1	Delayed effects of climate on vital rates lead to demographic
2	divergence in Amazonian forest fragments
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17 Abstract

18 Deforestation is a major threat to biodiversity in the Amazon, partly because it leaves 19 remaining forest habitat highly fragmented, with remnants of different sizes embedded in 20 an often highly contrasting matrix. Local extinction of species from individual fragments is 21 common, but the demographic mechanisms underlying these extinctions are poorly 22 understood. It is often hypothesized that the altered environmental conditions in 23 fragments drive declines in reproduction, recruitment, or survivorship. The Amazon basin, 24 in addition to experiencing continuing fragmentation, is warming and experiencing 25 changes in precipitation leading to altered frequency and intensity of droughts and 26 unusually wet periods. Whether plant populations in tropical forest fragments are 27 particularly susceptible to extremes in precipitation remains unclear. Most studies of 28 plants in fragments are relatively short (1-6 years), focus on a single life-history stage, and 29 often do not compare to populations in continuous forest. Even fewer studies consider 30 delayed effects of climate despite the importance of such delayed effects for plant 31 demographic vital rates in studies that look for them. Using a decade of demographic and 32 climate data from an experimentally fragmented landscape in the Central Amazon, we 33 assess the effects of climate on populations of an understory herb (*Heliconia acuminata*, Heliconiaceae). We used distributed lag non-linear models to understand the delayed 34 35 effects of temperature and precipitation on survival, growth, and flowering. We detected 36 delayed effects of climate up to 36 months. Drought two dry seasons prior to the February 37 census decreased survival and increased flowering probability while drought in the wet 38 season a year prior to the census decreased flowering probability and increased growth. 39 The effects of extremes in precipitation on survival and growth were more intense in forest

- 40 fragments compared to continuous forest. The complex delayed effects of climate and
- 41 habitat fragmentation on *H. acuminata* vital rates points to the importance of long-term
- 42 demography experiments in understanding the effects of anthropogenic change on plant
- 43 populations.
- 44

45 Introduction

46 The expansion of agriculture and other human activities is a primary driver of 47 deforestation in the tropics (Alroy, 2017; Haddad et al., 2015). It also results in landscapes 48 where the remaining forest can be highly fragmented, with patches of different sizes 49 embedded in a matrix of often contrasting habitat (Bianchi & Haig, 2013; Taubert et al., 50 2018). This fragmentation is associated with myriad ecological changes, including the local 51 and regional extinction of plant species (da Silva & Tabarelli, 2000; Laurance et al., 2006). 52 Although the demographic mechanisms responsible these extinctions are poorly 53 understood (Bruna et al., 2009), it is often hypothesized that the dramatically altered 54 environmental conditions in tropical forest fragments (Arroyo-Rodríguez et al., 2017; 55 Didham & Lawton, 1999; Ewers & Banks-Leite, 2013) drive declines in plant reproduction, 56 recruitment, or survivorship (Bruna, 1999; Laurance et al., 1998; Zartman et al., 2015). 57 Despite the prevalence of this hypothesis (Betts et al., 2019; Didham & Lawton, 1999; 58 Laurance et al., 2001), efforts to link population-level demographic responses with altered 59 environmental conditions in fragments remains scarce.

60 Studies in temperate systems have shown that the demography of species can also be 61 altered by climate change (Doak & Morris, 2010; Selwood et al., 2015; Sletvold, 2005; 62 Williams et al., 2015), and that the effects of climate change could be more pronounced 63 when habitat is fragmented (Holyoak & Heath, 2016; Oliver et al., 2015). While the 64 demographic consequences of climate change for tropical species are expected to be 65 similarly severe (Brodie et al., 2012; Scheffers et al., 2017), surprisingly little is known 66 about the responses of these species to climatic variability (Paniw et al., 2021). Tropical 67 plants may be particularly sensitive to climate change—they typically have narrow ranges of climatic tolerance (Feeley et al., 2012), and recent results suggest increases in the
frequency and severity of extreme precipitation events reduce survival and reproduction
(Esteban et al., 2021; Gaoue et al., 2019). This sensitivity to climatic fluctuations, coupled
with evidence that plant growth and survivorship are lower in fragments (Bruna et al.,
2002; Laurance et al., 1998; Zartman et al., 2015), has led to speculation that plants in
forest fragments will be especially susceptible to climate change (Laurance et al., 2001;
Opdam & Wascher, 2004; Selwood et al., 2015).

