

# Efficient Computation of Representative Weight Functions with Applications to Parameterized Counting (Extended Version)\*

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## Abstract

In this paper we prove an analogue of the classic Bollobás lemma for *approximate counting*. In fact, we match an analogous result of Fomin *et al.* [JACM 2016] for *decision*. This immediately yields, for a number of fundamental problems, parameterized approximate counting algorithms with the same running times as what is obtained for the decision variant using the representative family technique of Fomin *et al.* [JACM 2016]. For example, we devise an algorithm for approximately counting (a factor  $(1 \pm \epsilon)$  approximation algorithm)  $k$ -paths in an  $n$ -vertex directed graph ( $\#k$ -PATH) running in time  $\mathcal{O}((2.619^k + n^{o(1)}) \cdot \frac{1}{\epsilon^2} \cdot (n+m))$ . This improves over an earlier algorithm of Brand *et al.* [STOC 2018] that runs in time  $\mathcal{O}(4^k \cdot k^{o(1)} \cdot \frac{1}{\epsilon^2} \cdot (n+m))$ .

Additionally, we obtain an approximate counting analogue of the efficient computation of representative families for product families of Fomin *et al.* [TALG 2017], again essentially matching the running time for decision. This results in an algorithm with running time  $\mathcal{O}((3.841^k + |I|^{o(1)}) \cdot \frac{1}{\epsilon^8} \cdot |I|)$  for computing a  $(1 + \epsilon)$  approximation of the sum of the coefficients of the multilinear monomials in a degree- $k$  homogeneous  $n$ -variate polynomial encoded by a monotone circuit ( $\#$ MULTILINEAR MONOMIAL DETECTION). When restricted to monotone circuits (rather than polynomials of non-negative coefficients), this improves upon an earlier algorithm of Pratt [FOCS 2019] that runs in time  $4.075^k \cdot \frac{1}{\epsilon^2} \log \frac{1}{\epsilon} \cdot n^{o(1)}$ .

## 1 Introduction and Overview

The seminal paper of Valiant on counting problem [Val79] showed that although PERFECT MATCHING is solvable in polynomial time,  $\#$ PERFECT MATCHING is unlikely to be. This paper has since sparked vast interest in the study of counting problems. In this paper, we consider counting problems from the lens of Pa-

rameterized Complexity [CFK<sup>+</sup>15, DF13, FG06]. Our objective is twofold.

- Devise a general purpose algorithmic tool for parameterized counting problems.
- Use this tool to design state-of-the-art algorithms for several counting problems, including  $\#k$ -PATH.

The subfield of Parameterized Counting Complexity was initiated by Flum and Grohe [FG04], as early as 2002. Thus, this subfield has been around for the last 18 years, but until recently it has remained largely unexplored, with exceptions that are few and far between [AR02, Kou08, KW16b]. The last few years have seen a flurry of activities in this area resulting in the development of new tools and settlement of some old problems [Cur13, CM14, CDM17, CX15, BDH18, RW20, Bra19, DLM20]. We refer to the survey by Curticapean [Cur18] for a detailed exposition to parameterized counting problems.

As is the case with classical complexity, most of the natural counting problems are  $\#W[1]$ -hard [FG04] in the realm of Parameterized Complexity, which means that they are unlikely to be solvable in time  $f(k)n^{O(1)}$  for any computable function  $f$  of  $k$ . A problem admitting an algorithm with running time  $f(k)n^{O(1)}$  is called *fixed parameter tractable (FPT)* and the running time of the form  $f(k)n^{O(1)}$  is called FPT-time. For example, Flum and Grohe [FG04] showed that counting  $k$ -sized distinct (simple) paths in an undirected or directed graph ( $\#k$ -PATH) is  $\#W[1]$ -hard [FG04], although the decision version can be solved in FPT-time. In fact, until this day  $\#k$ -PATH is considered the most classical example of a problem solvable in FPT-time but which is  $\#W[1]$ -hard. Further, Flum and Grohe [FG04] conjectured the same for counting  $k$ -sized matchings ( $\#k$ -MATCHING) even on bipartite graphs. Curticapean [Cur13] and Curticapean and Marx [CM14] settled the parameterized complexity of  $\#k$ -MATCHING by showing that the problem is  $\#W[1]$ -hard.

The intractability of counting problems leads to the question of *approximately counting* in FPT-time. In par-

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ticular, there is a long history of FPT-approximation schemes (FPT-ASs), that is,  $f(k, \epsilon^{-1})n^{\mathcal{O}(1)}$ -time algorithms that approximate the number of certain combinatorial objects in the given input. Specifically, an FPT-AS for the  $\#k$ -PATH problem has been around for almost two decades [AR02] and is one of the fundamental problems driving the field of parameterized counting problems [AR02, ADH<sup>+</sup>08, AG10, BDH18, BLSZ19]. Recently, an approach based on representative families has been successful in the design of FPT-time algorithms for a wide-range of problems including  $k$ -PATH (the decision version of  $\#k$ -PATH), thus it is natural to consider a counting notion analogous to this notion. However, even just the existence of “small” representative families for counting purposes has not been known. In this paper, we develop a new technology that both asserts their existence and shows how to compute them efficiently.

**1.1 Representative Functions (or Counters) and Applications** Our starting point is the notion of *representative families* [Mon85, Mar09]. Let  $U$  be a universe and let  $\mathcal{S} = \{S_1, \dots, S_t\}$  be a family of subsets of  $U$  of size  $p$ . A subfamily  $\hat{\mathcal{S}} \subseteq \mathcal{S}$  is  $q$ -representative for  $\mathcal{S}$  if for every set  $Y \subseteq U$  of size at most  $q$ , if there is a set  $X \in \mathcal{S}$  disjoint from  $Y$ , then there exists a set  $X \in \hat{\mathcal{S}}$  disjoint from  $Y$ . By the classical combinatorial result of Bollobás, every family of sets of size  $p$  has a  $q$ -representative family with at most  $\binom{p+q}{p}$  sets [Bol65]. Given a family  $\mathcal{S}$  of sets of size  $p$ , and an integer  $q$ , an efficient algorithm computing a  $q$ -representative family  $\hat{\mathcal{S}} \subseteq \mathcal{S}$  was given in [Mar06, Mar09, FLPS16]. The fact that  $\hat{\mathcal{S}}$  can be efficiently computed from  $\mathcal{S}$  (and its generalizations to representative matroids) has found numerous applications in Parameterized and Exact Algorithms [Mon85, Mar09, FLPS16, SZ16, FGPS19, FLPS17, KW12, KS17, Mar06].

In this paper we prove an analogue of this result for *approximate counting*. More precisely, a function  $\hat{\mathcal{C}} : \mathcal{P} \rightarrow \mathbb{N}_0$  where  $\mathcal{P} \subseteq \binom{U}{p}$  (such a function is called a *counter*) is said to  $(\epsilon, q)$ -represent a function  $\mathcal{C} : \mathcal{P} \rightarrow \mathbb{N}_0$  with respect to  $\mathcal{Q} \subseteq \binom{U}{q}$  if for every set  $Q \in \mathcal{Q}$ , the following condition is satisfied:  $\sum_{P \in \mathcal{P}: P \cap Q = \emptyset} \mathcal{C}(P) \simeq (1 \pm \epsilon) \cdot \sum_{P \in \mathcal{P}: P \cap Q = \emptyset} \hat{\mathcal{C}}(P)$ . We prove that, when  $\mathcal{P}$  and  $\mathcal{Q}$  are “nice” (where the definition of “nice” is just the product of a minor technicality that can be ignored at the moment and to which we will return later), given any function  $\mathcal{C} : \mathcal{P} \rightarrow \mathbb{N}_0$ , a function  $\hat{\mathcal{C}} : \mathcal{P} \rightarrow \mathbb{N}_0$  that  $(\epsilon, q)$ -represents  $\mathcal{C}$  with respect to  $\mathcal{Q}$  and whose support (denoted by  $\text{supp}$ ) size is  $\binom{k}{p} \cdot 2^{o(k)} \cdot \frac{1}{\epsilon^2} \cdot n^{o(1)}$  where  $k = p + q$  and  $n = |U|$ , can be computed with success probability arbitrarily close to 1 and in time  $\mathcal{O}(|\text{supp}(\mathcal{C})| \cdot \binom{k}{q} \cdot 2^{o(k)} \cdot \frac{1}{\epsilon^2} \cdot n^{1+o(1)})$ .

We demonstrate how the efficient construction of representative functions can be a powerful tool in designing parameterized algorithms for counting problems.

**1.1.1 Applications** The  $k$ -PATH problem (on both directed and undirected graphs) is among the most extensively studied parameterized problems [CFK<sup>+</sup>15, FLSZ19]. This problem has played a pivotal role in the development of Parameterized Complexity and has led to several new tools and techniques in the area such as *color-coding* [AYZ95], *divide & color* [CKL<sup>+</sup>09b], *algebraic methods* [KW16b, BKKZ17, Wil09] and *representative families* [Mon85, Mar09, FLPS16]. After a long sequence of works in the past three decades, the current best known parameterized algorithms for  $k$ -PATH have running times  $1.657^k n^{\mathcal{O}(1)}$  (randomized, polynomial space, undirected only) [BHKK17, Bjö14] (extended in [BKKZ17]),  $2^k n^{\mathcal{O}(1)}$  (randomized, polynomial space) [Wil09],  $2.554^k n^{\mathcal{O}(1)}$  (deterministic, exponential space) [Tsu19, Zeh15, FLPS16, SZ16], and  $4^{k+o(k)} n^{\mathcal{O}(1)}$  (deterministic, polynomial space) [CKL<sup>+</sup>09a].

Similarly to  $k$ -PATH, the counting analogue  $\#k$ -PATH plays a significant role in the development of the field of parameterized counting. More than 15 years ago, Arvind and Raman [AR02] utilized the classic method of *color coding* [AYZ95] and *Karp-Luby approximate counting* technique to design a *randomized* exponential-space FPT-AS for  $\#k$ -PATH with running time  $k^{\mathcal{O}(k)} n^{\mathcal{O}(1)}$  whenever  $\epsilon^{-1} \leq k^{\mathcal{O}(k)}$ . A few years afterwards, the development and use of applications in computational biology to detect and analyze *network motifs* have already become common practice [SIKS06, SSRS06, SI06, DSG<sup>+</sup>08, HWZ08]. Roughly speaking, a network motif is a small pattern whose number of occurrences in a given network is substantially larger than its number of occurrences in a random network. Due to their tight relation to network motifs,  $\#k$ -PATH and other cases of the  $\#$ SUBGRAPH ISOMORPHISM problem became highly relevant to the study of gene transcription networks, protein-protein interaction (PPI) networks, neural networks and social networks [MSOI<sup>+</sup>02]. In light of these developments, Alon et al. [ADH<sup>+</sup>08] revisited the method of color coding to attain a running time whose dependency on  $k$  is *single-exponential* rather than *slightly super-exponential*. Specifically, they designed a *simple randomized*  $\mathcal{O}((2e)^k m \epsilon^{-2})$ -time exponential-space FPT-AS for  $\#k$ -PATH, which they employed to analyze PPI networks of unicellular organisms. In particular, their algorithm has running time  $2^{\mathcal{O}(k)} m$  whenever  $\epsilon^{-1} \leq 2^{\mathcal{O}(k)}$ . The first *deterministic* FPT-AS for  $\#k$ -PATH was found in 2007 by Alon and Gutner [AG10]; this algorithm has an exponential space complexity and run-

Table 1: History of  $\#k$ -PATH

Ref.	Time	Technique	Det.	Extension
[AR02]	$k^{\mathcal{O}(k)} n^{\mathcal{O}(1)}$	Karp-Luby	No	Treewidth $\mathcal{O}(1)$
[ADH <sup>+</sup> 08]	$(2e)^k n^{\mathcal{O}(1)}$	Color-Coding	No	No Extension
[AG09]	$(2e)^{k+o(k)} n^{\mathcal{O}(1)}$	Color-Coding	Yes	Treewidth $\mathcal{O}(1)$
[BDH18]	$4^k n^{\mathcal{O}(1)}$	Exterior Algebra	No	Pathwidth $\mathcal{O}(1)$
[Pra19]	$4.075^k n^{\mathcal{O}(1)}$	Waring Rank	No	Treewidth $\mathcal{O}(1)$
[BLSZ19]	$4^{k+o(k)} n^{\mathcal{O}(1)}$	Divide & Color	Yes	Treewidth $\mathcal{O}(1)$
<b>This Paper</b>	$2.619^k n^{\mathcal{O}(1)}$	<b>Representative Counters</b>	<b>No</b>	<b>Treewidth <math>\mathcal{O}(1)</math></b>

ning time  $2^{\mathcal{O}(k \log \log k)} m \log n$  whenever  $\epsilon^{-1} = 2^{o(\log k)}$ . Shortly afterwards, Alon and Gutner [AG09] improved upon their previous work, and designed a deterministic exponential-space FPT-AS for  $\#k$ -PATH with running time  $(2e)^{k+o(\log^3 k)} m \log n$  whenever  $\epsilon^{-1} = k^{\mathcal{O}(1)}$ . For close to a decade, this algorithm has remained the state-of-the-art. In 2016, Koutis and Williams [KW16a] made the following conjecture.

**Conjecture:**  $\#k$ -PATH admits an FPT-AS with running time  $2^{k(\frac{1}{\epsilon})^{\mathcal{O}(1)}} n^{\mathcal{O}(1)}$ .

After a decade, in 2018, Brand et al. [BDH18] provided a speed-up towards the resolution of this conjecture. Specifically, they gave an algebraic *randomized*  $\mathcal{O}(4^k m \epsilon^{-2})$ -time exponential-space algorithm. This was followed up by Björklund et al. [BLSZ19] who gave a *deterministic algorithm* with almost similar running time. However, this algorithm is still far away from resolving the conjecture of Koutis and Williams [KW16a].

As our first application we give an algorithm for  $\#k$ -PATH that runs in time  $\mathcal{O}((2.619^k + n^{o(1)}) \cdot \frac{1}{\epsilon^2} \cdot (n + m))$ . This results brings the gap between the known algorithm and the conjecture close. While on a superficial level, we make use of the notion of parsimonious universal families also present in [BLSZ19], our new result is centred around the efficient computation of representative counter functions (a concept introduced in this paper), which requires to develop a whole new machinery in general, and sampling primitives in particular.

The  $\#k$ -PATH problem is a special case of the  $\#k$ -SUBGRAPH ISOMORPHISM problem, where for a given  $n$ -vertex graph  $G$  and a given  $k$ -vertex graph  $F$ , the objective is to count the number of distinct subgraphs of  $G$  that are isomorphic to  $F$ . In addition to  $\#k$ -PATH, parameterized counting algorithms for two other variants of  $\#k$ -SUBGRAPH ISOMORPHISM, when  $F$  is a tree, and more generally, a graph of treewidth at most  $t$ , were studied in the literature. The algorithm of Björklund et al. [BLSZ19] can be

extended for these cases with running time similar to that for  $\#k$ -PATH. Independently, Pratt [Pra19] obtained an algorithm for these cases as an application of his algorithm for a more general problem, called  $\#$ MULTILINEAR DETECTION, which we discuss in more detail in the following subsection.

In particular, we obtain Theorem 1.1 ahead as an application of our first tool. Before we state it, let us give the definitions of the problems it addresses. In  $q$ -SET  $p$ -PACKING we are given a universe  $U$ , a family  $\mathcal{F}$  of subsets of size  $q$  of  $U$ , and  $p \in \mathbb{N}$ . Then, the objective is to determine whether there exist at least  $p$  pairwise-disjoint sets in  $\mathcal{F}$ . In  $q$ -DIMENSIONAL  $p$ -MATCHING, we are given a universe  $U$ , a partition  $(U_1, U_2, \dots, U_q)$  of  $U$ , a family  $\mathcal{F}$  of subsets of size  $q$  of  $U$  where each subset contains exactly one element from each part  $U_i$ , and  $p \in \mathbb{N}$ . Then, the objective is to determine whether there exist at least  $p$  pairwise-disjoint sets in  $\mathcal{F}$ . In GRAPH MOTIF, we are given a graph  $G$  where each vertex is assigned a set of colors, a multiset of colors  $M$ , and  $k \in \mathbb{N}$  (the sought motif size). Then, the objective is to determine whether there exist a subtree  $T$  of  $G$  on  $k$  vertices and a coloring of the vertices in  $T$  (each by a color from its set) so that no color is used more times than its number of occurrences in  $M$ .

