

# **Twitter reveals human mobility dynamics during the COVID-19 pandemic**

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## **Abstract:**

The current COVID-19 pandemic raises concerns worldwide, leading to serious health, economic, and social challenges. The rapid spread of the virus at a global scale highlights the need for a more harmonized, less privacy-concerning, easily accessible approach to monitoring the human mobility that has proven to be associated with viral transmission. In this study, we analyzed over 580 million tweets worldwide to see how global collaborative efforts in reducing human mobility are reflected from the user-generated information at the global, country, and U.S. state scale. Considering the multifaceted nature of mobility, we propose two types of distance: the single-day distance and the cross-day distance. To quantify the responsiveness in certain geographic regions, we further propose a mobility-based responsive index (MRI) that captures the overall degree of mobility changes within a time window. The results suggest that mobility patterns obtained from Twitter data are amenable to quantitatively reflect the mobility dynamics. Globally, the proposed two distances had greatly deviated from their baselines after March 11, 2020, when WHO declared COVID-19 as a pandemic. The considerably less periodicity after the declaration suggests that the protection measures have obviously affected people's travel routines. The country scale comparisons reveal the discrepancies in responsiveness, evidenced by the contrasting mobility patterns in different epidemic phases. We find that the triggers of mobility changes correspond well with the national announcements of mitigation measures, proving that Twitter-based mobility implies the effectiveness of those measures. In the U.S., the influence of the COVID-19 pandemic on mobility is distinct. However, the impacts vary substantially among states.

**Keywords:** Twitter, mobility, COVID-19, human movement, social media, big data

# 1. Introduction

The outbreak of Coronavirus disease (COVID-19) caused by the SARS-CoV-2 virus is a public health emergency that raises concerns worldwide, leading to serious health, economic, and social challenges. As of June 23, 2020, there had been a total of 8,993,659 infections and 469,587 deaths globally [1], and these figures are progressively increasing every day. On March 11, 2020, the World Health Organization (WHO) reassessed the situation and officially declared COVID-19 as a pandemic, urging countries and regions worldwide to join forces [2]. Since then, major behavioral, clinical, and intervention policies (both strict and loose) have been undertaken to reduce the spread and prevent the persistence of the virus in human populations.

An initial outbreak of COVID-19 was first declared in Wuhan, China in January 2020 [2], before cases were reported in European countries, most notably Italy, France, and the UK. In the United States, the first confirmed case occurred on January 19, 2020, in Snohomish County, Washington. Shortly after, the U.S. has become the new epicenter of the disease as it surpassed Italy in terms of confirmed cases on March 26, 2020 [3]. As of June 23, 2020, there had been a total of 2,268,753 confirmed cases (25.2% of global cases) and 119,761 deaths (25.5% of global deaths) in the U.S. alone [1]. To contain the COVID-19 pandemic, one of the non-pharmacological epidemic control measures is to reduce the transmission rate of SARS-COV-2 in the population via social distancing or other similar (self) quarantine measures [4], with the ultimate goal to reduce person-to-person interactions. Studies have found notable declines in transmission rates after the implementation of mobility-reducing policies in China, Korea, and many European countries [5-8]. Despite the success of these efforts, not all countries/regions chose to handle the pandemic in a similar manner [9, 10]. The discrepancies in policies and measures at different geographical levels urge an approach to monitoring the mobility

dynamics in response to the pandemic, as mobility patterns largely indicate how people respond to the pandemic and whether policies are implemented effectively [11-13].

Since the initial outbreak of COVID-19, numerous efforts have been made, by incorporating the emerging concept of “Web 2.0” [14], “Big Data” [15], and “Citizen as Sensors” [16], to obtain timely information regarding whether people are actively reducing their exposure to COVID-19 by reducing distances traveled, and by how much. Companies like Google and Apple have released their aggregated and anonymized community mobility reports based on data collected from their services (i.e., Google Location Services and Apple Maps). Those reports are updated on a daily basis and can be easily downloaded. In addition, authorities started to collaborate with mobile network operators to estimate and visualize the effectiveness of control measures [11, 17], in light of the previous success of mobile phone data in assisting the modeling of the spread of other epidemics [18-20]. Shortly after the outbreak in China, mobility data from Baidu, a famous Chinese online platform, have been put into use to evaluate the effectiveness of the lockdown measure in Wuhan [5]. Leading telecommunication firms also contribute by collaborating with local authorities to estimate the efficiency of travel restrictions as well as to identify the impact of other mobility-reducing related measures [21, 22]. In the U.S., Descarte Lab ([www.descarteslabs.com/mobility](http://www.descarteslabs.com/mobility)) has released mobility statistics derived from mobile devices, aiming to facilitate the acquisition of rapid situational awareness at the State- and County-level. City-level studies have also been conducted. For example, locational data from Cuebiq (<https://www.cuebiq.com/>), gathered via over 180 mobile applications, were used to monitor how social distancing guidelines are implemented on a daily basis in the city of Boston, MA [11]. However, privacy advocates have voiced concerns on whether sharing customer data is appropriate, even in a time of crisis [23, 24]. The rapid spread of the COVID-19 at the global level highlights the need

for a more harmonized, less privacy-concerning, easily accessible approach to monitoring human mobility.