75 Whether the demography of plant populations in tropical forest fragments is more 76 susceptible to climatic extremes remains unclear for three primary reasons. First, most 77 studies of plants in fragments have focused on a single life-history stage or process (Bruna 78 et al., 2009: Ehrlen et al., 2016), making it challenging to draw broader demographic 79 conclusions. Second, there is a growing literature on how tropical plants respond to 80 droughts (Esquivel-Muelbert et al., 2019; González-M et al., 2020; Uriarte et al., 2016), but 81 few studies have compared the responses of plants in continuous forest with those of 82 plants in forest fragments (Laurance et al., 2001). Finally, the multi-year data needed to 83 test population-level hypotheses about climate change and fragmentation are scant, 84 especially for tropical systems (Crone et al., 2011; Salguero-Gómez et al., 2015). These data 85 are critical not simply because they allow for capturing variation in climatic conditions and 86 the resulting demographic responses (Morris & Doak, 2002; Teller et al., 2016). They are 87 also essential because while some demographic effects of fragmentation or drought can be 88 detected immediately, others may take years to manifest (e.g., Gagnon et al., 2011). Indeed, 89 lagged responses of demographic vital rates to climate may be the rule rather than the

90 exception (Anderegg et al., 2015; Evers et al., 2021; Kannenberg et al., 2020; Schwalm et al.,
91 2017).

92 Herbaceous plants represent up to 25% of plant diversity in tropical forests (Gentry & 93 Dodson, 1987), are critical food and habitat for myriad species (Snow, 1981), and are 94 economically and culturally vital (Nakazono et al., 2004; Ticktin, 2003). Nevertheless, the 95 impacts of global change phenomena on their demography remain conspicuously 96 understudied (Bruna et al., 2009). We used a decade of demographic and climatic data from 97 an experimentally fragmented landscape in the Central Amazon to assess the effects of 98 climate on populations of a tropical understory herb (Heliconia acuminata, Heliconiaceae). 99 This time series, which included the severe droughts of 1997 (McPhaden, 1999) and 2005 100 (Marengo et al., 2008; Zeng et al., 2008), allowed us to address the following questions: (1) 101 Does drought increase or decrease the growth, survival, and fertility of plant populations in 102 continuous forest? (2) Are there delayed effects of drought on demographic vital rates, and 103 if so what lag times are most critical? (3) Are the effects of drought on the vital rates of 104 populations in fragments similar in direction and magnitude to those in continuous forest?

105 Methods

106 Study site

The Biological Dynamics of Forest Fragments Project (BDFFP) is located ~70 km north of
Manaus, Brazil (2º30' S, 60ºW). In addition to large areas of continuous forest, the BDFFP
has forest fragment reserves isolated from 1980–1984 by felling the trees surrounding the
area chosen for isolation and, in most cases, burning the downed trees once they dried
(Bierregaard et al., 1992). In subsequent decades the vegetation regenerating around

112 fragments has been periodically cleared to ensure fragment isolation (Bierregaard et al.,113 2001).

The BDFFP reserves are located in nonflooded (i.e., *terra firme*) tropical lowland forest
with a 30–37m tall canopy (Rankin-de-Mérona et al., 1992) and an understory dominated
by stemless palms (Scariot, 1999). The soils in the reserves are nutrient-poor xanthic
ferrosols; their water retention capacity is poor despite having a high clay content. Mean
annual temperature in the region is 26° C (range=19–39° C), and annual rainfall ranges
from 1900–2300 mm. There is a pronounced dry season from June to October (Figure S1).

120 Focal species

121 *Heliconia acuminata* (LC Rich.) (Heliconiaceae) is a perennial monocot distributed 122 throughout Central Amazonia (Kress, 1990) and is the most abundant understory herb at 123 the BDFFP (Ribeiro et al., 2010). While many *Heliconia* species grow in large patches in 124 treefall gaps and other disturbed areas, others, such as *H. acuminata*, are found at lower 125 densities in the darker and cooler forest understory (Rundel et al., 2020). These species 126 produce fewer inflorescences and are pollinated by traplining rather than territorial 127 hummingbirds (Bruna et al., 2004; Stouffer & Bierregaard, 1996). In our sites *H. acuminata* 128 is pollinated by *Phaeothornis superciliosus* and *P. bourcieri* (Bruna et al., 2004). Plants begin 129 flowering at the start of the rainy season; reproductive plants have $\bar{x} = 1.1$ flowering 130 shoots (range = 1-7), each of which has an inflorescence with 20-25 flowers (Bruna & 131 Kress, 2002). Fruits mature April-May, have 1–3 seeds per fruit ($\bar{x} = 2$), and are eaten by a 132 thrush and several species of manakin (Uriarte et al., 2011). Dispersed seeds germinate 133 approximately 6 months after dispersal at the onset of the subsequent rainy season, with

rates of germination and seedling establishment higher in continuous forest than forestfragments (Bruna, 1999; Bruna & Kress, 2002).

