association, sensitive to both monotonic and non-monotonic patterns. Strong¹² functional patterns provide evidence for causality. Applying the method on the¹³ FANTOM5 data covering the largest number of potential noncoding genes for many¹⁴ cancer types, we revealed putative cancer driver ncRNAs with a directional effect¹⁵ on cancer status stronger than 90% of all 719 curated cancer genes. Furthermore, we¹⁶ predicted 332 potential cancer biomarkers for lymphoid leukemia and 79 for myeloid¹⁷ leukemia, stronger than all known lymphoid or myeloid leukemia genes. Our study¹⁸ thus contributes a catalog of novel biomarker candidates that may signify a deeper¹⁹ understanding of cancer biology.²⁰

Methods

We used the normalized functional chi-squared test with an asymptotic normal null²⁴ distribution to discover directional association in contingency tables [8, 11]. The test²⁵ detects model-free functional dependency and does not need a prescribed functional²⁶ form. The directional functional dependency can potentially indicate the causal²⁷ direction of an interaction based on the causality-by-functionality principle [37].²⁸

An observed $r \times c$ contingency table O has r rows representing the discrete levels²⁹ for independent variable and c columns representing the discrete levels for dependent³⁰ variable. Let O_{ij} denote the sample counts at row i and column j . Let $O_{i\cdot}$ be the³¹ row sum of row i and $O_{\cdot j}$ be the column sum of column j , defined as³²

$$O_{i\cdot} = \sum_{j=1}^c O_{ij} \quad \text{and} \quad O_{\cdot j} = \sum_{i=1}^r O_{ij} \quad (1)$$

Let n represent the sample size of table O . The FunChisq statistic of observed table³⁷ O is defined by³⁸

$$\chi_f^2(O) = \left[\sum_{i=1}^r \sum_{j=1}^c \frac{(O_{ij} - O_{i\cdot}/c)^2}{O_{i\cdot}/c} \right] - \sum_{j=1}^c \frac{(O_{\cdot j} - n/c)^2}{n/c} \quad (2)$$

which asymptotically follows a chi-squared distribution with $\nu = (r-1)(c-1)$ ⁴³ degrees of freedom, under the null hypothesis of the row and column variables⁴⁴ being statistically independent and an assumption of the dependent variable being⁴⁵ uniformly distributed. We further define the normalized FunChisq by mean-shifting⁴⁶

and standard-deviation-scaling $\chi_f^2(O)$ to

$$\frac{\chi_f^2(O) - \nu}{\sqrt{2\nu}} \quad (\text{Normalized FunChisq})$$

which asymptotically follows a standard normal distribution when the degrees of freedom ν is high [38] under the null hypothesis. Our empirical evaluation in Figure 1 suggests that the normalized FunChisq is effective at detecting functional dependency even if ν is small.

We also introduce the function index ξ_f to measure the effect size of FunChisq test:

$$\xi_f = \sqrt{\frac{\chi_f^2(O)}{n(c-1) - \sum_{j=1}^c \frac{(O_j - n/c)^2}{n/c}}} \quad (4)$$

The index assesses the strength of functional dependency of column variable Y on row variable X . It ranges from 0 to 1, with greater values representing stronger non-constant functionality. The index should be used in conjunction with the p -value of the test statistic to ensure both a sufficient effect and an acceptable statistical significance.

Abbreviations

AML:	Acute myeloid leukemia
CLL:	Chronic lymphocytic leukemia
DREAM:	Dialogue for reverse engineering assessments and methods
FANTOM5:	Functional annotation of mammalian genome 5
FunChisq:	Functional chi-squared test
HPN:	Heritage Provider Network
ncRNA:	Noncoding RNA
PR:	Precision recall
ROC:	Receiver operating characteristic
TCGA:	The cancer genome atlas
TPM:	Tags per million
TSS:	Transcription start site

Declarations

Ethics approval and consent to participate	32
Not applicable.	33

Consent for publication	34
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Availability of data and material	36
See Additional files.	37

Competing interests	38
The authors declare that they have no competing interests.	39