The rise of social media platforms such as Twitter ([twitter.com](https://twitter.com)), Flickr ([www.flickr.com](https://www.flickr.com)), and Instagram ([www.instagram.com](https://www.instagram.com)) offers another possible solution to closely monitoring human mobility changes, thanks to the timely geospatial information from the enormous sensing network constituted by millions of users. The huge volume of user-generated content from social media platforms greatly facilitates the real-time or near real-time monitoring of human mobility, providing timely data of how people respond to the COVID-19 pandemic geographically, especially within different epidemic phases. The advantages of social media with respect to the aforementioned sources of digital information are that they are extensive (covering large spatial areas), easily accessible, with less privacy concern, and at low cost [25-28]. Extracting useful information from social media is not new, as the valuable geospatial insights from social media have been explored in a wide range of fields, including hazard mitigation [29-31], evacuation monitoring [27, 32, 33], urban analytics [34-37], and public health [38, 39], to list a few. Despite the existing applications, the potential of human mobility derived from social media data has not been fully explored. Questions like whether the mobility data from social media can quantitatively reflect the collaborative effort in fighting the COVID-19 pandemic and how it corresponds to the everchanging policies in different geographical regions deserve answers.

To answer the above questions, we focus on Twitter, a popular social media platform, and analyze over 580 million tweets from all over the world to see how the worldwide collaborative efforts in reducing mobility are reflected from this user-generated information in three different scales: global scale, country scale, and Conterminous U.S. (CONUS) state scale. We propose two types of distance, respectively

termed as single-day distance and cross-day distance, to quantify different aspects of public mobility observed from Twitter. We further normalize these distances by setting up their corresponding baselines. The baseline mobility values are calculated for each region separately (globe, countries, and U.S. states) using data collected before strict mobility-reducing measures are implemented. To quantify the responsiveness in certain geographic regions within different epidemic phases, we propose a mobility-based responsive index (MRI) to capture the overall degree of mobility changes in response to the COVID-19 pandemic within a specific time window. Finally, we contextualize the mobility dynamics derived from Twitter with detailed measures from local authorities to shed light on their effectiveness. The theoretical, methodological, and contextual knowledge in this study is expected to inspire future applications of these easily accessible, less privacy-concerning, highly spatiotemporal data.

## **2. Datasets and computing environment**

We collect a total of 583,748,902 geotagged tweets from 10,324,191 unique Twitter users using the official Twitter Streaming Application Programming Interface (API), comprising a five-month period from January 1, 2020 to May 31, 2020. These tweets are stored and queried in a tweet repository managed in an in-door Hadoop cluster with 13 servers using Apache Hive and Impala. A geotagged tweet is a Twitter post with embedded geolocation in the format of exact coordinates (latitude and longitude) from the device’s GPS or placenames (e.g., state, county, city). While the locational accuracy of a geotagged tweet varies, depending on the settings of the account and how a user chooses to share his/her location, we exclude the tweets that are geotagged with spatial resolution lower than the city level to increase the accuracy and credibility of the mobility pattern. Following Martin et al. [40], we filter out the non-human tweets (e.g., automated

weather reports, job offers, and advertising) by checking the tweet source from which application a tweet is posted. For example, tweets automatically posted for job offers from the source TweetMyJOBS and CareerArc are removed. After the filtering, a total of 496,068,100 tweets remain from 9,502,266 unique Twitter users (details of the user count with the number of tweets posted per day can be found in the supporting information S2 Fig). The computation of travel distance requires locational information from at least two positions. Thus, only users who post tweets on two consecutive days are included in the calculation of cross-day distance, a measure that quantifies displacement between two consecutive days (details in Section 3.1). For the single-day distance, a measure that highlights the daily travel pattern, only users who post at least twice a day are included in the calculation (details in Section 3.1). In our dataset, 53.7% of users tweet at least twice a day (for the calculation of single-day distance), while 49% of users tweet cross-day (for the calculation of cross-day distance). Note that all distances computed in this study are Great Circle distances.

### 3. Methods

#### 3.1 Single-day distance and cross-day distance

To quantify daily human mobility from collected Twitter data, we propose two different types of distance, respectively referred to as single-day distance ( $D_{sd}$ ) and cross-day distance ( $D_{cd}$ ). The concepts of the two distances are presented in Fig 1. To reduce the computational complexity, the calculation of  $D_{sd}$  is adopted and modified from Warren and Skillman [41]. In general,  $D_{sd}$  represents the users' daily maximum travel distance of all locations relative to the initial location. Its calculation is confined within a single day so that users' daily travel patterns can be revealed. Different from  $D_{sd}$ ,  $D_{cd}$  measures the mean center shift between two consecutive days.

**Fig 1. Conceptualization of single-day distance ( $D_{sd}$ ) and cross-day distance ( $D_{cd}$ ).**

For a selected Twitter user  $i$ , let  $P_{i,j}^m = \{P_{i,j}^1, P_{i,j}^2, \dots, P_{i,j}^n\}$  denote the collection of locations derived from his/her tweets within a certain day  $j$ . Among the total of  $n$  locations in day  $j$ ,  $P_{i,j}^1$  denotes the initial location and  $P_{i,j}^m$  always precedes  $P_{i,j}^{m+1}$  in time. To compute  $D_{sd}$ , a collection of location pairs (A) is first formed by coupling  $P_{i,j}^m$  with the initial location  $P_{i,j}^1$ , i.e.,  $A = \{(P_{i,j}^1, P_{i,j}^2), (P_{i,j}^1, P_{i,j}^3), \dots, (P_{i,j}^1, P_{i,j}^n)\}$ . The Great Circle Distance (GCD) is applied to compute the distance of each location pair within collection A. For a given location pair  $(P_{i,j}^1, P_{i,j}^m)$ , their GCD can be represented as  $GCD_{i,j}^{1,m}$ .  $D_{sd}$  for user  $i$  in day  $j$ , referred to as  $D_{sd_{i,j}}$ , is computed by selecting the maximum value of  $GCD_{i,j}^{1,m}$ , i.e.,  $D_{sd_{i,j}} = \max \{GCD_{i,j}^{1,2}, GCD_{i,j}^{1,3}, \dots, GCD_{i,j}^{1,n}\}$ . To compute  $D_{cd}$ , for a collection of locations from user  $i$  in day  $j$ , i.e.,  $\{P_{i,j}^1, P_{i,j}^2, \dots, P_{i,j}^n\}$ , a mean center ( $\overline{P_{i,j}}$ ) is first calculated by respectively averaging the coordinates of locations in  $\{P_{i,j}^1, P_{i,j}^2, \dots, P_{i,j}^n\}$ :

$$\overline{P_{i,j}} = \mu\{P_{i,j}^1, P_{i,j}^2, \dots, P_{i,j}^n\} \quad (1)$$

where  $\overline{P_{i,j}}$  denotes the mean center for user  $i$  in day  $j$  and  $\mu$  denotes the mean center operator.  $D_{cd}$  for user  $i$  in day  $j$ , referred to as  $D_{cd_{i,j}}$ , is the GCD between  $\overline{P_{i,j}}$  and  $P_{i,j+1}$ .