136 Demographic data collection

This study uses data collected in four 1-ha fragment reserves and six continuous forest 137 sites. In 1997–1998 we established a 5000 m² plots (50×100 m) in each of these sites in 138 139 which we marked and measured all *Heliconia acuminata*; plots in 1-ha fragments were on 140 one randomly selected half of the fragment, while plots in continuous forest were located 141 500–4000 m from the borders of secondary and mature forest. The distance between plots 142 ranged from 500 m-41 km. Our dataset comprised 4,083 plants in continuous forest and 143 1,010 plants in forest fragments. Plots in CF had on average 2.7-fold more plants than plots 144 in 1-ha fragments (CF = 681 ± 493 SD; 1-ha = 253 ± 30 SD).

145 Each plot was subdivided into 50 quadrats $(10 \times 10m)$ to simplify annual surveys, during 146 which we recorded the number of vegetative shoots each plant had, the height of each plant 147 to the tallest leaf, and whether each plant was flowering (height and shoot number are 148 correlated with leaf area, the probability of flowering, and rates of survivorship (Bruna, 149 2002; Bruna & Kress, 2002)). In this study, we used the product of shoot number and plant 150 height as our measure of plant size. Preliminary analysis showed that the product of shoot 151 number and height was a better predictor of total leaf area (which in turn is assumed to be 152 a strong predictor of aboveground biomass) than either shoot number or height alone 153 (Table S2). Plants that were not found for three consecutive surveys were considered 154 dead. We also surveyed plots regularly during the rainy season to identify any that

155 flowered after the survey. For additional details on the location of plots, survey methods,

and *H. acuminata* population structure see Bruna & Kress (2002).

157 Climate data

158 Data on precipitation and potential evapotranspiration in our sites were obtained from a 159 published gridded dataset ($0.25^{\circ} \times 0.25^{\circ}$ resolution) built using data from 3,625 ground-160 based weather stations across Brazil (Xavier et al., 2016). We used these data to calculate 161 the standardized precipitation evapotranspiration index (SPEI) (Vicente-Serrano et al., 162 2010). SPEI is a proxy for meteorological drought that integrates precipitation and 163 evapotranspiration anomalies over a specified time scale. Positive SPEI values for a given 164 month indicate conditions wetter than the historical average for that month, while negative 165 values of SPEI indicate droughts with intensity categorized as mild (0 to -1), moderate (-1 166 to -1.5), severe (-1.5 to -2), or extreme (< -2) (McKee et al., 1993). SPEI can be calculated to 167 represent different temporal scales of drought; we used 3-month SPEI because—given its 168 shallow roots and rhizome—*H. acuminata* relies primarily on soil moisture rather than 169 deeper water sources that can persist for longer timescales (Vicente-Serrano et al., 2010). 170 Note that 3-month SPEI is still monthly data—each month's SPEI value simply takes into 171 account precipitation and evapotranspiration of the previous three months. SPEI 172 calculations were made using the SPEI package (Beguería & Vicente-Serrano, 2017). The 173 timing of drought events based on these SPEI calculations is consistent with that resulting 174 from SPEI calculated with other data sources, though the magnitude of drought sometimes 175 differed (Figure S2; Table S1).

176 Statistical Modeling of Vital Rates

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177 To assess the effects of drought history on plant vital rates we used Distributed Lag Non-178 linear Models (DLNMs, Gasparrini et al., 2017). DLNMs capture how potentially delayed 179 effects of predictor variables (e.g. SPEI) affect an outcome (e.g. growth) well beyond the 180 event period. They do so by fitting a bi-dimensional predictor-lag-response association 181 spline, referred to as a crossbasis function. This models a non-linear relationship between 182 predictor and response (e.g. between SPEI and vital rates) and allows the shape of that 183 relationship to vary smoothly over lag time. Using the dlnm package (Gasparrini, 2011; R 184 Core Team, 2020), we created crossbasis functions with possible lags from 0–36 months. 185 We chose 36 months as a maximum lag because prior transplant experiments with *H*. 186 acuminata showed they typically recover from transplant shock in less than 36 months 187 (Bruna et al., 2002) so this is a reasonable upper bound for lagged effects of drought. 188 The crossbasis function was fit to the data in the context of a generalized additive model 189 (GAM) with restricted maximum likelihood using the mgcv package (Wood, 2017). The