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Authors' contributions	45
HZ designed the study, wrote software, and performed data analysis. HZ and MS wrote the manuscript. All authors have read and approved the final manuscript.	46

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Author details	3
¹ Department of Computer Science, New Mexico State University, University Ave, 88003, Las Cruces, NM, USA.	4
² Molecular Biology Graduate Program, New Mexico State University, University Ave, 88003, Las Cruces, NM, USA.	5
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Figure captions

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Figure 1 FunChisq outperformed six widely-used methods in detecting known cancer genes from FANTOM5 data. FunChisq test, Pearson's chi-squared test, Wilcoxon test, *t*-test and logistic regression used transformed expression data. DESeq2 and edgeR used raw read count data. (a) ROC curves of each method. (b) PR curves of each method. (c) AUROC distributions after repeating the randomized evaluation 100 times. (d) AUPR distributions after repeating the randomized evaluation 100 times.

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Figure 2 Non-monotonic directional interaction patterns from two known cancer genes to the cancer status of human samples. The horizontal axes are log-scaled abundance of the most expressed TSS of each gene from FANTOM5. The vertical axes of the two top plots represent tissue types. 'Cancer' indicates a sample is from a cancer cell-line, 'Normal' for a sample from a non-cancer tissue. The vertical axes of the two bottom plots are the probability density of gene expression level. FunChisq reported high statistical significance of both genes' directional association with cancer suggested by the low *p*-values, while *t*-test returned insignificant results indicated by large *p*-values. (a) *p10KAT6A*, the most transcribed TSS of known cancer gene *KAT6A*, is either up- or down-regulated in 527 non-cancer samples of various tissues, but has an intermediate level of expression in 225 samples of various cancers. (b) *p10BRAF*, the most transcribed TSS of known cancer gene *BRAF*, has a similar non-monotonic expression profile directionally associated with cancer status.

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Figure 3 Distributions of function index measuring the directional association from TSSs to cancer status. The red curve is the distribution of the index from representative TSSs of known cancer genes to cancer status. The blue curve is the distribution of representative TSSs of non-cancer genes to cancer status. Cancer gene TSSs apparently have more larger index values than non-cancer gene TSSs, implying that the former group is more powerful than the latter group at predicting cancer status. About 90% of known cancer gene representative TSSs have an index value of less than 0.40, as indicated by the vertical red dashed line.

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Figure 4 Two unannotated transcription start sites predicted as putative cancer driver ncRNAs. The horizontal axes are log-scaled TSS expression from FANTOM5. The vertical axes of the two top plots represent tissue types. 'Cancer' indicates a sample is from a cancer cell-line, 'Normal' for a sample from a non-cancer tissue. The vertical axes of the two bottom plots are the probability density of gene expression level. (a) Putative cancer ncRNA *hg_112446.1* has a monotonic pattern with cancer status. (b) Putative cancer ncRNA *hg_195085.1* exhibits a non-monotonic pattern with cancer status.

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Figure 5 Gene expression patterns of four potential leukemia biomarkers are nearly exclusively cancer-type specific. The horizontal axes are TSS levels of gene expression from FANTOM5. The vertical axes are sample types. (a) Putative lymphoid leukemia biomarker *SNX9*. (b) Putative lymphoid leukemia biomarker *hg_153880.1*. (c) Putative myeloid leukemia biomarker *LMO2*. (d) Putative myeloid leukemia biomarker *hg_35610.1*.

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Figure 6 Chromosomal locations of putative leukemia biomarkers. Chromosomal counts of putative biomarkers for (a) lymphoid and (b) myeloid leukemia. Genomic maps of putative biomarkers for (c) lymphoid and (d) myeloid leukemia.

Additional files

Additional file 1—Additional_file_1.xlsx

FunChisq predicted 1108 putative cancer driver ncRNAs with stronger directional³ effect to cancer than 90% of 719 known cancer driver genes.

Additional file 2—Additional_file_2.xlsx

FunChisq predicted 332 potential cancer biomarkers for lymphoid leukemia and 79⁷ for myeloid leukemia, which were stronger than 87 known lymphoid leukemia and⁸ 100 known myeloid leukemia driver genes.

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