Intuitively,  $D_{sd}$  and  $D_{cd}$  represent different aspects of mobility with  $D_{sd}$  measuring maximum single-day travel distance and  $D_{cd}$  measuring cross-day displacement. The dynamics of  $D_{sd}$  and  $D_{cd}$  are expected to reflect on how the COVID-19 pandemic affects people's mobility patterns geographically, presumably indicating the regional degree of responsiveness.



### 3.2 Normalized mobility index

Inspired by the methodological design in mobility reports from Google ([www.google.com/covid19/mobility](http://www.google.com/covid19/mobility)) and Apple ([www.apple.com/covid19/mobility](http://www.apple.com/covid19/mobility)), we set up baselines for  $D_{sd}$  and  $D_{cd}$  respectively. Unlike studies that utilize a single baseline value summarized from a fixed period, our mobility baselines are set for each corresponding day of a week, as a week has been widely recognized as an independent cycle in mobility [25, 42]. That is to say, we calculate a total of fourteen baseline values, seven for  $D_{sd}$  and seven for  $D_{cd}$ , corresponding to each day of a week. For a geographical region  $\mathbb{R}$  (globe, a country, or a state), let  $D_{sd_j}^{\mathbb{R}}$  and  $D_{cd_j}^{\mathbb{R}}$  represent the  $D_{sd}$  and  $D_{cd}$  of  $\mathbb{R}$  in day  $j$ , respectively. We define that  $D_{sd_j}^{\mathbb{R}}$  is the mean value of all  $D_{sd_{i,j}}^{\mathbb{R}}$  in day  $j$ , i.e.,

$$D_{sd_j}^{\mathbb{R}} = \frac{\sum_i D_{sd_{i,j}}^{\mathbb{R}}}{N}, \text{ where } N \text{ denotes the total number of selected users in day } j \text{ within } \mathbb{R} \text{ and}$$

$$P_{i,j}^1 \in \mathbb{R}. \text{ Similarly, } D_{cd_j}^{\mathbb{R}} \text{ is the mean value of all } D_{cd_{i,j}}^{\mathbb{R}} \text{ in day } j, \text{ i.e., } D_{cd_j}^{\mathbb{R}} = \frac{\sum_i D_{cd_{i,j}}^{\mathbb{R}}}{N},$$

where  $\overline{P_{i,j}} \in \mathbb{R}$ . Consequently, the normalized mobility index of region  $\mathbb{R}$  in day  $j$  for

single-day distance ( $NMI_{sd_j}^{\mathbb{R}}$ ) and cross-day distance ( $NMI_{cd_j}^{\mathbb{R}}$ ) are respectively defined

as the ratios of  $D_{sd_j}^{\mathbb{R}}$  and  $D_{cd_{i,j}}^{\mathbb{R}}$  to their baseline values of a corresponding day in a week.

Given their calculations,  $NMI_{sd_j}^{\mathbb{R}}$  and  $NMI_{cd_j}^{\mathbb{R}}$  both have a range of  $[0, +\infty)$ , with 1 being

the critical value. When the  $NMI_{sd_j}^{\mathbb{R}}$  (or  $NMI_{cd_j}^{\mathbb{R}}$ ) is less than 1, it suggests that within

region  $\mathbb{R}$  in day  $j$ , reduced mobility is observed compared with the baseline mobility

when measuring single-day distance (or cross-day distance).

### 3.3 Mobility-based responsive index

After the normalization in the previous section, a baseline of  $NMI$  (i.e.,  $NMI = 1$ ) that

separates patterns of increased mobility and reduced mobility is formed. Intuitively, for a

time series of *NMI* values, the size of the area under the *NMI* baseline ( $S_{AUB}$ ) represents the degree of positive responses (i.e., reduce in mobility) for a given period, while the size of the area above the *NMI* baseline ( $S_{AAB}$ ) indicates otherwise (Fig 2). Hypothetically, the area in perfect condition ( $S_{APC}$ ) represents a perfect scenario where mobility instantly reduced to 0 from the beginning and remains 0 until the time series ends. Apparently, such a scenario is purely theoretical and certainly does not exist in the real world. However, it provides a baseline where other scenarios are compared against, facilitating the quantification of how close other scenarios are compared to the perfect scenario. Conceptually, the mobility-based responsive index we propose is the ratio between the net positive response to the perfect scenario, i.e.,  $\frac{\sum S_{AUB} - \sum S_{AAB}}{S_{APC}}$ , where  $\sum S_{AUB}$  and  $\sum S_{AAB}$  respectively denote the summation of areas under the curve and the summation of areas above the curve, given a specific period.

