

1 This submission is intended as a Letter

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3 **New methods to calculate concordance factors for phylogenomic datasets**

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5 Bui Quang Minh^{1,2}, Matthew W. Hahn^{3,4}, and Robert Lanfear^{2,*}

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7 ¹ Research School of Computer Science, Australian National University, Canberra, ACT,

8 Australia

9 ² Department of Ecology and Evolution, Research School of Biology, Australian National

10 University, Canberra, ACT, Australia

11 ³ Department of Biology, Indiana University, Bloomington, Indiana, USA

12 ⁴ Department of Computer Science, Indiana University, Bloomington, Indiana, USA

13 * To whom correspondence should be addressed (rob.lanfear@anu.edu.au)

14

15 **Abstract**

16 We implement two measures for quantifying genealogical concordance in phylogenomic
17 datasets: the gene concordance factor (gCF) and the novel site concordance factor (sCF). For
18 every branch of a reference tree, gCF is defined as the percentage of “decisive” gene trees
19 containing that branch. This measure is already in wide usage, but here we introduce a
20 package that calculates it while accounting for variable taxon coverage among gene trees.
21 sCF is a new measure defined as the percentage of decisive sites supporting a branch in the
22 reference tree. gCF and sCF complement classical measures of branch support in
23 phylogenetics by providing a full description of underlying disagreement among loci and
24 sites. An easy to use implementation and tutorial is freely available in the IQ-TREE software
25 package (<http://www.iqtree.org>).

26 **Keywords:** phylogenetic inference, concordance factor

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29 Measures of branch support such as the bootstrap (Felsenstein 1985; Minh et al. 2013)
30 and Bayesian posterior probability (Ronquist and Huelsenbeck 2003) are important for
31 making robust inferences from phylogenetic trees. However, while these measures provide
32 useful information about the statistical support for a given branch, neither captures the
33 topological variation present in the underlying data (Kumar et al. 2012).

34 A complementary approach involves calculating the fraction of loci consistent with a
35 particular branch, thus capturing underlying agreement and disagreement in the data. Various
36 approaches have been suggested, including the gene concordance factor (gCF) (Baum 2007)
37 — sometimes also referred to as the ‘gene support frequency’ (Gadagkar et al. 2005; Salichos
38 and Rokas 2013) — and internode certainty (Salichos and Rokas 2013), with the former
39 being the most widely used. The gCF describes for each branch in a reference tree the
40 proportion of inferred single-locus trees that contain that branch. While intuitive, the gCF
41 suffers from three limitations: first, there are few software implementations (Ané et al. 2007);
42 second, while issues with incomplete taxon sampling of gene trees have been addressed by
43 some authors (Smith et al. 2015; Koberst et al. 2016), these fixes are not available in
44 implementations that allow for the calculation of gCF values; and third, low gCF values are
45 hard to interpret because they may result from strongly supported discordance among
46 individual gene trees or weak phylogenetic signal in individual loci.

47 Here, we resolve these issues by implementing the calculation of gCF while
48 accounting for unequal taxon sampling in the popular IQ-TREE package (Nguyen et al. 2015;
49 Minh et al. 2020), and by introducing the site concordance factor (sCF), a measure that
50 estimates concordance at the level of individual sites. We argue, and show by detailed
51 analyses of 10 empirical datasets, that concordance factors complement commonly used
52 metrics like the bootstrap by providing additional information and insights about topological
53 variation.

55 **Gene Concordance Factor (gCF)**

56 For a given *unrooted* bifurcating reference tree, T (e.g. an estimate of the species
 57 tree), and a set of *unrooted* bifurcating input trees, $S = \{T_1, \dots, T_n\}$ (e.g. individual gene trees),
 58 we can calculate the gCF for every internal branch x in T . Each gene tree, T_i , must contain a
 59 subset of the taxa in T , but need not include all of the same taxa.

