

1                   **Temporal coherence between lake and landscape primary productivity**  
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28                   **Highlights:**

29                   • Lake and watershed primary production can vary coherently  
30                   • The magnitude, direction and time lag of coherence relationships vary widely  
31                   • Hydrologic connectedness is a key feature mediating the lake-landscape coherence  
32                   relationship

34     **Abstract:** Understanding the patterns and drivers of primary productivity is a major goal of  
35     ecology, but little is known about whether the primary productivities of different types of  
36     ecosystems—here, lakes and the landscapes in which they are embedded—fluctuate in related  
37     ways through time. Due to shared climatic variation and well-known connections between lake  
38     and terrestrial ecosystems, such as nutrient and resource subsidies, we hypothesized that  
39     interannual fluctuations in aquatic and terrestrial primary productivity indices could be coherent.  
40     We also expected that lake and watershed characteristics could modify the strength and nature of  
41     primary productivity relationships. We applied wavelet coherence analyses to time series of lake  
42     chlorophyll-a and satellite-derived NDVI to examine coherence between lakes and land, and  
43     used random forest regression and generalized additive models to evaluate why coherence varies  
44     among lakes. There can be substantial coherence between lake and terrestrial primary  
45     productivity, but the strength and phase (direction and time lag) of this relationship varies  
46     widely, and there were marked differences between short (2-4 year periods of oscillation) and  
47     long (>4 year periods of oscillation) timescales. Across all timescales, variables associated with  
48     the connectedness of lakes to their watersheds were consistently the important explanatory  
49     variables of the strength and phase of coherence. The patterns observed in this study suggest the  
50     importance of cross-ecosystem flows, as opposed to shared climatic variation, in determining  
51     temporal coherence between lakes and the landscape.

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53     **Keywords:** synchrony, compensation, chlorophyll-a, NDVI, resource subsidies, hydrologic  
54     connectivity

55 **Introduction**

56        Quantifying the primary productivity of ecosystems, and how and by what mechanisms  
57    this process varies spatiotemporally, is a fundamental goal of ecology, especially in light of  
58    global climate change (Schlesinger and Bernhardt 2013). There are rich, but separate, literatures  
59    on the rates and patterns of primary productivity in lakes (Carpenter and others 1985; Dodson  
60    and others 2000; Downing 2009; Seekell and others 2018) and in terrestrial landscapes (Melillo  
61    and others 1993; Running and others 2000; Schimel and others 2001; Anav and others 2015).  
62    Less is known about the relatedness of temporal fluctuations in primary productivity between  
63    lakes and the landscape. Aquatic and terrestrial ecosystems are linked through substantial and  
64    temporally variable fluxes of matter and energy (Cole and others 2007; Harrison and others  
65    2009; Butman and others 2016), which plays an important role in regional and global carbon  
66    cycles (Tranvik and others 2009; Buffam and others 2011). Given these factors, whether there  
67    are persistent relationships in temporal fluctuations (i.e., coherence) in primary productivity  
68    between lakes and the landscape has substantial implications for understanding the dynamic  
69    linkages between these systems. Coupling of temporal patterns in primary productivity across  
70    different ecosystem types has received little attention with the exception of studies focused on  
71    synchronization of terrestrial and coastal marine ecosystems (Ong and others 2016; Black and  
72    others 2018; Lara and others 2019).

73        Temporally coherent fluctuations in terrestrial and aquatic primary productivity might  
74    plausibly arise from two general mechanisms. The systems might both be influenced, directly  
75    and independently, by the same exogenous environmental drivers (e.g., climate and weather),  
76    bringing them into coherence. This mechanism is analogous to the Moran effect, in which a  
77    shared environmental driver synchronizes spatially disjunct populations (Moran 1953). While

78 spatial synchrony among lakes has been observed for variables like surface water temperature  
79 (Magnuson and others 1990), it is unclear if exogenous environmental drivers also result in  
80 coherence between different ecosystem types. Coherence could also arise from flows of carbon  
81 and nutrients across ecosystem boundaries. One possible scenario is that terrestrially-fixed  
82 carbon enters a lake as dissolved organic matter, reducing light availability and therefore  
83 photosynthesis (Karlsson and others 2009; Solomon and others 2015). Alternatively, nutrients  
84 entering a lake along with organic matter can stimulate primary productivity (Thrane and others  
85 2014; Corman and others 2018; Kelly and others 2018), or terrestrially-fixed carbon could affect  
86 aquatic primary productivity by providing subsidies to consumers (Tanentzap and others 2017).

87 These two mechanisms (Moran-like effects and cross-ecosystem flows) may not be  
88 mutually exclusive, as climate and weather may be important drivers of temporal variability in  
89 the magnitude of cross-ecosystem flows. For example, anthropogenic nutrient enrichment of  
90 terrestrial ecosystems enhances terrestrial primary production and could also drive variability in  
91 aquatic primary production through episodic nutrient loading from the watershed, for example in  
92 runoff from precipitation events (Kelly and others 2019; Stockwell and others 2020). Although  
93 nutrient loading is an exogenous forcing that may be shaped by weather (e.g., precipitation), we  
94 distinguish between this and the pure Moran-like effect because of the likelihood of vegetation  
95 and watershed processes determining the magnitude and timing of inputs to the lake, and  
96 creating time lags between nutrients' stimulating effects on primary productivity on the  
97 landscape and in lakes. While the Moran-like independent effects of weather could be positive in  
98 one system and negative in the other, we do not expect them to be time-lagged given the short  
99 response times of plant photosynthetic rates and high turnover rates of phytoplankton.

100 Properties of lakes and watersheds likely shape the strength and direction of relationships  
101 between aquatic and terrestrial primary productivity, and time-lags between these variables.  
102 Lakes are differentially affected by catchment processes based on their position in the watershed,  
103 with lower lakes tending to be more strongly influenced (Kratz and others 1997; Martin and  
104 Soranno 2006). Lake morphology could also play an important role, with shallower lakes (Qin  
105 and others 2020) or those with greater shoreline development indices (Scheuerell and Schindler  
106 2004) potentially experiencing stronger influences of the adjacent terrestrial ecosystem. What  
107 dominates cross-ecosystem flows could also help determine whether lakes and the landscape  
108 tend to be positively or negatively related. If the dominant mechanism is nutrient inputs from  
109 land to lake, then the two systems may be positively related; whereas if inputs of dissolved  
110 organic carbon (DOC) from land to lake dominate, then reductions in water clarity could inhibit  
111 lake primary productivity (Karlsson and others 2009; Solomon and others 2015), resulting in a  
112 negative relationship. Time lags between lake and landscape primary productivity could reflect  
113 time differences between carbon fixation and release (Kuzyakov and Gavrichkova 2010), and  
114 lags associated with processing and transport through the watershed (Harman 2015). Thus,  
115 temporal relationships in primary productivity between the lake and the landscape provide  
116 insight into the drivers of those dynamics and the strength of the linkages among these  
117 ecosystems.