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$$g[E(y_i)] = \alpha_j + s_1(z_i) + s_2(d_{i,t}, \dots, d_{i,t-36}) + \beta x_i$$
$$\alpha_j \sim N\left(\mu_{\alpha_j}, \sigma_{\alpha_j}^2\right), \text{ for plot } j$$
(1)

general form of the vital rate (y) models was as follows:

where $s_1(z_i)$ is a smooth function of plant size (natural log of height × shoot number), fit using a penalized cubic regression spline, $s_2(\cdot)$ is the crossbasis function in which $d_{i,t}$ is the SPEI value during the census month of an observation (February) and $d_{i,t-l}$ is the SPEI lmonths prior (see Gasparrini et al. 2017 for details). The crossbasis function, $s_2(\cdot)$ can also be written:

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$$s_2(d_{i,t},\ldots,d_{i,t-24}) = \sum_{l=l_0}^{L} f \cdot w(d_{i,t-l},l)$$
(2)

198 where the crossbasis function, $f \cdot w(d, l)$, is composed of two marginal basis functions: the 199 standard predictor-response function f(d), and the additional lag-response function w(l). 200 These marginal functions are combined as a tensor product smooth such that the shape of 201 one marginal function varies smoothly along the other dimension (see chapter 5 of Wood 202 (2017) and Gasparrini et al. (2017) for more detail). Penalized cubic regression splines 203 were used for both marginal bases of the crossbasis function, with 35 knots for the lag 204 dimension (i.e. number of lagged SPEI values for each observation with 36 months as a 205 maximum lag) and 3 knots for the drought response dimension to restrict the shape of the 206 fitted response to drought to bimodal when most complex. Because of penalization, the 207 number of knots is generally not important as long it is large enough to allow the smooth to 208 represent the 'true' relationship (Wood, 2017). Estimated degrees of freedom (edf) 209 represent the 'true' complexity of the smooth after penalization with edf = 1 being 210 equivalent to a straight line and larger numbers representing more complex curves.

To determine if plot characteristics influenced average vital rates we included a random effect of plot ID on the intercept; this was represented by a_j in eq. 1. We determined the effects of SPEI on plant growth using plant size in year t+1 as a response variable. This was modeled with a scaled t family error distribution because residuals were leptokurtic with a Gaussian error structure. Because number of inflorescences was highly zero-inflated, we converted this to a binary response to model reproduction (i.e., 1 for \geq 1 inflorescence, 0 for no inflorescences). We modeled both reproduction and survival (i.e., from year t to year 218t+1) using a binomial family error distribution with a logit link function. We modeled a219potential cost of reproduction by including flowering in the previous year as covariate, x_i ,220in eq. 1. Additionally, in models for flowering probability and size, we included plant ID as a221random effect to account for variation among individuals. Preliminary analyses showed222that this random effect was not significant in the growth models (edf ~ 0, p > 0.2) and as223such it was dropped to improve computational efficiency. 1 A random effect of plant ID was224not included in survival models since each plant only dies once.

In the process of fitting the models, the penalty on the crossbasis smooth (and other smoothed terms) is optimized such that more linear shapes are favored unless the data supports non-linearity (Wood, 2017). We applied an additional penalty to shrink linear portions toward zero with the select=TRUE option to the gam() function, and inferred statistical significance of model terms with p-values from the summary.gam() function as recommended in Marra & Wood (2011).

The dlnm package does not currently allow the modeling of interaction terms, which means we could not asses the interaction of habitat type and lagged effects. We therefore fit separate models for plants in fragments and in continuous forest to allow the shape of the crossbasis function to differ between habitats. Significant main effects of habitat type were assessed by looking for overlap in the 84% confidence intervals of model intercepts; the 84% CIs of two samples drawn from the same population overlap about 95% of the time (Payton et al., 2003).

All analyses were conducted in R version 4.0.2 (2020-06-22) (R Core Team, 2020).

239 **Results**

The meteorological droughts in our focal region indicated by SPEI are generally consistent with those reported in the literature (Table S1). For example, the drought associated with the 1997 El Niño Southern Oscillation (ENSO) event was one of the most severe on record for the Amazon (Williamson et al., 2000); correspondingly, 1997 has the lowest SPEI values in our timeseries (Figure 1d). The 2005 dry season (June–October) was also reported as an exceptionally dry year, although this drought mostly affected the southwestern Amazon (Marengo et al., 2008; Zeng et al., 2008). Our SPEI data show the 2005 dry season to be a

247 moderate drought (-1 > SPEI > -1.5).