**THEOREM 1.1.** *For any  $0 < \epsilon < 1$ , the  $\#k$ -PATH,  $\#q$ -SET  $p$ -PACKING with  $k = qp$ ,  $\#q$ -DIMENSIONAL  $p$ -MATCHING with  $k = (q-1)p$  and  $\#$ GRAPH MOTIF with  $k$  being twice the sought motif size problems can be approximated with factor  $(1 \pm \epsilon)$  and success probability at least  $\frac{9}{10}$  in time  $\mathcal{O}((2.619^k + |I|^{o(1)}) \cdot \frac{1}{\epsilon^2} \cdot |I|)$ , where  $k$  is the parameter and  $|I|$  is the input size. Moreover, for any  $0 < \epsilon < 1$ , the  $\#k$ -TREE (or, more generally,  $\#$ SUBGRAPH ISOMORPHISM where the treewidth of pattern graph is bounded by a fixed constant) can be approximated with factor  $(1 \pm \epsilon)$  and success probability at least  $\frac{9}{10}$  in time  $2.619^k \cdot \frac{1}{\epsilon^2} \cdot |I|^{\mathcal{O}(1)}$ .*

**1.2 Representation for Product Functions (or Counters) and Applications** Let  $\mathcal{P} \subseteq \binom{U}{p}$ . Given

two functions  $\mathfrak{C}_1 : \mathcal{P}_1 \rightarrow \mathbb{N}_0$  and  $\mathfrak{C}_2 : \mathcal{P}_2 \rightarrow \mathbb{N}_0$  where  $\mathcal{P}_1 \subseteq \binom{U}{p_1}$ ,  $\mathcal{P}_2 \subseteq \binom{U}{p_2}$  and  $p_1 + p_2 = p$ , the product  $\mathfrak{C}_1 \times \mathfrak{C}_2$  (with respect to  $\mathcal{P}$ ) is the function  $\mathfrak{C}_1 \times \mathfrak{C}_2 : \mathcal{P} \rightarrow \mathbb{N}_0$  defined as follows: For each  $P \in \mathcal{P}$ ,

$$(\mathfrak{C}_1 \times \mathfrak{C}_2)(P) = \sum_{\substack{P_1 \in \mathcal{P}_1, P_2 \in \mathcal{P}_2: \\ P_1 \cap P_2 = \emptyset, P_1 \cup P_2 = P}} \mathfrak{C}_1(P_1) \cdot \mathfrak{C}_2(P_2).$$

We prove that, given that  $\mathcal{P}_1, \mathcal{P}_2, \mathcal{P}$  and  $\mathcal{Q} \subseteq \binom{U}{q}$  are “nice”, given any two functions  $\mathfrak{C}_1 : \mathcal{P}_1 \rightarrow \mathbb{N}_0$  and  $\mathfrak{C}_2 : \mathcal{P}_2 \rightarrow \mathbb{N}_0$ , a function  $\hat{\mathfrak{C}} : \mathcal{P} \rightarrow \mathbb{N}_0$  that  $(\epsilon, q)$ -represents  $\mathfrak{C} = \mathfrak{C}_1 \times \mathfrak{C}_2$  with respect to  $\mathcal{Q}$  and whose support size is  $\binom{k}{p} \cdot 2^{o(k)} \cdot \frac{1}{\epsilon^2} \cdot n^{o(1)}$  where  $k = p + q$ , can be computed with success probability arbitrarily close to 1 and in time

$$\mathcal{O}((3.841^k + |\text{supp}(\mathfrak{C}_1)| \cdot (\frac{k}{q + p_2})^{q + p_2} + |\text{supp}(\mathfrak{C}_2)| \cdot (\frac{k}{q + p_1})^{q + p_1}) \cdot 2^{o(k)} \cdot \frac{1}{\epsilon^2} \cdot n^{1 + o(1)}).$$

A more exact expression of the upper bound on the time complexity that precisely describes the dependence on the sizes of the supports of  $\mathfrak{C}_1$  and  $\mathfrak{C}_2$  rather than the term  $3.841^k$  is given in the paper. However, the crux here is that the time complexity to compute the output function can be substantially smaller than even just the time to explicitly write up the function  $\mathfrak{C}_1 \times \mathfrak{C}_2$  that it represents (even if both  $\mathfrak{C}_1 \times \mathfrak{C}_2$  have already been reduced to have support size  $\binom{k}{p_1}$  and  $\binom{k}{p_2}$ )! For example, if both  $p_1$  and  $p_2$  are close to  $k/2$ , then the support of their product is already  $4^k$ .

Our main application is a randomized algorithm for the #MULTILINEAR MONOMIAL DETECTION problem, mentioned in the previous subsection. In this problem, the objective is to compute a  $(1 + \epsilon)$  approximation of the sum of the coefficients of the multilinear monomials in a degree- $k$  homogeneous  $n$ -variate polynomial encoded by an arithmetic circuit with non-negative coefficients (i.e., a *monotone circuit*). Recently, Pratt [Pra19] developed a randomized  $(1 + \epsilon)$ -approximation algorithm for this problem with time complexity  $\mathcal{O}(4.075^k \cdot \frac{1}{\epsilon^2} \log \frac{1}{\epsilon} \cdot s(C)^{\mathcal{O}(1)})$ . In fact, the result of Pratt [Pra19] is for a notably more general—it deals with the #MULTILINEAR MONOMIAL DETECTION problem extended to only requiring the polynomial to have nonnegative coefficients, thus allowing the arithmetic circuit to have negative coefficients, though not the polynomial that it encodes. Improving upon this result for the case of monotone circuits, we get the following.

**THEOREM 1.2.** *For any  $0 < \epsilon < 1$ , the #MULTILINEAR MONOMIAL DETECTION problem (on monotone*

*circuits) can be approximated with factor  $(1 \pm \epsilon)$  and success probability at least  $\frac{9}{10}$  in time  $\mathcal{O}((3.841^k + s(C)^{o(1)}) \cdot \frac{1}{\epsilon^2} \cdot s(C))$ .*

The decision version of #MULTILINEAR MONOMIAL DETECTION is the central problem in the algebraic approach of Koutis and Williams for designing fast parameterized algorithms [Kou08, KW16b, Wil09]. Here, the objective is to decide whether there exists a multilinear monomial of degree- $k$  with non-zero coefficient (rather than to compute the sum of coefficients of such monomials). Let  $s(C)$  denote the size of  $C$ . Williams [Wil09] gave a *randomized* algorithm solving  $k$ -MULTILINEAR MONOMIAL DETECTION in time  $2^k \cdot s(C)^{\mathcal{O}(1)}$  (over monotone circuits). The only known algorithm for the problem when there is no restriction on circuits is by Brand et al. [BDH18], who gave an algorithm with running time  $4.32^k \cdot s(C)^{\mathcal{O}(1)}$  (with exponential space complexity). (Recently, further (yet unpublished) developments were given in the preprint [BP20].) Afterwards, Arvind et al. [ACDM19] obtained an algorithm with the same running time and with polynomial space complexity. The algorithms based on the algebraic method of Koutis-Williams provide a dramatic improvement for a number of fundamental problems. See the survey by Koutis and Williams [KW16a] for further details. The idea behind the approach is to translate a given problem into the language of algebra by reducing it to the problem of deciding whether a constructed polynomial has a multilinear monomial of degree  $k$ .

We note that # $k$ -SUBGRAPH ISOMORPHISM can be reduced to the #MULTILINEAR MONOMIAL DETECTION problem and thus one can obtain an algorithm (that is efficient when the sought graph is of constant treewidth) for it as an application of Theorem 1.2. In fact, # $k$ -SUBGRAPH ISOMORPHISM reduces to #MULTILINEAR MONOMIAL DETECTION on special circuits where we can obtain a faster algorithm. This is what is exploited in the proof of Theorem 1.1. More precisely, the aforementioned special circuits are “ $d$ -skewed circuits” (mostly, for  $d = \mathcal{O}(1)$ ), where every multiplication gate has at most one child whose polynomial can consist of more than  $d$  monomials. Specifically, we have the following theorem, where we are particularly interested in the case where  $\ell = 0$ . This theorem is also our intermediate step to derive Theorem 1.1.

**THEOREM 1.3.** *For any  $0 < \epsilon < 1$  and  $\ell \in \mathbb{N}_0$ , the #MULTILINEAR MONOMIAL DETECTION problem on  $2^{o(k)} s(C)^\ell$ -skewed circuits can be approximated with factor  $(1 \pm \epsilon)$  and success probability at least  $\frac{9}{10}$  in time  $\mathcal{O}((2.619^k + s(C)^{o(1)}) \cdot \frac{1}{\epsilon^2} \cdot s(C)^{\ell+1})$ .*



**1.3 Additional Related Works** The algorithms by Alon et al. [ADH<sup>+</sup>08] and Alon and Gutner [AG10, AG09], just like our algorithms, extend to approximate counting of graphs of bounded treewidth. (This remark is also made by Alon and Gutner [AG10, AG09].) In what follows, we briefly review works related to exact counting and decision from the viewpoint of Parameterized Complexity. Since these topics are not the focus of our work, the survey is illustrative rather than comprehensive.

The problem of counting the number of subgraphs of a graph  $G$  that are isomorphic to a graph  $H$ —that is, #SUBGRAPH ISOMORPHISM WITH PATTERN  $H$ —admits a dichotomy: If the vertex cover number of  $H$  is bounded, then it is FPT [WW13], and otherwise it is #W[1]-hard [CM14]. The #W[1]-hardness of # $k$ -PATH, originally shown by Flum and Grohe [FG04], follows from this dichotomy. By using the “meet in the middle” approach, the # $k$ -PATH problem and, more generally, #SUBGRAPH ISOMORPHISM WITH PATTERN  $H$  where  $H$  has bounded *pathwidth* and  $k$  vertices, was shown to admit an  $n^{\frac{k}{2} + \mathcal{O}(1)}$ -time algorithm [BHK09]. Later, Björklund et al. [BKK17] showed that  $\frac{k}{2}$  is not a barrier (which was considered to be the case at that time) by designing an  $n^{0.455k + \mathcal{O}(1)}$ -time algorithm. A breakthrough that resulted in substantially faster running times took place: Curticapean et al. [CDM17] showed that #SUBGRAPH ISOMORPHISM WITH PATTERN  $H$  is solvable in time  $\ell^{\mathcal{O}(\ell)} n^{0.174\ell}$  where  $\ell$  is the number of edges in  $H$ ; in particular, this algorithm solves # $k$ -PATH in time  $k^{\mathcal{O}(k)} n^{0.174k}$ . Recently, Arvind et al. [ACDM19] obtained an algorithm for #MULTILINEAR MONOMIAL DETECTION with time complexity  $n^{k/2 + \mathcal{O}(\log k)}$ . Also recently, Dell et al. [DLM20] gave “black box” results for turning algorithms which decide whether or not a witness exists into algorithms to approximately count the number of witnesses (with overheads of  $k^{\mathcal{O}(k)}$  that are prohibitive for our settings).

## 2 Preliminaries

Let  $U$  be a universe, and let  $p, q \in \mathbb{N}_0$  be non-negative integers. Then, let  $\binom{U}{p}$  be the collection of subsets of  $U$  of size exactly  $p$ , and denote  $\binom{U}{\leq p} = \bigcup_{i=0}^p \binom{U}{i}$ . Given two subsets  $P, Q$  of  $U$  and a family  $\mathcal{F} \subseteq 2^U$  of subsets of  $U$ , denote  $\mathcal{F}[P, Q] \triangleq \{F \in \mathcal{F} : P \subseteq F, Q \cap F = \emptyset\}$ . Given a function  $f : U \rightarrow \mathbb{R}$ , let  $\text{supp}(f) = \{u \in U : f(u) \neq 0\}$  denote the support of  $f$ . Given two functions  $f : U \rightarrow \mathbb{R}$  and  $g : U \rightarrow \mathbb{R}$  such that for every  $a \in A$ , it holds that  $g(a) \leq f(a)$ , we denote  $g \leq f$ .

A central notion in our proofs is of parsimonious universal families, defined as follows.

**DEFINITION 1. ( $\epsilon$ -Parsimonious Universal Family)**

Let  $n, p, q \in \mathbb{N}$  and  $0 < \epsilon < 1$ . Let  $U$  be a universe of size  $n$ , and let  $\mathcal{P} \subseteq \binom{U}{p}$  and  $\mathcal{Q} \subseteq \binom{U}{q}$ . A family  $\mathcal{F} \subseteq 2^U$  is an  $\epsilon$ -parsimonious  $(n, p, q)$ -universal family with respect to  $(\mathcal{P}, \mathcal{Q})$  if there exists  $T = T(n, p, q, \epsilon) > 0$ , called a correction factor, such that for each pair of disjoint sets  $P \in \mathcal{P}$  and  $Q \in \mathcal{Q}$ , it holds that  $(1 - \epsilon) \cdot T \leq |\mathcal{F}[P, Q]| \leq (1 + \epsilon) \cdot T$ .

The special case of Definition 1 where  $\mathcal{P} = \binom{U}{p}$  and  $\mathcal{Q} = \binom{U}{q}$  is the definition of an  $\epsilon$ -parsimonious universal family in [BLSZ19]. For parsimonious universal families, the following proposition is known, based on a straightforward sampling argument.

**PROPOSITION 2.1. ([BLSZ19])** Let  $c \in \mathbb{N}$  be a fixed constant. Let  $n, p, q \in \mathbb{N}$  and  $0 < \epsilon < 1$ , and denote  $k = p + q$ . Let  $U$  be a universe of size  $n$ . An  $\epsilon$ -parsimonious  $(n, p, q)$ -universal family  $\mathcal{F} \subseteq 2^U$  of size  $t = \mathcal{O}\left(\frac{k^k}{p^p q^q} \cdot k \log n \cdot \frac{1}{\epsilon^2}\right)$ , can be computed with success probability at least  $1 - 1/n^{ck}$  in time  $\mathcal{O}(t \cdot n)$ .

We will need more sophisticated parsimonious universal families, constructed in a manner to enable having an efficient “membership query” procedure—that is, a procedure that given any set  $P \in \binom{U}{p}$ , outputs all the sets in the family that contain  $P$ . We will address the computation of such families and procedures in Section 3.2. Formally, they are defined as follows.

**DEFINITION 2. (Membership Query Procedure)**

Let  $n, p, q \in \mathbb{N}$  and  $0 < \epsilon < 1$ . Let  $U$  be a universe of size  $n$ , and let  $\mathcal{P} \subseteq \binom{U}{p}$  and  $\mathcal{Q} \subseteq \binom{U}{q}$ . Let  $\mathcal{F} \subseteq 2^U$  be an  $\epsilon$ -parsimonious  $(n, p, q)$ -universal family with respect to  $(\mathcal{P}, \mathcal{Q})$ . A  $T$ -membership query procedure is a procedure that given any set  $P \in \mathcal{P}$  as input, outputs the subfamily  $\{F \in \mathcal{F} : P \subseteq F\}$  in time  $\mathcal{O}(T)$ .

We will also make use of the following well known inequality to bound probabilities.

**PROPOSITION 2.2. (Chernoff Bound)** Let

$X_1, \dots, X_\ell$  be independent random variables bounded by the interval  $[0, 1]$ . Let  $X = \sum_{i=1}^{\ell} X_i$ . For any  $\epsilon \geq 0$ ,

$$\Pr(|X - E[X]| > \epsilon E[X]) \leq 2e^{-\frac{\epsilon^2 E[X]}{2}}.$$

Lastly, we define the notion of an arithmetic circuit. An *arithmetic circuit*  $C$  over a commutative ring  $R$  is a simple labelled directed acyclic graph whose internal nodes are labeled by  $+$  or  $\times$  and whose leaves (in-degree zero nodes) are labeled from  $X$  where  $X = \{x_1, x_2, \dots, x_n\}$  is a set of variables. There is a node

of out-degree zero, called the *root* node or the output gate. The size of  $C$ , denoted by  $s(C)$ , is the number of nodes,  $s_V(C)$ , plus the number of arcs,  $s_A(C)$ , in the digraph.

### 3 Representative Counters

In this section, we will be working with counters, defined as follows.

**DEFINITION 3. (Counter)** Let  $U$  be a universe. Let  $p \in \mathbb{N}_0$ , and let  $\mathcal{P} \subseteq \binom{U}{p}$ . A function  $\mathfrak{C} : \mathcal{P} \rightarrow \mathbb{N}_0$  is called a counter. A counter is encoded as a collection of pairs, where each pair consists of an element  $P \in \text{supp}(\mathfrak{C})$  and its value  $\mathfrak{C}(P)$ .

The main objective of this section is to compute representative counters, defined as follows.

**DEFINITION 4. ((Representative))** Let  $U$  be a universe. Let  $\alpha \leq 1, \beta \geq 1$ , and let  $p, q \in \mathbb{N}_0$ . Let  $\mathcal{P} \subseteq \binom{U}{p}$  and  $\mathcal{Q} \subseteq \binom{U}{q}$ . A counter  $\hat{\mathfrak{C}} : \mathcal{P} \rightarrow \mathbb{N}_0$  is said to  $(\alpha, \beta, q)$ -represent a counter  $\mathfrak{C} : \mathcal{P} \rightarrow \mathbb{N}_0$  with respect to  $\mathcal{Q}$  if for every set  $Q \in \mathcal{Q}$ , the following condition is satisfied.

$$\alpha \cdot \sum_{P \in \mathcal{P} : P \cap Q = \emptyset} \mathfrak{C}(P) \leq \sum_{P \in \mathcal{P} : P \cap Q = \emptyset} \hat{\mathfrak{C}}(P) \leq \beta \cdot \sum_{P \in \mathcal{P} : P \cap Q = \emptyset} \mathfrak{C}(P).$$

When  $\alpha = 1 - \epsilon$  and  $\beta = 1 + \epsilon$  for some  $0 < \epsilon < 1$ ,  $\hat{\mathfrak{C}}$  is said to  $(\epsilon, q)$ -represent  $\mathfrak{C}$ .

