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Mental Distress during the COVID-19 Pandemic among US Adults without a Pre-existing Mental Health Condition: Findings from American Trend Panel Survey

Calliope Holingue, PhD[1,2], **Elena Badillo-Goicoechea, MS**[1], **Kira E. Riehm, MSc**[1], **Cindy Veldhuis, PhD**[3], **Johannes Thrul, PhD**[1], **Renee M. Johnson, PhD**[1], **M. Daniele Fallin, PhD**[1], **Frauke Kreuter, PhD**[4,5,6], **Elizabeth A. Stuart, PhD**[1], **Luther G. Kalb, PhD**[1,2]

[1]Department of Mental Health, Johns Hopkins Bloomberg School of Public Health

[2]Department of Neuropsychology, Kennedy Krieger Institute

[3]Columbia University

[4]University of Maryland, College Park

[5]University of Mannheim, Germany

[6]Institute for Employment Research, Nuremberg, Germany

Abstract

Most individuals in the United States have no history of a mental health condition yet are at risk for psychological distress due to the COVID-19 pandemic. The objective of this study was to assess the frequency and risk and protective factors of psychological distress, during the beginning of the COVID-19 pandemic, in this group. Data comes from the Pew Research Center's American Trends Panel (ATP), a probability-based online survey panel representative of the US adult population. The analytic sample consisted of 9,687 individuals with no prior history of a mental health condition who completed the survey between March 19–24, 2020. Explanatory variables included sociodemographic factors and items related to behavior, perceptions, and experiences surrounding the pandemic. The outcome was psychological distress, measured by five items on symptoms of anxiety, depression, loneliness, sleep difficulties, and hyperarousal. A multivariable linear regression model was used to identify risk and protective factors for psychological distress. Fifteen percent of the sample experienced 2 psychological distress symptoms for at least 3 days over the past week; 13% had three or more symptoms. Risk factors for higher distress included searching online or using social media to post about coronavirus, reporting that the outbreak caused major changes to personal life, and perception that the virus was a threat to the US economy, the individual's personal health or finances. This has important implications for mental health service delivery.

Keywords

COVID-19; pandemic; psychological distress; epidemiology

Corresponding author: Calliope Holingue, Office 3050A, 1750 E Fairmount Ave, Baltimore, MD, 21231, choling1@jhu.edu, (443) 691-1125.

The COVID-19 pandemic has dramatically changed the day-to-day life of all Americans. A robust body of work has demonstrated that smaller-scale public health disasters may negatively impact mental health (Bonanno et al., 2007; Douglas et al., 2009; Havenaar et al., 1997; Henriksen et al., 2010; Huremović, 2019; Kunii et al., 2016). Previous outbreaks indicate acute mental health decrements due to the effects of quarantining, lack of clear information, fears of infection, boredom, and frustration. Longer-term mental health impacts were concentrated among those who were impacted financially, felt stigmatized due to having been ill and/or quarantined, and those who reported having a history of mental health concerns (Brooks et al., 2020; Cao et al., 2020). However, the global scope of the COVID-19 pandemic is unprecedented. Social distancing measures, needed to prevent the spread of infections, are severely limiting interpersonal interactions (Unacast, 2020). Widespread job losses and furloughs of uncertain duration have contributed to rising anxiety over financial well-being (Shiller, 2020). Finally, the force of morbidity and mortality associated with COVID-19 is both substantial and wide-spread; as of July 11, 2020, there were over 3 million cases and 130,000 deaths in the United States (US) alone (Johns Hopkins University, 2020).

The path of devastation laid by COVID-19 raises serious concerns about mental health in the US population, both in the midst of the pandemic and beyond. Recent data from China indicate that the impact of the COVID-19 pandemic on psychological distress has been significant (Qiu et al., 2020; Wang et al., 2020). In the United States, data gathered from the nationally representative Understanding America Study (UAS), collected during the very beginning of the outbreak in early March 2020, suggest a linear relationship between time (as cases increased) and psychological distress among individuals living in states with the most severe outbreaks (Holingue et al., In Press). Data gathered from the Kaiser Family Foundation provides additional evidence that the COVID-19 pandemic is impacting the public's mental health (Kirzinger et al., 2020).

It is expected that the COVID-19 pandemic will exacerbate latent or existing mental health symptoms among individuals with a history of a mental condition (Druss, 2020). However, the influence of the pandemic on individuals with *no prior history* of a mental condition is still unknown. It is likely this group is experiencing psychological distress, given the unique and severe impacts of the pandemic on daily life. Understanding the impact on this group is critical for several reasons. First, this group represents the majority of adults in the US (Substance Abuse and Mental Health Services Administration, 2019). The mental health system is likely ill-prepared to accommodate such a large increase in mental health disorders (Auerbach & Miller, 2020). Second, those with no history of psychological distress are likely wholly unfamiliar with the mental health system and may need specific public mental health messaging to help them navigate this system. Third, this population may experience confusion and self-stigmatization since these symptoms or degree of distress may be unprecedented (Bathje & Pryor, 2011). Fourth and finally, there are many activities (World Health Organization, 2020), such as meditation (Behan, 2020), exercise (Stathopoulou et al., 2006), and establishing regular sleep patterns (Altena et al., 2020) which can be used to prevent further exacerbation of symptoms. These activities may be more amenable and

perhaps appropriate for those with emerging, but not significantly impairing, psychological distress.