118 Temporal relationships have often been studied using correlation, but approaches based  
119 on the wavelet transform (Torrence and Compo 1998) have grown in popularity because of their  
120 ability to resolve common patterns that confound standard correlation (Vasseur and Gaedke  
121 2007; Downing and others 2008; Sheppard and others 2016; Walter and others 2017). Ecology is  
122 multi-causal and mechanisms often have specific timescales of variation, so timescale-specific

123 relationships can reflect specific drivers (Defriez and Reuman 2017; Sheppard and others 2019;  
124 Wilkinson and others 2020). Standard correlation can fail to resolve timescale-specific and  
125 temporally-lagged effects (Sheppard and others 2016; Walter and others 2019). Time lags result  
126 from common processes including transport times between observation points, intervals between  
127 generations, and certain predator-prey relationships. Wavelet coherence overcomes both  
128 problems by indicating the strength of relatedness between two variables, and the time lag (phase  
129 difference) between them, as a function of timescale (Grinsted and others 2004; Sheppard and  
130 others 2017). Figure 1 illustrates a timescale-specific relationship and examples of phase  
131 differences between two variables.

132 To investigate patterns of temporal coherence in primary productivity between lakes and  
133 the landscapes in which they are embedded, we analyzed 135 long term ( $\geq 20$  year) paired lake  
134 and land time series in the northeastern USA. We focus specifically on the following questions.  
135 Q1) To what extent are multi-annual [*sensu* Wilkinson *et al.* (2020)] patterns of lake primary  
136 productivity coherent with primary productivity in the surrounding landscape? Q2) What are the  
137 phase differences between them, and what does this imply about the predominant mechanisms  
138 coupling primary productivity in lakes and the landscape? Q3) What factors explain variability in  
139 the magnitude and phase of coherence between lake and landscape primary productivity? Q4) Do  
140 the answers to questions Q1 through Q3 depend on timescale? We found that on multi-annual  
141 timescales, landscape and lake productivity vary widely in their coherence and the phase  
142 differences between landscape and lake. The coupling relationships appear to mainly reflect  
143 flows of carbon and nutrients between the ecosystem types.

144

145 **Methods**

146 *Data acquisition and processing*

147 We obtained time series data on chlorophyll-a, an indicator of primary production, from  
148 the LAGOS-NE LIMNO database, version 1.087.3 (Soranno and others 2017; Soranno and  
149 Cheruvilil 2019). The LAGOS-NE database contains time series of physical, chemical, and  
150 biological parameters from lakes located in seventeen states of the northeastern and mid-western  
151 portions of the conterminous United States. Data were aggregated to annual intervals by  
152 averaging measurements from May-September, the period of most active growth. If a lake had  
153 fewer than three observations during a growing season, data for that year were discarded. Lake  
154 time series selected for analysis spanned a minimum of 20 years, could have a maximum of two  
155 years with missing data, and years with missing data were non-consecutive. Because our  
156 analyses require complete time series, missing data were filled with the median of the time  
157 series. A number of lakes in LAGOS-NE narrowly missed satisfying these criteria, but because  
158 this version of LAGOS-NE contains data through the year 2013, we augmented the dataset by  
159 obtaining additional, freely available data on a total of 38 lakes from the Wisconsin Department  
160 of Natural Resources (14 lakes) and Minnesota Pollution Control Agency lake monitoring  
161 programs (24 lakes). These data were processed to match the format of the LAGOS-NE data. In  
162 total, 135 lake time series were analyzed. A map of lakes and selected examples of coherent and  
163 non-coherent time series is provided in Figure S1.

164 We also obtained from LAGOS-NE, or derived from data contained therein, a suite of 53  
165 variables describing lake morphometry, watershed land cover, hydrologic connections, landscape  
166 position, atmospheric deposition, glaciation history, and climate for each lake (Soranno and  
167 Cheruvilil 2017). Surficial geology was also considered but could not be adequately addressed

168 due to the lack of data for most variables in this category. A complete list and a criterion for  
169 culling the list of potential variables are given in Supplementary Material S1. These variables  
170 reflect conditions either of the lake itself, in the watershed as delineated in LAGOS-NE (Soranno  
171 and others 2017), or within the Hydrologic Unit Code (HUC) Level-12 unit. Hydrologic Unit  
172 Codes are a hierarchical system for identifying watersheds, of which level-12 (i.e., a 12-digit  
173 identifying code) is the finest. In our dataset, there was only one lake in 101 of 114 HUC-12  
174 units, and at most 5 lakes. LAGOS-NE was accessed through the LAGOSNE R package  
175 (Stachelek and Oliver 2019).

176 We used growing season accumulated normalized difference vegetation index (NDVI)  
177 data as a proxy for annual primary productivity in the terrestrial landscape surrounding each  
178 lake. To develop this dataset, start-of-season (SOS) and end-of-season (EOS) dates were  
179 estimated annually using the 30-year (1989-2018) Advanced Very High Resolution Radiometer  
180 (AVHRR) NDVI time series dataset available for the conterminous U.S. (United States Geologic  
181 Survey). These raster image data, which have 1-km spatial resolution, consist of weekly issued,  
182 biweekly maximum value composite NDVI scores (Eidenshink 1992, 2006). Source imagery  
183 was collected almost daily, so that the biweekly maximum value composite NDVI scenes are  
184 largely cloud-free (clouds, ice, and snow tend to suppress NDVI values). AVHRR sensors on  
185 multiple satellite platforms have been utilized over the years to provide a gapless time series.

186 For each pixel and each year, SOS was determined using the midpoint of the methods  
187 described in Zhang and others (2003) and Yu and others (2004). Considering the temporal  
188 uncertainties associated with the NDVI time series (e.g. the precise date of acquisition for each  
189 pixel's value in a composite image is not known), time steps were assigned to integers in [1:52]  
190 representing calendar week of issue. Both SOS approaches are somewhat sensitive to high

191 frequency noise, so to facilitate consistent SOS estimation, NDVI time series were initially  
192 smoothed using time series tail and minima treatment methods adapted from Wardlow and others  
193 (2006). To obtain EOS, NDVI time series were reversed prior to processing, with the result  
194 subtracted from 53 to obtain the correct temporal position. Growing season accumulated NDVI  
195 was then determined by summing NDVI values from SOS to EOS. Some years for some pixels  
196 representing water or barren lands did not satisfy NDVI threshold-based criteria for vegetation  
197 presence, and in those cases no growing season accumulated NDVI was assigned.

198 Using only AVHRR pixels consisting of <5% water based on the National Land Cover  
199 Database (NLCD 2011; Homer and others 2015), we averaged annual terrestrial NDVI within a  
200 radius dependent on the surface area of the lake. Prior research established that there is a power-  
201 law relationship between lake surface area and watershed area, and that the intercept (but not the  
202 scaling parameter) differs between lakes and reservoirs (Walter and others 2020). We set a  
203 minimum radius of 2.5 km to ensure an adequate sample of terrestrial pixels. Otherwise, we used  
204 the simplifying assumption that lakes are circular to derive the following equation to scale the  
205 search radius to the surface area of the lake:

$$206 \quad r = \sqrt{\frac{10^{\beta_0 + \beta_1 \log_{10} A_l}}{\pi}}$$

207 Here,  $\beta_0$  and  $\beta_1$  are empirically estimated regression coefficients that differed between lakes and  
208 reservoirs, and  $A_l$  is the surface area of the lake. For reservoirs,  $\beta_0 = 1.8$  and  $\beta_1 = 1.05$ , and for  
209 lakes  $\beta_0 = 1.4$  and  $\beta_1 = 1.05$  (Walter and others 2020). We also considered using watershed  
210 boundaries as spatial units for NDVI time series; however, a number of small lakes had too few  
211 suitable AVHRR pixels within their watersheds and so were discarded. For the remaining lake-  
212 landscape pairs, wavelet coherence test results (see *Analyses*) were entirely consistent whether

213 NDVI time series were determined from circular buffers or from watershed boundaries, and so  
214 we present results using circular buffers.