248 Survival, growth, and flowering in continuous forest vs. fragments

249 Survival: Across all plots, the proportion of plants surviving was lowest in the 2003–2004 250 transition year ($P_{surv} = 0.92$). This coincided with droughts in both the 2003 and 2004 251 rainy seasons (Figure 1b,d) and was preceded by a drop in average plant size in the 2002– 2003 transition year (Figure 1a). The lowest survival for 1-ha fragment plots ($P_{surv} = 0.93$) 252 253 was for the 2005–2006 transition year, which encompassed a moderate drought in October 254 2005 and and wetter than average conditions (SPEI > 0.5) in December 2005 and January 2006 (Figure 1b,d). The lowest survival for continuous forest was in 2004 ($P_{surv} = 0.91$). 255 256 When summarizing across years, plots, and plant sizes, the survival probability of *Heliconia acuminata* was similarly high in both continuous forest and fragments ($P_{surv} = 0.95$; Figure 257 258 1b; the overlapping 85% CI of model intercepts indicate no significant difference). 259 However, survival in both habitats was size dependent (p < 0.001 for the effect of log-260 transformed plant size in year *t* on survival in year *t*+1 in both habitats). The survival

probability of large plants approached 1 in both habitat types (Figure 3b), but the smallestplants had higher survival in 1-ha fragments.

263 Growth: Plants in continuous forest had an average of 2.9 shoots (± 1.8 SD) and were on 264 average 40.6 cm tall (± 26.5 SD). Plants in 1-ha fragments had on average 13.8% fewer 265 shoots (2.5 ± 1.5 SD) and were 10.8% shorter ($36.3 \text{ cm} \pm 24.1$ SD). Because our proxy for 266 plant size was the product of these two metrics, plants in continuous forest were on 267 average 34% larger than those in forest fragments (150 ± 175 SD vs. 112 ± 141 SD, 268 respectively), with fragments having proportionately fewer large plants (Figure 3d). This 269 difference was not significant, however (overlap in 84% CI of model intercepts), and the 270 disparity in plant size—which was most pronounced in the initial years of our surveys— 271 diminished over time (Figure 1a).

272 Mean plant size dropped dramatically in 2003 in both habitat types, corresponding with a

severe drought during the February census (SPEI = -1.39) (Figure 1d). As with survival,

size in year t was a significant predictor of size in year t+1 (p < 0.001 in both habitats).

275 While the effect was generally similar across size classes and habitat types, the impact of

276 plant size on growth was greatest for mid-sized plants in continuous forest (Figure 3a).

277 *Flowering:* The overall proportion of plants that flowered was very low. While it was almost

40% higher in continuous forest than 1-ha fragments (0.05 ± 0.21 vs. 0.04 ± 0.19 ,

279 respectively), this difference was not statistically significant (84% CIs of model intercepts

overlapped). The observed disparity was largely due to the fact that flowering is also size-

dependent (p < 0.001 in both habitats), with the probability of flowering increasing

dramatically once plants reached the threshold size of about 148 (i.e., log(size) > 5 in

Figure 3c). Despite the flowering probability of large plants being greater in fragments than
continuous forest, populations in fragments had proportionately fewer plants above the
reproductive size threshold (Figure 3d). The most striking difference between habitat
types coincided with a severe drought in 2003, when the percentage of flowering
reproductive-sized plants was 28% in continuous forest vs. only 13.6% in 1-ha fragments
(Figure 1c).

289 Delayed effects of drought on demographic vital rates

Drought history had a significant (p < 0.001) effect on the survival, growth, and flowering of plants in both habitats. Comparing the respective crossbasis surfaces, however, reveals that the specific climatic drivers, their timing, and their impact on individual vital rates all differed among habitats.

294 *Survival:* For 1-ha fragments, there was a significant effect on survival of SPEI in the 295 preceding 13 months. The highest survival was near SPEI of 0, with mortality increasing as 296 conditions became either drier or wetter (i.e., as SPEI values became increasingly negative 297 or positive, respectively; Figure 4b). Wet conditions in the preceding 23 months (i.e., SPEI >298 1) also had a significant negative effect on survival in fragments (Figure 4b). In contrast, 299 the effect of recent SPEI in continuous forest was weaker, with only the 6 months preceding 300 a census having a significant effect on survival (Figure 4a). These short-term effects of SPEI 301 on survival were also unidirectional—the probability of survival declined, albeit only 302 slightly, with increasingly negative values of SPEI (i.e., as droughts became more severe; 303 Figure 4a). In contrast, the most pronounced negative effects of SPEI on the survival of 304 plants in continuous forest were at lag times of 15–20 months and 32–36 months. Drought

305 15–20 months prior to a census (i.e. two dry seasons prior to a census) was significantly 306 associated with reduced survival, while high precipitation (i.e., SPEI > 1) was significantly 307 associated with higher survival. Finally, plants in both habitat types showed an increase in 308 survival probability with very high SPEI values (i.e., extremely high precipitation) at a lag 309 time of 32–36 months. It should be noted, however, that only the first year of census data 310 (1999) met these conditions. We compared the effects of SPEI history in continuous forest 311 and fragments by subtracting the fitted values in Figure 4b from Figure 4a to produce 312 Figure 4c. This shows that in average conditions (SPEI = 0), there is little difference in 313 survival probability between continuous forest and forest fragments (Figure 4c). However, 314 under extreme conditions, survival probability is higher in continuous forest by up to 315 0.025.