Further, we will need the representative counter to be, in expectation, not just similar, but identical to the given counter.

**DEFINITION 5. (Representative in Expectation)** Let  $U$  be a universe. Let  $p \in \mathbb{N}_0$ , and let  $\mathcal{P} \subseteq \binom{U}{p}$ . A sampled counter  $\hat{\mathfrak{C}} : \mathcal{P} \rightarrow \mathbb{N}_0$  is said to represent in expectation a counter  $\mathfrak{C} : \mathcal{P} \rightarrow \mathbb{N}_0$  if for every set  $P \in \mathcal{P}$ , the following condition is satisfied.

$$E_{\hat{\mathfrak{C}}}[\hat{\mathfrak{C}}(P)] = \mathfrak{C}(P).$$

We will first show how to efficiently compute representative counters under the assumption that we can compute parsimonious universal families equipped with efficient membership query procedures. Next, we will show how to compute a parsimonious universal family equipped with efficient membership query procedure for specific choices of  $(\mathcal{P}, \mathcal{Q})$ . We remark that in what follows, we implicitly suppose that the counter to represent has non-empty support, because otherwise representation is trivial.

### 3.1 Computation of Representative Counters of Small Support

We first extend the notion of a counter to also assign values to sets of size larger than  $p$ .

**DEFINITION 6. (Domain Extension)** Let  $U$  be a universe. Let  $p \in \mathbb{N}_0$ , and let  $\mathcal{P} \subseteq \binom{U}{p}$ . Let  $\mathfrak{C} : \mathcal{P} \rightarrow \mathbb{N}_0$  be a counter. The extender  $\mathfrak{C}_{\text{ext}} : 2^U \rightarrow \mathbb{N}_0$  is defined as follows. For any set  $F \subseteq U$ , define

$$\mathfrak{C}_{\text{ext}}(F) \triangleq \sum_{P \in \binom{F}{p} \cap \mathcal{P}} \mathfrak{C}(P).$$

Notice that for any set  $P \in \mathcal{P}$ , we have that  $\mathfrak{C}_{\text{ext}}(P) = \mathfrak{C}(P)$ . Now, we present an alternative (to Definition 4) notion of similarity between counters, based on a given family  $\mathcal{F}$  (that will, when used ahead, be a parsimonious universal family). In particular, it makes similarly, in a sense, be more focused, considering only sets in  $\mathcal{F}$  rather than all possible choices of  $P \in \mathcal{P}$  and  $Q \in \mathcal{Q}$  in order to measure similarity. Being more focused, working with this definition for the computation of representative counters will also yield efficiency. Notice that this definition does not replace Definition 4—the usage of representative counters for applications will require Definition 4.

**DEFINITION 7. (( $(\epsilon, \mathcal{F})$ -Similarly))** Let  $U$  be a universe, and let  $\mathcal{F} \subseteq 2^U$ . Let  $0 < \epsilon < 1$ , and let  $p \in \mathbb{N}_0$  and  $\mathcal{P} \subseteq \binom{U}{p}$ . Let  $\mathfrak{C} : \mathcal{P} \rightarrow \mathbb{N}_0$  and  $\hat{\mathfrak{C}} : \mathcal{P} \rightarrow \mathbb{N}_0$  be two counters. We say that  $\mathfrak{C}$  and  $\hat{\mathfrak{C}}$  are  $(\epsilon, \mathcal{F})$ -similar if for every set  $F \in \mathcal{F}$ ,  $(1 - \epsilon) \cdot \mathfrak{C}_{\text{ext}}(F) \leq \hat{\mathfrak{C}}_{\text{ext}}(F) \leq (1 + \epsilon) \cdot \mathfrak{C}_{\text{ext}}(F)$ .

We now prove that for the sake of efficient computation of representative counters, we can indeed work with the new definition.

**LEMMA 3.1.** Let  $n, p, q \in \mathbb{N}$ ,  $0 < \epsilon < 1$  and  $0 < \delta < 1$ . Let  $\mathcal{P} \subseteq \binom{U}{p}$  and  $\mathcal{Q} \subseteq \binom{U}{q}$ . Let  $\mathcal{F} \subseteq 2^U$  be an  $\epsilon$ -parsimonious  $(n, p, q)$ -universal family with respect to  $(\mathcal{P}, \mathcal{Q})$ . Let  $\mathfrak{C} : \mathcal{P} \rightarrow \mathbb{N}_0$  and  $\hat{\mathfrak{C}} : \mathcal{P} \rightarrow \mathbb{N}_0$  be  $(\delta, \mathcal{F})$ -similar counters. Then,  $\hat{\mathfrak{C}}$   $(4\epsilon + \delta, q)$ -represents  $\mathfrak{C}$  with respect to  $\mathcal{Q}$ .

*Proof.* To prove that  $\hat{\mathfrak{C}}$   $(\epsilon, q)$ -represents  $\mathfrak{C}$  with respect to  $\mathcal{Q}$ , consider some set  $Q \in \mathcal{Q}$ . First, observe that

$$\begin{aligned} (*) \quad & \sum_{P \in \mathcal{P} : P \cap Q = \emptyset} |\mathcal{F}[P, Q]| \cdot \mathfrak{C}(P) = \sum_{P \in \mathcal{P} : P \cap Q = \emptyset} \sum_{F \in \mathcal{F}[P, Q]} \mathfrak{C}(P) \\ &= \sum_{F \in \mathcal{F} : Q \cap F = \emptyset} \sum_{P \in \mathcal{P} : P \subseteq F} \mathfrak{C}(P) \\ &= \sum_{F \in \mathcal{F} : Q \cap F = \emptyset} \mathfrak{C}_{\text{ext}}(F). \end{aligned}$$

Let  $T$  be the correction factor of  $\mathcal{F}$ . Then, for any set  $P \in \mathcal{P}$ , we have that  $(1 - \epsilon)T \leq |\mathcal{F}[P, Q]| \leq (1 + \epsilon)T$ . On the one hand, this implies that

$$\begin{aligned} \text{(I)} \quad & \sum_{P \in \mathcal{P}: P \cap Q = \emptyset} \mathfrak{C}(P) \\ &= \frac{1}{(1 - \epsilon)T} \cdot \sum_{P \in \mathcal{P}: P \cap Q = \emptyset} (1 - \epsilon)T \cdot \mathfrak{C}(P) \\ &\leq \frac{1}{(1 - \epsilon)T} \cdot \sum_{P \in \mathcal{P}: P \cap Q = \emptyset} |\mathcal{F}[P, Q]| \cdot \mathfrak{C}(P) \\ &= \frac{1}{(1 - \epsilon)T} \cdot \sum_{F \in \mathcal{F}: Q \cap F = \emptyset} \mathfrak{C}_{\text{ext}}(F). \end{aligned}$$

Here, the last equality was derived from equality (\*). Symmetrically, we have that

$$\text{(II)} \quad \sum_{P \in \mathcal{P}: P \cap Q = \emptyset} \widehat{\mathfrak{C}}(P) \leq \frac{1}{(1 - \epsilon)T} \cdot \sum_{F \in \mathcal{F}: Q \cap F = \emptyset} \widehat{\mathfrak{C}}_{\text{ext}}(F).$$

On the other hand, this implies that

$$\begin{aligned} \text{(III)} \quad & \sum_{P \in \mathcal{P}: P \cap Q = \emptyset} \mathfrak{C}(P) \\ &= \frac{1}{(1 + \epsilon)T} \cdot \sum_{P \in \mathcal{P}: P \cap Q = \emptyset} (1 + \epsilon)T \cdot \mathfrak{C}(P) \\ &\geq \frac{1}{(1 + \epsilon)T} \cdot \sum_{P \in \mathcal{P}: P \cap Q = \emptyset} |\mathcal{F}[P, Q]| \cdot \mathfrak{C}(P) \\ &= \frac{1}{(1 + \epsilon)T} \cdot \sum_{F \in \mathcal{F}: Q \cap F = \emptyset} \mathfrak{C}_{\text{ext}}(F). \end{aligned}$$

Again, the last equality was derived from equality (\*). Symmetrically, we have that

$$\text{(IV)} \quad \sum_{P \in \mathcal{P}: P \cap Q = \emptyset} \widehat{\mathfrak{C}}(P) \geq \frac{1}{(1 + \epsilon)T} \cdot \sum_{F \in \mathcal{F}: Q \cap F = \emptyset} \widehat{\mathfrak{C}}_{\text{ext}}(F).$$

Because  $\mathfrak{C}$  and  $\widehat{\mathfrak{C}}$  are  $(\delta, \mathcal{F})$ -similar, for any set  $F \in \mathcal{F}$ , we have that  $(1 - \delta) \cdot \mathfrak{C}_{\text{ext}}(F) \leq \widehat{\mathfrak{C}}_{\text{ext}}(F) \leq (1 + \delta) \cdot \mathfrak{C}_{\text{ext}}(F)$ . On the one hand, combined with inequalities (II) and (III), this implies that

$$\begin{aligned} & \sum_{P \in \mathcal{P}: P \cap Q = \emptyset} \widehat{\mathfrak{C}}(P) \\ &\leq \frac{1}{(1 - \epsilon)T} \cdot \sum_{F \in \mathcal{F}: Q \cap F = \emptyset} \widehat{\mathfrak{C}}_{\text{ext}}(F) \\ &\leq \frac{(1 + \delta)}{(1 - \epsilon)T} \cdot \sum_{F \in \mathcal{F}: Q \cap F = \emptyset} \mathfrak{C}_{\text{ext}}(F) \\ &= \frac{(1 + \delta)(1 + \epsilon)}{(1 - \epsilon)} \cdot \frac{1}{(1 + \epsilon)T} \cdot \sum_{F \in \mathcal{F}: Q \cap F = \emptyset} \mathfrak{C}_{\text{ext}}(F) \\ &\leq \frac{(1 + \delta)(1 + \epsilon)}{(1 - \epsilon)} \cdot \sum_{P \in \mathcal{P}: P \cap Q = \emptyset} \mathfrak{C}(P). \end{aligned}$$

On the other hand, combined with inequalities (I) and (IV), this implies that

$$\begin{aligned} & \sum_{P \in \mathcal{P}: P \cap Q = \emptyset} \widehat{\mathfrak{C}}(P) \\ &\geq \frac{1}{(1 + \epsilon)T} \cdot \sum_{F \in \mathcal{F}: Q \cap F = \emptyset} \widehat{\mathfrak{C}}_{\text{ext}}(F) \\ &\geq \frac{(1 - \delta)}{(1 + \epsilon)T} \cdot \sum_{F \in \mathcal{F}: Q \cap F = \emptyset} \mathfrak{C}_{\text{ext}}(F) \\ &= \frac{(1 - \delta)(1 - \epsilon)}{(1 + \epsilon)} \cdot \frac{1}{(1 - \epsilon)T} \cdot \sum_{F \in \mathcal{F}: Q \cap F = \emptyset} \mathfrak{C}_{\text{ext}}(F) \\ &\geq \frac{(1 - \delta)(1 - \epsilon)}{(1 + \epsilon)} \cdot \sum_{P \in \mathcal{P}: P \cap Q = \emptyset} \mathfrak{C}(P). \end{aligned}$$

Overall, we have that  $\frac{(1 + \delta)(1 - \epsilon)}{(1 + \epsilon)} \cdot \sum_{P \in \mathcal{P}: P \cap Q = \emptyset} \widehat{\mathfrak{C}}(P) \leq \sum_{P \in \mathcal{P}: P \cap Q = \emptyset} \mathfrak{C}(P) \leq \frac{(1 + \delta)(1 + \epsilon)}{(1 - \epsilon)} \cdot \sum_{P \in \mathcal{P}: P \cap Q = \emptyset} \widehat{\mathfrak{C}}(P)$ . Notice that  $1 - \epsilon > 1 - \epsilon - 2\epsilon^2 = (1 - 2\epsilon)(1 + \epsilon)$ , and hence  $(1 - \epsilon)/(1 + \epsilon) > 1 - 2\epsilon$ ; similarly,  $1 + \epsilon > 1 + \epsilon - 2\epsilon^2 = (1 + 2\epsilon)(1 - \epsilon)$ , and hence  $(1 + \epsilon)/(1 - \epsilon) > 1 + 2\epsilon$ . Moreover, because  $0 < \epsilon, \delta < 1$ ,  $(1 - \delta)(1 - 2\epsilon) = 1 - (2\epsilon + \delta - 2\epsilon\delta) > 1 - (4\epsilon + \delta)$ , and  $(1 + \delta)(1 + 2\epsilon) = 1 + (2\epsilon + \delta + 2\epsilon\delta) < 1 + (4\epsilon + \delta)$ . Thus,  $\sum_{P \in \mathcal{P}: P \cap Q = \emptyset} \widehat{\mathfrak{C}}(P) \leq (1 + (4\epsilon + \delta)) \cdot \sum_{P \in \mathcal{P}: P \cap Q = \emptyset} \mathfrak{C}(P)$ . Since the choice of  $Q \in \mathcal{Q}$  was arbitrary, the proof is complete.  $\square$

Our computation of representative counters will be done in a sampling procedure defined as follows. (Some explanation of the intuition behind it is given ahead.)

**DEFINITION 8. (( $\mathfrak{C}, \mathcal{F}$ )-Counter Sampling)** Let  $U$  be a universe, and let  $\mathcal{F} \subseteq 2^U$  with  $U \in \mathcal{F}$ . Let  $p, L \in \mathbb{N}_0$  and  $\mathcal{P} \subseteq \binom{U}{p}$ . Let  $\mathfrak{C} : \mathcal{P} \rightarrow \mathbb{N}_0$  be a counter. Then,  $(\mathfrak{C}, \mathcal{F}, L)$ -counter sampling is the randomized procedure that constructs a counter  $\widehat{\mathfrak{C}} : \mathcal{P} \rightarrow \mathbb{N}_0$  as follows. For any set  $P \in \mathcal{P}$ , define

$$\text{assoc}_{\mathfrak{C}, \mathcal{F}, L}(P) \triangleq \min_{F \in \mathcal{F}: P \subseteq F} \mathfrak{C}_{\text{ext}}(F),$$

$$\text{prob}_{\mathfrak{C}, \mathcal{F}, L}(P) \triangleq \min(1, L \cdot \frac{\mathfrak{C}(P)}{\text{assoc}_{\mathfrak{C}, \mathcal{F}, L}(P)}), \text{ and}$$

$$\text{count}_{\mathfrak{C}, \mathcal{F}, L}(P) \triangleq \frac{\mathfrak{C}(P)}{\text{prob}_{\mathfrak{C}, \mathcal{F}, L}(P)}.$$

Then, for any set  $P \in \mathcal{P}$ , set  $\widehat{\mathfrak{C}}(P)$  to  $\text{count}_{\mathfrak{C}, \mathcal{F}, L}(P)$  with probability  $\text{prob}_{\mathfrak{C}, \mathcal{F}, L}(P)$  and to 0 with probability

$$1 - \text{prob}_{\mathfrak{C}, \mathcal{F}, L}(P).^1$$

Firstly, observe that the support of any counter that can be potentially output is contained in the support of the input counter. Essentially, with this sampling procedure we aim to discard as many sets as possible from the support of the input counter while modifying the values of those that are kept so that we obtain a representative counter of small support. Intuitively, each set  $P \in \mathcal{P}$  is associated, among the sets in  $\mathcal{F}$  that contain it and hence whose value is effected by the value of  $P$  (as assigned by the counter), with a set  $F$  having minimum value. Thus,  $P$  is associated with a set  $F$  for which  $P$  is most significant among all sets in  $\mathcal{F}$ —that is, in which the fraction of the value of  $P$  from the entire value of  $F$  is largest. In a sense, this means that the value of  $F$  is most “vulnerable” in case  $P$  will be dropped from the support of the counter. Next, the probability of keeping  $P$  in the support is chosen to be proportional to its fraction of value within  $F$ —the larger  $\mathfrak{C}(P)$  is, the larger is the probability to choose it, but at the same time, the larger  $\text{assoc}_{\mathfrak{C}, \mathcal{F}, L}(P)$  is (which means that the set  $F$  associated with  $P$ , and hence all other sets in  $\mathcal{F}$  as well, are less vulnerable to  $P$  being dropped out), the smaller is the probability to choose  $P$ . The factor  $L$  (whose exact value will be determined later) is meant to boost up the probability to be larger than just the fraction of the value of  $P$  within  $F$  (else we may drop “too many” sets from the support, and hence the output counter will not represent the input counter). Due to this boosting factor, we also need to trim down the boosted fraction to be 1 so that it will indeed represent a probability. Lastly, the new value of  $P$  when decided to be kept in the support, is chosen in a way as to ensure that its expected value (being the probability to keep it times its new value when it is kept) will be equal to its original value.