215

216 *Analyses*

217 We tested for wavelet coherence (Grinsted and others 2004; Sheppard and others 2017)  
218 between time series of lake chlorophyll and terrestrial NDVI to determine the strength and phase  
219 of temporally persistent relationships between lake and land primary productivity. Wavelet  
220 coherence quantifies the degree to which two time series have correlated magnitudes of  
221 oscillation and consistent phase differences through time, as a function of timescale. Its  
222 magnitude ranges from 0 (no relationship) to 1 (perfect coherence). As we focus on consistent, as  
223 opposed to transient or episodic, relationships, we considered coherence over the full time series.  
224 Significance testing was performed by comparing the empirical coherence to a distribution of  
225 surrogate coherences generated under a null hypothesis of no coherence that retained the spectral  
226 (i.e., temporal autocorrelation) properties of the empirical time series (Sheppard and others  
227 2017). This test is conservative because in the procedure by which surrogates are generated only  
228 the phase of oscillations varies, while the power spectrum of each time series is entirely  
229 preserved. In reality, if one variable drives fluctuations in a second variable, the power spectrum  
230 of the second variable is influenced by the first. However, incorporating such effects is difficult,  
231 and failure to do so appropriately would result in potentially identifying false relationships.

232 To determine to what extent interannual patterns of lake and terrestrial primary  
233 productivity are coherent (Q1), we examined the distribution of coherence magnitudes at short (2  
234 to 4-year periods of oscillation) and long (>4-year periods of oscillation) timescale bands, and  
235 compared the number of statistically significant coherences to the number expected under a

236 false-positive error rate of  $\alpha = 0.05$ , assuming independent tests. The period length of 4 years  
237 was chosen to demarcate short from long timescales because it separates anti-persistent patterns  
238 (i.e., successive observations tend to be negatively correlated) from persistent patterns (i.e.,  
239 successive observations tend to be positively correlated) (Sheppard and others 2016). Choosing  
240 focal timescale bands *a priori* facilitates significance testing (Sheppard and others 2016).

241 To assess the prevalence of different phase relationships between lake and terrestrial  
242 primary productivity (Q2), we examined the distribution of coherence phases for coherence  
243 relationships exceeding the 70<sup>th</sup> percentile of a distribution of surrogate coherences generated  
244 under a null hypothesis of no coherence but preserving the spectral properties of each time series.  
245 Only relatively strong coherences were used because for incoherent variables the phase  
246 difference is essentially a random variable uniformly distributed between  $-\pi$  and  $\pi$ , and hence is  
247 not meaningful. Short and long timescales were again considered separately.

248 To assess what factors explain variability in the magnitude and phase of coherence  
249 between lake and landscape primary productivity (Q3), we combined “feature selection” using  
250 conditional random forest regression (Hothorn and others 2006) with generalized additive  
251 models (Wood 2006). Because phases are angular measurements, we performed analyses on the  
252 sine- or cosine-transformed phase difference between lake and landscape primary productivity  
253 indices. Taking the cosine assigns in-phase relationships ( $\phi = 0$ ) to a value of 1, anti-phase  
254 relationships ( $\phi = \pm \pi$ ) to a value of -1, and quarter-phase relationships ( $\phi = \pm \pi/2$ ) to a value of  
255 0. Consequently, analysis of  $\cos(\phi)$  focuses on how close the relationship is to being in-phase.  
256 Taking the sine transforms to a value of 0 for both in-phase and anti-phase relationships; to -1 for  
257 a relationship in which peaks in chlorophyll-a lag NDVI by  $1/2$  a cycle length, or in other words a  
258 time-lagged positive relationship; and to 1 for a relationship in which peaks in chlorophyll-a lead

259 NDVI by  $\frac{1}{2}$  a cycle length, or in other words a time-lagged negative relationship, assuming that  
260 the dominant direction of flows is from land to lake. Consequently, analysis of  $\sin(\phi)$  focuses on  
261 whether the time-lagged relationship between NDVI and chlorophyll-a tends to be positive or  
262 negative. Whether to apply the sine or cosine transformation depended on the dominant mode of  
263 variability in  $\phi$  for a timescale band.

264 Conditional random forests were used to select the most important predictors from the  
265 suite of 53 predictor variables described above and in Supplementary Material S1. Conditional  
266 random forests are an ensemble machine learning technique based on classification and  
267 regression trees (Hothorn and others 2006). We fit conditional random forests consisting of  
268 50,000 trees for each predictor variable, and quantified variable importance values for each  
269 predictor. Variable importance values and the Pearson correlation between empirical values and  
270 model predictions were stable at this forest size.

271 We used generalized additive models (GAMs) to investigate statistical effects and overall  
272 explanatory power of selected predictors on the timescale-specific coherence and phase between  
273 chlorophyll-a and NDVI. GAMs replace regression coefficients with penalized regression  
274 splines, thereby identifying nonlinear relationships while balancing parsimony and model  
275 complexity (Wood 2006). We built GAMs with 3 predictors for the response variables short  
276 timescale coherence magnitude ( $n = 135$ ), long timescale coherence magnitude ( $n = 135$ ), short  
277 timescale phase ( $n = 39$ ) and long timescale phase ( $n = 45$ ). Predictors were chosen for inclusion  
278 in rank order of variable importance, skipping variables with strong concurvity with a higher-  
279 ranked predictor. Concurvity is a generalization of collinearity used with GAMs. We deemed  
280 estimated concurvity values  $< 0.6$  to be acceptable; GAMs are highly robust to concurvity (Wood  
281 2008). Observations were weighted by time series length to give greater weight to longer time

282 series, for which we have greater certainty in the nature of lake-landscape coherence. Since  
283 coherence magnitudes are bounded between 0 and 1, we used a beta distribution with the GAM  
284 models for these variables. Some predictors were  $\log_{10}$  or square-root transformed to reduce the  
285 influence of extreme values. Because our goal for this analysis was to explore relationships  
286 explaining variability in temporal coherence between lake and land primary productivity indices,  
287 as opposed to testing *a priori* hypotheses about these potential drivers, we did not apply  
288 significance testing or further model selection/model parsimony methods to our GAMs.

289 The timescale specificity of the magnitude, phase, and predictors of coherence (Q4) was  
290 determined by comparing results from short timescales (2 to 4-year periods of oscillation) versus  
291 from long timescales (>4-year periods). Analyses were carried out in R version 3.5.1 (R Core  
292 Team 2018) using the “wsyn” (Reuman and others 2019), “party” (Hothorn and others 2019),  
293 and “mgcv” (Wood 2006) packages.

294

## 295 **Results**

296 Coherences between lakes and landscape primary productivity varied widely, spanning  
297 nearly 0 to 1, the entire range of the statistic (Q1). Example time series from a particularly  
298 coherent lake-landscape pair and a particularly non-coherent lake-landscape pair are shown in  
299 Fig. S1. At short timescales, coherence ranged 0.06 to 0.92 with a median of 0.37 (Fig. 2a). Nine  
300 lake-landscape pairs were significantly coherent at short timescales using  $\alpha = 0.05$  significance  
301 level. At long timescales, coherence magnitudes ranged 0.08 to 0.95, with a median of 0.53, and  
302 fifteen lake-landscape pairs were significantly coherent using  $\alpha = 0.05$  significance level (Fig.  
303 2b). In each case, the number of significant coherences is greater than the number expected by  
304 chance given the selected type-1 error rate (i.e.,  $135*0.05 = 6.75$ ) assuming independent tests.