### Fig 2. Mobility-based responsive index

To remove noises and reveal the general trend, we smooth the time series using a one-dimensional Gaussian filter ( $\sigma = 2$ ), one of the most popular filters widely applied in many temporal smoothing tasks [43, 44]. Further calculations regarding the size of the areas are all based on the smoothed time series. Given the different nature of the two proposed distances, we calculate their *MRI* separately:

$$MRI_{sd} = \frac{\sum S_{AUB_{sd}} - \sum S_{AAB_{sd}}}{S_{APC}} \quad (2)$$

$$MRI_{cd} = \frac{\sum S_{AUB_{cd}} - \sum S_{AAB_{cd}}}{S_{APC}} \quad (3)$$

where  $MRI_{sd}$  and  $MRI_{cd}$  denote the *MRI* with  $D_{sd}$  and  $D_{cd}$  being measured, respectively. We further compute an integrated *MRI* by weighting  $MRI_{sd}$  and  $MRI_{cd}$

using their total sample sizes:

$$MRI = \frac{MRI_{sd} \times u_{sd} + MRI_{cd} \times u_{cd}}{u_{sd} + u_{cd}} \quad (4)$$

where  $u_{sd}$  and  $u_{cd}$  denote the total sample sizes used to calculate  $MRI_{sd}$  and  $MRI_{cd}$ , respectively. Intuitively,  $MRI_{sd}$  captures the mobility responsiveness confined in a single day, revealing the dynamics of daily travel patterns while  $MRI_{cd}$  captures the mobility responsiveness between two consecutive days, revealing the dynamics of cross-day travel patterns. The rationale of deriving an integrated  $MRI$  by fusing  $MRI_{sd}$  and  $MRI_{cd}$  is that, despite their different calculations, they reflect human mobility from diverse perspectives, and therefore their integration serves as an overall index that better summarizes the general degree of mobility-based responsiveness geographically. The derived  $MRI$  has a range of  $(-\infty, 1]$ . In general, the higher the value, the better responsiveness a region has, with  $MRI = 1$  suggesting hypothetically perfect responsiveness. A positive  $MRI$  ( $MRI > 0$ ) suggests positive responsiveness (reduce in mobility) for a region, while a negative one suggests otherwise.

## 4. Results

### 4.1 Global scale

As most countries in the world started to aggressively respond to the COVID-19 pandemic after March, 2020, we set our baselines in a temporal period from January 13, 2020 (to exclude abnormal mobility patterns due to the New Year holiday season) to February 29, 2020. Since the outbreak in China and the dramatic increase in cases in Europe, many countries have imposed and continue to impose travel bans and lockdowns [45]. As a result, both  $D_{sd}$  and  $D_{cd}$  have greatly deviated from their corresponding baselines, especially after March 11, 2020, when WHO declared COVID-19 as a

pandemic (Fig 3). Because both our baselines are set for the individual day in a week, their projections exhibit a clear weekly pattern. In comparison, the time series of  $D_{sd}$  and  $D_{cd}$ , especially after the declaration of COVID-19 as a pandemic, show considerably less periodicity (Fig 3), suggesting that the protection measures (e.g., travel restrictions, social distancing policies, stay-at-home orders) have obviously affected people's weekly routines. The gap between baselines proves the different nature of  $D_{sd}$  and  $D_{cd}$ , well explaining our rationale of normalizing  $D_{sd}$  and  $D_{cd}$  separately. We further observe that, throughout the entire time series, the daily value of  $D_{cd}$  is considerably lower than the daily value of  $D_{sd}$ . This phenomenon can be explained by the existence of a large amount of Twitter users who, despite their large single-day travel distance (high  $D_{sd}$  value), keep a similar daily posting routine, which leads to no significant shift of mean centers between two consecutive days (low  $D_{cd}$  value).

**Fig 3. Temporal distribution of global  $D_{sd}$  and  $D_{cd}$  in the four-month period (February, March, April, and May).**

We observe similar mobility dynamics when  $D_{sd}$  and  $D_{cd}$  are respectively normalized to  $NMI_{sd}$  and  $NMI_{cd}$  according to their baselines (Fig 4). Both  $NMI_{sd}$  and  $NMI_{cd}$  started to deviate from the baseline ( $NMI = 1$ ) around ten days before the pandemic declaration from the WHO, suggesting that strong mobility-reducing measures had been taken before the declaration on March 11, 2020. This mobility pattern coincides with strong early travel restrictions implemented in Europe and Asia at the beginning of March [6, 46]. At the end of March, both  $NMI_{sd}$  and  $NMI_{cd}$  reached the bottom with the lowest  $NMI_{sd} = 0.70$  and the lowest  $NMI_{cd} = 0.45$ , indicating that single-day distance and cross-day distance respectively reduced to 70% and 45% of the ones in the normal

situation. Starting from the end of April, however, both  $NMI_{sd}$  and  $NMI_{cd}$  started to bounce back, and the increasing trend continued to the end of May, presumably resulting from the gradually lifted quarantine measures [47]. Compared with the hypothetically perfect scenario ( $MRI = 1$ ) where mobility instantly halts and remains 0 throughout the time series, the overall  $MRI$  for the three-month combined is 0.32, and the  $MRI$ s for the March, April, and May, respectively are 0.24, 0.39, and 0.33, revealing the less responsiveness in May compared with April.