We first show the the size of the support of the output counter is expected to be “small” (in case the size of the family  $\mathcal{F}$  and the boosting factor  $L$  are both “small”).

**LEMMA 3.2.** *Let  $U$  be a universe, and let  $\mathcal{F} \subseteq 2^U$  with  $U \in \mathcal{F}$ . Let  $p, L \in \mathbb{N}_0$  and  $\mathcal{P} \subseteq \binom{U}{p}$ . Let  $\mathfrak{C} : \mathcal{P} \rightarrow \mathbb{N}_0$  be a counter. Then, the expected size of the support of the output counter  $\hat{\mathfrak{C}}$  of  $(\mathfrak{C}, \mathcal{F}, L)$ -counter sampling is upper bounded as follows.*

$$E[|\text{supp}(\hat{\mathfrak{C}})|] \leq |\mathcal{F}| \cdot L.$$

Moreover, for any  $\hat{c} \geq 0$ , we have that  $\Pr(|\text{supp}(\hat{\mathfrak{C}})| > (\hat{c} + 1) \cdot |\mathcal{F}| \cdot L) \leq 2e^{-\frac{\hat{c}^2}{2}}$ .

<sup>1</sup>Since  $U \in \mathcal{F}$ , there exists  $F \in \mathcal{F}$  such that  $P \subseteq F$ , hence  $\text{assoc}_{\mathfrak{C}, \mathcal{F}, L}(P)$  is well defined.

*Proof.* Observe that

$$E[|\text{supp}(\hat{\mathfrak{C}})|] = \sum_{P \in \mathcal{P}} \text{prob}_{\mathfrak{C}, \mathcal{F}, L}(P) \quad (1)$$

$$= \sum_{P \in \mathcal{P}} \min(1, L \cdot \frac{\mathfrak{C}(P)}{\text{assoc}_{\mathfrak{C}, \mathcal{F}, L}(P)}) \quad (2)$$

$$\leq L \cdot \sum_{P \in \mathcal{P}} \frac{\mathfrak{C}(P)}{\text{assoc}_{\mathfrak{C}, \mathcal{F}, L}(P)} \quad (3)$$

$$\leq L \cdot \sum_{F \in \mathcal{F}} \sum_{P \in \mathcal{P} : \text{assoc}_{\mathfrak{C}, \mathcal{F}, L}(P) = \mathfrak{C}_{\text{ext}}(F)} \frac{\mathfrak{C}(P)}{\text{assoc}_{\mathfrak{C}, \mathcal{F}, L}(P)} \quad (4)$$

$$= L \cdot \sum_{F \in \mathcal{F}} \sum_{P \in \mathcal{P} : \text{assoc}_{\mathfrak{C}, \mathcal{F}, L}(P) = \mathfrak{C}_{\text{ext}}(F)} \frac{\mathfrak{C}(P)}{\mathfrak{C}_{\text{ext}}(F)} \quad (5)$$

$$\leq L \cdot \sum_{F \in \mathcal{F}} \left( \frac{1}{\mathfrak{C}_{\text{ext}}(F)} \cdot \sum_{P \in \mathcal{P} : \text{assoc}_{\mathfrak{C}, \mathcal{F}, L}(P) = \mathfrak{C}_{\text{ext}}(F)} \mathfrak{C}(P) \right) \quad (6)$$

$$\leq L \cdot \sum_{F \in \mathcal{F}} \left( \frac{1}{\mathfrak{C}_{\text{ext}}(F)} \cdot \sum_{P \in \mathcal{P} : P \subseteq F} \mathfrak{C}(P) \right) \quad (7)$$

$$\leq L \cdot \sum_{F \in \mathcal{F}} \frac{1}{\mathfrak{C}_{\text{ext}}(F)} \cdot \mathfrak{C}_{\text{ext}}(F) = L \cdot |\mathcal{F}|. \quad (8)$$

Here, (1), (3), (5), (6) and the equality at (8) are immediate. The equality (2) follows from the definition of  $\text{prob}_{\mathfrak{C}, \mathcal{F}, L}(P)$ . The inequality (4) follows from the observation that for each set  $P \in \mathcal{P}$ , there exists a (not necessarily unique) set  $F \in \mathcal{F}$  such that  $\text{assoc}_{\mathfrak{C}, \mathcal{F}, L}(P) = \mathfrak{C}(F)$ . The inequality (7) follows from the definition of  $\text{assoc}_{\mathfrak{C}, \mathcal{F}, L}$ , and the inequality at (8) follows from the definition of  $\mathfrak{C}_{\text{ext}}$ .

For the second claim in the proof, let  $\hat{c} \geq 1$ . Because  $E[|\text{supp}(\hat{\mathfrak{C}})|] \leq |\mathcal{F}| \cdot L$ , we have that  $\Pr(|\text{supp}(\hat{\mathfrak{C}})| > (\hat{c} + 1) \cdot |\mathcal{F}| \cdot L) \leq \Pr(|\text{supp}(\hat{\mathfrak{C}})| - E[|\text{supp}(\hat{\mathfrak{C}})|] > \hat{c} \cdot E[|\text{supp}(\hat{\mathfrak{C}})|])$ . By Chernoff bound (Proposition 2.2), the aforementioned term is upper bounded by  $2e^{-\frac{\hat{c}^2 E[|\text{supp}(\hat{\mathfrak{C}})|]^2}{2}}$ . In case  $E[|\text{supp}(\hat{\mathfrak{C}})|] \geq 1$ , then the aforementioned term is upper bounded by  $2e^{-\frac{\hat{c}^2}{2}}$ , which completes the proof. Else, in case  $E[|\text{supp}(\hat{\mathfrak{C}})|] < 1$ , we have that

$$\begin{aligned} E[|\text{supp}(\hat{\mathfrak{C}})|] &= \sum_{P \in \mathcal{P}} \text{prob}_{\mathfrak{C}, \mathcal{F}, L}(P) \\ &= \sum_{P \in \mathcal{P}} \min(1, L \cdot \frac{\mathfrak{C}(P)}{\text{assoc}_{\mathfrak{C}, \mathcal{F}, L}(P)}) < 1. \end{aligned}$$

Thus, for every  $P \in \mathcal{P}$ , we have that  $\text{prob}_{\mathfrak{C}, \mathcal{F}, L}(P) = L \cdot \frac{\mathfrak{C}(P)}{\text{assoc}_{\mathfrak{C}, \mathcal{F}, L}(P)}$ . Moreover, for every  $P \in \mathcal{P}$ , we have that  $\text{assoc}_{\mathfrak{C}, \mathcal{F}, L}(P) \leq \mathfrak{C}_{\text{ext}}(U)$ , and therefore  $\text{prob}_{\mathfrak{C}, \mathcal{F}, L}(P) \geq L \cdot \frac{\mathfrak{C}(P)}{\mathfrak{C}_{\text{ext}}(U)}$ . However, we thus derive that  $E[|\text{supp}(\hat{\mathfrak{C}})|] =$



$\sum_{P \in \mathcal{P}} \text{prob}_{\mathfrak{C}, \mathcal{F}, L}(P) \geq L \cdot \frac{\sum_{P \in \mathcal{P}} \mathfrak{C}(P)}{\mathfrak{C}_{\text{ext}}(U)} = L \geq 1$ , which is a contradiction.  $\square$

Now, we present a statement regarding the new values and probabilities assigned by the sampling procedure. This statement will be used soon together with the observation ahead, towards the proof that the output counter is likely to represent the input one.

**LEMMA 3.3.** *Let  $U$  be a universe, and let  $\mathcal{F} \subseteq 2^U$  with  $U \in \mathcal{F}$ . Let  $p, L \in \mathbb{N}_0$  and  $\mathcal{P} \subseteq \binom{U}{p}$ . Let  $\mathfrak{C} : \mathcal{P} \rightarrow \mathbb{N}_0$  be a counter. Then, for all  $P \in \mathcal{P}$  and  $F \in \mathcal{F}$  such that  $P \subseteq F$ , at least one of the following two conditions is satisfied.*

- $\text{count}_{\mathfrak{C}, \mathcal{F}, L}(P) \leq \frac{\mathfrak{C}_{\text{ext}}(F)}{L}$ .
- $\text{prob}_{\mathfrak{C}, \mathcal{F}, L}(P) = 1$ .

*Proof.* Consider some  $P \in \mathcal{P}$  and  $F \in \mathcal{F}$  such that  $P \subseteq F$ . We need to prove that at least one of the two conditions in the lemma is satisfied. In case  $\text{prob}_{\mathfrak{C}, \mathcal{F}, L}(P) = 1$ , we are done. Thus, we next suppose that  $\text{prob}_{\mathfrak{C}, \mathcal{F}, L}(P) \neq 1$ . Then, by the definition of  $\text{prob}_{\mathfrak{C}, \mathcal{F}, L}(P)$ , we have that  $\text{prob}_{\mathfrak{C}, \mathcal{F}, L}(P) = L \cdot \frac{\mathfrak{C}(P)}{\text{assoc}_{\mathfrak{C}, \mathcal{F}, L}(P)}$ . By the definition of  $\text{assoc}_{\mathfrak{C}, \mathcal{F}, L}(P)$ , we have that  $\text{assoc}_{\mathfrak{C}, \mathcal{F}, L}(P) \leq \mathfrak{C}_{\text{ext}}(F)$ . Thus,  $\text{prob}_{\mathfrak{C}, \mathcal{F}, L}(P) \geq L \cdot \frac{\mathfrak{C}(P)}{\mathfrak{C}_{\text{ext}}(F)}$ . From this inequality and the definition of  $\text{count}_{\mathfrak{C}, \mathcal{F}, L}(P)$ , we derive that

$$\text{count}_{\mathfrak{C}, \mathcal{F}, L}(P) = \frac{\mathfrak{C}(P)}{\text{prob}_{\mathfrak{C}, \mathcal{F}, L}(P)} \leq \frac{\mathfrak{C}(P)}{L \cdot \frac{\mathfrak{C}(P)}{\mathfrak{C}_{\text{ext}}(F)}} = \frac{\mathfrak{C}_{\text{ext}}(F)}{L}.$$

This completes the proof.  $\square$

We will also need the following two simple observations where the first asserts representation in expectation and the second, which is an immediate consequence of the first, concerns the expected output value of each set in  $\mathcal{F}$ .

**OBSERVATION 1.** *Let  $U$  be a universe, and let  $\mathcal{F} \subseteq 2^U$  with  $U \in \mathcal{F}$ . Let  $p, L \in \mathbb{N}_0$  and  $\mathcal{P} \subseteq \binom{U}{p}$ . Let  $\mathfrak{C} : \mathcal{P} \rightarrow \mathbb{N}_0$  be a counter. The output counter  $\hat{\mathfrak{C}}$  of  $(\mathfrak{C}, \mathcal{F}, L)$ -counter sampling represents in expectation  $\mathfrak{C}$ .*

*Proof.* Consider some set  $P \in \mathcal{P}$ . Then, the definition of  $(\mathfrak{C}, \mathcal{F}, L)$ -counter sampling yields that

$$E[\hat{\mathfrak{C}}(P)] = \text{prob}_{\mathfrak{C}, \mathcal{F}, L}(P) \cdot \text{count}_{\mathfrak{C}, \mathcal{F}, L}(P) = \mathfrak{C}(P).$$

Since the choice of  $P$  was arbitrary, we derive that  $\hat{\mathfrak{C}}$  represents in expectation  $\mathfrak{C}$ .  $\square$

**OBSERVATION 2.** *Let  $U$  be a universe, and let  $\mathcal{F} \subseteq 2^U$  with  $U \in \mathcal{F}$ . Let  $p, L \in \mathbb{N}_0$  and  $\mathcal{P} \subseteq \binom{U}{p}$ . Let  $\mathfrak{C} : \mathcal{P} \rightarrow \mathbb{N}_0$  be a counter. For any set  $F \subseteq U$ , for the output counter  $\hat{\mathfrak{C}}$  of  $(\mathfrak{C}, \mathcal{F}, L)$ -counter sampling, we have that  $E[\hat{\mathfrak{C}}_{\text{ext}}(F)] = \mathfrak{C}_{\text{ext}}(F)$ .*

*Proof.* Consider some set  $F \subseteq U$ . Then,

$$E[\hat{\mathfrak{C}}_{\text{ext}}(F)] = E\left[\sum_{P \in \binom{F}{p} \cap \mathcal{P}} \hat{\mathfrak{C}}(P)\right] \quad (1)$$

$$= \sum_{P \in \binom{F}{p} \cap \mathcal{P}} E[\hat{\mathfrak{C}}(P)] \quad (2)$$

$$= \sum_{P \in \binom{F}{p} \cap \mathcal{P}} \mathfrak{C}(P) = \mathfrak{C}_{\text{ext}}(F). \quad (3)$$

Here, equality (1) and the second equality at (3) follow from the definition of domain extension, equality (2) follows from the linearity of expectation, and the first equality at (3) follows from Observation 1.  $\square$

From Lemma 3.3 and Observation 2, we derive the following corollary.

**COROLLARY 3.1.** *Let  $U$  be a universe, and let  $\mathcal{F} \subseteq 2^U$  with  $U \in \mathcal{F}$ . Let  $p, L \in \mathbb{N}_0$  and  $\mathcal{P} \subseteq \binom{U}{p}$ . Let  $\mathfrak{C} : \mathcal{P} \rightarrow \mathbb{N}_0$  be a counter. For any  $F \in \mathcal{F}$ , for the output counter  $\hat{\mathfrak{C}}$  of  $(\mathfrak{C}, \mathcal{F}, L)$ -counter sampling, we have that  $E\left[\frac{\hat{\mathfrak{C}}_{\text{ext}}(F)}{W}\right] \geq L$  where  $W = \max\{\text{count}_{\mathfrak{C}, \mathcal{F}, L}(P) : P \in \mathcal{P}, \text{prob}_{\mathfrak{C}, \mathcal{F}, L}(P) < 1\}$ .*

*Proof.* Consider some  $F \in \mathcal{F}$ . By Lemma 3.3, for all  $P \in \binom{F}{p} \cap \mathcal{P}$  such that  $\text{prob}_{\mathfrak{C}, \mathcal{F}, L}(P) < 1$ , it must hold that  $\text{count}_{\mathfrak{C}, \mathcal{F}, L}(P) \leq \frac{\mathfrak{C}_{\text{ext}}(F)}{L}$ , implying that necessarily  $W \leq \frac{\mathfrak{C}_{\text{ext}}(F)}{L}$ . Therefore,

$$E\left[\frac{\hat{\mathfrak{C}}_{\text{ext}}(F)}{W}\right] = \frac{1}{W} \cdot E[\hat{\mathfrak{C}}_{\text{ext}}(F)] \geq \frac{L}{\mathfrak{C}_{\text{ext}}(F)} \cdot E[\hat{\mathfrak{C}}_{\text{ext}}(F)] = L.$$

Here, the last equality follows from Observation 2.  $\square$

We are now ready to prove that the output counter is likely to represent the input one.