305 This result highlights that there is robust evidence that *some* lake-landscape pairs are strongly  
306 coherent, but we emphasize that it is also meaningful that there is a wide range of observed  
307 coherences. The spatial distribution of coherences at short (Figure 3) and long (Figure S2)  
308 timescales indicates no apparent regional patterns in lake-landscape coherence. Lake-landscape  
309 coherences were substantially timescale-specific (Q4). Although coherence magnitudes were  
310 somewhat greater at long timescales than short, wavelet coherence suffers from a bias wherein  
311 greater values tend to be returned at long timescales, so care should be taken in comparing  
312 coherence magnitudes across timescales. Our significance tests do not suffer the same bias,  
313 however. There was no correlation between coherence magnitudes at short timescales and  
314 coherence at long timescales (Pearson correlation = 0.03), and only three lakes were significantly  
315 coherent at both short and long timescales.

316 Phase relationships between lakes and the landscape also spanned the range of possible  
317 values (0 to  $\pm\pi$ ), but certain phase relationships were more common than others (Q2). At short  
318 timescales, most coherence relationships were approximately in-phase (Fig. 2c). At long  
319 timescales, coherence relationships were bimodally distributed with most coherent lakes  
320 exhibiting either phase-lagged positive ( $-\pi/4 < \phi < -3\pi/4$ ) or negative ( $3\pi/4 < \phi < -\pi/4$ )  
321 relationships with terrestrial NDVI (Figure 2d).

322 The coefficient of variation in NDVI and variables associated with modulation of flows  
323 between lake and the landscape (herein termed “hydrologic connectedness”) tended to be the  
324 best predictors of lake-landscape coherence and phase relationships (Q3). Such variables  
325 included wetland cover and shoreline, precipitation and runoff, and groundwater recharge.  
326 Variation in the magnitude of coherence at short timescales was best explained by total nitrogen  
327 deposition, the temporal coefficient of variation in terrestrial NDVI [cv(NDVI)], and several

328 variables related to wetland cover in the watershed and adjacent to the lake (Fig. 4a). A GAM  
329 model including the top 3 predictors explained 14.6% of deviance in short timescale coherence,  
330 and featured a negative effect of total N deposition, and positive effects of cv(NDVI) and  
331 herbaceous wetland cover (Fig. 5a-c). Among lakes exhibiting substantial coherence with the  
332 landscape, for which computed phase relationships are reliable, lakes with more open water  
333 wetlands on their shoreline and with higher average chlorophyll-a concentrations (i.e., more  
334 eutrophic lakes) were more likely to be in-phase with the landscape, but those with a high  
335 percentage of shrub/scrub landcover were less likely to be in-phase with the landscape (Fig. 5d-  
336 f). A GAM model with these predictors explained 20.9% of deviance in  $\cos(\phi)$ .

337 At long timescales, coherence was best explained by the percentage of shrub/scrub land  
338 cover in the watershed, cv(NDVI), annual precipitation, and the percentage of watershed area  
339 composed of semi-permanently flooded (regime f) wetlands (Fig. 4c). Regime f wetlands are  
340 semi-permanently flooded where surface water persists throughout the growing season in most  
341 years. Long-timescale coherence decreased with increasing percentages of shrub/scrub landcover  
342 in the watershed and with the percentage of watershed area covered by semi-permanently  
343 flooded wetlands, and increased with increasing cv(NDVI) (Fig. 5g-i). This combination of  
344 predictors explained 13.5% of deviance in coherence. Phase relationships at long timescales were  
345 most strongly influenced by wetland shoreline and area, and other land cover types (Fig. 4d).  
346  $\text{Sin}(\phi)$  tended to decline with increasing open-wetland shoreline, notwithstanding a few outliers,  
347 indicating a tendency toward time-lagged negative effects of NDVI on chlorophyll-a fluctuations  
348 (Fig. 5j). Increases in the areal percentage of semi-permanently flooded wetlands in the  
349 watershed were also associated with time-lagged negative effects, and the percentage of

350 cultivated crops was associated with time-lagged positive effects (Fig. 5i, l). These predictors  
351 explained 31.9% of deviance in  $\sin(\phi)$ .

352

### 353 **Discussion**

354 There is wide variability in the coherence of indices of primary productivity between  
355 lakes and the surrounding landscape, from complete incoherence to near-perfect coupling across  
356 interannual timescales. Although the presence of links between terrestrial and aquatic systems is  
357 well-known (Wilkinson and others 2013; Tanentzap and others 2017; Tranvik and others 2018),  
358 our findings underscore the importance of temporally dynamic links between aquatic and  
359 terrestrial ecosystems. Although lakes typically make proportionally small contributions to  
360 regional primary productivity, except in particularly lake-rich regions, aquatic ecosystems  
361 process large amounts of terrestrially-derived nutrients and organic matter (Cole and others  
362 2007; Downing 2009). Understanding whether these inputs stimulate or inhibit lake primary  
363 productivity, and the mechanisms that couple these systems, is important especially on  
364 timescales from a few years to decades [i.e., multi-annual timescales *sensu* Wilkinson and others  
365 (2020)]. For example, episodic nutrient inputs may stimulate algal blooms while precipitation  
366 driven dissolved organic matter inputs may have longer term effects on lake primary producers.

367 We proposed that coherence between lakes and the landscape could be driven by shared  
368 responses to weather, potentially mediated by flows of carbon and nutrients across ecosystem  
369 boundaries, and by lake characteristics (Figure 6). Our results provide some inference into the  
370 importance of these mechanisms in the lakes we studied, specifically that lake-landscape  
371 coherence largely arises through cross-ecosystem flows of carbon and nutrients, with wetlands  
372 playing a key mediating role. If coherence was driven predominantly by shared, independent

373 Moran-like effects of exogenous drivers like weather, we should expect strong coherences to  
374 mainly be in-phase or anti-phase; instead, we see many time-lagged relationships, although at  
375 short timescales coherences were more often in-phase or anti-phase. Further evidence for this  
376 interpretation is that variables associated with hydrologic connections between lakes and the  
377 landscape—for example, wetlands area and shoreline, runoff, groundwater recharge, and  
378 headwater stream density—were among the most important predictors of the magnitude and  
379 phase of lake-landscape coherence. Hydrologic connectivity, via both surface waters and sub-  
380 surface flows, has a fundamental role in the spatiotemporal dynamics of lake ecosystems  
381 (Canham and others 2004; Martin and Soranno 2006; Fergus and others 2017). Wetlands are an  
382 important component of these connections, processing and exporting large amounts of carbon  
383 and nutrients (Detenbeck and others 1993; Gergel and others 1999; Martin and Soranno 2006).  
384 Many of these variables are also correlates of lake landscape position (Kratz and others 1997;  
385 Martin and Soranno 2006). The area of wetlands around the lake tends to increase moving from  
386 high in the watershed to low; these lakes also tend to accumulate inputs from larger areas, and  
387 also are more likely to have inflowing streams. However other correlates of lake landscape  
388 position including stream density, lake connection, and upstream lake area were also candidate  
389 predictor variables but were less influential.