**Fig 4. Global  $NMI_{sd}$  (normalized  $D_{sd}$ ) and  $NMI_{cd}$  (normalized  $D_{cd}$ ) in the four-month period, and the monthly  $MRI$  for March, April, and May.**

## 4.2 Country scale

For country scale study, we set the mobility baseline in a period from January 13, 2020 to February 15, 2020, as some countries (e.g., Italy and South Korea) already imposed strict or voluntary mobility-reducing policies as early as in late-February. To ensure that Twitter records are sufficient enough to generate a reasonable and stable time series, we mainly target the top 20 countries with most Twitter users, according to the Digital 2020 April Global Statshot Report [48]. The selection of those countries mostly agrees with the Twitter data we collected

In general, the impact of the COVID-19 pandemic on mobility derived from Twitter is obvious, as the mobility of the selected 20 countries, measured by single-day distance and cross-day distance, is mostly below the mobility baseline in March, April, and May (Fig 5), suggesting that mobility-reducing measures have been suggested and adopted in those countries. However, the country-level discrepancies in the time series of  $NMI_{sd}$  and  $NMI_{cd}$  can be clearly observed. The mobility in Japan started to drop in late-

February (Fig 5), presumably in response to the announcement by Prime Minister Shinzo Abe on February 27, 2020 to close all Japanese elementary, junior high, and high school [49]. The further decline of mobility from early-April to late-April can be explained by the proclamation of the State of Emergency for Tokyo (April 7) and for the rest of the country (April 16) [50]. The mobility of Japan is expected to bounce back, as Japan ended the state of emergency in all of Japan On May 25, 2020 [51]. Given the limited temporal coverage of our data, however, its impact on mobility remains unknown. The mobility of the United States started to drop in mid-March when a series of statements were announced, including the declaration of COVID-19 as a pandemic by the WHO (March 11) and the declaration of National Emergency by the White House (March 13, 2020). The mobility remained consistently low in April, then gained an upward momentum in May, largely due to the gradually loosened measures [52]. A similar mobility pattern can also be observed in India, where mobility reduced following the WHO's declaration in mid-March and gradually rose in May. Mobility in Malaysia was slightly below the baseline in late-February and early-March. The sudden mobility drop appeared on March 18, which coincides with the date when the Movement Control Order (MCO) from the federal government took effect [53]. The rapid mobility reduction in Malaysia demonstrates that the MCO was effectively and efficiently executed. In Saudi Arabia, mobility started to reduce as early as March 2, when the first case was confirmed [54]. However,  $NMI_{sd}$  and  $NMI_{cd}$  gradually diverged as  $NMI_{cd}$  remained stably low in April and May, while  $NMI_{sd}$  became unstable and eventually recovered and even surpassed baseline mobility in mid-May. The divergence in trends of the two types of distances can be partially explained by the suspension of flights and mass land transport (trains, buses, and taxis) that took effect on March 21 [55]. The lack of public transit is responsible for the consistently low cross-day distance.

**Fig 5. Temporal distribution of  $NMI_{sd}$  and  $NMI_{cd}$  for the top 20 countries with most Twitter users in February, March, April, and May.**

Compared with the hypothetically perfect scenario, i.e.,  $MRI = 1$ , Turkey has the highest three-month  $MRI$  (0.49), followed by Spain (0.43), Japan (0.42), Malaysia (0.41), and the U.K. (0.41) (Table 1). Russia has the lowest three-month  $MRI$  (0.18), followed by Australia (0.22), Indonesia (0.25), Philippines (0.25), Canada (0.26), and the U.S. (0.29) (Table 1). The high  $MRI$  (0.44) of March in South Korea, a country that suffered from the initial spread of the epidemic in its early stage besides China, indicates that the early and strong mitigation measures were announced and implemented effectively. In light of the gradually easing situation [56], the social distancing measures in South Korea started to be lifted, evidenced by the fact that its  $MRI$  decreased respectively by 0.09 and 0.15 in April and May. In the U.S., the mobility-based responsiveness in March (0.20) is among the weakest in the 20 selected countries (Table 1). In April, the  $MRI$  of the U.S. reached 0.38, a net gain of 0.18 compared to the  $MRI$  in March. The strong responsiveness of mobility in April is largely due to the gradually issued statewide stay-at-home orders since late-March that eventually affected at least 316 million people in at least 42 states [57]. With the lifting of orders in late-April and May, however, the U.S. showed reduced responsiveness, evidenced by its 0.09 loss in  $MRI$  of May compared to April. As the U.S. has become the new COVID-19 epicenter, the reduced mobility responsiveness, along with the rocketing number of confirmed cases, deserves more attention.

**Table 1. Mobility-based Responsive index (*MRI*) for the top 20 countries with most Twitter users**

Country names	<i>MRI</i>					
	Mar	Apr	May	Three-month average	$\nabla(\text{Apr-Mar})$	$\nabla(\text{May-Apr})$
Argentina	0.32	0.39	0.35	0.35	0.07	-0.04
Australia	0.23	0.22	0.20	0.22	-0.01	-0.02
Brazil	0.25	0.38	0.27	0.30	0.13	-0.11
Canada	0.21	0.35	0.23	0.26	0.14	-0.12
Germany	0.32	0.39	0.37	0.36	0.07	-0.02
Spain	0.32	0.49	0.45	0.42	0.17	-0.04
France	0.31	0.41	0.29	0.34	0.1	-0.12
The United Kingdom	0.27	0.48	0.47	0.41	0.21	-0.01
Indonesia	0.24	0.25	0.26	0.25	0.01	0.01
India	0.21	0.40	0.35	0.32	0.19	-0.05
Japan	0.25	0.49	0.53	0.42	0.24	0.04
South Korea	0.44	0.35	0.20	0.33	-0.09	-0.15
Mexico	0.16	0.38	0.38	0.31	0.22	0.00
Malaysia	0.34	0.50	0.40	0.41	0.16	-0.1
Philippines	0.28	0.28	0.20	0.25	0.00	-0.08
Russia	0.12	0.23	0.20	0.18	0.11	-0.03
Saudi Arabia	0.36	0.41	0.23	0.33	0.05	-0.18
Thailand	0.30	0.34	0.39	0.34	0.04	0.05
Turkey	0.29	0.57	0.60	0.49	0.28	0.03
The United States	0.20	0.38	0.29	0.29	0.18	-0.09

Besides the countries presented in Fig 5, the temporal distribution of  $NMI_{sd}$  and  $NMI_{cd}$  for the other 16 countries with relatively fewer Twitter samples can be found in the Fig in S1 Fig. Information regarding the accumulated user count for distance calculation (both  $D_{sd}$  and  $D_{cd}$ ) in selected countries is presented in the Table in S1 Table.