**LEMMA 3.4.** *Let  $U$  be a universe, and let  $\mathcal{F} \subseteq 2^U$  with  $U \in \mathcal{F}$ . Let  $p, c, L \in \mathbb{N}_0$  such that  $L \geq 2^{\frac{1}{c}} \ln(2c|\mathcal{F}|)$ . Let  $\mathcal{P} \subseteq \binom{U}{p}$ . Let  $\mathfrak{C} : \mathcal{P} \rightarrow \mathbb{N}_0$  be a counter. Then, the probability that  $\mathfrak{C}$  and the output counter  $\hat{\mathfrak{C}}$  of  $(\mathfrak{C}, \mathcal{F}, L)$ -counter sampling are  $(\epsilon, \mathcal{F})$ -similar is at least  $1 - \frac{1}{c}$ .*

*Proof.* Consider some  $F \in \mathcal{F}$ , and denote  $W = \max\{\text{count}_{\mathfrak{C}, \mathcal{F}, L}(P) : P \in \mathcal{P}, \text{prob}_{\mathfrak{C}, \mathcal{F}, L}(P) < 1\}$ . For any  $P \in \binom{F}{p} \cap \mathcal{P}$  such that  $\text{prob}_{\mathfrak{C}, \mathcal{F}, L}(P) < 1$ , define

$\ell_P = 1$  and the random variable  $X_{P,1} = \frac{\hat{\mathfrak{C}}(P)}{W}$ . For any  $P \in \binom{F}{p} \cap \mathcal{P}$  such that  $\text{prob}_{\mathfrak{C},\mathcal{F},L}(P) = 1$ , define  $\ell_P = \lceil \frac{\hat{\mathfrak{C}}(P)}{W} \rceil$  and the deterministic variable  $X_{P,\ell_P} = \frac{\hat{\mathfrak{C}}(P)}{W} - (\ell_P - 1)$ , as well as for any  $i \in \{1, \dots, \ell_P - 1\}$ , the deterministic variable  $X_{P,i} = 1$ , and notice that  $\sum_{i=1}^{\ell_P} X_{P,i} = \frac{\hat{\mathfrak{C}}(P)}{W}$ . Then,  $\{X_{P,i} : P \in \binom{F}{p} \cap \mathcal{P}, i \in \{1, \dots, \ell_P\}\}$  is a collection of independent random variables bounded by the interval  $[0, 1]$ . Here, independence among variables  $X_{P,i}$  corresponding to the same set  $P \in \binom{F}{p} \cap \mathcal{P}$  follows because these variables are deterministic. Let  $\bar{X} = \sum_{P \in \binom{F}{p} \cap \mathcal{P}} \sum_{i=1}^{\ell_P} X_{P,i}$ . By Chernoff bound (Proposition 2.2),

$$Pr(|\bar{X} - E[\bar{X}]| > \epsilon E[\bar{X}]) \leq 2e^{-\frac{\epsilon^2 E[\bar{X}]}{2}}.$$

Observe that  $\bar{X} = \sum_{P \in \binom{F}{p} \cap \mathcal{P}} \sum_{i=1}^{\ell_P} X_{P,i} = \sum_{P \in \binom{F}{p} \cap \mathcal{P}} \frac{\hat{\mathfrak{C}}(P)}{W} = \frac{\hat{\mathfrak{C}}_{\text{ext}}(F)}{W}$ . Thus,  $E[\bar{X}] = E[\frac{\hat{\mathfrak{C}}_{\text{ext}}(F)}{W}]$ , hence by Observation 2,  $E[\bar{X}] = \frac{\mathfrak{C}_{\text{ext}}(F)}{W}$ . This means that  $|\bar{X} - E[\bar{X}]| > \epsilon E[\bar{X}]$  is true if and only if  $|\hat{\mathfrak{C}}_{\text{ext}}(F) - \mathfrak{C}_{\text{ext}}(F)| > \epsilon \cdot \mathfrak{C}_{\text{ext}}(F)$  is true. Thus, by this equivalence between events,

$$Pr(|\hat{\mathfrak{C}}_{\text{ext}}(F) - \mathfrak{C}_{\text{ext}}(F)| > \epsilon \cdot \mathfrak{C}_{\text{ext}}(F)) \leq 2e^{-\frac{\epsilon^2 E[\bar{X}]}{2}}.$$

Recall that  $E[\bar{X}] = E[\frac{\hat{\mathfrak{C}}_{\text{ext}}(F)}{W}]$ , hence by Corollary 3.1 and the given lower bound on  $L$ ,  $E[\bar{X}] \geq L \geq 2\frac{1}{\epsilon^2} \ln(2c|\mathcal{F}|)$ . Thus,

$$\begin{aligned} Pr(|\hat{\mathfrak{C}}_{\text{ext}}(F) - \mathfrak{C}_{\text{ext}}(F)| > \epsilon \cdot \mathfrak{C}_{\text{ext}}(F)) &\leq 2e^{-\frac{\epsilon^2 \cdot (2\frac{1}{\epsilon^2} \ln(2c|\mathcal{F}|))}{2}} \\ &= 2e^{-\ln(2c|\mathcal{F}|)} = \frac{2}{2c|\mathcal{F}|} = \frac{1}{c|\mathcal{F}|}. \end{aligned}$$

As the choice of  $F \in \mathcal{F}$  was arbitrary, union bound implies that the probability that there exists  $F \in \mathcal{F}$  such that  $(1 - \epsilon) \cdot \hat{\mathfrak{C}}_{\text{ext}}(F) > \mathfrak{C}_{\text{ext}}(F)$  or  $\mathfrak{C}_{\text{ext}}(F) > (1 + \epsilon) \cdot \hat{\mathfrak{C}}_{\text{ext}}(F)$  is upper bounded by  $|\mathcal{F}| \cdot \frac{1}{c|\mathcal{F}|} = \frac{1}{c}$ . Thus, the probability that  $\mathfrak{C}$  and  $\hat{\mathfrak{C}}$  are  $(\epsilon, \mathcal{F})$ -similar is at least  $1 - \frac{1}{c}$ .  $\square$

We now turn to analyze the time complexity of the sampling procedure.

**LEMMA 3.5.** *Let  $U$  be a universe. Let  $p, L \in \mathbb{N}_0$  and  $\mathcal{P} \subseteq \binom{U}{p}$ . Let  $\mathcal{F} \subseteq 2^U$  with  $U \in \mathcal{F}$  be an  $\epsilon$ -parsimonious  $(n, p, q)$ -universal family  $\mathcal{F}$ , equipped with a  $T$ -membership query procedure. Let  $\mathfrak{C} : \mathcal{P} \rightarrow \mathbb{N}_0$  be a counter. Then, the time complexity of  $(\mathfrak{C}, \mathcal{F}, L)$ -counter sampling is bounded by  $\mathcal{O}(|\text{supp}(\mathfrak{C})| \cdot T)$ .*

*Proof.* First, we initialize the value  $\mathfrak{C}_{\text{ext}}(F)$  of each set  $F \in \mathcal{F}$  to be 0. Then, for every set  $P \in \text{supp}(\mathfrak{C})$ , we compute  $\mathcal{F}' = \{F \in \mathcal{F} : P \subseteq F\}$  in time  $\mathcal{O}(T)$  using the membership query procedure (which implies that  $|\mathcal{F}'| = \mathcal{O}(T)$ ), and then for each set  $F \in \mathcal{F}'$  we update  $\mathfrak{C}_{\text{ext}}(F)$  by adding  $\mathfrak{C}(P)$  to it. Thus, in time  $\mathcal{O}(|\text{supp}(\mathfrak{C})| \cdot T)$  we correctly compute  $\mathfrak{C}_{\text{ext}}(F)$  for all  $F \in \mathcal{F}$ . Now, for each set  $P \in \text{supp}(\mathfrak{C})$ , we can compute  $\text{assoc}_{\mathfrak{C},\mathcal{F},L}(P)$  in time  $\mathcal{O}(T)$ , then  $\text{prob}_{\mathfrak{C},\mathcal{F},L}(P)$  in time  $\mathcal{O}(1)$ , and lastly  $\text{count}_{\mathfrak{C},\mathcal{F},L}(P)$  in time  $\mathcal{O}(1)$ . Overall, we have so far spent time  $\mathcal{O}(|\text{supp}(\mathfrak{C})| \cdot T)$ . Finally, picking up sets using their probabilities and new values is done in time  $\mathcal{O}(|\text{supp}(\mathfrak{C})|)$ .  $\square$

We conclude this subsection with the following theorem.

**THEOREM 3.1.** *Let  $U$  be a universe. Let  $0 < \epsilon < 1$ ,  $p, q, c \in \mathbb{N}_0$ ,  $\mathcal{P} \subseteq \binom{U}{p}$  and  $\mathcal{Q} \subseteq \binom{U}{q}$ . Let  $\mathcal{F} \subseteq 2^U$  be an  $\frac{1}{5}\epsilon$ -parsimonious  $(n, p, q)$ -universal family with respect to  $(\mathcal{P}, \mathcal{Q})$  of size  $S$ , equipped with a  $T$ -membership query procedure. Let  $\mathfrak{C} : \mathcal{P} \rightarrow \mathbb{N}_0$  be a counter. Then, a counter  $\hat{\mathfrak{C}} : \mathcal{P} \rightarrow \mathbb{N}_0$  such that*

1.  $\hat{\mathfrak{C}}$  necessarily (with probability 1) represents in expectation  $\mathfrak{C}$ , and
2. with success probability at least  $1 - \frac{1}{c}$ ,  $\hat{\mathfrak{C}}$   $(\epsilon, q)$ -represents  $\mathfrak{C}$  with respect to  $\mathcal{Q}$  and satisfies  $|\text{supp}(\hat{\mathfrak{C}})| \leq \mathcal{O}((\frac{1}{\epsilon})^2 S \log c (\log c + \log S))$ ,

can be computed in time  $\mathcal{O}(|\text{supp}(\mathfrak{C})| \cdot T)$ .

*Proof.* Without loss of generality, we suppose that  $U \in \mathcal{F}$ , else we just add  $U$  to  $\mathcal{F}$ . By Lemma 3.1, to prove the theorem, it suffices to compute in time  $\mathcal{O}(|\text{supp}(\mathfrak{C})| \cdot T)$  a counter  $\hat{\mathfrak{C}} : \binom{U}{p} \rightarrow \mathbb{N}_0$  that necessarily represents in expectation  $\mathfrak{C}$ , and that with probability at least  $1 - \frac{1}{c}$  is  $(\frac{\epsilon}{5}, \mathcal{F})$ -similar to  $\mathfrak{C}$  and satisfies  $|\text{supp}(\hat{\mathfrak{C}})| \leq \mathcal{O}((\frac{1}{\epsilon})^2 S \log c \log(cS))$ .

Fix  $L = \lceil 2\frac{1}{(\frac{\epsilon}{5})^2} \ln(4cS) \rceil = \mathcal{O}((\frac{1}{\epsilon})^2 \log(cS))$ . By Lemma 3.2 with  $\hat{c} = \sqrt{2\ln(4c)}$ , with probability at most  $2e^{-\frac{\hat{c}^2}{2}} = 2e^{-\ln(4c)} = \frac{1}{2c}$ , we have that the expected size of the support of the output counter  $\hat{\mathfrak{C}}$  of  $(\mathfrak{C}, \mathcal{F}, L)$ -counter sampling is upper bounded as follows.

$$E[|\text{supp}(\hat{\mathfrak{C}})|] \leq (\hat{c} + 1)SL.$$

Moreover, by Lemma 3.4,  $\mathfrak{C}$  and  $\widehat{\mathfrak{C}}$  are  $(\frac{\epsilon}{5}, \mathcal{F})$ -similar with probability at least  $1 - \frac{1}{2c}$ . By union bound, the probability that  $|\text{supp}(\widehat{\mathfrak{C}})| \geq (\widehat{c} + 1)SL$  or that  $\mathfrak{C}$  and  $\widehat{\mathfrak{C}}$  are not  $(\frac{\epsilon}{5}, \mathcal{F})$ -similar is at most  $\frac{1}{2c} + \frac{1}{2c} = \frac{1}{c}$ . Thus, with probability at least  $1 - \frac{1}{c}$ , both  $|\text{supp}(\mathfrak{C})| \leq (\widetilde{c} + 1)SL = \mathcal{O}(\log c \cdot S \cdot (\frac{1}{\epsilon})^2 \log(cS)) = \mathcal{O}((\frac{1}{\epsilon})^2 S \log c (\log c + \log S))$  and  $\mathfrak{C}$  and  $\widehat{\mathfrak{C}}$  are  $(\frac{\epsilon}{5}, \mathcal{F})$ -similar. Further, by Lemma 3.5,  $\widehat{\mathfrak{C}}$  is computed in time  $\mathcal{O}(|\text{supp}(\mathfrak{C})| \cdot T)$ . Lastly, by Observation 1,  $\widehat{\mathfrak{C}}$  necessarily represents in expectation  $\mathfrak{C}$ . This completes the proof.  $\square$

We remark that as a corollary to this theorem (with  $c = 2$  and where the membership query procedure is simply brute-force) and Proposition 2.1, we can already assert the *existence* of representative counters of small support.

**3.2 Parsimonious Universal Families with Membership Query Procedures** We will be able to equip our parsimonious universal families with efficient membership query procedures only when we deal with  $\mathcal{P}$  and  $\mathcal{Q}$  that are “balancedly split”. Towards the definition of this term, we first present the following definition.

**DEFINITION 9.** Let  $t, k, p, b \in \mathbb{N}$ . A tuple  $\overline{U} = (U_1, U_2, \dots, U_t)$  where  $U_1, U_2, \dots, U_t$  are pairwise-disjoint universes is called a  $t$ -partitioned universe. Moreover, a function  $f : \{1, 2, \dots, t\} \rightarrow \{0, 1, \dots, \lfloor bk/t \rfloor\}$  that satisfies  $\sum_{i=1}^t f(i) = k$  is called a  $(t, k, b)$ -splitting function. Lastly, a pair  $(f, g)$  of a  $(t, k, b)$ -splitting function  $f : \{1, 2, \dots, t\} \rightarrow \{0, 1, \dots, \lfloor bk/t \rfloor\}$  and a function  $g : \{1, 2, \dots, t\} \rightarrow \{0, 1, \dots, \lfloor bk/t \rfloor\}$  that satisfies  $g \leq f$  and  $\sum_{i=1}^t g(i) = p$ , is called a  $(t, k, p, b)$ -splitting function pair.

When  $t$  or  $(t, k, p, b)$  is clear from context, we do not mention it explicitly. Notice that when  $p = k$ , necessarily  $g = f$ . We now present a definition which will be useful only for product counters; by considering it already here, we will be able to avoid repetition of arguments.

Now, we define the notion of balancedly split sets.

**DEFINITION 10. (Balancedly Split Sets I)** Let  $t, k, p, b \in \mathbb{N}$  with  $p \leq k$ . Let  $\overline{U} = (U_1, U_2, \dots, U_t)$  be a partitioned universe with  $U = \bigcup_{i=1}^t U_i$ , and let  $(f, g)$  be a splitting function pair. Then,  $P \in \binom{U}{p}$  is  $(\overline{U}, f, g)$ -balancedly split if for every  $i \in \{1, 2, \dots, t\}$ , it holds that  $|P \cap U_i| = g(i)$ ; in case  $k = p$ ,  $P$  is  $(\overline{U}, f)$ -balancedly split. Further,  $\mathcal{P}_{\overline{U}, f, g}^{\text{BAL}} \subseteq \binom{U}{p}$  denotes the collection of all  $(\overline{U}, f, g)$ -balancedly split sets. Moreover,  $Q \in \binom{U}{k-p}$  is complementary  $(\overline{U}, f, g)$ -balancedly

split if for every  $i \in \{1, 2, \dots, t\}$ , it holds that  $|Q \cap U_i| = f(i) - g(i)$ . Further,  $\mathcal{Q}_{\overline{U}, f, g}^{\text{CBAL}} \subseteq \binom{U}{k-p}$  denotes the collection of all complementary  $(\overline{U}, f, g)$ -balancedly split sets.

When  $\overline{U}, f$  and  $g$  are clear from context, we do not mention it explicitly.

Our computation of universal families will be done in a sampling procedure defined as follows.

**DEFINITION 11. (Universal Family Sampling)** Let  $t, k, p, b \in \mathbb{N}$  with  $p \leq k$ ,  $0 < \epsilon < 1$  and  $c, d \geq 1$ . Let  $\overline{U} = (U_1, U_2, \dots, U_t)$  be a partitioned universe with  $U = \bigcup_{i=1}^t U_i$  of size  $n$ , and let  $(f, g)$  be a splitting function pair. Then,  $(\overline{U}, f, g, \epsilon, c, d)$ -universal family sampling is the randomized procedure that constructs a family  $\mathcal{F} \subseteq 2^U$  as follows.

- For  $i \in \{1, 2, \dots, t\}$ :
  - For  $j \in \{1, 2, \dots, s_i\}$  with  $s_i$  being
 
$$\frac{(d \cdot f(i))^{f(i)}}{g(i)^{g(i)}(d \cdot f(i) - g(i))^{f(i) - g(i)}} \cdot \frac{1}{\widehat{c}^2} 10k \ln(nc),$$
 where  $\widehat{c} = \frac{\ln(1+\epsilon)}{t}$ , construct a set  $F_{i,j} \subseteq U_i$  as follows. Each element in  $U_i$  is inserted independently with probability  $\frac{g(i)}{d \cdot f(i)}$  into  $F_{i,j}$ .
  - Denote  $\mathcal{F}_i = \{F_{i,j} : j \in \{1, 2, \dots, s_i\}\}$ .
- Then, construct  $\mathcal{F} = \{F_{1,j_1} \cup F_{2,j_2} \cup \dots \cup F_{t,j_t} : F_{1,j_1} \in \mathcal{F}_1, F_{i,j_2} \in \mathcal{F}_2, \dots, F_{t,j_t} \in \mathcal{F}_t\}$ .

We remark that  $d$  can depend on any argument of interest (e.g.,  $k$  and  $p$ ). We begin the analysis of the sampling procedure by an observation concerning its time complexity and by giving an upper bound on the size of the family it produces.

**OBSERVATION 3.** Let  $t, k, p, b \in \mathbb{N}$  with  $p \leq k$ ,  $0 < \epsilon < 1$  and  $c, d \geq 1$ . Let  $\overline{U} = (U_1, U_2, \dots, U_t)$  be a partitioned universe with  $U = \bigcup_{i=1}^t U_i$  of size  $n$ , and let  $(f, g)$  be a splitting function pair. Then, the time complexity of  $(\overline{U}, b, f, g, \epsilon, c, d)$ -universal is  $\mathcal{O}(|\mathcal{F}|n)$ , where  $\mathcal{F} \subseteq 2^U$  is the output family.

For lack of space, we omit the proof of the following lemma.