60 Each internal branch x of T is associated with four clades of T representing four taxon
 61 subsets, A, B, C , and D , such that the bipartition $A \cup B | C \cup D$ corresponds to the branch x
 62 (Fig. 1). The reference tree T contains branch x by definition, but branch x may or may not
 63 be present in a gene tree of the same taxa (or a subset of them). Let A_i, B_i, C_i , and D_i be the
 64 set of taxa in A, B, C , and D that are also present in T_i . We say T_i is *decisive* for x if
 65 A_i, B_i, C_i, D_i are non-empty (we use ‘decisive’ in the sense of (Sanderson et al. 2011), but
 66 note for clarity that Dell'Ampio et al. (2014) use this term in a different way): i.e. a gene tree
 67 is *decisive* as long as it contains at least one of the taxa in each of the taxon subsets
 68 A, B, C , and D . A *decisive* gene tree can potentially contain the branch x . We say T_i is
 69 *concordant* with x if T_i is *decisive* for x and the bipartition $A_i \cup B_i | C_i \cup D_i$ corresponds to a
 70 branch in T_i . In other words, a gene tree is *concordant* with the reference tree if it could have
 71 contained branch x (i.e. it is *decisive*) and it does contain branch x . The gene concordance
 72 factor for branch x is now defined as:

$$73 \quad \text{gCF}(x) = \frac{|\{i: T_i \text{ is concordant with } x\}|}{|\{i: T_i \text{ is decisive for } x\}|}$$

74 In other words, the gCF is the proportion of input trees decisive for x that are concordant
 75 with x ; i.e. the proportion of all input trees that *could have* contained x that *do* contain x . We

76 note that gCF is also used to calculate internode certainty and related measures (Salichos et
77 al. 2014; Kobert et al. 2016).

78 To further understand *concordant* and *discordant* gene trees, we categorise the
79 *discordant* input trees (i.e. those that do not contain branch x) into three groups, and calculate
80 three *discordance* factors: gDF₁, gDF₂, and gDF_P. The first two groups are obtained
81 by applying a nearest neighbour interchange (NNI) around branch x to result in two
82 alternative topologies (Fig. 1) with two branches not appearing in T : $y = A \cup C | B \cup D$ and
83 $z = A \cup D | B \cup C$. Accordingly, we then calculate discordance factors as $gCF(y) = gDF_1(x)$
84 and $gCF(z) = gDF_2(x)$. Some input trees may not contain any of the branches x , y , or z . This
85 will occur for input trees in which one or more of clades A , B , C , or D are not monophyletic.
86 We define the proportion of *decisive* input trees that fall into this category as $gDF_P(x)$, where
87 the ‘P’ stands for ‘paraphyly’. The four proportions $gCF(x)$, $gDF_1(x)$, $gDF_2(x)$, and
88 $gDF_P(x)$ will sum to 1, as every *decisive* input tree must be included in one of the four
89 categories.

90

91 **Site Concordance Factor (sCF)**

92 To calculate the sCF for branch x , we randomly sample m quartets of taxa $q =$
93 $\{a, b, c, d\}$, where a, b, c , and d are in A, B, C , and D , respectively. For each quartet q , we
94 examine the sub-alignment of taxa a, b, c, d . For every site j in this alignment, we call j
95 *decisive* for x if the characters a_j, b_j, c_j , and d_j are all present and j is parsimony informative
96 when restricted to this quartet of taxa. Decisive sites can be concordant or discordant with x .
97 We say that site j is *concordant* with x if $a_j = b_j \neq c_j = d_j$ (i.e., j supports the bipartition
98 $\{a, b\} | \{c, d\}$). The concordance factor for q is defined as:

99
$$\text{CF}_q(x) = \frac{|\{j: j \text{ is concordant with } x\}|}{|\{j: j \text{ is decisive for } x\}|}$$

100 Since many such quartets exist around branch x (when sampling individual tips from
 101 within A, B, C , and D), we define $\text{sCF}(x)$ as the mean $\text{CF}_q(x)$ over m random quartets:

102
$$\text{sCF}(x) = \frac{1}{m} \left(\sum_q \text{CF}_q(x) \right)$$

103 Thus, the sCF is the average proportion of sites decisive for x that are concordant with x . In
 104 effect, the sCF is a measure of concordance for sites that is directly comparable to the
 105 measure of concordance for single-locus trees provided by the gCF . Unlike related quartet
 106 approaches that are based on maximum likelihood tree inference and that calculate
 107 discordance at the level of the whole alignment (Strimmer and von Haeseler 1997; Pease et
 108 al. 2018), the sCF uses parsimony criteria to calculate discordance at the level of individual
 109 sites. We note that the sCF is closely related to the values derived from spectral analysis
 110 (Hendy and Penny 1993; Charleston 1998). Spectral analysis is a tree-independent method of
 111 investigating phylogenetic signal. In the simple case of binary data, such that each site
 112 corresponds to a bipartition of taxa, the spectral support for a single branch counts the
 113 number of sites that correspond to the bipartition defined by that branch. In the case of binary
 114 data, this is the same as the value $\text{CF}_q(x)$ that we define above. The sCF in this case differs
 115 from the spectral support insofar as it is calculated by averaging over a large set of $\text{CF}_q(x)$
 116 values calculated by repeatedly subsampling quartets of taxa from an alignment.

117 Similarly to gene discordance factors, we also define the site discordance factors
 118 $\text{sCF}(y) = \text{sDF}_1(x)$ and $\text{sCF}(z) = \text{sDF}_2(x)$. There is no sDF_P category, because unlike the
 119 gDF values, every *decisive* site must contribute to one of the three proportions $\text{sCF}(x)$,

120 $sDF_1(x)$, or $sDF_2(x)$; any quartet of taxa resolves into exactly three tree topologies shown in
121 Figure 1. In other words, the sum of $sCF(x)$, $sDF_1(x)$, and $sDF_2(x)$ will always be 1.

122

123 **gCF and sCF for rooted trees**

124 The above definitions of gCF and sCF apply only to unrooted trees. However, in the
125 case that users have a rooted reference tree as well as rooted gene trees, we can extend the
126 calculation of the gCF to allow us to calculate different gCF values on either side of the root.
127 To do this we first add a virtual root node into the rooted reference tree T , resulting in an
128 unrooted tree T' . Similarly, we convert the rooted gene trees T_i into unrooted trees T'_i by
129 adding the same virtual root. This allows us to then follow the same procedure as above for
130 calculating gCF values on unrooted trees, with the sole difference that the output is a rooted
131 instead of an unrooted reference tree. In this case, it is possible to calculate gCF values on a
132 rooted reference tree where the gCF values on either side of the root may differ. In all other
133 cases (i.e. for sCF values, and if either the reference tree or the gene trees are unrooted), it is
134 not possible to calculate different concordance factors on either side of the root.

135

136 **Implementation in IQ-TREE 2**

137 We provide two new options, `--gcf` and `--scf`, in IQ-TREE version 2 {Minh 2020} to
138 compute gCF and sCF respectively. A tutorial for how to use these options is provided at
139 <http://www.iqtree.org/doc/Concordance-Factor>. Both options can be combined in a single
140 run, which will calculate the gCF and sCF for every branch in the input reference tree. While
141 sCFs can be calculated on any alignment, the calculation of gCFs requires individual gene
142 trees. We have therefore implemented a convenient option, `-S`, to specify a partition file or a

143 directory of single-locus alignments in which IQ-TREE will infer separate trees for each
144 partition or alignment (Minh et al. 2020).

145 IQ-TREE provides a suite of output files to assist users in understanding and
146 investigating gCF and sCF values. It provides tree files that can be viewed in most tree
147 viewers, which contain information on both the proportional data (gCF, gDF₁, gDF₂, gDF_P;
148 sCF, sDF₁, sDF₂) and their corresponding absolute count data (gCF_N, gDF₁_N, gDF₂_N,
149 gDF_P_N; sCF_N, sDF₁_N, sDF₂_N), as well as the number of decisive genes and sites for
150 each branch (gN and sN, respectively). It also provides a tree file that combines information
151 on the gCF, sCF, and any bootstrap values that have been calculated (e.g. Figure 2). In
152 addition to this, IQ-TREE provides a ‘.cf.stat’ file that contains all 16 concordance and
153 discordance values listed above for every branch in the reference tree in a machine-readable
154 tabulated format, and through the --cf-verbose option it provides tabulated files that detail for
155 every branch in the reference tree whether each gene tree was concordant with that branch
156 (the ‘.cf.stat_tree’ file produced from a --gcf analysis), and the average number of sites in
157 each locus that were concordant with that branch (the ‘.cf.stat_loci’ file produced from a --scf
158 analysis). Together these output files provide both a convenient overview of the data, and the
159 opportunity to understand in much more detail the extent to which each locus is concordant
160 or discordant with the reference tree.