316 *Growth:* The effects of drought history on trends in plant size were generally similar for 317 continuous forest and fragments. However, the crossbasis function for 1-ha fragments 318 indicated more complex responses in some situations (edf = 17.8 for 1-ha fragments; edf = 319 13.0 for continuous forest; see also Figure 5). For example, under average conditions (i.e., 320 SPEI = 0), growth is similar or slightly higher in continuous forest over all lag periods 321 (Figure 5c). However, the growth of plants in fragments is reduced when the current wet 322 season is unusually wet (i.e., SPEI>2), and as a result the plants in continuous forests will 323 be larger by up to $\log(size) = 0.57$. In contrast, drought at lags of 8–11 months (i.e., the end 324 of the preceding year's wet season) led to increased growth in both habitats, with a more 325 pronounced response in 1-ha fragments.

326 *Flowering:* Overall, the probability of flowering was higher in continuous forest than in 327 1-ha fragments for all values of SPEI (Figure 6), although this difference was not significant 328 (84% CIs of intercepts overlap). The responses in 1-ha fragments were also more muted as 329 indicated by the shape of the crossbasis function (1-ha edf = 8.3, continuous forest edf =330 10.6) (Figure 6). This led to some important inter-habitat differences in plant responses to 331 prior droughts. In continuous forests, recent drought (i.e., at lag = 0-2 with SPEI < -1), 332 drought two dry seasons prior (lags 15-20) and in the wet season 34-36 months prior all 333 increased the probability of flowering. However, drought at the end of the rainy season one 334 year prior (lags 7–13) significantly reduced flowering probability (Figure 6a). In fragments, 335 recent drought had no significant effect on flowering probability and only drought at two 336 dry seasons prior (lags 16-20) and in the wet season 33-36 months prior increased 337 flowering probability (Figure 6b). The effects of drought on flowering probability were 338 strong in continuous forest compared to 1 ha fragments (Figure 6c). We found no evidence 339 for a cost of reproduction: in both forest and fragments, plants that had flowered in the 340 previous year were on average more likely to be larger (CF: p = 0.048; 1-ha: p = 0.030) 341 and flower again (CF: p < 0.001; 1-ha: p = 0.004). The random effect of plant ID on 342 flowering probability was also significant (CF: p < 0.001; 1-ha: p = 0.003), indicating 343 significant individual-level variation in flowering probability. 344 Finally, with the exception of survival in 1-ha fragments (p = 0.253), the delayed effects of

345 SPEI on all three vital rates varied significantly among plots (p < 0.01 for the random effect 346 of plot).

347 **Discussion**

348 Understanding how landscape structure and abiotic conditions act to influence population 349 dynamics is central to many conceptual frameworks for studying and conserving 350 fragmented landscapes (Didham et al., 2012; Driscoll et al., 2013). Our results support the 351 emerging consensus that the effects of climatic extremes on demographic vital rates can be 352 delayed for months or even years (Evers et al., 2021; Teller et al., 2016; Tenhumberg et al., 353 2018). We also found that the delayed responses of populations in fragments can differ 354 significantly in magnitude, direction, and lag time from those of populations in continuous 355 forest. This suggests that the hypothesized synergies between climate and fragmentation 356 on population dynamics (Laurance & Williamson, 2001; Opdam & Wascher, 2004; Selwood 357 et al., 2015) are likely to be pervasive, but also far more complex than previously thought.

358 Temporal variation in demographic responses to forest fragmentation

359 Many studies investigating the biological consequences of habitat fragmentation on plant 360 growth, survival, and reproduction comprise short-term (\leq 3 year) experiments and 361 observations. Our results underscore the difficulty in extrapolating long-term trends from 362 such short-term studies, particularly when studying long-lived organisms or when the 363 responses of interest can vary with size or age. For instance, one would have reached a very 364 different conclusion regarding the effect of fragmentation on annual survival if the study 365 windows were 1999–2002 (i.e., higher survival in continuous forest), 2002–2005 (i.e., 366 higher survival in fragments), or 2004–2007 (i.e, no clear effect of fragmentation) (Figure 367 1b). It is only when evaluating over longer time windows that it becomes apparent 368 mortality is elevated in fragments relative to continuous forest (Figure 2), and that the 369 observed interannual variation is largely driven by dynamic patterns of recruitment

370 (Bruna, 2002) coupled with low mortality for plants beyond the smallest size classes371 (Bruna, 2003).