The primary goal of this study was to describe the frequency of psychological distress during the COVID-19 pandemic within a nationally-representative sample of adults in the US with no reported history of a pre-existing mental health condition. In addition, multivariable analyses were conducted to identify independent risk factors, such as sociodemographic variables and behavior, perceptions, and experiences surrounding the coronavirus outbreak, for psychological distress in this group.

Methods

Study Sample

Data for this study come from the Pew Research Center's American Trends Panel (ATP), an online survey panel recruited through random sampling of residential addresses throughout the US, in accordance with the best practice guideline set forth by the American Association for Public Opinion Research (AAPOR) (Task Force on Address-based Sampling, 2016). Panelists who do not have internet access are provided a tablet and wireless internet connection in order to complete self-administered web surveys. The current study drew from the panel wave conducted from March 19–24, 2020. Of the 15,433 individuals sampled, 11,537 completed the survey (75% response rate). Further details on panel construction can be found here: https://www.pewresearch.org/methods/u-s-survey-research/american-trends-panel/.

The sample was restricted to individuals who reported no prior history of a mental health condition (n=9,687) (reported have never "been told by a doctor or other healthcare provider that [they] have a mental health condition"). The proportion of individuals with a history of a mental health condition (16%) is consistent with prior population-based studies (Substance Abuse and Mental Health Services Administration, 2019). Survey weights, used in all statistical analyses in this study, allow for inferences to the US adult population (Pew Research Center, 2020).

Mental Health Measures

The primary outcome was psychological distress, measured by a score derived from five questionnaire items in the survey. These items were adapted from the Center for Epidemiologic Studies Depression (CES-D) Scale (Radloff, 1977), the Generalized Anxiety Disorder-7 (GAD-7) (Spitzer et al., 2006), and the Impact of Event Scale – Revised (IES-R) (Weiss, 2007). Participants were asked how often in the past 7 days they: 1) felt nervous, anxious, or on edge; 2) felt depressed; 3) felt lonely; 4) had trouble sleeping; and 5) had physical reactions, such as sweating, trouble breathing, nausea, or a pounding heart when thinking about their experience with the coronavirus outbreak (hereafter referred to as hyperarousal). Response options for these five items were "rarely or none of the time (less than 1 day)"; "some or a little of the time (1–2 days)"; "occasionally or a moderate amount of time (3–4 days)"; or "most or all of the time (5–7 days)". These four response options were given a weight of 1, 2, 3, and 4, respectively, to reflect increasing frequency of

symptoms in the past week. A summary score for each person was derived by taking the sum of response weights across each of the five items (possible range 5–20). This scoring procedure yielded a scale with adequate internal consistency (Cronbach's $\alpha = 0.73$).

Demographic Variables

Demographic variables included age categories (18–29, 30–49, 50–64, 65+), sex, education level (high school graduate or less, some college, associate degree, college graduate, or post graduate), race/ethnicity (non-Hispanic white, non-Hispanic Black, Hispanic, other), marital status (married, never married, widowed, divorced or separated, living with partner), and income ranges (<40k, 40–49k, 50–74k, 75–99k, 100k+). Participants were asked how often they attended religious services (seldom or never, a few times a year, once or twice a month, once a week or more).

COVID-19 Context Variables

Panelists were also asked questions about their behavior, perceptions, and experiences surrounding the coronavirus outbreak. Participants reported whether or not they searched online or used social media to share or post information about the coronavirus (*yes/no*). They were asked how closely they had been following news of the outbreak (4-point response scale from *not at all closely* to *very closely*). Next, participants were asked whether they felt that the coronavirus was a threat to "the health of the US population as a whole," to their "personal health," "the US economy," and to their "personal financial situation" (response options: *not a threat*, *minor threat*, *major threat*), and whether their personal life changed due to the coronavirus (response options: *same*, *minor change*, *major change*). Lastly, they reported whether or not anyone in their household had "been laid off or lost a job" or "had to take a cut in pay due to reduced hours or demand for [their] work," because of the coronavirus pandemic (*yes/no*).