161

162 **Application to empirical datasets**

163 To demonstrate the use of gCF and sCF values, we first analysed a dataset containing
164 3,220 ultra-conserved elements (UCEs) from lizards, totalling 1,301,107 bp for 43 species
165 (Rodriguez et al. 2018). To do this, we estimated a concatenated maximum likelihood (ML)
166 tree, 3,220 UCE trees, and the gCF, sCF, and bootstrap values with IQ-TREE (Figure 2).

167 To investigate the relationships between dataset size, concordance factors, and
168 bootstrap values, we also analysed a collection of nine additional empirical phylogenomic
169 datasets that represent a range of clades and data types (Table 1). For each dataset, we
170 performed 20 analyses across a 20-fold range of dataset sizes: the first analysis included 10
171 randomly selected loci from the complete dataset, and each subsequent analysis added 10
172 more randomly selected loci up to a maximum of 200 loci. For each such analysis, we
173 calculated the gCF, sCF, and bootstrap values for every branch of the concatenated ML tree
174 estimated from the 200-locus dataset, with the same approach as above. For the lizard dataset,
175 we calculated both standard bootstrap (StdBoot) values (Felsenstein 1985) and ultrafast
176 bootstrap (UFBoot) values (Hoang et al. 2018). As expected (Minh et al. 2013), we observed
177 that StdBoot and UFBoot values were very similar (Figure 3B, supplementary figure 8D);
178 therefore, because of the high computational cost of calculating StdBoot values, we
179 calculated only UFBoot values for the remaining nine datasets. Since the results of all 10
180 analyses are very similar, we focus here on the lizard dataset, and present the results for the
181 other 9 datasets in the supplementary figures.

182 Reproducible analyses are provided in the supplementary material available at
183 <https://doi.org/10.5281/zenodo.1949287>.

184

185 **Interpretation of gCF and sCF values**

186 Our analyses highlight a number of important differences between concordance
187 factors and bootstrap values. First, concordance factors of any type tend to be much lower
188 than bootstrap values or other measures of statistical support for branches in the species tree
189 (e.g. Fig 3, supplementary figures 1 to 10). This is because bootstrap values measure the
190 sampling variance in support of a focal branch, while gCF and sCF values measure the

191 underlying variance in support of that branch at the gene- and site-level respectively. In other
192 words, resampled datasets may always return the same tree (i.e. 100% bootstrap support),
193 even though incomplete lineage sorting or other processes that lead to genealogical
194 discordance are at work (e.g. gCF and sCF values <<100%). Of particular note is that a very
195 high bootstrap value (e.g. of 100%) does not predict or require a similarly high concordance
196 factor. For example, although all but one bootstrap value on the tree in Figure 1 is 100%, the
197 smallest gCF and sCF values are 4.5% and 33.2% respectively. This pattern is repeated
198 across all 200 analyses we performed on all 10 empirical datasets (Figure 3, supplementary
199 figures 1 to 10).

200 Because they measure different things, concordance factors and bootstrap values are
201 affected very differently by the addition of loci to a dataset (Figure 3, supplementary figures
202 1 to 10). In all of the 10 empirical datasets that we analysed, adding loci to the dataset tended
203 to increase the bootstrap values of individual branches (e.g. Figure 3A, see also
204 supplementary figures) and concomitantly the proportion of branches with a bootstrap value
205 of 100% (Figure 3B, see also supplementary information). This is expected: the bootstrap is
206 effectively measuring the standard error of the mean in a dataset (where the mean represents
207 the tree inferred from the full dataset), and this standard error will go down with more
208 samples. The same is not true of concordance factors, which display some estimation error
209 when datasets are small (e.g. 30 or fewer loci in Figure 3), but subsequently remain almost
210 completely insensitive to the addition of loci to the dataset (Figure 3, supplementary figures
211 1-10). This is also expected: any measure of the underlying variance (or standard deviation)
212 in a distribution will have some estimation error when the sample size is small, but should not
213 change monotonically as the sample size increases. These differences highlight the
214 complementary information that can be gained from calculating both bootstrap values and
215 concordance factors for individual branches. One measure is not better or worse than the

216 other, rather, concordance factors provide useful information that bootstraps do not, and vice
217 versa. Note that bootstrap values or posterior probabilities calculated by resampling gene
218 trees (e.g., Sayyari and Mirarab 2016) have the same behaviour as the site-wise bootstrap
219 carried out here, and are also not equivalent to concordance factors.