## 4.2 States in the CONUS

Given that the first State of Emergency related to COVID-19 in the U.S. was declared by Washington State (WA) on February 29, 2020, while the majority of the states started to react aggressively after mid-March, we set the U.S. mobility baseline in a period from January 13 to February 29, 2020. In general, the influence of the COVID-19 pandemic



on mobility is distinct, as the drop of mobility in most of the states happened in mid-March (Fig 6), potentially triggered by the events that include the pandemic declaration (March 11) and the National Emergency declaration (March 13). Although social distancing guidelines that aim to curb the spread have been suggested in the entire nation, the impacts varied substantially among states (Fig 6). Heavily hit states, e.g., NY, NJ, IL, CA, MA, and PA, generally experienced sharp mobility reduction, and their mobility remained stably low since mid-March. States with low numbers of cases, e.g., DE, MT, ME, WV, SD, and WY, despite the fluctuations in their time series, exhibited relatively marginal mobility reduction compared with heavily hit states. As the first state to announce the State of Emergency at the end of February, the mobility in WA remained close to the baseline in early March. It was not until mid-March that the mobility of WA started to noticeably decrease, which potentially indicates that the early mitigation policies in WA were not implemented effectively. The time series of mobility in states that include KS, MN, MS, AL, WV, SC, and WY, presents a bowl-shaped pattern, suggesting the strong recovery of mobility with some even bouncing beyond the baseline due to the gradually loosened measures. In response to the COVID-19 pandemic, eight states, including AR, IA, ND, NE, SD, UT, OK, and WY, decline to impose statewide stay-at-home orders by favoring other restrictions [57]. Without the orders, however, the aforementioned states still present considerable mobility reduction amid the pandemic, indicating the effectiveness of the federal guidelines and other mitigation approaches from the local government. Given the insufficient samples in the calculation of the baseline mobility, the mobility pattern in VT is not presented in Fig 6. The state names associated with their abbreviations and the accumulated user count for distance calculation in each state are presented in the supporting information S2 Table.

**Fig 6. Temporal distribution of  $NMI_{sd}$  and  $NMI_{cd}$  for states in CONUS (DC included; VT not included) in March, April, and May.**

In late-May, the risk of transmission in the U.S. was further complicated by the protests demanding justice after Mr. George Floyd died following an altercation with police. A noticeable mobility increase following the incident can be found in MN, where the incident happened (Fig 7). Carried by the existed mobility recovering momentum in mid-May, MN saw a significant increasing trend in both  $NMI_{sd}$  and  $NMI_{cd}$  at the end of the time series. A distinct spike can be found on May 29, 2020, when the raw  $NMI_{sd}$  and raw  $NMI_{cd}$  all went beyond the baseline, with the  $NMI_{sd}$  (representing single-day maximum travel distance) reaching about 2.5 times than usual as a consequence of the increased activity during the protests. The divergent functionality of the smoothed  $NMI$  and the raw  $NMI$  is well illustrated, as the former highlights the general trend while the latter is able to capture the spikes caused by disruptive events. At the time of writing, the protests have gradually spread across the U.S. and even overseas. The increase in mobility resulting from the protests deserves close monitoring, as standing in a crowd for long periods undoubtedly raises the risk of increased transmission and further worsens the situation.

**Fig 7. Temporal distribution of  $NMI_{sd}$  and  $NMI_{cd}$  for Minnesota**

The monthly  $MRI$  at the state level further highlights the responsiveness of each state in the three-month period. As expected, states with early spikes of cases and early strong mitigation policies tend to have a higher  $MRI$  in March (Fig 8). WA (0.41) leads the  $MRI$  in March, followed by NY (0.33), NH (0.32), and MA (0.32) (Table 2). The high responsiveness at the early stage suggests that the mobility-reducing guidelines were

implemented timely and efficiently in those states. In April, the responsiveness in all the states continued to strengthen (Fig 8), given the rising of cases and gradually tightened measures (Table 2). From March to April, MD, FL, and MI are the top three states with the most increase of *MRI*, respectively by 0.42, 0.41, and 0.39 (Table 2). The significant boost of mobility-based responsiveness reflects not only the severity of the situation but also the strong implementation of the mitigation measures. However, with the lifting of orders, 47 states (except MT) have shown reduced responsiveness in May compared to April (Fig 8). In light of the increasing number of cases in the U.S. with no sign of slowing down (at the time of writing), the reduced mobility responsiveness can potentially foster a second wave of infections. Because of the insufficient samples in the baseline calculation, the *MRI* for VT is not presented in Fig 8 and Table 2.

**Fig 8. Mobility-based Responsive index (*MRI*) for CONUS states in March, April, May, and the difference between two consecutive months. State boundaries are retrieved from the U.S. Census Bureau (<https://www.census.gov/geographies/mapping-files/time-series/geo/cartoboundary-file.html>).**

**Table 2. Mobility-based Responsive index (*MRI*) for states in the CONUS.**

State abbreviations	<i>MRI</i>					
	Mar	Apr	May	Three-month average	$\nabla(\text{Apr-Mar})$	$\nabla(\text{May-Apr})$
AL	0.09	0.45	0.15	0.23	0.37	-0.30
AR	0.13	0.44	0.23	0.27	0.32	-0.21
AZ	0.19	0.57	0.46	0.40	0.38	-0.11
CA	0.30	0.55	0.45	0.43	0.25	-0.10
CO	0.25	0.57	0.46	0.42	0.32	-0.11