Statistical Analysis

Missingness ranged between 0–4% per variable. Because complete case analysis would have reduced the sample by 12% overall, multiple imputation with chained equations was used to impute missing data. All analytic variables were used for imputation and ten imputations, with ten iterations per imputation, were generated. In addition to examining descriptive statistics, a survey-weighted multivariable linear regression model was used to assess associations between explanatory variables and psychological distress summary score, among individuals with no reported history of mental condition. All statistical analyses were performed in RStudio (version 1.1.383) (RStudio Team, 2016) using the "pewmethods" (version 1.0) (Lau et al., 2020), "survey" (version 3.37) (Lumley, 2004, 2019), and "mice" (version 3.8.0) (Buuren & Groothuis-Oudshoorn, 2011) packages.

Ethical Considerations

As with all studies by Pew Research Center, the ethical implications of the research were evaluated and approved by the Vice President of Research at Pew Research Center and Senior Vice President for government and nonprofit research at the data collection firm, Ipsos. Both organizations have standing approval for general population surveys that (1) are

limited to administering questions and do not feature more invasive data collections (e.g. biospecimens) and (2) explicitly inform participants that they may decline to answer any questions they wish. Pew Research Center also has extensive protocols in place to prevent accidental or malicious disclosure of personally identifying information about their panelists. The authors of this paper are not affiliated with the Pew Research Center, though they did consult on the design of the psychological distress items that were collected through this survey. Local IRB approval was not sought given these were secondary analyses of de-identified, publicly available data.

Results

Frequency of psychological distress symptoms

The weighted population reflects the distribution of characteristics of individuals in the United States with no reported history of a mental condition (Table 1). The most common symptom of psychological distress experienced for at least 3 days in the past week was feeling nervous, anxious, or on edge (mean 39% [95% CI: 0.37-.40]). The next most common symptom was trouble sleeping (27% [95% CI: 0.26–0.29]), followed by feeling depressed (19% [95% CI: 0.17–0.20]), feeling lonely (15% [95% CI: 0.14–0.17]), and having hyperarousal symptoms (5% [95% CI: 0.04–0.07]) (Figure 1). The mean score for psychological distress (range 5–20) was 8.69 (SD=5.3) (Figure 2). Over the past week, over seventy percent (72%) of individuals had 0–1 symptoms, 15% had two symptoms, and 13% had three or more symptoms for at least 3 days.

Risk factors for Psychological Distress

Results of the survey-weighted multivariable linear regression model showed several independent risk factors significantly associated with higher psychological distress scores among individuals with no reported prior history of a mental condition. These included being female, never having been married, being divorced or separated, searching online for coronavirus information, using social media to post about the coronavirus, reporting that the coronavirus caused major changes to personal life, or that it was a minor or major threat to their own personal health. Additionally, having someone in the household with a pay cut or reduced hours was a risk factor for psychological distress, as were perceptions that the coronavirus is a minor or major threat to the US economy, and a major threat to personal finances

Protective factors against Psychological Distress

Conversely, old age (50–64 and 65+) was protective against levels of psychological distress, as was being Hispanic, or having an “other” race/ethnicity, having an income of 50k or higher, and attending religious services on average once a week or more prior to this pandemic (Figure 3 and Table 3). The distributions of explanatory variables, stratified by number of psychological distress symptoms occurring for at least 3 days in the past week, are shown in Tables 1 and 2.

Discussion

In this nationally-representative sample of US adults with no reported prior history of a mental condition, more than 1 in 4 are experiencing psychological distress in the early phases of the COVID-19 pandemic. Specifically, 15% experienced some psychological distress (2 symptoms for at least 3 days in the past week) and 13% experienced significant psychological distress (3–5 symptoms for at least 3 days) during the beginning of the COVID-19 pandemic (March 19–24, 2020). These findings raise concerns about a generalized burden of psychological distress, and potential risks for emergent psychological disorders, in a population with no reported history of a mental health condition.

Symptoms of anxiety and depression were the most prevalent symptoms of psychological distress. We found that 39% of individuals reported symptoms of anxiety and 19% reported depressive symptoms for at least 3 days in the past week. Recent benchmarks for the U.S. population can help contextualize our findings at the symptom level. According to January–June 2019 estimates from the National Health Interview Survey, 8.2% of the U.S. adult population experienced symptoms of anxiety disorder and 6.6% experienced symptoms of depressive disorder. The above suggests elevated levels of psychological distress among the U.S. population at the beginning of the COVID-19 pandemic, compared to the months before the pandemic began.

Searching online and using social media to post information about the coronavirus were both independent risk factors associated with higher psychological distress. Although the media represents an important avenue for disseminating information during public health disasters and causality cannot be inferred from our analysis, this finding is consistent with prior studies demonstrating a positive association between increased exposure to disaster-related media content and poor mental health (Lubens & Holman, 2017), including within the COVID-19 pandemic (Gao et al., 2020; Riehm et al., 2020). Other researchers have provided guidance on responsible media communication during the COVID-19 pandemic (Garfin et al., 2020), such as avoiding the use of sensationalized headlines and refraining from publishing disturbing images. At the individual level, it is recommended that social media users avoid passive activities (Verduyn et al., 2015), such as repetitive “scrolling” and observation of others; instead, social media can be used to obtain accurate information from trusted sources, engage in purposeful connections with others, and obtain social support throughout periods of social distancing.