**LEMMA 3.6.** Let  $t, k, p, b \in \mathbb{N}$  with  $p \leq k$ ,  $0 < \epsilon < 1$  and  $c, d \geq 1$ . Let  $\overline{U} = (U_1, U_2, \dots, U_t)$  be a partitioned universe with  $U = \bigcup_{i=1}^t U_i$  of size  $n$ , and let  $(f, g)$  be a splitting function pair. Then, the output family  $\mathcal{F} \subseteq 2^U$  of  $(\overline{U}, b, f, g, \epsilon, c, d)$ -universal family sampling necessarily satisfies  $|\mathcal{F}| \leq \frac{(dk)^k}{p^p(dk - p)^{k-p}} \cdot \left(\frac{1}{\ln^2(1 + \epsilon)} \cdot 10k^3 \cdot \ln(nc)\right)^t$ .

We proceed by giving a lower bound for the probability of failure of the procedure to produce a parsimonious universal family with respect to a balancedly split pair.

**LEMMA 3.7.** *Let  $t, k, p, b \in \mathbb{N}$  with  $p \leq k$ ,  $0 < \epsilon < 1$  and  $c, d \geq 1$ . Let  $\bar{U} = (U_1, U_2, \dots, U_t)$  be a partitioned universe with  $U = \bigcup_{i=1}^t U_i$  of size  $n$ , and let  $(f, g)$  be a splitting function pair. With probability at least  $1 - \frac{1}{2c}$ , the output family  $\mathcal{F} \subseteq 2^U$  of  $(\bar{U}, b, f, g, \epsilon, c, d)$ -universal family sampling is an  $\epsilon$ -parsimonious  $(n, p, q)$ -universal family with respect to  $(\mathcal{P}_{k,f,g}^{\text{BAL}}, \mathcal{Q}_{k,f,g}^{\text{CBAL}})$  with correction factor upper bounded by  $\left(\frac{1}{\left(\frac{\ln(1+\epsilon)}{t}\right)^2} \cdot 10k \cdot \ln(nc)\right)^t$ .*

*Proof.* Towards the proof of the lemma, we first show that the following claim is correct.

**CLAIM 1.** *With probability at least  $1 - \frac{1}{2c}$ , for every  $i \in \{1, 2, \dots, t\}$ , we have that  $\mathcal{F}_i$  is an  $\hat{\epsilon}$ -parsimonious  $(|U_i|, g(i), f(i) - g(i))$ -universal family with respect to  $(\binom{U_i}{g(i)}, \binom{U_i}{f(i)-g(i)})$  with correction factor  $T_i = \frac{1}{\hat{\epsilon}^2} \cdot 10k \cdot \ln(nc)$ .*

*Proof.* By union bound, it suffices to choose some  $i \in \{1, 2, \dots, t\}$ , and prove that with failure probability at most  $\frac{1}{2ct}$ , we have that  $\mathcal{F}_i$  is an  $\hat{\epsilon}$ -parsimonious  $(|U_i|, g(i), f(i) - g(i))$ -universal family with respect to  $(\binom{U_i}{g(i)}, \binom{U_i}{f(i)-g(i)})$  with correction factor  $T_i = \frac{1}{\hat{\epsilon}^2} \cdot 10k \cdot \ln(nc)$ . Further, by union bound, because there are at most  $|U_i|^{f(i)} \leq n^k$  pairs of disjoint sets  $P \in \binom{U_i}{g(i)}$  and  $Q \in \binom{U_i}{f(i)-g(i)}$ , it suffices to choose some such pair of disjoint sets  $P \in \binom{U_i}{g(i)}$  and  $Q \in \binom{U_i}{f(i)-g(i)}$ , and prove that with failure probability at most  $\frac{1}{2ctn^k}$ , it holds that  $(1 - \hat{\epsilon})T_i \leq |\mathcal{F}_i[P, Q]| \leq (1 + \hat{\epsilon})T_i$ .

Towards the proof of the above, observe that each set  $F_{i,j} \in \mathcal{F}_i$  contains  $P$  and is disjoint from  $Q$  with probability  $\frac{g(i)^{g(i)}(d \cdot f(i) - g(i))^{f(i)-g(i)}}{(d \cdot f(i))^{f(i)}}$ . Thus, the expected number of sets in  $\mathcal{F}_i$  that contain  $P$  and are disjoint from  $Q$  is  $T_i$ . Because the sets in  $\mathcal{F}_i$  are sampled independently from one another, by Chernoff bound (Proposition 2.2), we have that

$$\begin{aligned} & Pr(|\mathcal{F}_i[P, Q]| - T_i > \hat{\epsilon}T_i) \\ & \leq 2e^{-\frac{\hat{\epsilon}^2 T_i}{2}} \\ & = 2e^{-5k \cdot \ln(nc)} = \frac{2}{(nc)^{5k}} \leq \frac{2}{n^4 \cdot n^k \cdot c} \leq \frac{1}{2ct} \end{aligned}$$

Here, the last inequality follows since  $n \geq \max(2, t)$ . This completes the proof of the claim.  $\square$

We now return to the proof of the lemma. Let  $T = \prod_{i=1}^t T_i$  where  $T_i$  is the correction factor of  $\mathcal{F}_i$ .

Then,  $T = \left(\frac{1}{\hat{\epsilon}^2} \cdot 10k \cdot \ln(nc)\right)^t$ . Due to Claim 1, to prove the lemma it suffices to show that, under the assumption that for every  $i \in \{1, 2, \dots, t\}$ , we have that  $\mathcal{F}_i \subseteq 2^{U_i}$  is an  $\hat{\epsilon}$ -parsimonious  $(|U_i|, g(i), f(i) - g(i))$ -universal family with respect to  $(\binom{U_i}{g(i)}, \binom{U_i}{f(i)-g(i)})$ , it holds that  $\mathcal{F}$  is an  $\epsilon$ -parsimonious  $(n, p, q)$ -universal family with respect to  $(\mathcal{P}^{\text{BAL}}, \mathcal{Q}^{\text{CBAL}})$  with correction factor  $T$ . Towards the proof of this, consider some pair of disjoint sets  $P \in \mathcal{P}^{\text{BAL}}$  and  $Q \in \mathcal{Q}^{\text{CBAL}}$ . Then,

$$|\mathcal{F}[P, Q]| = \prod_{i=1}^t |\mathcal{F}_i[P \cap U_i, Q \cap U_i]|.$$

Because  $P \in \mathcal{P}^{\text{BAL}}$  and  $Q \in \mathcal{Q}^{\text{CBAL}}$ , it holds that for every  $i \in \{1, 2, \dots, t\}$ ,  $P \cap U_i \in \binom{U_i}{g(i)}$  and  $Q \cap U_i \in \binom{U_i}{f(i)-g(i)}$ . Thus, for every  $i \in \{1, 2, \dots, t\}$ , because  $\mathcal{F}_i$  is an  $\hat{\epsilon}$ -parsimonious  $(|U_i|, g(i), f(i) - g(i))$ -universal family with respect to  $(\binom{U_i}{g(i)}, \binom{U_i}{f(i)-g(i)})$ , it holds that

$$(1 - \hat{\epsilon})T_i \leq |\mathcal{F}_i[P \cap U_i, Q \cap U_i]| \leq (1 + \hat{\epsilon})T_i$$

Therefore, on the one hand,  $|\mathcal{F}[P, Q]| \leq \prod_{i=1}^t (1 + \hat{\epsilon})T_i = (1 + \hat{\epsilon})^t \cdot T = (1 + \frac{\ln(1+\epsilon)}{t})^t \cdot T \leq e^{\ln(1+\epsilon) \cdot T} = (1 + \epsilon) \cdot T$ . On

the other hand,  $|\mathcal{F}[P, Q]| \geq \prod_{i=1}^t (1 - \hat{\epsilon})T_i = (1 - \hat{\epsilon})^t \cdot T = (1 - \frac{\ln(1+\epsilon)}{t})^t \cdot T \geq (1 - \ln(1 + \epsilon)) \cdot T \geq (1 - \epsilon) \cdot T$ .

Here, the inequality  $(1 - \frac{\ln(1+\epsilon)}{t})^t \geq (1 - \ln(1 + \epsilon))$  follows since the larger  $t$  is (starting at 1), the larger the value of  $(1 - \frac{\ln(1+\epsilon)}{t})^t$  (approaching  $e^{-\ln(1+\epsilon)}$ ), and the inequality  $\ln(1 + \epsilon) \leq \epsilon$  follows from Taylor series. Because the choice of the disjoint sets  $P \in \mathcal{P}^{\text{BAL}}$  and  $Q \in \mathcal{Q}^{\text{CBAL}}$  was arbitrary, the proof is complete.  $\square$

To devise an efficient membership query procedure, we also need to upper bound, for any set  $P$ , the number of sets in  $\mathcal{F}$  that contain  $P$ . We consider any choice of  $P$  of size  $p' \leq p$  rather than just any choice of  $P$  of size exactly  $p$  as that is required for product counters.

**LEMMA 3.8.** *Let  $t, k, p, b \in \mathbb{N}$  with  $p \leq k$ ,  $0 < \epsilon < 1$  and  $c, d \geq 1$ . Let  $\bar{U} = (U_1, U_2, \dots, U_t)$  be a partitioned universe with  $U = \bigcup_{i=1}^t U_i$  of size  $n$ , and let  $(f, g)$  be a  $(t, k, p, b)$ -splitting function pair.*



With probability at least  $1 - \frac{1}{2c}$ , the output family  $\mathcal{F} \subseteq 2^U$  of  $(\bar{U}, b, f, g, \epsilon, c, d)$ -universal family sampling has the following property: For every  $g'$  be such that  $(f, g')$  is a  $(t, k, p', b)$ -splitting function pair (for some  $p' \leq p$ ) where  $g' \leq g$  and set  $P \in \mathcal{P}_{\bar{U}, f, g'}^{\text{BAL}}$ , we have that  $|\{F \in \mathcal{F} : P \subseteq F\}| \leq \left(\frac{d \cdot k}{d \cdot k - p}\right)^{k-p} \cdot \left(\frac{d \cdot k}{p}\right)^{p-p'} \cdot \left(\frac{1}{\ln^2(1+\epsilon)} \cdot 20k^3 \cdot \ln(nc)\right)^t$ .

*Proof.* Towards the proof of the lemma, we first show that the following claim is correct.

**CLAIM 2.** With probability at least  $1 - \frac{1}{2c}$ , for every  $i \in \{1, 2, \dots, t\}$ ,  $g'(i) \leq g(i)$  and  $P \in \binom{U_i}{g'(i)}$ , we have that  $|\{F \in \mathcal{F}_i : P \subseteq F\}| \leq \left(\frac{d \cdot f(i)}{d \cdot f(i) - g(i)}\right)^{f(i)-g(i)} \cdot \left(\frac{d \cdot f(i)}{g(i)}\right)^{g(i)-g'(i)} \cdot \frac{1}{\ln^2(1+\epsilon)} \cdot 20k^3 \cdot \ln(nc)$ .

*Proof.* Let  $E_i = \left(\frac{d \cdot f(i)}{d \cdot f(i) - g(i)}\right)^{f(i)-g(i)} \cdot \left(\frac{d \cdot f(i)}{g(i)}\right)^{g(i)-g'(i)} \cdot \frac{1}{\ln^2(1+\epsilon)} \cdot 10k^3 \cdot \ln(nc)$ . By union bound and because  $|\binom{U_i}{\leq p}| \leq n^k$ , it suffices to choose some  $i \in \{1, 2, \dots, t\}$ ,  $g'(i) \leq g(i)$  and  $P \in \binom{U_i}{g'(i)}$ , and prove that with failure probability at most  $\frac{1}{2ctn^k}$ , we have that  $|\{F \in \mathcal{F}_i : P \subseteq F\}| \leq E_i$ . To this end, observe that each set  $F_{i,j} \in \mathcal{F}_i$  contains  $P$  with probability  $\left(\frac{g(i)}{d \cdot f(i)}\right)^{g'(i)}$ . Thus, the expected number of sets in  $\mathcal{F}_i$  that contain  $P$  is  $E_i$ . Because the sets in  $\mathcal{F}_i$  are sampled independently from one another, by Chernoff bound (Proposition 2.2), we have that

$$\begin{aligned} \Pr(|\mathcal{F}_i[P, Q]| - E_i > E_i) &\leq 2e^{-\frac{E_i}{2}} \\ &\leq 2e^{-5k \cdot \ln(nc)} = \frac{2}{(nc)^{5k}} \leq \frac{1}{n^4 \cdot n^k \cdot c} \leq \frac{1}{2ct} \end{aligned}$$

Here, the last inequality follows since  $n \geq \max(2, t)$ . This completes the proof of the claim.  $\square$

We now return to the proof of the lemma. Due to Claim 1, to prove the lemma it suffices to show that, under the assumption that for every  $i \in \{1, 2, \dots, t\}$ ,  $g'(i) \leq g(i)$  and  $P \in \binom{U_i}{g'(i)}$ , we have that  $|\{F \in \mathcal{F}_i : P \subseteq F\}| \leq \left(\frac{d \cdot f(i)}{d \cdot f(i) - g(i)}\right)^{f(i)-g(i)} \cdot \left(\frac{d \cdot f(i)}{g(i)}\right)^{g(i)-g'(i)} \cdot \frac{1}{\ln^2(1+\epsilon)} \cdot 20k^3 \cdot \ln(nc)$ , it holds that for every  $g'$  be such that  $(f, g')$  is a  $(t, k, p', b)$ -splitting function pair where  $g' \leq g$  and set  $P \in \mathcal{P}_{\bar{U}, f, g'}^{\text{BAL}}$ , we have that  $|\{F \in \mathcal{F} : P \subseteq F\}| \leq \left(\frac{d \cdot k}{d \cdot k - p}\right)^{k-p} \cdot \left(\frac{d \cdot k}{p}\right)^{p-p'} \cdot \left(\frac{1}{\ln^2(1+\epsilon)} \cdot 20k^3 \cdot \ln(nc)\right)^t$ . Towards the proof of this, consider some set  $P \in \mathcal{P}_{\bar{U}, f, g'}^{\text{BAL}}$ .

Then,

$$|\{F \in \mathcal{F} : P \subseteq F\}| = \prod_{i=1}^t |\{F \in \mathcal{F}_i : P \cap U_i \subseteq F\}|.$$

Because  $P \in \mathcal{P}_{\bar{U}, f, g'}^{\text{BAL}}$ , it holds that for every  $i \in \{1, \dots, t\}$ ,  $P \cap U_i \in \binom{U_i}{g'(i)}$ , and therefore  $|\{F \in \mathcal{F}_i : P \cap U_i \subseteq F\}| \leq \left(\frac{d \cdot f(i)}{d \cdot f(i) - g(i)}\right)^{f(i)-g(i)} \cdot \left(\frac{d \cdot f(i)}{g(i)}\right)^{g(i)-g'(i)} \cdot \frac{1}{\ln^2(1+\epsilon)} \cdot 20k^3 \cdot \ln(nc)$ . Thus,

$$\begin{aligned} |\{F \in \mathcal{F} : P \subseteq F\}| &\leq \prod_{i=1}^t \left(\left(\frac{d \cdot f(i)}{d \cdot f(i) - g(i)}\right)^{f(i)-g(i)} \cdot \left(\frac{d \cdot f(i)}{g(i)}\right)^{g(i)-g'(i)} \cdot \frac{1}{\ln^2(1+\epsilon)} \cdot 20k^3 \cdot \ln(nc)\right) \\ &\leq \left(\prod_{i=1}^t \left(\frac{d \cdot f(i)}{d \cdot f(i) - g(i)}\right)^{f(i)-g(i)} \cdot \left(\frac{d \cdot f(i)}{g(i)}\right)^{g(i)-g'(i)}\right) \cdot \left(\frac{1}{\ln^2(1+\epsilon)} \cdot 20k^3 \cdot \ln(nc)\right)^t \end{aligned}$$

Recall that  $f : \{1, 2, \dots, t\} \rightarrow \{1, 2, \dots, \lceil bk/t \rceil\}$  and  $g' \leq g \leq f$  satisfy  $\sum_{i=1}^t f(i) = k$ ,  $\sum_{i=1}^t g(i) = p$  and  $\sum_{i=1}^t g'(i) = p'$ . Relaxing the supposition  $f : \{1, 2, \dots, t\} \rightarrow \{1, 2, \dots, \lceil bk/t \rceil\}$  to  $f : \{1, 2, \dots, t\} \rightarrow \{1, 2, \dots, k\}$ , the maximum of  $\prod_{i=1}^t \left(\frac{d \cdot f(i)}{d \cdot f(i) - g(i)}\right)^{f(i)-g(i)} \cdot \left(\frac{d \cdot f(i)}{g(i)}\right)^{g(i)-g'(i)}$  is attained when  $f(i) = k$ ,  $g(i) = p$  and  $g'(i) = p'$  for some  $i \in \{1, 2, \dots, t\}$ , and  $f(i') = g(i') = g'(i') = 0$  for all other  $i' \in \{1, 2, \dots, t\} \setminus \{i\}$ . Then, the value is  $\frac{(d \cdot k)^{k-p}}{(d \cdot k - p)^{k-p}} \cdot \left(\frac{d \cdot k}{p}\right)^{p-p'}$ . This completes the proof.  $\square$

The property in Lemma 3.7 together with the product-like manner in which we construct  $\mathcal{F}$  yields an efficient membership query procedure as follows.