372 Similarly, conclusions regarding the effects of fragmentation on flowering—which is also 373 both rare and size-dependent (Brooks et al., 2019)—would also differ based on the year in 374 which they were investigated. This could lead to erroneous extrapolations regarding the 375 effects of fragmentation on reproductive mutualists or population genetic structure (Côrtes 376 et al., 2013; Uriarte et al., 2010; Uriarte et al., 2011). Conclusions based on short-term 377 observations of temporally variable vital rates could lead to conservation and management 378 practices that are ineffective or even counterproductive, especially when when failing to 379 consider how the consequences of this variation might be modulated by organismal life 380 history (Morris et al., 2008).

381 It is important to emphasize, however, that the overall effects of SPEI on survival and 382 growth are more severe in fragments than continuous forest (Figures 4, 5). Furthermore, 383 the magnitude of plant responses to climatic extremes is also greater in habitat 384 fragments—extreme drought in dry seasons and extreme precipitation in during rainy 385 seasons are most detrimental to growth and survival in fragments. While intact forest and 386 its canopy buffer populations from climatic extremes, populations in fragments—especially 387 near edges with high contrast matrix—likely lack this protection (Didham & Lawton, 1999; 388 Ewers & Banks-Leite, 2013). We suggest it is these climate extremes, rather than trends in 389 average temperature, precipitation, or SPEI (Laurance et al., 2014), that that are the causal 390 mechanism underlying reduced plant growth and survival in forest fragments.

391 Delayed effects of climate on demographic vital rates

392 Climate anomalies are known to have immediate effects on the growth, survival, or 393 reproduction of plants (Esteban et al., 2021; Wright & Calderon, 2006), including Heliconia 394 (Stiles, 1975; Westerband et al., 2017) and other tropical herbs (Wright, 1992). These 395 effects can be complex or even contradictory—mild droughts can increase the growth rates 396 of tropical trees and seedling survival, perhaps due to reductions in cloud cover and 397 concomitant increases in solar radiation (Alfaro-Sánchez et al., 2017; Condit et al., 2004; 398 Huete et al., 2006; Jones et al., 2014; Uriarte et al., 2018), but in severe drought years 399 growth can be extremely low and mortality can be sharply elevated (Connell & Green, 400 2000; Edwards & Krockenberger, 2006; Engelbrecht et al., 2002). There is also evidence 401 that the effects can persist for multiple years (Phillips et al., 2010), such as a boom in 402 drought-year fruit production followed by severe post-drought "famine" (Pau et al., 2013; 403 Wright et al., 1999).

404 Despite these insights, models of plant population dynamics rarely include the effects of 405 environmental drivers [but see Williams et al. (2015); Tenhumberg et al. (2018); Molowny-406 Horas et al. (2017)). This has largely been due to the challenge (both ecologically and 407 statistically) of detecting any demographic responses to climatic extremes that are delayed 408 for multiple growing seasons. To address this, researchers have begun to use a number of 409 statistical methods that test for time lags in demographic responses without *a priori* 410 assumptions about the influence of any particular climate window (Evers et al., 2021; 411 Teller et al., 2016; Tenhumberg et al., 2018). Our expansion of this approach, which offers 412 an unbiased way of identifying these delayed effects without overfitting (but see Pierre et 413 al. (2020) and Ogle et al. (2015) for alternative methods) yielded results consistent with

this emerging literature—that the effects of precipitation extremes on the demography of *Heliconia acuminata* could be delayed for up to 3 growing seasons.

416 While it appears that delayed effects of climate on demographic vital rates may be 417 ubiquitous (Evers et al., 2021), the extent to which they vary spatially or with habitat 418 remains an open question. Our results clearly indicate that they can, with habitat-specific 419 differences in how environmental conditions influenced future vital rates. For example, 420 extreme values of SPEI—both positive (unusually high precipitation) and negative 421 (drought conditions)—led to declines in the probability of individual survival in both 422 habitat types. However, the magnitude of these declines was far greater in forest 423 fragments. Similarly, the detrimental effects of extremes in SPEI on growth rates were also 424 more pronounced in fragments. In contrast, variation in SPEI had far stronger effects on the 425 probability of flowering in continuous forest than fragments. These results should be 426 interpreted with some caution, however, as the relatively low number of plants in 427 fragments that are above the threshold-size for flowering could limit the power to detect 428 delayed effects.