Individuals who reported that the virus had caused a major change to their life showed significantly higher psychological distress. Importantly, this variable was associated with distress even after adjusting for sociodemographic factors and lost employment or wages secondary to COVID-19, suggesting it may be capturing changes and stressors independent of financial ones. Although information on the specific changes people underwent is not available, these could plausibly include homeschooling children and increased parenting responsibilities (Brown et al., 2020; Canady, 2020; Kantamneni, 2020), working from home, moving, caretaking for sick relatives (Kantamneni, 2020), increased social isolation or loneliness(Usher et al., 2020), and major changes in work-related habits, among others. These changes are likely to be concomitant with reduced time for self-care and disrupted

work-life balance, which can contribute to increased distress (Ensel & Lin, 1991; Haines et al., 2015; Kessler, 1979; Pfefferbaum & North, 2020; Sprang & Silman, 2013; Tausig & Fenwick, 2001; Vinokur & Selzer, 1975).

We also found that those who reported that the virus was a threat to the US economy or to their own personal finances had significantly higher distress, consistent with an initial report by Pew Research Center (Keeter, 2020). As millions of Americans continue to experience financial and health insurance loss due to the pandemic (Woolhandler & Himmelstein, 2020), there may be a continued deterioration in public mental health. Policies and interventions, such as access to healthcare (including mental health care) irrespective of employment status, may significantly alleviate this burden of psychological distress.

Individuals who perceived that the virus was a threat to their personal health also had significantly elevated psychological distress, even after adjusting for age. This might represent individuals with a physical health condition that makes them more susceptible to COVID-19. Alternatively, individuals who perceive the virus is a greater threat to their personal health may be in occupations that do not permit them to practice social distancing (e.g., grocers) or that expose them to higher risk of contracting the virus (e.g., healthcare workers). Research from China shows that healthcare workers have experienced significant distress due to the pandemic (Chen et al., 2020; Dai et al., 2020; Huang & rong Liu, 2020; Qi et al., 2020; Zhu et al., 2020), and this is expected to be occurring in the US as well (Greenberg et al., 2020; World Health Organization, 2020). Lastly, higher perception of threat could simply be a proxy for higher symptoms of anxiety or worries about the virus.

Beyond COVID-19 specific factors, sociodemographic variables were also associated with higher distress. Being female, never married, divorced/separated, or white non-Hispanic race were all risk factors, while older age and higher income were protective. These findings are consistent with decades of psychiatric epidemiology studies (Breslau et al., 2005; Kessler, Berglund, et al., 2005; Kessler, Chiu, et al., 2005; Manderscheid & Sonnenschein, 1998; Regier et al., 1988, 1993). While examining correlates of distress among those with a previous history of a mental health was out-of-scope for this study, our findings were consistent with previous psychiatric epidemiologic research among this population (see citations above).

It is important to recognize that the group under study--those with no reported prior mental health condition--may not be well equipped to handle incident mental health symptoms, especially in the context of a pandemic. Typical buffers against psychological distress such as in-person social supports, attending religious services, outdoor time, and gym attendance may be harder to access. Further, treatment-naïve individuals may not know how to navigate the mental health services system, and financial distress related to the pandemic may act as yet another barrier to accessing treatment. A mental health care system already struggling to meet the treatment need prior to this pandemic (Kohn et al., 2004) may now need to accommodate these individuals. Thus, measures that improve accessibility such as increased availability of telehealth, insurance coverage for phone and video psychotherapy, co-pay waivers, and license reciprocity to enable telehealth across state lines should be continued. It also suggests bolstering access to mental health care through primary care (Ross et al.,

2019), when appropriate. Use of psychological care needs to not just be made more accessible, but also less stigmatized. Public health campaigns that normalize distress in reaction to highly stressful events, promote self-care resources and strategies, and disseminate simple, useful information on how to access mental health services will be critical over the course of and following this pandemic. Together these measures could increase access to those newly experiencing mental health concerns, who are experiencing relapses or increased symptomatology, and to those who live in areas with few psychological healthcare providers (e.g., rural areas) and ensure continued access and support long after the acute needs have subsided and longer-term impacts are felt.