**DEFINITION 12. (Membership Query Procedure)** Let  $t, k, p, b \in \mathbb{N}$  with  $p \leq k$ ,  $0 < \epsilon < 1$  and  $c, d \geq 1$ . Let  $\bar{U} = (U_1, U_2, \dots, U_t)$  be a partitioned universe with  $U = \bigcup_{i=1}^t U_i$  of size  $n$ . Let  $(f, g)$  be a  $(t, k, p, b)$ -splitting function pair. Let  $\mathcal{F} \subseteq 2^U$  be the output family of  $(\bar{U}, b, f, g, \epsilon, c, d)$ -universal family sampling. Then, the procedure **MEMBERSHIP** is defined as follows. Let  $\{\mathcal{F}_i\}_{i=1}^t$  be the collection of families sampled to construct  $\mathcal{F}$  (see Definition 11). Given  $g'$  such that  $(f, g')$  is a  $(t, k, p', b)$ -splitting function pair (for some  $p' \leq p$ ) where  $g' \leq g$  and  $P \in \mathcal{P}_{\bar{U}, f, g'}^{\text{BAL}}$ , **MEMBERSHIP**

naively computes  $\mathcal{F}'_i = \{F_{i,j_i} \in \mathcal{F}_i : P \cap U_i \subseteq F_{i,j_i}\}$  by iterating over every set in  $\mathcal{F}_i$ ; then, it outputs  $\{F_{1,j_1} \cup F_{2,j_2} \cup \dots \cup F_{t,j_t} : F_{1,j_1} \in \mathcal{F}'_1, F_{2,j_2} \in \mathcal{F}'_2, \dots, F_{t,j_t} \in \mathcal{F}'_t\}$ , computed using naive enumeration.

We now assert that our procedure is indeed an efficient membership query procedure as a corollary of Lemma 3.8. For lack of space, we omit the proof of this corollary.

**COROLLARY 3.2.** *Let  $t, k, p, b \in \mathbb{N}$  with  $p \leq k$ ,  $0 < \epsilon < 1$  and  $c, d \geq 1$ . Let  $\bar{U} = (U_1, U_2, \dots, U_t)$  be a partitioned universe with  $U = \bigcup_{i=1}^t U_i$  of size  $n$ . Let  $(f, g)$  be a  $(t, k, p, b)$ -splitting function pair. Let  $\mathcal{F} \subseteq 2^U$  be the output family of  $(\bar{U}, b, f, g, \epsilon, c, d)$ -universal family sampling. Then, with probability at least  $1 - \frac{1}{2c}$ , for every  $g'$  be such that  $(f, g')$  is a  $(t, k, p', b)$ -splitting function pair (for some  $p' \leq p$ ) where  $g' \leq g$ , the procedure MEMBERSHIP is a  $T$ -membership query procedure with respect to  $\mathcal{P}_{\bar{U}, f, g'}^{\text{BAL}}$  for*

$$T = \left( (d \cdot bk)^{bk/t} + \left( \frac{d \cdot k}{d \cdot k - p} \right)^{k-p} \cdot \left( \frac{d \cdot k}{p} \right)^{p-p'} \right) \cdot \left( \frac{1}{\ln^2(1+\epsilon)} \cdot 20k^3 \cdot \ln(nc) \right)^t.$$

By putting together Observation 3, Lemma 3.6, Lemma 3.7 and Corollary 3.2, we derive our main statement regarding the produced family  $\mathcal{F}$ .

**THEOREM 3.2.** *Let  $t, k, p, b \in \mathbb{N}$  with  $p \leq k$ ,  $0 < \epsilon < 1$  and  $c, d \geq 1$ . Let  $\bar{U} = (U_1, U_2, \dots, U_t)$  be a partitioned universe with  $U = \bigcup_{i=1}^t U_i$  of size  $n$ , and let  $(f, g)$  be a splitting function pair. With probability at least  $1 - \frac{1}{c}$ , the output family  $\mathcal{F} \subseteq 2^U$  of  $(\bar{U}, b, f, g, \epsilon, c, d)$ -universal family sampling, computed in time  $\mathcal{O}(|\mathcal{F}|n)$ , satisfies all of the following conditions.*

1.  $|\mathcal{F}| \leq \frac{(dk)^k}{p^p(dk-p)^{k-p}} \cdot \left( \frac{1}{\ln^2(1+\epsilon)} \cdot 10k^3 \cdot \ln(nc) \right)^t$ .
2.  $\mathcal{F}$  is an  $\epsilon$ -parsimonious  $(n, p, k-p)$ -universal family with respect to  $(\mathcal{P}_{\bar{U}, f, g}^{\text{BAL}}, \mathcal{Q}_{\bar{U}, f, g}^{\text{CBAL}})$ , whose correction factor is upper bounded by  $\left( \frac{1}{\ln^2(1+\epsilon)} \cdot 10k^3 \cdot \ln(nc) \right)^t$ .
3. With respect to  $\mathcal{F}$  and any  $g'$  be such that  $(f, g')$  is a  $(t, k, p', b)$ -splitting function pair (for some  $p' \leq p$ ) where  $g' \leq g$ , MEMBERSHIP is a  $T$ -membership query procedure with respect to  $\mathcal{P}_{\bar{U}, f, g'}^{\text{BAL}}$  for

$$T = \left( (d \cdot bk)^{bk/t} + \left( \frac{dk}{dk-p} \right)^{k-p} \left( \frac{dk}{p} \right)^{p-p'} \right)$$

$$\cdot \left( \frac{1}{\ln^2(1+\epsilon)} \cdot 20k^3 \cdot \ln(nc) \right)^t.$$

In particular, we will be interested in the case where  $b = 2$ ,  $\epsilon = \frac{\ln \frac{3}{5k^2}}{5k^2}$ ,  $p' = p$  and  $d = 1.447 = \mathcal{O}(1)$ ; later, we will run our entire process multiple times to enable having arbitrarily small error. Then,  $\left( \frac{1}{\ln^2(1+\epsilon)} \right)^t \leq (\epsilon - \epsilon^2/2)^t$  (by Taylor series), upper bounded by  $2^{\mathcal{O}(\sqrt{k} \log k)}$ . Further, we will choose  $c \geq n$  and  $t = \lceil \sqrt{k} \rceil$ . By these substitutions, we obtain the following corollary of Theorem 3.2.

**COROLLARY 3.3.** *Let  $k, p \in \mathbb{N}$  with  $p \leq k$ , and  $c \geq 1$ . Let  $\bar{U} = (U_1, U_2, \dots, U_{\lceil \sqrt{k} \rceil})$  be a partitioned universe with  $U = \bigcup_{i=1}^{\lceil \sqrt{k} \rceil} U_i$  of size  $n \leq c$ , and let  $(f, g)$  be a splitting function pair. With probability at least  $1 - \frac{1}{c}$ , the output family  $\mathcal{F} \subseteq 2^U$  of  $(\bar{U}, 2, f, g, \frac{\ln \frac{3}{5k^2}}{5k^2}, c, 1.447)$ -universal family sampling, computed in time  $\mathcal{O}(|\mathcal{F}|n)$ , satisfies all of the following conditions.*

1.  $|\mathcal{F}| \leq \frac{(1.447k)^k}{p^p(1.447k-p)^{k-p}} \cdot 2^{\mathcal{O}(\sqrt{k} \log k)} \cdot \log^{\sqrt{k}} c$ .
2.  $\mathcal{F}$  is a  $\frac{\ln \frac{3}{5k^2}}{5k^2}$ -parsimonious  $(n, p, k-p)$ -universal family with respect to  $(\mathcal{P}^{\text{BAL}}, \mathcal{Q}^{\text{CBAL}})$ .
3. With respect to  $\mathcal{F}$ , MEMBERSHIP is a  $T$ -membership query procedure for

$$T = \left( \frac{1.447k}{1.447k-p} \right)^{k-p} \cdot 2^{\mathcal{O}(\sqrt{k} \log k)} \cdot \log^{\sqrt{k}} c.$$

### 3.3 Reducing a Problem to Its Split Version

Because we only deal with balancedly split sets, we now develop a simple procedure whose employment will allow us to reduce the general case to one focused only on balancedly split sets. To this end, we need the following definition.

**DEFINITION 13. (Balancedly Split Sets II)** *Let  $t, k, b \in \mathbb{N}$ . Let  $\bar{U} = (U_1, U_2, \dots, U_t)$  be a partitioned universe with  $U = \bigcup_{i=1}^t U_i$ . Then,  $P \in \binom{U}{k}$  is  $(\bar{U}, k, b)$ -balancedly split if for every  $i \in \{1, 2, \dots, t\}$ , it holds that  $|P \cap U_i| \leq \lceil bk/t \rceil$ .*

We now present the procedure.

**LEMMA 3.9.** *Given  $t, k, b \in \mathbb{N}$  and  $c \geq 1$ , a universe  $U$  of size  $n$ , and  $0 < \delta < 1$  with  $b^2 \frac{k}{2t} \geq \ln(4t)$ , a collection  $\mathcal{U}$  of  $\frac{4}{\delta^2} k \ln(2nc)$   $t$ -partitioned universes over  $U$  such that the following property holds with probability at least  $1 - \frac{1}{c}$  (resp. 1) can be computed in time  $\mathcal{O}(n \frac{1}{\delta^2} k \ln(nc))$ : for every set  $P \in \binom{U}{k}$ , the (resp. expected) number of*

partitioned universes  $\bar{U} \in \mathcal{U}$  such that  $P$  is  $(\bar{U}, k, b)$ -balancedly split is between  $(1 - \delta)X$  and  $(1 + \delta)X$  (resp. exactly  $X$ ) for some  $X = X(n, k, t, b, \delta) > 0$ . (We note that  $X$  can be computed in time  $\mathcal{O}(|\mathcal{U}| \cdot (\frac{2bk}{t})^t)$ ).

*Proof.* Denote  $r = \frac{4}{\delta^2} k \ln(2nc)$ . Given the input  $t, k, c, U, b, \delta$ , the algorithm constructs  $\mathcal{U} = \{\bar{U}_1, \bar{U}_2, \dots, \bar{U}_r\}$  as follows. For  $i = 1, 2, \dots, r$ , the partitioned universe  $\bar{U}_i = (U_{i,1}, U_{i,2}, \dots, U_{i,t})$  is constructed as follows. Each element  $u \in U$  is inserted into exactly one part  $U_{i,j}$  where the choices of  $j \in \{1, 2, \dots, t\}$  are made independently and uniformly at random. Clearly, the time complexity of the algorithm is  $\mathcal{O}(nr)$ .

Let  $X$  denote the expected number of partitioned universes  $\bar{U} \in \mathcal{U}$  such that any set  $P \in \binom{U}{k}$  is  $(\bar{U}, k, b)$ -balancedly split. Note that  $X$  is the same for all sets  $P \in \binom{U}{k}$ , thus it is well defined. The exact value of  $X$  will be calculated later.

Now, arbitrarily choose some set  $P \in \binom{U}{k}$ . Additionally, consider some  $i \in \{1, 2, \dots, r\}$ . Notice that for any  $j \in \{1, 2, \dots, t\}$ , the expected number of elements in  $P$  contained in  $U_{i,j}$  is  $k/t$ , therefore Chernoff bound (Proposition 2.2) implies that the probability that the number of elements in  $P$  contained in  $U_{i,j}$  is not upper bounded by  $\lceil bk/t \rceil$  is at most  $2e^{-\frac{b^2(k/t)}{2}} \leq 2e^{-\ln(4t)} = \frac{1}{2t}$  where the inequality follows from the supposition  $b^2 \frac{k}{2t} \geq \ln(4t)$  in the lemma. Then, by union bound, the probability that  $P$  is not  $(\bar{U}, k, b)$ -balancedly split is at most  $t \cdot \frac{1}{2t} = \frac{1}{2}$ , hence the probability that it is  $(\bar{U}, k, b)$ -balancedly split is at least  $\frac{1}{2}$ . Therefore,  $X \geq \frac{r}{2}$ . In turn, by Chernoff bound (Proposition 2.2) and this lower bound on  $X$ , the probability that the number of partitioned universes  $\bar{U} \in \mathcal{U}$  such that  $P$  is  $(\bar{U}, k, b)$ -balancedly split is not between  $(1 - \delta)X$  and  $(1 + \delta)X$  is at most  $2e^{-\frac{\delta^2 X}{2}} \leq 2e^{-\frac{\delta^2 r}{4}} = 2e^{-k \ln(2nc)} = \frac{2}{(2nc)^k} \leq \frac{1}{n^k c}$ .

Since the choice of  $P \in \binom{U}{k}$  was arbitrary and by union bound, the probability that there exists  $P \in \binom{U}{k}$  such that the number of partitioned universes  $\bar{U} \in \mathcal{U}$  such that  $P$  is  $(\bar{U}, k, \epsilon)$ -balancedly split is not between  $(1 - \delta)X$  and  $(1 + \delta)X$  is upper bounded by  $\binom{n}{k} \cdot \frac{1}{n^k c} \leq \frac{1}{c}$ . Thus, with probability at least  $1 - \frac{1}{c}$ , for every set  $P \in \binom{U}{k}$  the number of partitioned universes  $\bar{U} \in \mathcal{U}$  such that  $P$  is  $(\bar{U}, k, b)$ -balancedly split is between  $(1 - \delta)X$  and  $(1 + \delta)X$ .

It remains to calculate  $X$ . To this end, arbitrarily choose some set  $P \in \binom{U}{k}$  and  $i \in \{1, 2, \dots, r\}$ . Clearly,  $X = r \cdot Y$ , where  $Y$  is the probability that  $P$  is  $(\bar{U}_i, k, b)$ -balancedly split. Now, observe that  $Y = \sum_{\substack{\ell_1, \ell_2, \dots, \ell_t \in \{1, 2, \dots, \lceil bk/t \rceil\} \\ \text{s.t. } \sum_{j=1}^t \ell_j = k}} \binom{k}{\ell_1} \binom{k - \ell_1}{\ell_2} \dots \binom{k - \sum_{j=1}^{t-1} \ell_j}{\ell_t} \cdot (1/t)^k$ . This completes the proof.  $\square$

We now present the our main utility of this procedure, which is a reduction of a problem to a “split” version of itself. To this end, we first define the notion of a split version of a problem.

**DEFINITION 14. (Splittable Problem)** Let  $\Pi$  be a problem whose input consists, among possibly other components, of a universe  $U$  of size  $n$  and  $k \in \mathbb{N}$ , and whose solutions are subsets (resp. ordered subsets) of  $U$  of size  $k$ . Such a problem  $\Pi$  is said to be splittable. Then, the general split version of  $\Pi$  is defined as follows. Its input consists of the same components as the input of  $\Pi$ , and in addition, of a  $t$ -partitioned universe  $\bar{U}$  for some  $t \in \mathbb{N}$ ,  $b \in \mathbb{N}$  and a  $(t, k, b)$ -splitting function  $f$ , and whose solutions are all the subsets (resp. ordered subsets) of  $U$  that are both solutions of  $\Pi$  and are  $(\bar{U}, f)$ -balancedly split. When  $t = \sqrt{k}$  and  $b = 2$ , the general split version is called the split version in short.

Next, we present the reduction.

**LEMMA 3.10.** Let  $\Pi$  be a splittable problem such that the number of solutions of the general split version of  $\Pi$  can be approximately counted with multiplicative error  $(1 \pm \alpha)$  (resp. and the expectation equals the exact number of solutions) in time  $T = T(\alpha, t, b)$  (where  $t, b$  are input to the split version) and with success probability at least  $1 - \frac{1}{c}$ . Then, for any  $c \in \mathbb{N}$  such that  $(2bk/t)^t \cdot \frac{1}{\beta^2} k \ln(nc) \cdot \frac{1}{c} \leq \frac{1}{2c}$  and  $0 < \beta < 1$ , the number of solutions of  $\Pi$  can be approximately counted with multiplicative error  $(1 \pm \alpha)(1 \pm \beta)$  (resp. and the expectation equals the exact number of solutions) in time  $\mathcal{O}(((2bk/t)^t \cdot T + n) \cdot \frac{1}{\beta^2} k \ln(nc))$  where  $b^2 \frac{k}{2t} \geq \ln(4t)$  and with success probability at least  $1 - \frac{1}{c}$ .

*Proof.* Let ALG1 be the algorithm supposed to approximately count solutions of the general split version of  $\Pi$  with multiplicative error  $(1 \pm \alpha)$  where the expectation equals the exact number of solutions in time  $T$  and with success probability at least  $1 - \frac{1}{c}$ . We remark that if the condition regarding the expectation is not assumed to hold, then disregard the arguments below concerning its satisfaction for the output. Then, we design an algorithm ALG2 as follows. Given an instance  $I$  of  $\Pi$ ,  $c \in \mathbb{N}$  and  $0 < \beta < 1$ , ALG2 executes the following operations.