390 Interestingly, the dominant mode of variability in phase differences among lakes that  
391 were coherent with the landscape on long timescales was between a time-lagged positive and a  
392 time-lagged negative effect. Negative effects of terrestrial primary productivity on aquatic  
393 primary productivity could reflect dissolved organic carbon (DOC) loading; DOC may inhibit  
394 photosynthesis by decreasing water clarity (Karlsson and others 2009; Solomon and others  
395 2015). In our analyses time-lagged negative relationships were associated with increasing

396 amounts of wetlands in the watershed and on the shoreline, and time-lagged positive  
397 relationships were associated with the more agricultural watersheds. DOC export to aquatic  
398 systems is associated with forests and wetlands (Gergel and others 1999; Canham and others  
399 2004), while agriculture can be a considerable source of nutrient runoff. The time lags between  
400 terrestrial and aquatic primary productivity may reflect both transport time, e.g., through slow  
401 pathways like groundwater, and the timing of biogeochemical transformations (Cardille and  
402 others 2007). For example, a typical pathway for DOC originating on land is for carbon to be  
403 fixed into leaves during the growing season, fall during autumn senescence, and decompose  
404 before entering the DOC pool.

405 Surprisingly, we found little evidence that lake-landscape coherence depended on  
406 characteristics of lakes themselves. We considered a number of predictors characterizing lake  
407 morphometry (maximum depth, shoreline development ratio) and biogeochemistry (mean  
408 chlorophyll-a), but only mean chlorophyll-a was a relatively important predictor of any response  
409 variable. Taken together with our results on the importance of hydrologic connectedness and  
410  $cv(NDVI)$  for lake-landscape coherence, it seems that lake-landscape coherence is largely  
411 imposed by the terrestrial landscape and mediated through flows of carbon and nutrients across  
412 the terrestrial-aquatic interface (Carpenter and others 1998; Buffam and others 2011).

413 Insights from the study of spatial synchrony in population dynamics (Liebhold and others  
414 2004; Walter and others 2017) suggest that weather variation and climate oscillations should be  
415 important in coupling these dynamics, in a sort of cross-ecosystem “Moran effect” (Moran  
416 1953). Indeed, spatially synchronous weather has been inferred to synchronize dynamics across  
417 lakes (Magnuson and others 1990; Baines and others 2000; Pace and Cole 2002; Pham and  
418 others 2008; Rusak and others 2008) and drive spatial synchrony in terrestrial primary

419 production (Koenig and Knops 1998; Wettstein and others 2011; Shestakova and others 2016;  
420 Defriez and Reuman 2017). However, weather may not synchronize lakes at the regional spatial  
421 extent of our study (Soranno and others 2019), and weather does not seem commonly to directly  
422 synchronize lakes with the landscape, at least on multi-annual timescales. While we did not  
423 explicitly test for weather drivers of coherence, the relative predominance of phase-lagged  
424 coherences—particularly at long timescales—and the relatedness of hydrologic connectedness to  
425 spatial variation in coherence between lakes and the landscape suggest that the proximal  
426 mechanisms of lake-landscape coherence likely have more to do with nutrient and organic matter  
427 subsidies from the landscape to the lake, than shared environmental forcing operating similarly,  
428 but separately and simultaneously on the lake and the landscape. This is in contrast to spatial  
429 synchrony in population dynamics, where climate can synchronize populations over large areas,  
430 even with little or no dispersal between them (Liebhold and others 2004). Other studies of cross-  
431 ecosystem synchrony between terrestrial and marine ecosystems have found climate to be an  
432 important driver (Ong and others 2016; Black and others 2018), but the relative magnitude of  
433 inputs from the land to lakes is likely larger than from the land to the ocean.

434 The magnitude of lake-landscape coherence was entirely uncorrelated across timescales,  
435 affirming the power of our timescale-specific approach. Standard approaches based on  
436 correlation or regression would not have uncovered this pattern, and would have been  
437 confounded by the strong differences between short and long timescales. Importantly, this also  
438 suggests that different underlying mechanisms are responsible for short- versus long-timescale  
439 coherence between lakes and the landscape. Particular mechanisms have characteristic  
440 timescales of variation and tend manifest on those same timescales, as has been shown in studies  
441 of population spatial synchrony (Sheppard and others 2016; Anderson and others 2019) and

442 inferred for some studies of synchrony in lakes (Baines and others 2000; Pace and Cole 2002).  
443 Although variables associated with hydrologic connectedness tended to explain among-lake  
444 variation in lake-landscape coherence, specifically what mechanisms are at play are not yet  
445 known. For example, the role of “flashy” events, such as storms resulting in overland flows and  
446 relatively high flow rates through riverine systems, versus slow-but-consistent processes like  
447 groundwater recharge, is unclear, but could be resolved by studying time series with sub-annual  
448 intervals between observations.

449 Our conclusions are limited by data drawn from a biased set of lakes (Stanley and others  
450 2019) with relatively short time series of error-prone indicator variables. Despite these  
451 limitations, we believe it would be inappropriate to dismiss evidence for lake-landscape  
452 coherence on these grounds. We chose NDVI and lake chlorophyll-a as indices of primary  
453 productivity because of their relative availability. Other measures, such as from eddy covariance  
454 for terrestrial ecosystems (Vesala and others 2006; Aubinet and others 2012) and in-situ  
455 measurements of oxygen dynamics and carbon fixation for lakes provide more direct  
456 measurements of primary productivity. However, their limited availability made them ill-suited  
457 to the goals of this study. Satellite remote sensing of lake water quality can enable study of  
458 additional lakes, but the need for calibration data (Ross and others 2019) and cloud cover mean  
459 that remotely-sensed time series would have similar limitations to our dataset. Additionally,  
460 while our time series are near the lower length limits where wavelet analyses can be reasonably  
461 applied, our data are among the longest records that currently exist, and requiring longer time  
462 series would have substantially reduced the number of lakes in the study, hindering our goals.  
463 Considering the limited statistical power afforded by our short time series and potential for  
464 observation error, the detection of strong relationships is meaningful.

465 This study provides evidence of wide variation in the coupling of temporal ecosystem  
466 dynamics between lakes and the landscapes they are embedded in, and provides a springboard  
467 for future work leveraging temporal pattern to understand causes and consequences of coupling  
468 between lakes and the landscape. Although further research is needed to clarify the specific  
469 mechanisms of temporal coherence between lakes and the landscape, variables associated with  
470 hydrologic connectedness explained substantial spatial variation in lake-landscape coherence,  
471 suggesting that cross-ecosystem flows may commonly be a proximal mechanism. Based on  
472 analogs with synchrony in populations and communities, and observations of coastal systems  
473 (Ong and others 2016; Black and others 2018; Lara and others 2019) we expected that shared  
474 climatic fluctuations could induce synchrony between lakes and the landscapes, but this does not  
475 seem prevalent. Finally, this study affirms that multi-annual dynamics, which particularly in  
476 aquatic ecosystems are less well understood compared to seasonal and short-term trends  
477 (Wilkinson and others 2020), exhibit rich patterns, and that their investigation can lead to new  
478 insights into ecosystem dynamics.