This study is not without limitations. First, while items to assess psychological distress were derived from established, well-validated measures, including the CES-D (Radloff, 1977), GAD-7 (Spitzer et al., 2006), and IES-R (Weiss, 2007), their response options had to be adjusted to the last week to reflect the recent onset and highly dynamic course of the pandemic, which could have altered the psychometric properties and may limit comparability with other samples. Since the wording of the questions themselves did not change substantially, it is less likely that the face, construct, or content validity of the scale was negatively affected, but more likely than the reliability may be compromised. The Cronbach's α of 0.73 suggests this scale has adequate but not excellent internal reliability. Second, all data were self-reported by respondents and may be susceptible to misclassification errors or biases. In addition, participants who do not belong to the white, Black, or Hispanic racial/ethnic groups were combined together into an "other" group in this publicly available dataset. Therefore, the associations between other racial groups, such as Asian American individuals, and degree of psychological distress could not be assessed. Lastly, while a strength of this study is that the data come from a large, nationally-representative sample of US adults, the descriptive and analytic inferences made from this analysis are generalizable to the adult US population under the assumption that non-response is unrelated to any factors that are not included in the construction of the survey weights (sex, age, race and ethnicity, educational attainment, and geographic region). This survey also remains a web-based survey of thousands of individuals, which limits the depth of information that can be obtained, and may influence the composition of the sample.

Nevertheless, this study provides a unique opportunity to analyze the risk and protective factors that were collected during this time, allowing us to identify at-risk groups and suggest interventions for reducing the burden of this pandemic on the public's mental health. In sum, this study reports evidence of substantial levels of psychological distress among US adults with no prior diagnosis of a mental condition during the rise of the COVID-19 pandemic. Given the scope and expected duration of this pandemic, there are numerous opportunities to address public mental health through campaigns that de-stigmatize mental health, primary care screenings, expansion of and continued funding for telehealth, and promoting self-care through nutrition, exercise, and contact with friends and families, while observing recommended measures of social distancing. It is essential that prevention efforts, interventions, and policies be put in place to mitigate the potential fallout.

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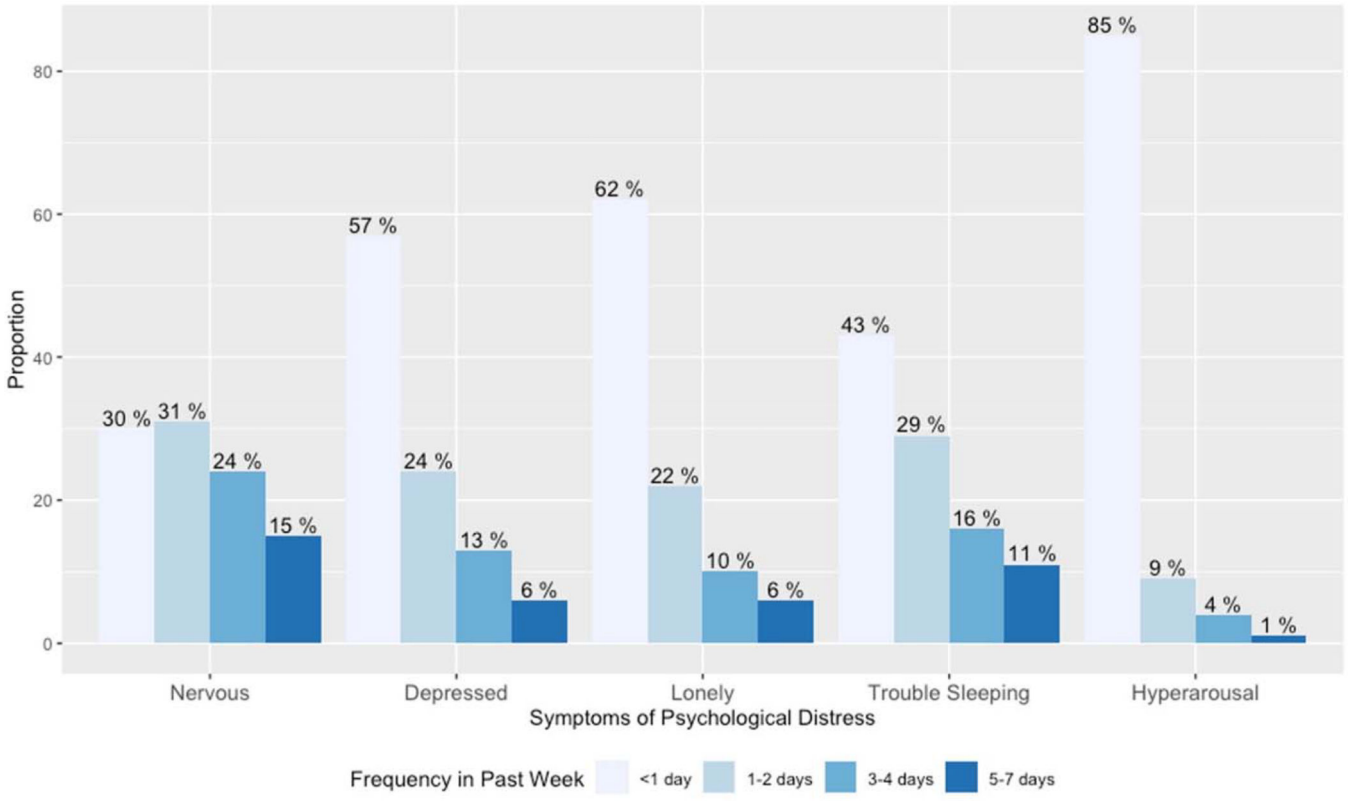

Figure 1.