1. Use the algorithm in Lemma 3.9 to compute a collection  $\mathcal{U}$  of  $\frac{4}{\beta^2} k \ln(4nc)$   $t$ -partitioned universes over  $U$  such that the following property holds with probability at least  $1 - \frac{1}{2c}$  (resp. 1): for every set  $P \in \binom{U}{k}$ , the (resp. expected) number of partitioned universes  $\bar{U} \in \mathcal{U}$  such that  $P$  is  $(\bar{U}, k, b)$ -balancedly split is between  $(1 - \beta)X$  and  $(1 + \beta)X$  (resp. exactly  $X$ ) for some  $X = X(n, k, t, b, \beta) > 0$ .

2. Let  $\mathcal{F}$  be the family of all  $(t, k, b)$ -splitting functions.
3. For every partitioned universe  $\bar{\mathbf{U}} \in \mathcal{U}$ :
  - (a) For every  $f \in \mathcal{F}$ :
    - i. Run **ALG1** on  $(I, \bar{\mathbf{U}}, b, f)$  as input, and denote its output by  $O_{\bar{\mathbf{U}}, f}$ .
  - (b) Let  $O_{\bar{\mathbf{U}}} = \sum_{f \in \mathcal{F}} O_{\bar{\mathbf{U}}, f}$ .
4. Output  $O = \frac{1}{X} \cdot \sum_{\bar{\mathbf{U}} \in \mathcal{U}} O_{\bar{\mathbf{U}}}$ .

By Lemma 3.9, Step 1 is performed in time  $\mathcal{O}(n^{\frac{1}{\beta^2}} k \ln(nc))$ . Now, observe that  $|\mathcal{F}| \leq (\lceil bk/t \rceil + 1)^t = \mathcal{O}((2bk/t)^t)$ . Thus, we perform Step 3(a)i  $|\mathcal{U}| \cdot |\mathcal{F}| = \mathcal{O}(\frac{1}{\beta^2} k \ln(nc) \cdot (2bk/t)^t)$  times, where each single performance is done in time  $\mathcal{O}(T)$ . Thus, the total running time is indeed  $\mathcal{O}(((2bk/t)^t \cdot T + n) \cdot \frac{1}{\beta^2} k \ln(nc))$ .

By union bound, with probability at least  $1 - |\mathcal{U}| |\mathcal{F}| \cdot \frac{1}{c^t} - \frac{1}{2c}$ , which is lower bounded by  $1 - (2bk/t)^t \cdot \frac{1}{\beta^2} k \ln(nc) \cdot \frac{1}{c^t} - \frac{1}{2c} \geq 1 - \frac{1}{c}$ , the call to the algorithm in Lemma 3.9 as well as all calls to **ALG2** are successful. Thus, to prove the lemma, it suffices to prove that  $E[O]$  is the exact number of solutions, and that under the aforementioned condition (of all calls being successful), the number of solutions of  $\Pi$  is necessarily approximated by  $O$  with multiplicative error  $(1 \pm \alpha)(1 \pm \beta)$ .

First, observe that for any  $\bar{\mathbf{U}} \in \mathcal{U}$ , the number of  $(\bar{\mathbf{U}}, k, b)$ -balancedly split solutions is exactly the sum over all  $f \in \mathcal{F}$  of the number of  $(\bar{\mathbf{U}}, f)$ -balancedly split solutions. Thus, because the approximation factor of **ALG2** is  $(1 \pm \alpha)$  and the expectation is exact, we have that for any  $\bar{\mathbf{U}} \in \mathcal{U}$ , the number of  $(\bar{\mathbf{U}}, k, b)$ -balancedly split solutions is exactly  $E[O_{\bar{\mathbf{U}}}]$ , and (under the aforementioned condition) it is approximated by  $O_{\bar{\mathbf{U}}}$  with multiplicative error  $(1 \pm \alpha)$ . Now, recall that for every set  $P \in \binom{U}{k}$  (and, in particular, for every solution of  $\Pi$ ), the number of partitioned universes  $\bar{\mathbf{U}} \in \mathcal{U}$  such that  $P$  is  $(\bar{\mathbf{U}}, k, b)$ -balancedly split is in expectation  $X$ , and (under the aforementioned condition) it is between  $(1 - \beta)X$  and  $(1 + \beta)X$ . Since  $O = \frac{1}{X} \cdot \sum_{\bar{\mathbf{U}} \in \mathcal{U}} O_{\bar{\mathbf{U}}}$ , we

conclude that indeed the number of solutions of  $\Pi$  is  $E[O]$ , and that (under the aforementioned condition) it is necessarily approximated by  $O$  with multiplicative error  $(1 \pm \alpha)(1 \pm \beta)$ .  $\square$

For the (non-general) split version and  $\alpha = \beta = \frac{1}{2}$ , in which we will be specifically interested, we obtain the following corollary.

**COROLLARY 3.4.** *Let  $\Pi$  be a splittable problem such that the number of solutions of the split version of  $\Pi$  can be approximately counted with multiplicative error  $(1 \pm \frac{1}{2})$  where the expectation equals the exact number of solutions in time  $T$  and with success probability at least  $1 - \frac{1}{c}$ . Then, for any  $c \in \mathbb{N}$  such that  $4k(4\sqrt{k})^{\sqrt{k}} \ln(nc) \cdot \frac{1}{c^t} \leq \frac{1}{c}$ , the number of solutions of  $\Pi$  can be approximately counted with multiplicative error between  $\frac{1}{4}$  and  $2\frac{1}{4}$  where the expectation equals the exact number of solutions in time  $\mathcal{O}((2^{\mathcal{O}(\sqrt{k} \log k)} \cdot T + n) \cdot k \ln(nc))$  and with success probability at least  $1 - \frac{1}{c}$ .*

Lastly, we give a lemma that can be considered folklore (but whose proof is given for completeness), whose utility is to enable us to focus on achieving some small constant multiplicative error for a counting problem, as this can be boosted to an arbitrarily small error as follows.

**LEMMA 3.11.** *Let  $\Pi$  be a problem that admits a randomized algorithm that, given an instance of  $\Pi$  whose number of solutions is  $X$ , returns a number  $Y$  such that  $E[Y] = X$  and  $\alpha X \leq Y \leq \beta X$  for some  $0 < \alpha \leq 1$  and  $\beta \geq 1$  in time  $T$  with success probability  $1 - \frac{1}{c^t}$ . Then, for any  $0 < \epsilon < 1$  and  $c \geq 1$  such that  $\frac{t}{c^t} \leq \frac{1}{2c}$  where  $t = \frac{2\beta}{\epsilon^2} \lceil \ln(4c) \rceil$ ,  $\Pi$  also admits an algorithm that, given an instance of  $\Pi$  whose number of solutions is  $X$ , returns a number  $Z$  such that  $(1 - \epsilon)X \leq Z \leq (1 + \epsilon)X$  in time  $\mathcal{O}(\frac{\beta}{\epsilon^2} \log c \cdot T)$  with success probability at least  $1 - \frac{1}{c}$ .*

*Proof.* Let **ALG1** denote the algorithm given in the supposition of the lemma. Let  $0 < \epsilon < 1$ . Then, we design an algorithm **ALG2** as follows. Given an instance  $I$  of  $\Pi$ , **ALG2** executes the following operations.

1. For  $i = 1, 2, \dots, t$ : Call **ALG1** with  $I$  as input and let  $Y_i$  denote the result.
2. Output  $Z = \frac{1}{t} \cdot \sum_{i=1}^t Y_i$ .

First, notice that the time complexity of **ALG2** is  $\mathcal{O}(t \cdot T) = \mathcal{O}(\frac{\beta}{\epsilon^2} \log c \cdot T)$ . Second, by union bound, with success probability at least  $1 - \frac{t}{c^t} \geq 1 - \frac{1}{2c}$ , all the calls it makes to **ALG1** are successful. Thus, by union bound, to prove the lemma, it suffices to prove that under the assumption that all the calls made to **ALG1** are successful, with probability at least  $1 - \frac{1}{2c}$ , it holds that  $(1 - \epsilon)X \leq Z \leq (1 + \epsilon)X$ .

For all  $i \in \{1, 2, \dots, t\}$ , denote  $Y'_i = \frac{Y_i}{\beta X}$ . Moreover, denote  $Z' = \sum_{i=1}^t Y'_i$ . Notice that  $(1 - \epsilon)X \leq Z \leq (1 + \epsilon)X$  if and only if  $(1 - \epsilon)\frac{t}{\beta} \leq Z' \leq (1 + \epsilon)\frac{t}{\beta}$ , and thus it suffices to consider the probability that the latter event occurs. Since all calls are assumed to be successful,



we have that  $0 \leq Y'_i \leq 1$ . Moreover, by linearity of expectation,  $E[Z'] = \sum_{i=1}^t E[Y'_i] = \sum_{i=1}^t \frac{E[Y_i]}{\beta X} = t/\beta$ . Therefore,  $(1 - \epsilon)\frac{t}{\beta} \leq Z' \leq (1 + \epsilon)\frac{t}{\beta}$  if and only if  $|Z' - E[Z']| \leq \epsilon E[Z']$ , and thus it further suffices to consider the probability that the latter event occurs. By Chernoff Bound (Proposition 2.2), we have that

$$\begin{aligned} \Pr(|Z' - E[Z']| > \epsilon E[Z']) &\leq 2e^{-\frac{\epsilon^2 E[Z']}{2}} \\ &= 2e^{-\frac{\epsilon^2 t}{2\beta}} \\ &= 2e^{-\ln(4c)} = \frac{1}{2c}. \end{aligned}$$

Thus,  $|Z' - E[Z']| \leq \epsilon E[Z']$  with probability at least  $1 - \frac{1}{2c}$ . As claimed above, this completes the proof.  $\square$

Combining Corollary 3.4 and Lemma 3.11, we have the following read-to-use corollary. We did not make any attempt to optimize the lower bound on  $c'$ , but just give a short expression. Clearly, the success probability can be boosted to any constant close to 1. To simplify notation, we will work with  $\frac{9}{10}$ .

**COROLLARY 3.5.** *Let  $\Pi$  be a splittable problem such that the number of solutions of the split version of  $\Pi$  can be approximated with multiplicative error  $(1 \pm \frac{1}{2})$  in time  $T \geq n$  where the expectation equals the exact number of solutions, and with success probability at least  $1 - \frac{1}{c'}$ . Then, for any  $0 < \epsilon < 1$  such that  $c' \geq \frac{1}{\epsilon^2} \cdot (1000\sqrt{k})^{\sqrt{k}} \cdot \ln(n\frac{1}{\epsilon})$ , the number of solutions of  $\Pi$  can be approximated with multiplicative error  $(1 \pm \epsilon)$  in time  $2^{\mathcal{O}(\sqrt{k} \log k)} \cdot T \cdot \frac{1}{\epsilon^2} (\log n + \log \frac{1}{\epsilon})$  and with success probability at least  $\frac{9}{10}$ .*

*Proof.* Denote  $c'' = 2ct$  where  $c = 10$ ,  $\alpha = \frac{1}{4}$ ,  $\beta = 2\frac{1}{4}$  and  $t = \frac{2\beta}{\epsilon^2} \ln(4c)$ . Then,  $4k(4\sqrt{k})^{\sqrt{k}} \ln(nc'') \cdot \frac{1}{c'} \leq \frac{1}{c''}$ . Thus, by Corollary 3.4, the number of solutions of  $\Pi$  can be approximately counted with multiplicative error between  $\frac{1}{4}$  and  $2\frac{1}{4}$  in time  $T' = \mathcal{O}((2^{\mathcal{O}(\sqrt{k} \log k)} \cdot T + n) \cdot k \ln(nc''))$  and with success probability at least  $1 - \frac{1}{c''}$ . Therefore, by Lemma 3.11, the number of solutions of  $\Pi$  can be approximated with multiplicative error  $(1 \pm \epsilon)$  in time  $\mathcal{O}(\frac{\beta}{\epsilon^2} \log c \cdot T') = 2^{\mathcal{O}(\sqrt{k} \log k)} \cdot T \cdot \frac{1}{\epsilon^2} (\log n + \log \frac{1}{\epsilon})$  and with success probability at least  $\frac{9}{10}$ .  $\square$

## 4 Product Functions and Applications

Due to lack of space, these details are relegated to the full version of the papers. Here, we only briefly discuss them.

**4.1 Extension to Product Functions.** The computation of a representative function for a product function is technically involved. Among the main difficulties

being faced here is the fact that we cannot even iterate over the support of the input product function (since that in itself is too costly) and decide for each set in the support whether to insert it to the support of the output function (with some probability and new assigned value). Instead, we pre-determine how many sets to pick up, and devise a somewhat complex mechanism that allows us to efficiently sample sets from the support according to some distribution without ever computing the support! In particular, we now have two approximately parsimonious families rather than one (where one is meant to separate between sets in  $\mathcal{P}$  and sets in  $\mathcal{Q}$ , and the other is meant to separate between sets in  $\mathcal{P}_1$  and sets in  $\mathcal{P}_2$ ), and the sampling is done in three stages after some critical preprocessing to efficiently determine (in part) the probability distributions used in these stages. The first stage involves sampling a set  $P_1$  from the support of  $\mathfrak{C}_1$ , the second (which depends on the outcome of the first) involves sampling a pair of sets from our approximately parsimonious families, and the third (which depends on the outcome of the first and second) involves sampling a set  $P_2$  disjoint from  $P_1$  from the support of  $\mathfrak{C}_2$ , so as to pick up  $P_1 \cup P_2$ . We defer further technical details on the extension to product functions to the full version.

**4.2 Applications** Our algorithm for #MULTILINEAR DETECTION on skewed circuits is based on dynamic programming over the nodes of the input circuit. For each node, we store a counter that assigns to each monomial (encoded by the set containing its variables) of the polynomial of the subcircuit rooted at the current node its coefficient with “small error”. (More precisely, for each node together with a combination of other arguments, we store one such counter, but for the sake of simplicity of this overview, we ignore these other arguments here.) When we consider a node, we have already computed the aforementioned counters for all its outgoing neighbours. So, as the circuit is skewed, we can explicitly compute the counter for the current node, and then compute a representative counter for it and store the representative counter instead of it (else, even though the circuit is skewed, after several levels just writing the polynomial via a counter explicitly may take time  $\binom{n}{k}$ ). When we reach the root, we can solve the problem.

On general (monotone) circuits we cannot write the polynomial (and hence the counter) of a node that results from the multiplication of the polynomials stored for its outgoing neighbours (within the desired time complexity) even after their sizes have already been reduced by representation. So, instead, here we use our computation for product counters that sidesteps

this. Having attained algorithms for #MULTILINEAR DETECTION on skewed and general circuits, all our other applications, including the algorithm for # $k$ -PATH, follow just by using reductions known in the literature and observing that they are parsimonious.

## 5 Conclusion and Open Problems

In this paper, we presented a general tool to design FPT-approximation schemes for counting problems. Specifically, we introduced the notion of a representative function where our main contribution is a novel sampling procedure to compute representative functions of small support efficiently. Along the way, we developed a data structure to efficiently query membership and disjointness in approximately universal families, which is of independent interest. We have demonstrated the wide applicability of our tool by developing a  $\mathcal{O}((2.619^k + |I|^{o(1)}) \cdot \frac{1}{\epsilon^2} \cdot |I|)$ -time algorithm for #MULTILINEAR MONOMIAL DETECTION on skewed circuits, # $k$ -PATH, and several other problems as well (including # $q$ -SET  $p$ -PACKING with  $k = qp$ , # $q$ -DIMENSIONAL  $p$ -MATCHING with  $k = (q - 1)p$ , # GRAPH MOTIF, and #SUBGRAPH ISOMORPHISM for pattern graphs of constant treewidth). Additionally, we developed a  $\mathcal{O}((3.841^k + |I|^{o(1)}) \cdot \frac{1}{\epsilon^6} \cdot |I|)$ -time algorithm for #MULTILINEAR MONOMIAL DETECTION on general (monotone) circuits.

We conclude our paper with a few open problems.

- Does the # $k$ -PATH problem admit an FPT-approximation scheme with running time  $2^k(\frac{1}{\epsilon})^{\mathcal{O}(1)}n^{\mathcal{O}(1)}$ ?
- Does the #MULTILINEAR MONOMIAL DETECTION problem admit an FPT-approximation scheme with running time substantially better than  $3.841^k(\frac{1}{\epsilon})^{\mathcal{O}(1)}n^{\mathcal{O}(1)}$ ? In particular, can the time bound  $2.619^k(\frac{1}{\epsilon})^{\mathcal{O}(1)}n^{\mathcal{O}(1)}$  given for # $k$ -PATH be matched?
- Can our result for the #MULTILINEAR MONOMIAL DETECTION problem be extended to non-monotone arithmetic circuits (where subtraction is allowed)?
- Are there relations between techniques based on exterior algebra, Hadamard product, warring rank and representative functions?
- Can we compute representative functions efficiently with respect to linear matroids rather than only set systems (i.e., uniform matroids)? For more information on representation with respect to a matroid, we refer to [FLPS16].
- We remark that results on bounded skewness by themselves may be of interest in this context. Can we derive a general theorem about the problems

that admits them? What can be said in this context on VP-circuits and homomorphism polynomials?

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