220 In principle, both gCF and sCF values can range from 0% (no genes / sites are
221 concordant with the focal branch) to 100% (all genes / sites are concordant with the focal
222 branch). In practice however, as exemplified in Figure 3 and supplementary figures 1 to 10,
223 empirical gCF values tend to range from 0% to 100%, while empirical sCF values are rarely
224 lower than 33% (represented by the dashed line in the sCF panel of Figure 3A). This is due to
225 an important underlying difference in the way that the two values are calculated. The sCF is
226 calculated from quartets, so a single site can only support one of three topologies (Figure 1).
227 Because of this, if there is no consistent information in an alignment (e.g. if a long alignment
228 were generated at random) we expect a roughly equal proportion of sites supporting each of
229 the three trees, leading to an sCF value of approximately 33% (for the same reason, sCF
230 values for very long branches will approach 33% due to saturation). The same is not true for
231 gCF values, because a gene tree can support not only the three possible relationships shown
232 in Figure 1, but any other relationship in which one or more of clades *A*, *B*, *C*, or *D* is not
233 monophyletic. The higher the number of gene trees in this latter group, the closer the gCF
234 value will be to 0%. Because of this, we should expect gCF values to be particularly low
235 when gene trees are estimated from alignments with limited information or where branch *x* is
236 extremely short; in such cases either technical or biological processes may increase the
237 proportion of gene trees that fail to recover the monophyly of clades *A*, *B*, *C*, or *D* found in
238 the reference tree. Missing data may also impact gCF and sCF values, however, the
239 relationship here is less clear. The requirement that a gene tree or a site is decisive (i.e. could
240 in principle contain branch *x*, see above) should limit the impact of missing data on gCF and

241 sCF estimates. Nevertheless, since phylogenetic estimates are known to worsen as the
242 proportion of missing data increases (e.g., Roure et al. 2013; Xi et al. 2016), it is plausible
243 that concordance factors may systematically decrease as the proportion of missing data
244 increases.

245 Finally, cases where the sCF value is lower than 33% may be of particular interest.
246 These cases are, by definition, those in which maximum parsimony (MP) would favour a
247 different resolution of a split found in the reference tree. If the reference tree was calculated
248 from any method other than MP, there are at least two explanations for an sCF value lower
249 than 33%. First, the branch of interest may be in an area of parameter space in which high
250 levels of incomplete lineage sorting are known to mislead concatenated ML analyses
251 (Kubatko and Degnan 2007) but not MP analyses (Mendes and Hahn 2018). Misleading
252 reference trees may therefore be produced by either concatenated ML of the entire dataset, or
253 by gene tree methods that use shorter sets of concatenated loci as their input (because most
254 protein-coding genes are themselves made up of multiple topologies). Second, and more
255 generally, there are multiple reasons why likelihood, Bayesian, or gene tree methods for
256 producing a species tree will differ from MP resolutions. For instance, the branch of interest
257 may be unduly affected by a small number of highly influential sites in a concatenated ML
258 analysis (Shen et al. 2017). In this case, the influential sites can have an outsized influence on
259 the ML resolution of a split because they have extreme differences in likelihood between
260 different resolutions of that split. Because MP does not account for likelihood differences—it
261 instead weights all sites equally—MP analyses remain unaffected by such outliers. Thus,
262 cases in which the sCF is much lower than 33% may merit further investigation.

263 We hope that the user-friendly implementation of gene- and site-concordance factors
264 in IQ-TREE will assist researchers in gaining additional insights into their phylogenetic
265 reconstructions. In particular, we encourage phylogeneticists to calculate both bootstrap

266 values and concordance factors for the branches on their trees, as the two measures provide
267 complementary information that may help to improve the accuracy of our interpretations of
268 phylogenetic reconstructions. Indeed, the use of concordance factors may help to alleviate the
269 commonly cited problem in phylogenomics that bootstrap values provide relatively little
270 information when they are all 100% (Kumar et al. 2012).