Frequency of psychological distress symptoms in past week.

Bar plot shows frequency of five symptoms of psychological distress in past week, in a nationally representative sample of individuals in the United States with no history of a mental health condition (N=9687). Surveys were taken between March 19–24, 2020.

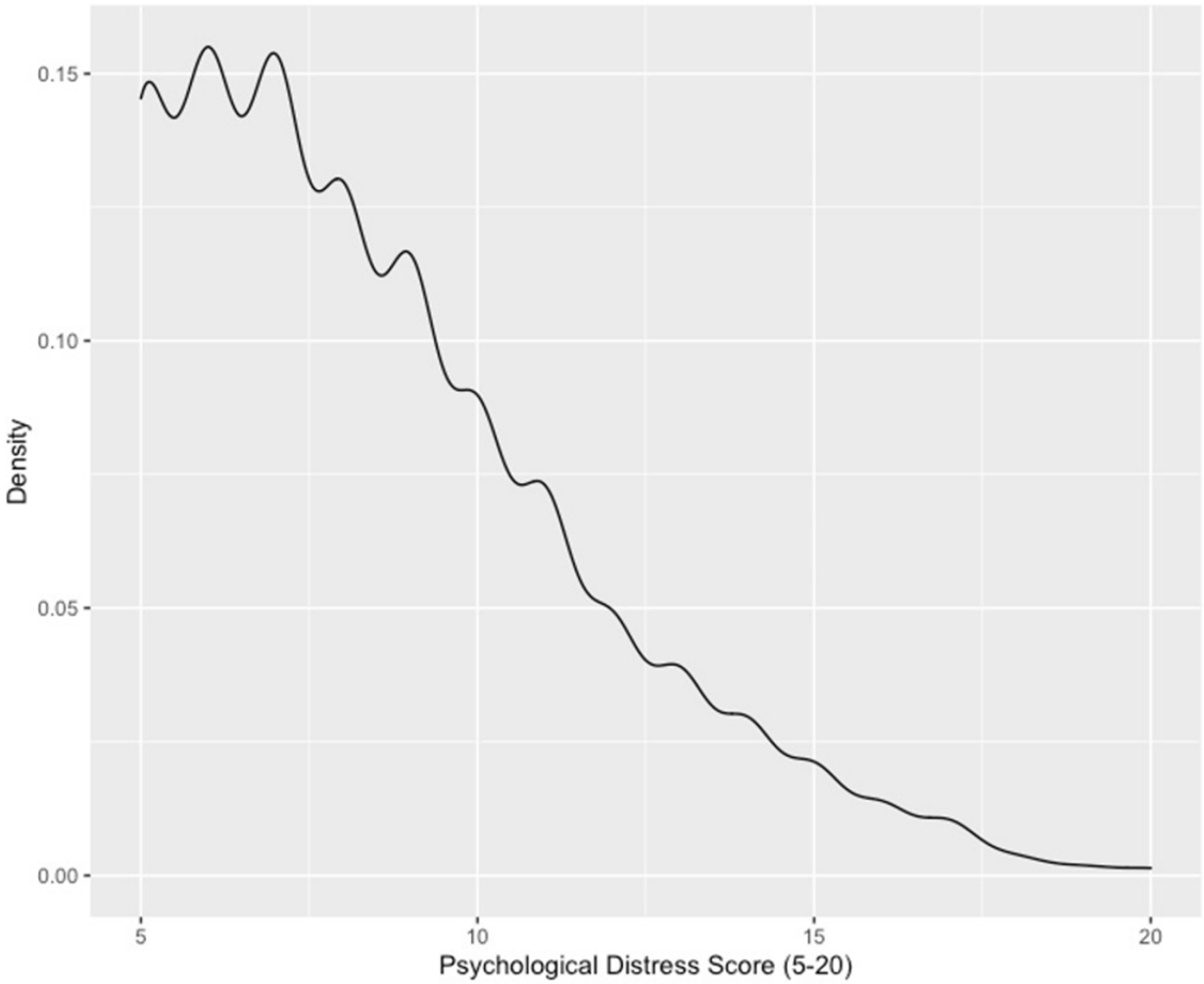

Figure 2.

Distribution of psychological distress score.

Plot shows distribution of psychological distress score in past week, in a nationally representative sample of individuals in the United States with no history of a mental health condition (N=9687). Surveys were taken between March 19–24, 2020.