479

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484

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718 **Figure Captions**

719 Figure 1: Illustration of timescale-specific relationships between two variables (a) and different  
720 phase relationships (b-d). In (a) the blue and green signals are perfectly positively correlated on  
721 short timescales and perfectly negatively correlated at long timescales; this relationship would be  
722 confounded by standard correlation. In (b), fluctuations are in-phase ( $\phi = 0$ ), corresponding to  
723 positive correlation; in (c), fluctuations are temporally lagged, with the green signal peaking  
724 ahead of the blue signal ( $\phi = \pi/2$ ); in (d) fluctuations are anti-phase ( $\phi = \pi$ ), corresponding to  
725 negative correlation.

726

727 Figure 2: Distributions of coherence magnitudes (a, b) and phase differences (c, d) at short  
728 timescales (2-4 years; panels a, c) and at long timescales (>4 years; panels b, d). In (c, d),  
729 frequency is proportional to radius length.

730

731 Figure 3: Map of lakes included in this study by short-timescale coherence between chlorophyll-  
732 a and NDVI. Black outlines indicate statistically significant coherence. See Figure S2 for long  
733 timescales.

734

735 Figure 4: Variable importance values (in descending order) from conditional random forest  
736 analyses of among lake variability in the coherence magnitude (a, c) and phase difference (b, d)  
737 between primary productivity in lakes and the landscape. Results for short timescales (2-4 years)  
738 are shown in panels a, b; those for long timescales (>4 years) in panels c, d. Wetland variables  
739 are drawn in brown; lake variables are in blue; terrestrial vegetation variables are colored in blue;  
740 climate and atmospheric deposition variables are in white; and variables not fitting these

741 categories are in grey.

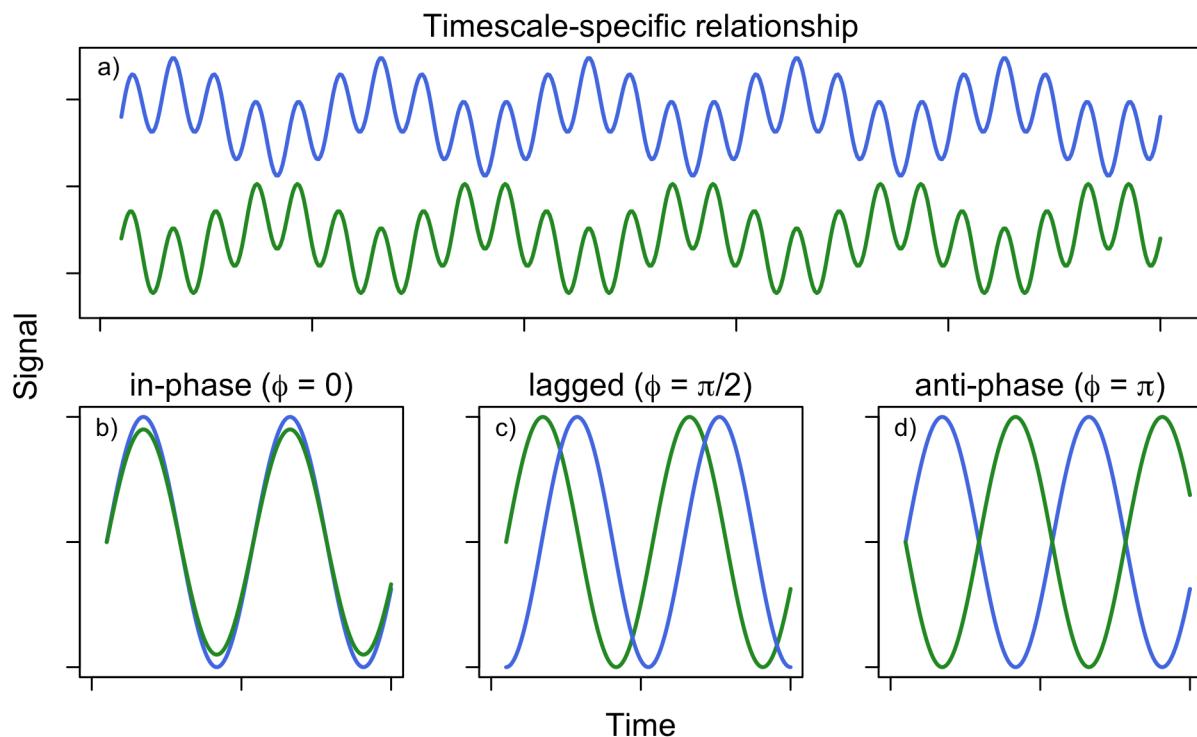
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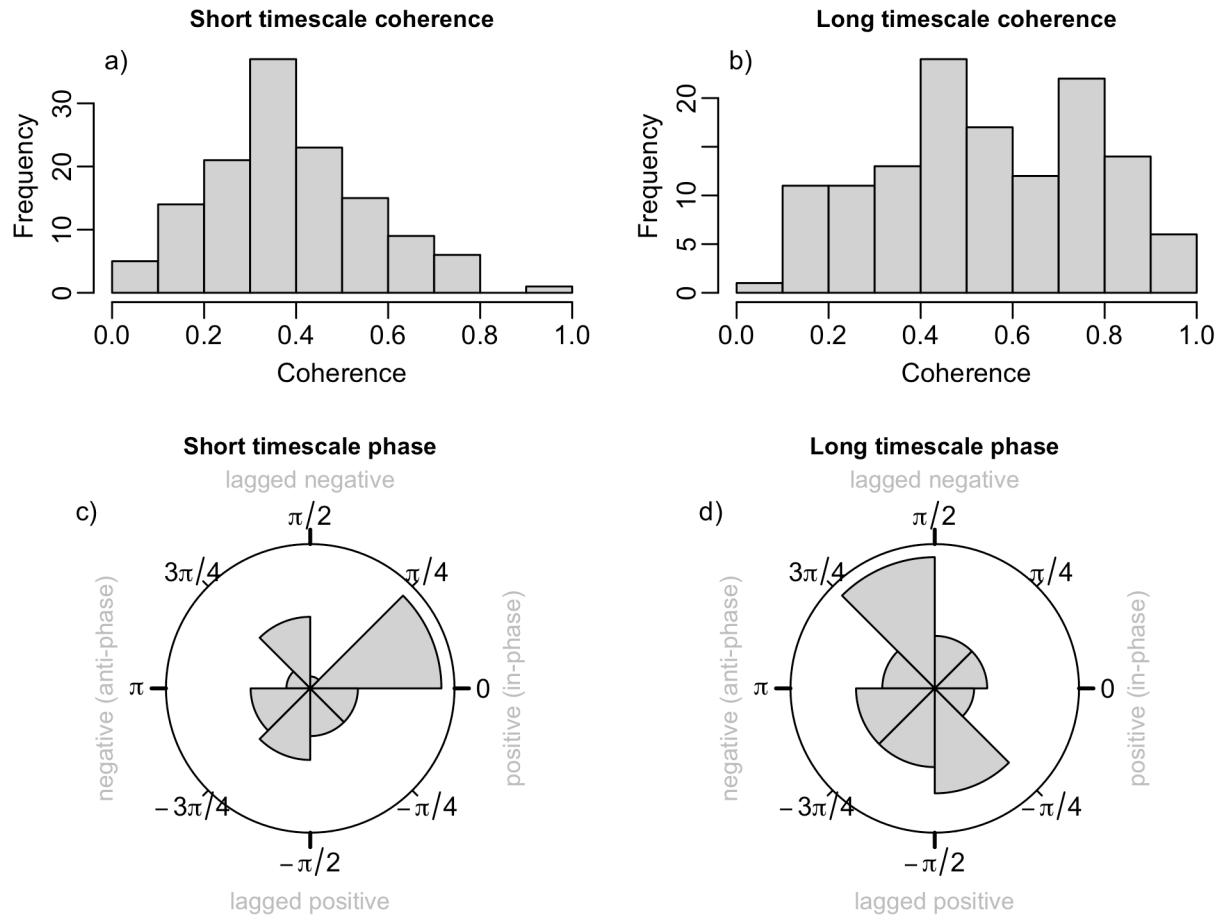
743 Figure 5: GAM partial effect plots depicting effects of top predictors on (a-c) short timescale  
744 coherence; (d-f) cosine-transformed short-timescale phase difference; (g-i) long timescale  
745 coherence; (j-l); sine-transformed long-timescale phase difference. Models explained,  
746 respectively, 14.6%, 20.9%, 13.5%, and 31.9% of deviance in the response variable. Grey  
747 regions indicate  $\pm 2$  standard errors.