271

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279 **Supplementary Materials**

280 Data are available at <https://doi.org/10.5281/zenodo.1949290>

281 Supplementary figures are available from the journal website.

282

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365

366 **Table 1. The five DNA and five amino acid (AA) datasets analysed in this study.**

No	Reference	Type	Clade	Taxa	Loci	Sites
1	(Ballesteros and Sharma 2019)	AA	Chelicerata	53	3,534	1,484,206
2	(Branstetter et al. 2017)	DNA	Aculeata	187	807	183,747
3	(Cannon et al. 2016)	DNA	Metazoa	78	424	89,792
4	(Jarvis et al. 2015)	AA	Aves	52	8,295	4,519,041
5	(Misof et al. 2014)	AA	Insecta	144	2,868	595,033
6	(Ran et al. 2018)	AA	Spermatophyta	38	1,308	432,014
7	(Ran et al. 2018)	DNA	Spermatophyta	38	3,924	1,296,042
8	(Rodriguez et al. 2018)	DNA	Prasinohaema	43	3,220	1,301,107
9	(Wu et al. 2018)	AA	Mammalia	90	5,162	3,050,199
10	(Wu et al. 2018)	DNA	Mammalia	90	15,486	9,150,597

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369 **Figure legends**

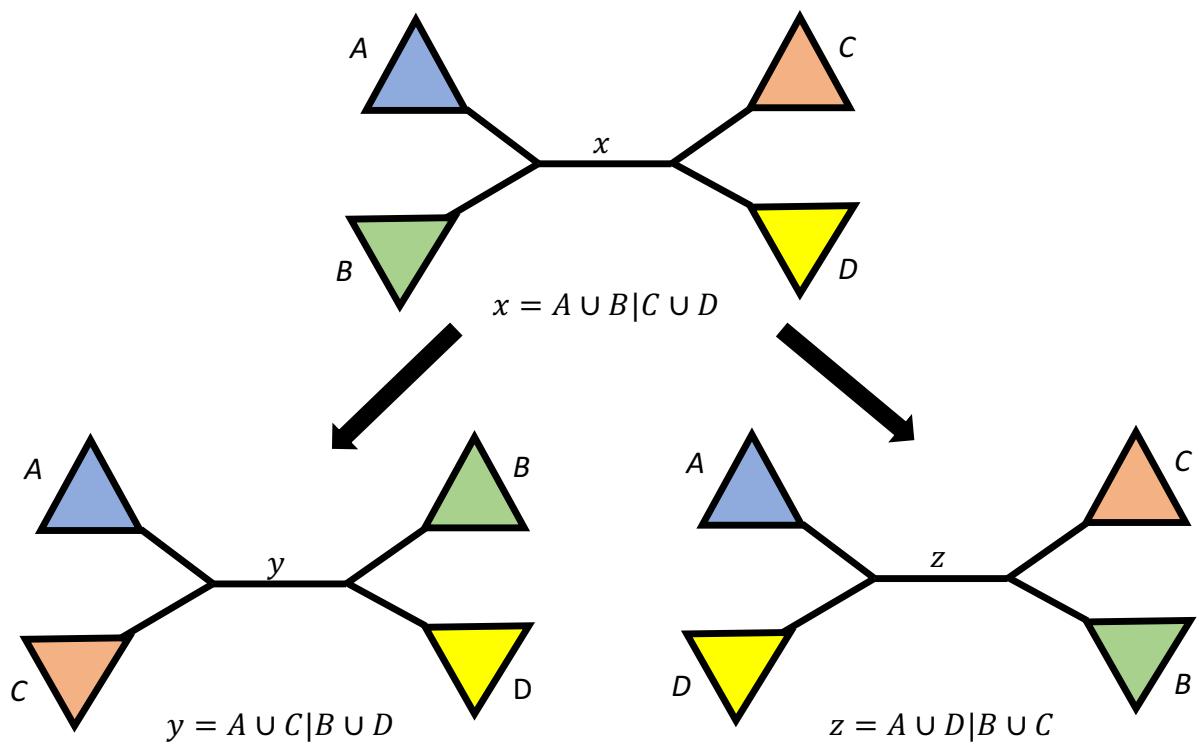
370 **Figure 1.** Schematic view of a bifurcating tree at an internal branch x with four surrounding
371 sub-trees containing the taxon sets A , B , C , and D . Branch x corresponds to the bipartition
372 $A \cup B | C \cup D$. In our definition, branch x is always present in the reference tree, but it may or
373 may not be present in each of the input trees. By applying the two nearest neighbour
374 interchanges (denoted as arrows) one can produce the two other trees that contain taxon sets
375 A , B , C , and D , but with internal branches y and z .