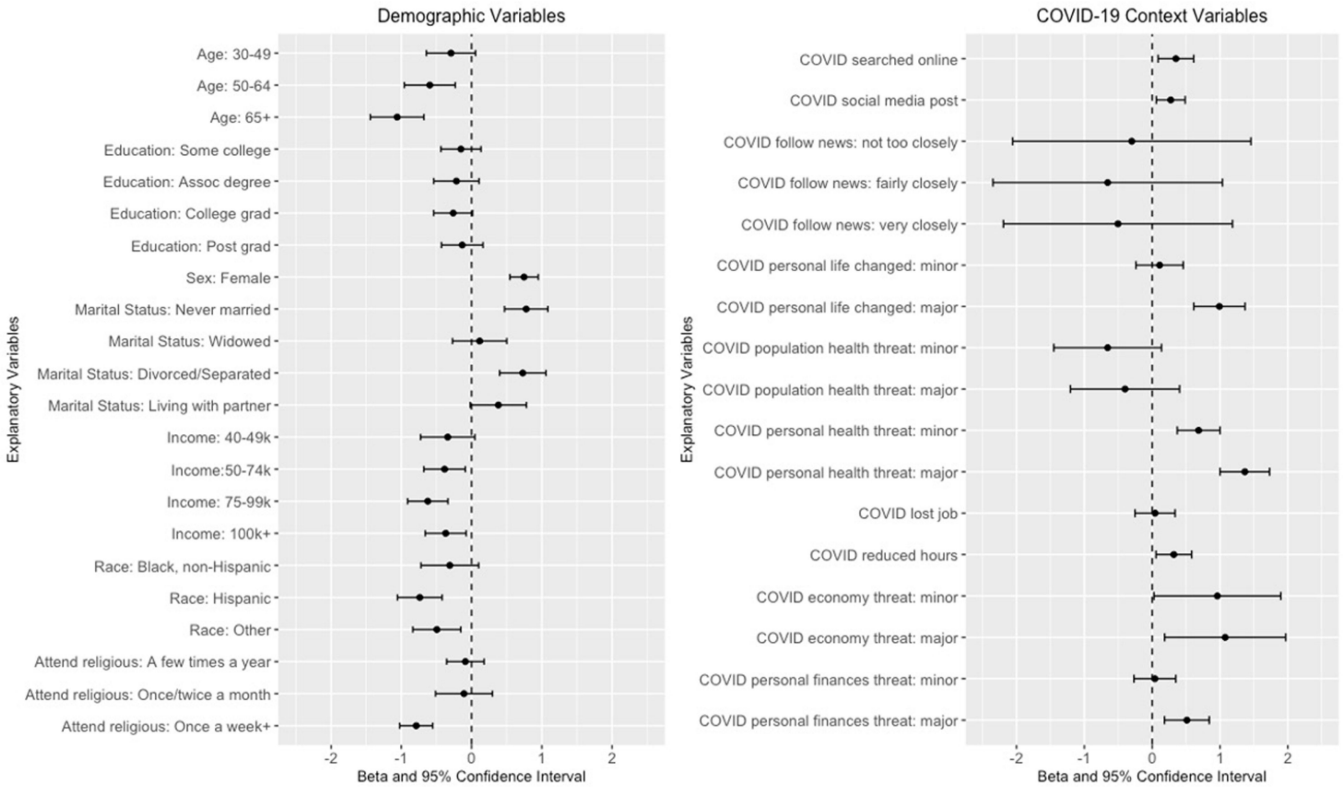

Figure 3.

Multivariable linear regression model estimating association between explanatory variables and psychological distress score.

Plots show results of single multivariable linear regression model estimating association between explanatory variables and psychological distress score, in a nationally representative sample of individuals in the United States with no history of mental disorder (N=9687). Vertical line represents null of Beta=0. Bars denote 95% Confidence Intervals. Both panels together represent one single model. Surveys were taken between March 19–24, 2020.

Table 1.

Demographic characteristics in a nationally representative sample of individuals in the United States with no history of mental disorder (N=9687).

	Number of Distress Symptoms		
	0–1 (n=7008)	**2 (n=1402)**	**3–5 (n=1227)**
Age (%)			
18–29	16	23	25
30–49	31	33	34
50–64	27	25	25
65+	26	19	16
Sex (%)			
Male	54	44	39
Female	46	56	61
Education (%)			
High school graduate or less	36	34	42
Some college	21	23	22
Associate degree	9	10	9
College graduate	19	19	16
Post graduate	14	15	11
Race (%)			
Non-Hispanic White	65	62	57
Non-Hispanic Black	11	13	14
Hispanic	15	17	21
Other	9	9	8
Marital Status (%)			
Married	53	46	37
Never married	21	26	31
Widowed	7	5	4
Divorced/Separated	12	13	16
Living with partner	8	10	13
Income (%)			
<40k	34	37	50
40–49k	11	8	10
50–74k	17	17	14
75–99k	14	14	8
100k+	24	24	18
Attends religious services (%)			
Seldom or Never	50	55	60
A few times a year	15	18	17
Once or twice a month	7	9	8
Once a week+	28	18	15

Demographic characteristics are stratified by number of symptoms occurring 3+ days in past week (0–1, 2, 3–5). Chi-squared differences for each variable by number of symptoms occurring 3+ days in past week are all statistically significant at $p<0.001$.