748

749 Figure 6: Diagram of hypothesized mechanisms underpinning lake-landscape coherence results.  
750 (1) Weather may cause shared, independent effects on terrestrial and aquatic primary production  
751 leading to in-phase or anti-phase coherence. (2) Dissolved organic matter from terrestrial  
752 primary production may carry nutrients that stimulate aquatic primary production or decrease  
753 light availability in lakes. Land cover shapes the quantity and quality of terrestrial dissolved  
754 organic matter delivered to lakes; flow paths and terrestrial decomposition contribute to time  
755 (phase) lagged effects. (3) The form and magnitude of nutrient loading from the watershed in to  
756 lakes is influenced by land use. Nutrients that stimulate production in the watershed can also  
757 stimulate primary production in the lake at a phase lag based on the bioavailability of the  
758 exogenous inputs and the timing of delivery to the receiving aquatic ecosystem. (4) Wetlands  
759 modify the coherence between lakes and the landscape by altering flow paths. Wetlands retain  
760 water that would otherwise be immediately delivered downstream and process nutrients and  
761 dissolved organic matter, altering the quality and quantity of material delivered downstream. (5)  
762 Lake characteristics such as the size, depth, and food web structure potentially modify these  
763 influences. Artwork attribution: graphics from Integration and Application Network, University

764 of Maryland Center for Environmental Science ([ian.umces.edu/imagelibrary/](http://ian.umces.edu/imagelibrary/)); pine tree by  
765 Tracey Saxby, corn stalk by Jane Thomas, pondweed by Dieter Tracy, carp by Kate Moore, and  
766 largemouth bass and *Daphnia* by Kim Kraeer and Lucy Van Essen-Fishman.  
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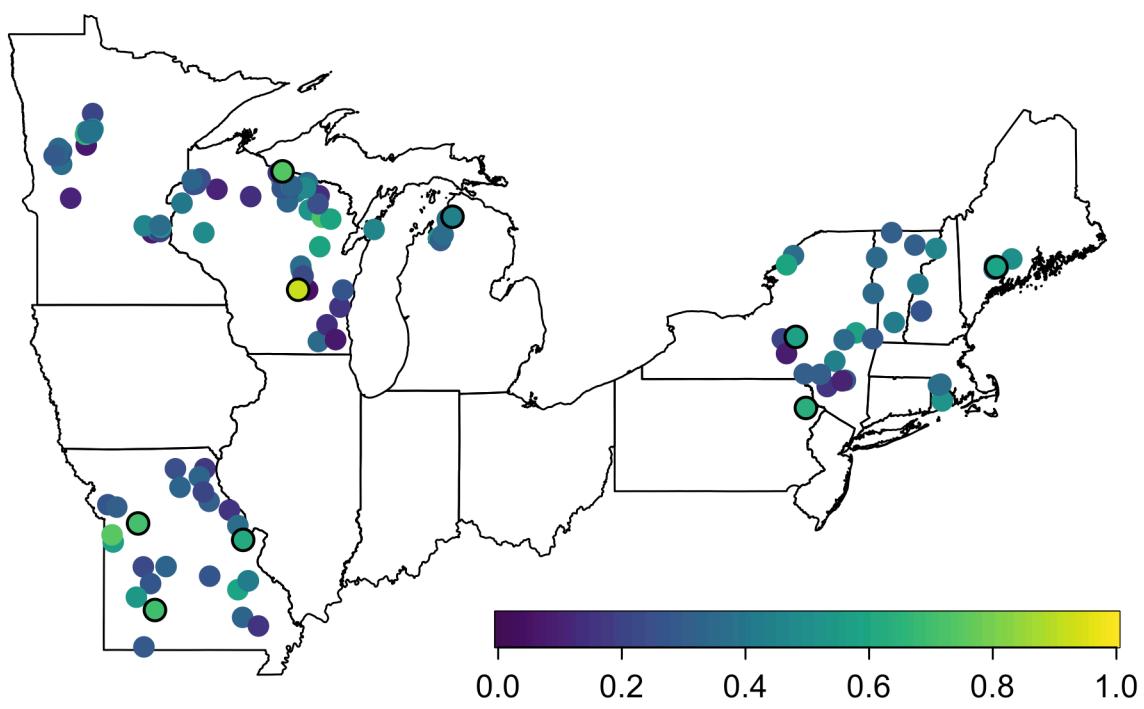




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Figure 2.

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Figure 3.