376 **Figure 2.** An example of concordance factors on a dataset of lizards (Rodriguez et al. 2018).
377 A cladogram is shown to facilitate the plotting of concordance factors on branches. Numbers
378 on each branch show the site concordance factor (sCF) above the branch (e.g. s73) and the
379 gene concordance factor (gCF) below the branch (e.g. g37). Bootstrap values are 100% on
380 every branch, except for the branch leading to *Lipinia rouxi* and *L. pulchella*, which has a
381 bootstrap value of 62%. The inset shows a scatter plot of gCF values against sCF values for
382 all branches, revealing the large range of gCF and sCF values as well as the fact that for this
383 dataset sCF values are always at least as large as gCF values. This is likely because of the
384 short length of the UCE loci used to infer gene trees (Rodriguez et al. 2018).

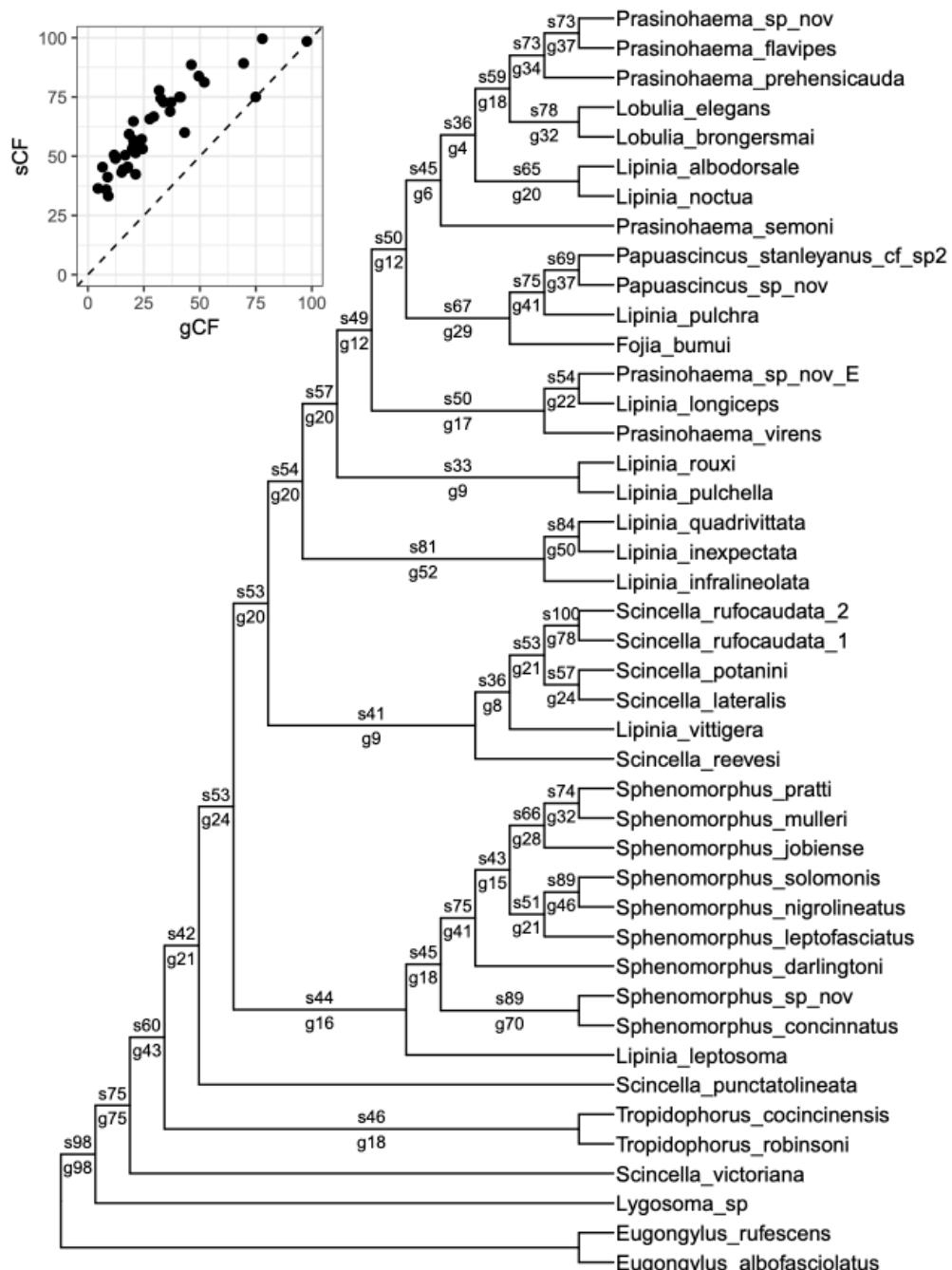
385 **Figure 3.** Concordance factors remain relatively stable as loci are added to an analysis, while
386 bootstrap values continue to increase towards 100%. A) The results of 20 reanalyses of the
387 lizard dataset, each of which adds a further ten loci to the analysis, up to a maximum of 200
388 loci (x-axis). Each coloured line represents a different branch in the tree. The dashed line on
389 the sCF panel shows a value of 33%. B) For each of the four metrics considered here
390 (represented by different coloured lines), the number of loci included in the analysis affects
391 the proportion of the branches that have a value of 100%.