CT	0.27	0.60	0.50	0.46	0.34	-0.10
DC	0.30	0.56	0.47	0.44	0.26	-0.09
DE	0.10	0.48	0.25	0.28	0.38	-0.23
FL	0.22	0.63	0.48	0.44	0.41	-0.14
GA	0.15	0.51	0.29	0.31	0.36	-0.22
IA	0.26	0.58	0.44	0.43	0.32	-0.14
ID	0.07	0.36	0.15	0.19	0.29	-0.21
IL	0.24	0.60	0.50	0.45	0.35	-0.10
IN	0.19	0.55	0.36	0.37	0.36	-0.19
KS	0.14	0.45	0.15	0.25	0.32	-0.30
KY	0.16	0.51	0.30	0.32	0.35	-0.20
LA	0.18	0.55	0.37	0.37	0.37	-0.18
MA	0.32	0.65	0.57	0.51	0.34	-0.08
MD	0.23	0.65	0.49	0.46	0.42	-0.17
ME	0.19	0.42	0.39	0.33	0.23	-0.03
MI	0.20	0.60	0.40	0.40	0.39	-0.19
MN	0.30	0.64	0.44	0.46	0.35	-0.20
MO	0.26	0.60	0.44	0.43	0.33	-0.15
MS	0.07	0.42	0.16	0.22	0.35	-0.26
MT	0.11	0.16	0.19	0.15	0.06	0.03
NC	0.11	0.47	0.36	0.31	0.36	-0.12
ND	0.10	0.36	0.25	0.24	0.26	-0.11
NE	0.23	0.56	0.40	0.40	0.34	-0.16
NH	0.32	0.64	0.57	0.51	0.31	-0.07
NJ	0.23	0.57	0.46	0.42	0.33	-0.11
NM	0.15	0.45	0.32	0.31	0.30	-0.13
NV	0.30	0.63	0.53	0.48	0.33	-0.10
NY	0.33	0.62	0.55	0.50	0.29	-0.07
OH	0.16	0.53	0.38	0.36	0.37	-0.16
OK	0.18	0.45	0.23	0.29	0.27	-0.22
OR	0.30	0.54	0.43	0.42	0.24	-0.10
PA	0.22	0.57	0.46	0.42	0.35	-0.12
RI	0.30	0.63	0.62	0.52	0.33	-0.01
SC	0.11	0.45	0.18	0.25	0.34	-0.27
SD	0.11	0.42	0.21	0.24	0.31	-0.20
TN	0.16	0.54	0.31	0.34	0.37	-0.22
TX	0.20	0.55	0.34	0.36	0.35	-0.20
UT	0.27	0.59	0.47	0.44	0.32	-0.12
VA	0.22	0.58	0.43	0.41	0.35	-0.14
WA	0.41	0.66	0.61	0.56	0.25	-0.05
WI	0.22	0.51	0.42	0.38	0.30	-0.10
WV	0.07	0.41	0.23	0.24	0.34	-0.18
WY	0.13	0.34	0.05	0.18	0.22	-0.29

*Note.* VT (Vermont) is not included due to the insufficient samples in baseline calculation.

## **5. Discussion**

### **5.1 Merits of social media data in gauging human mobility dynamics**

The rise of social media platforms in recent years offers a potential solution to closely monitoring human mobility dynamics, given their real-time high-volume user-generated content. Public health crises like the COVID-19 pandemic uniquely highlight several merits of social media data. First, social media data are a more harmonized source compared to cellphone records from certain providers that differ geographically. Twitter, for example, has 330 million monthly active users and 500 million daily posts worldwide [58]. Its popularity allows it to serve as a valuable venue where derived mobility dynamics can be cross-compared in different regions, especially for a global epidemic event like the COVID-19. Second, social media data offer both immediacy and spatially explicit geo-information that traditional approaches like surveys and censuses are often not capable of. The rapid spread of the SARS-CoV-2 virus and the everchanging mitigation policies greatly magnify the merit of timeliness in real-time crowdsourcing platforms, including social media. Third, social media data are relatively less privacy-concerning compared to passive data-collecting approaches that include phone calls, cellular records, and smart cards. The privacy issues in the above passive methods preclude the analysis in a more spatial explicit manner, as data collected via those methods are usually de-identified and aggregated before application. Finally, despite the required computational resources and storage that are essential to handle the large volume and velocity of Twitter data, such data are easily accessible and cost-efficient. In these respects, geotagged tweets can and should be considered as a valuable proxy for human mobility, especially during times of crisis (like the COVID-19 pandemic we are facing) that usually cause dramatic

mobility changes.

## 5.2 Limitations

The results of this study should be interpreted in light of several important limitations. First, the representativeness of Twitter data may not reflect the characteristics of the population as a whole in terms of socioeconomic status, age, gender, or race. Furthermore, the representativeness may vary geographically. Despite the attempts to improve the understanding of the demographics of Twitter users via profile scrutiny and tweets mining [27, 59], the intrinsic biases in Twitter samples should be considered when the results of this study are interpreted. The problem of representativeness, however, exists in all digital services. Mobility patterns derived from phone calls and cell phone applications (e.g., Google Maps and Apple Maps) also have to face the criticisms that people left behind by the “Digital Divide” [60] are underrepresented.

Second, the Twitter API allows unrestricted access to only about 1% of the total records [61]. From the tweets that streamed down via the Twitter API, we only use tweets that are geotagged with spatial resolution lower than the city level. Despite the “Big Data” nature of Twitter as a data source, the available records that can be used to derive human mobility patterns are still insufficient in some regions at a temporal resolution of daily. Mobility time series computed from insufficient samples tend to have more fluctuations, making the general pattern less recognizable and less reliable. In this respect, the mobility dynamics identified in the study only account for the reaction of Twitter in response to the COVID-19 pandemic and should not be generalized to infer the mobility of the total population without caution.