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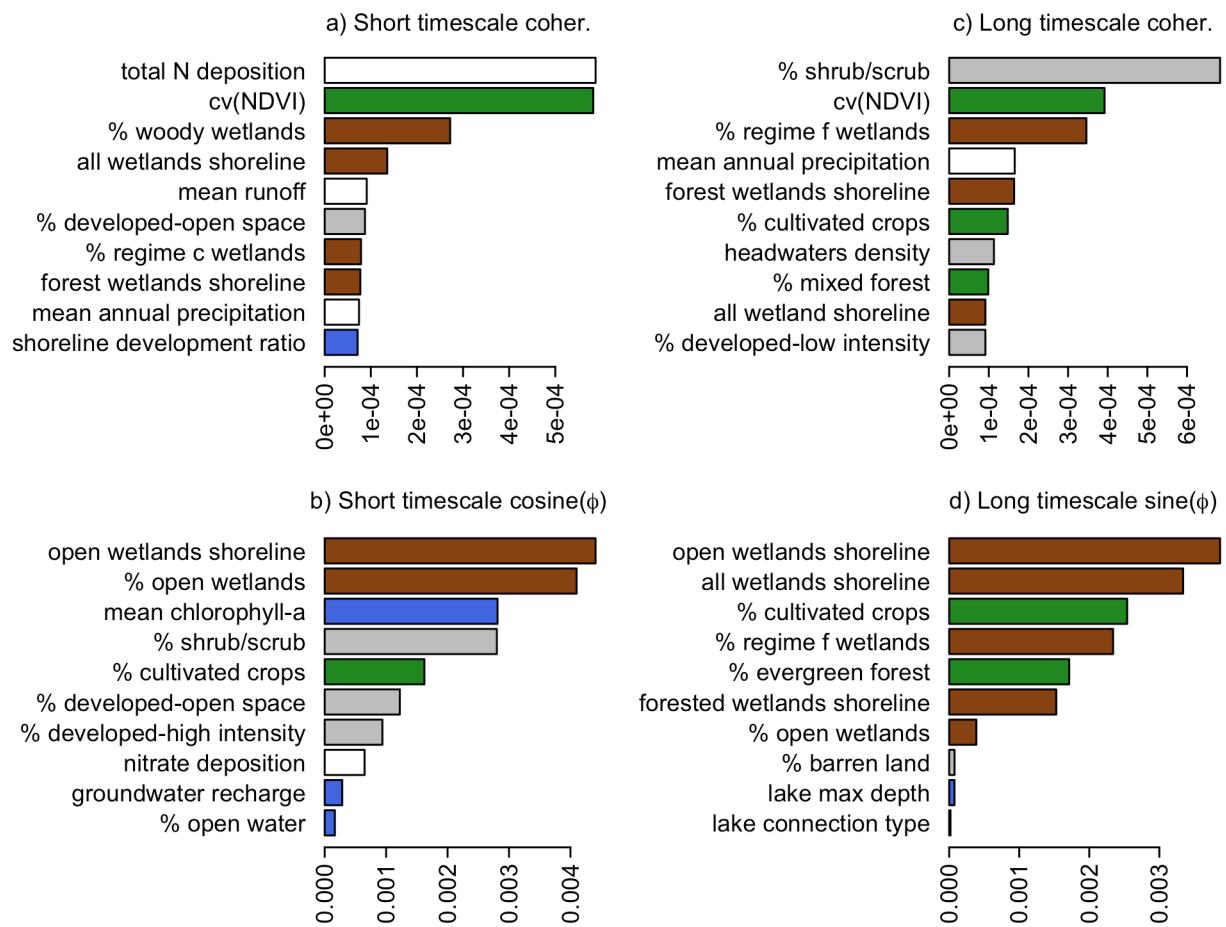
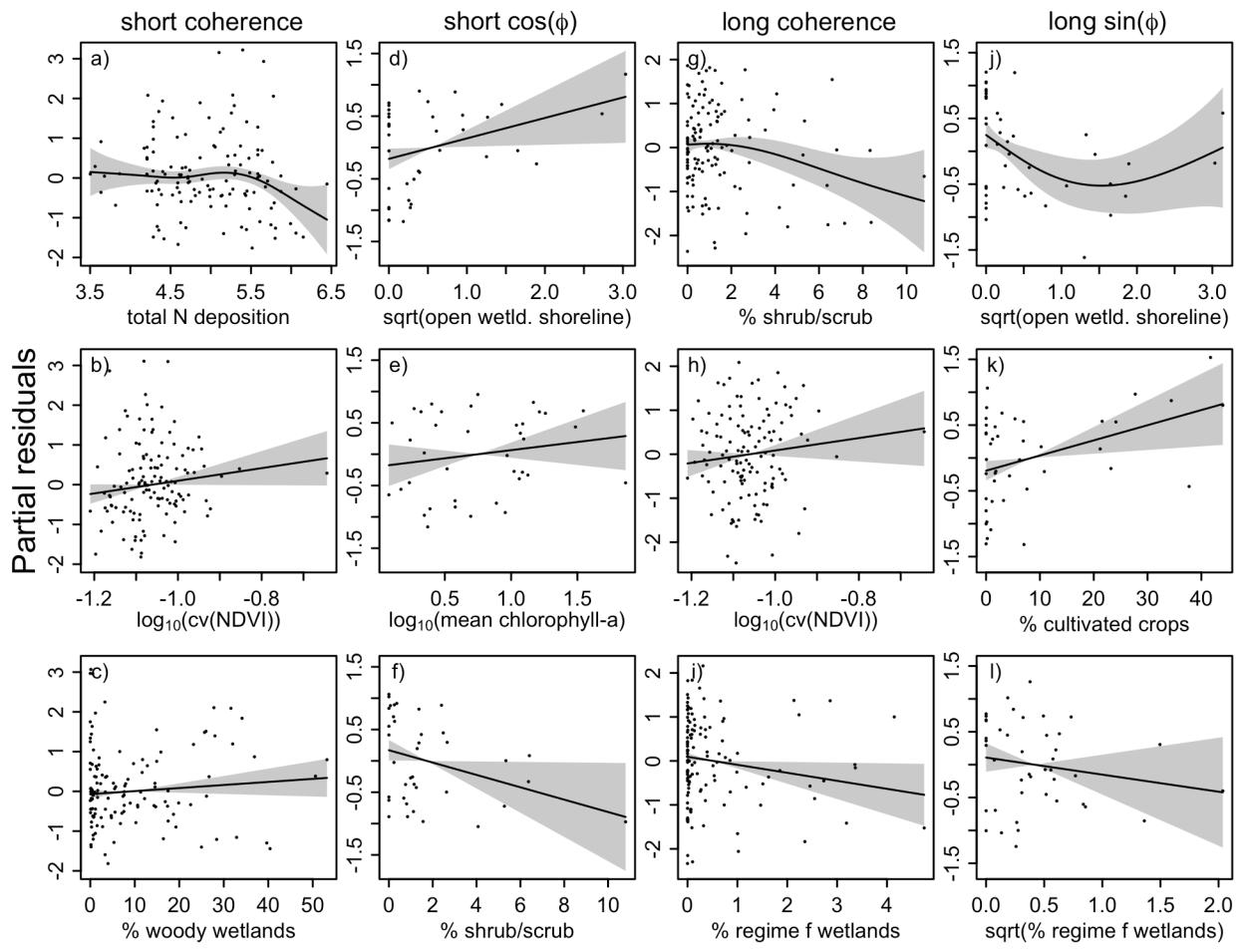
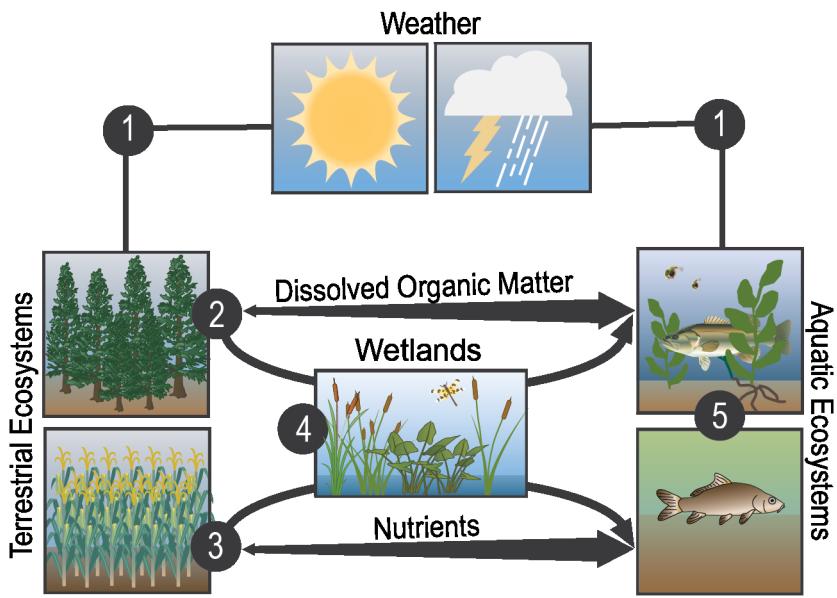


Figure 4.





786  
787 Figure 6.