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Figure 1



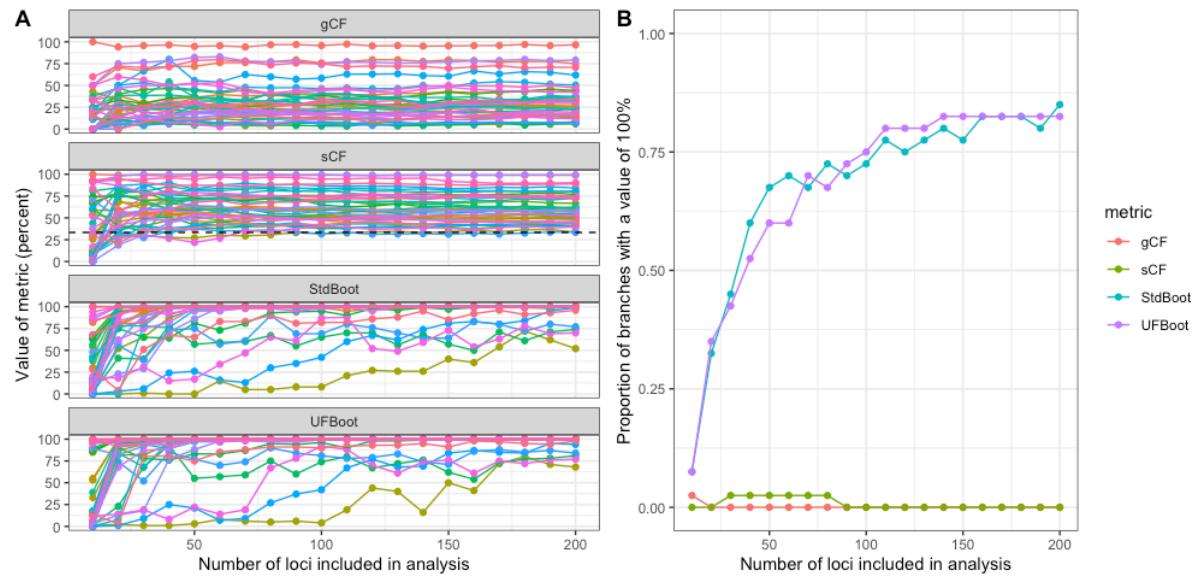
395 **Figure 2**



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398 **Figure 3**



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