Third, the less privacy-concerning nature of social media data also creates many challenges. Unlikely the passively collected data from mobile telephone records, smart

cards, and wireless networks, social media data own intrinsic active nature, as users must grant permission to share their data and determine the locational accuracy of their posts, all depending on their personal settings. Thus, the two types of distances proposed in this study, single-day distance and cross-day distance, only reflect the travel behaviors that users are willing to share. This active nature protects privacy to some degree. At the same time, however, it dilutes the total amount of available trajectory data both spatially and temporally, potentially causing skewness in the extracted origin-destination information.

Finally, our mobility baseline, where mobility patterns from other periods are compared against, is derived from a one-month period that starts from mid-January. We further compute baselines for each corresponding day of a week by recognizing a week as an independent mobility cycle, without considering the monthly discrepancies that mobility patterns may present. Studies have shown that mobility may vary regularly on a monthly basis [62, 63], and the variations differ geographically due to the different cultural and societal settings. The uncertainty resulting from the short baseline period that specifically covers late-January and February needs to be acknowledged.

### **5.3 Future directions**

Despite these limitations, we believe the strengths and valuable findings in this study outweigh the shortcomings. However, several lines of future studies are still in need. First, future work should investigate the representativeness of Twitter data by delving into the demographics of Twitter users. The mobility patterns documented in this study only reflect Twitter users' collective activities responding to the COVID-19 pandemic. Thus, the representativeness of the findings largely depends on the demographics of the local users in relation to the demographics of the local population. Another research direction is to examine the similarity and dissimilarity in mobility patterns derived from various

sources (social media, phone calls, cellular records, smart cards, etc.), as they reflect human mobility from different yet valuable perspectives. Following the mathematical design in this study, we compared Twitter mobility with the Apple mobility report, the Google mobility report, and mobility data from Descartes Labs [64]. Future studies are needed to investigate the characteristics of other heterogeneous mobility sources to understand the strengths and pitfalls of each source. The third line of research is to explore the potential in the integration of mobility indices from heterogeneous data sources. An integrated mobility index from multiple sources is expected to better reflect the multifaceted nature of human mobility, thus greatly facilitating comprehensive mobility monitoring. Fourth, research on more extensive Twitter datasets is needed to investigate the possible improvement in mobility that can be captured. Although millions of spatially explicit Twitter posts collected in this study are sufficient to quantitatively reflect the human mobility dynamics during the pandemic, an increasing amount of tweets are expected to generate more stable and reliable trends with fewer random fluctuations. Despite the fact that the licenses for other sample sizes, such as Gecahose (returning 10% of the public data) and Firehose (returning 100% of the public data), are costly, difficult to obtain, and requiring a demanding computational environment [65], their potential in obtaining reliable mobility dynamics at much finer spatiotemporal resolutions deserves attention. Finally, as human population movement is among the critical dimension that drives the spatial spread of COVID-19, how to leverage such Twitter derived mobility information for better predicting the future infectious risk of a state, county, or community warrants investigation.

## **6. Conclusion**

As the whole world is now fighting the COVID-19 pandemic, the effectiveness of



mobility-reducing measures (e.g., social/physical distancing) at varying scales needs rapid investigation. This article examines the reaction in social media, specifically Twitter, spatially and temporally in response to the COVID-19 pandemic as a more harmonized, less privacy-concerning, and cost-efficient approach to assessing human mobility dynamics promptly. Through analyzing more than 580 million tweets worldwide, we present how our collaborative efforts in mobility reduction are reflected from this user-generated information in three different geographic scales: global scale, country scale, and U.S. state scale. To quantify various aspects of mobility from Twitter, we propose two types of distance, i.e., the single-day distance that highlights daily travel behavior and the cross-day distance that highlights the displacement between two consecutive days. To facilitate the comparison with normal situations, we further normalize these distances by separately setting up their baselines for each corresponding day of a week. We also propose a mobility-based responsive index (*MRI*) to capture the overall degree of mobility-related responsiveness of particular geographic regions in response to the COVID-19 pandemic.

The results suggest that mobility patterns obtained from Twitter data are amenable to quantitatively reflect the mobility dynamics in COVID-19 pandemic at various geographic scales. Globally, the proposed two distances measured from Twitter had greatly deviated from their baselines after March 11, 2020, when WHO declared COVID-19 as a pandemic. The considerably less periodicity after the declaration suggests that the protection measures have obviously affected people's weekly routines. The global *MRI* reveals less responsiveness in May compared with April. At the country scale, the country-level discrepancies in responsiveness are obvious, evidenced by the contrasting mobility patterns in different epidemic phases. We further find that the triggers of mobility changes correspond well with the announcements of mitigation measures, which

in return proves that Twitter-based mobility, to some degree, implies the effectiveness of those measures. At the U.S. state scale, the influence of the COVID-19 pandemic on mobility is distinct, as the drop of mobility in most of the states happened in mid-March following the National Emergency declaration on March 13. However, the impacts varied substantially among states. Heavily hit states generally experienced sharp mobility reduction while states with low numbers of cases exhibited relatively marginal mobility reduction. With orders gradually being lifted since late-April, 45 states (except MT, NH, and WA) have shown reduced responsiveness in May compared to April. The methodological knowledge and contextual findings in this study seed future applications of the easily accessible, less privacy-concerning, highly spatiotemporal Twitter data in monitoring multi-scale mobility dynamics during disaster events.

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## Supporting information

**S1 Fig. Temporal distribution of  $NMI_{sd}$  and  $NMI_{cd}$  for the other 16 selected countries in February, March, April, and May.**



**S2 Fig. The user count with the number of tweets posted per day.**

**S1 Table. Country names and accumulated user count for distance calculation in selected countries.**

**S2 Table. State abbreviations, state names, and accumulated user count for distance calculation in the CONUS states.**