Table 2.

COVID-19 context variables in a nationally representative sample of individuals in the United States with no history of mental disorder (N=9687).

	Number of Distress Symptoms		
	0–1 (n=7008)	**2 (n=1402)**	**3–5 (n=1277)**
Search online for info (%)			
No	31	22	22
Yes	69	78	78
Used social media to post (%)			
No	65	59	55
Yes	35	41	45
How closely following news (%)			
Not at all closely	1	0	2
Not too closely	6	7	7
Fairly closely	37	30	28
Very closely	56	63	62
Personal life change coronavirus (%)			
Same	13	8	8
Major change	40	54	60
Minor change	48	38	31
Threat for US population health (%)			
Not a threat	3	1	3
Major threat	62	75	78
Minor threat	35	25	19
Threat for personal health (%)			
Not a threat	13	7	6
Major threat	32	41	53
Minor threat	55	52	42
Laid off/lost job (household) (%)			
No	83	79	72
Yes	17	21	28
Pay cut/reduced hours (household) (%)			
No	77	68	63
Yes	23	32	37
Threat to US economy (%)			
Not a threat	1	1	1
Major threat	87	93	92
Minor threat	12	6	8
Threat for personal finances (%)			
Not a threat	13	9	7
Major threat	44	58	66
Minor threat	44	33	27

COVID-19 context variables are stratified by number of symptoms occurring 3+ days in past week (0–1, 2, 3–5). Chi-squared differences for each variable by number of symptoms occurring 3+ days in past week are all statistically significant at $p<0.001$.

Table 3.

Multivariable linear regression model estimating association between explanatory variables and psychological distress score.

	Beta (95% CI)	
Age		
30–40	−0.29	(−0.64, 0.06)
50–64	−0.59	(−0.96, −0.23)***
65+	−1.06	(−1.44, −0.68) ***
Sex		
Female	0.75	(0.55, 0.95)***
Education		
Some college	−0.15	(−0.44, 0.13)
Associate degree	−0.22	(−0.54, 0.11)
College graduate	−0.26	(−0.54, 0.02)
Post-graduate	−0.13	(−0.43, 0.16)
Race		
Black, non-hispanic	−0.31	(−0.72, 0.10)
Hispanic	−0.74	(−1.06, −0.42)***
Other	−0.49	(−0.84, −0.15) ***
Marital Status		
Never married	0.78	(0.47, 1.09)***
Widowed	0.11	(−0.27, 0.50)
Divorced/Separated	0.73	(0.40, 1.06) ***
Living with partner	0.38	(−0.02, 0.78)
Income		
40–49k	−0.34	(−0.73, 0.05)
50–74k	−0.38	(−0.68, −0.09)**
75–99k	−0.62	(−0.91, −0.34)***
100k	−0.37	(−0.66, −0.08) **
Attends religious services		
A few times a year	−0.09	(−0.35, 0.18)
Once or twice per month	−0.11	(−0.51, 0.30)
Once a week or more	−0.79	(−1.02, −0.55)***
Search online for coronavirus info		
Yes	0.35	(0.09, 0.61)*
Used social media to post about coronavirus		
Yes	0.27	(0.06, 0.49)*
How closely following news on COVID-19		
Not too closely	−0.30	(−2.06, 1.46)
Fairly closely	−0.66	(−2.35, 1.04)
Very closely	−0.50	(−2.19, 1.18)
Personal life change as result of coronavirus		

		Beta (95% CI)
Major change	0.99	(0.61, 1.37)***
Minor change	0.11	(−0.24, 0.46)
Threat of coronavirus for US population health		
Major threat	−0.40	(−1.21, 0.4)
Minor threat	−0.66	(−1.45, 0.14)
Threat of coronavirus for personal health		
Major threat	1.37	(1.00, 1.73)***
Minor threat	0.69	(0.37, 1.00) ***
Laid off/lost job due to COVID (household)		
Yes	0.04	(−0.25, 0.34)
Pay cut, reduced hours due to COVID (household)		
Yes	0.32	(0.06, 0.58)*
Threat of coronavirus to US economy		
Major threat	1.08	(0.18, 1.97)*
Minor threat	0.96	(0.03, 1.90) *
Threat of coronavirus for personal finances		
Major threat	0.51	(0.18, 0.84)***
Minor threat	0.04	(−0.27, 0.35)

Multivariable linear regression models estimating association between explanatory variables and psychological distress score, in a nationally representative sample of individuals in the United States with no history of mental disorder (N=9687).