There are several, non-mutually-exclusive explanations for delayed effects of SPEI on demography. The first is that the physiological processes underlying vital rates might be initiated long before they are demographically apparent (Evers et al., 2021), and hence be shaped by climatic events at any point in that physiological window. For example, the flowering shoots of *Heliconia chartacea* begin to develop 6–10 months prior to the appearance of inflorescences (Criley & Lekawatana, 1994). Adverse conditions during the 6 months following initiation, rather than the months when inflorescences are starting expand, leads to the aborted production of flowering shoots. Interestingly, we observed the
opposite effect—drought conditions increased the probability of flowering two years later.
While this could reflect bet-hedging in response to stress (Nihad et al., 2018), this does not
appear to be the case, as growth or survival do not appear to decrease following
reproduction (see also (Horvitz & Schemske, 1988). In fact, flowering in one year is
associated with increased reproduction and growth in the next.

442 Demographic responses will also be delayed if abiotic stress causes plants to invest in 443 belowground rhizomes (sensu Pumisutapon et al., 2012). The carbohydrates stored in 444 rhizomes allow Heliconia to regenerate aboveground biomass following damage (Rundel et 445 al., 1998) and protect the buds that give rise to new shoots from stressful conditions 446 (Klimešová et al., 2018). This may be why drought led to delayed increases in growth—by 447 shedding shoots and leaves (Bruna et al., 2002) and investing in rhizomes, plants are 448 generating proportionately more buds with which to regenerate when conditions improve. 449 This would also be consistent with the results of prior experiments, in which the growth 450 rates of *H. acuminata* 8 months after they were mechanically damaged far exceeded those 451 of control plants (Bruna & Ribeiro, 2005).

Third, it may be that the delayed demographic effects we observed are indirectly mediated by the effect of SPEI on other species rather than the direct effects on individual physiology (Evers et al., 2021). For example, topical trees may not die until three or more years after a drought (Criley & Lekawatana, 1994). When they finally do, the resulting leaf drop (Janssen et al., 2021) and treefalls allow for light penetration to the forest understory (Canham et al., 1990; Leitold et al., 2018), triggering a boom in the growth and flowering of understory 458 plants (Bruna & Oli, 2005). Similar delayed changes in the local environment could also 459 influence the foraging behavior of a plant's pollinators (Bruna et al., 2004; Stouffer & 460 Bierregaard, 1996), seed dispersers (Uriarte et al., 2011), or herbivores (Scott et al., 2021). 461 While more work is needed to explain why the (delayed) effects of SPEI on *H. acuminata* 462 survival and growth are greater in fragments than forest interiors, one hypothesis, 463 motivated by recent intriguing results from other systems (Sapsford et al., 2017), is that the 464 greater litterfall on edges (Vasconcelos & Luizão, 2004) may be altering the abundance of 465 pathogens or mycorrhizae.

466 Finally, demographic delays could be an artifact of the timing of responses in relation to the 467 census date. If extreme drought in the dry season before the census increased plant 468 mortality during that season, for example, this would nevertheless appear in models as a 469 delayed effect (e.g. in Figure 4b). In our case, this potential explanation for delayed effects 470 applies only to plant size and survival, as plots were surveyed regularly throughout the 471 reproductive season to identify flowering plants. This possibility is not unique to our study, 472 rather it is a consequence of conducting demographic censuses on an annual scale while 473 the climate is quantified monthly or seasonally. While the very slow growth and extremely 474 low mortality rates of *H. acuminata* mean this effect is unlikely to be acting in our system, it 475 may be that for some species it will be important to conduct demographic surveys at the 476 same temporal scale at which climate is aggregated.

477 Conclusions & Future Directions

478 Over 24 million ha of the Brazilian Amazon have been cleared in the last two decades (Silva

479 Junior et al., 2021), resulting in their extensive fragmentation (Broadbent et al., 2008).

480 Climate models predict a future of extremes for these forests—increases in the frequency 481 and geographic extent of droughts, but also increases in the frequency and area affected by 482 periods of unusual wetness (Duffy et al., 2015). Our results support the hypothesis that 483 populations in Amazonian forest fragments could be more susceptible to the effects of 484 changing climate than those in continuous forest (Laurance et al., 2014). However, they 485 also indicate that the demographic responses to climate change of populations in 486 fragmented landscapes may be far more complex than previously appreciated. Multi-487 factorial, multi-season experiments (Aguirre et al., 2021; sensu Bruna & Ribeiro, 2005; 488 Markewitz et al., 2010; Westerband et al., 2017), ideally manipulating multiple 489 combinations of climatic variables (Mundim & Bruna, 2016), are needed to determine how 490 and why habitat-specific differences in environmental conditions interact to delay the 491 demographic responses of plants to climatic variability. Also needed are statistical tools 492 that can test for synergistic effects of fragmentation and climate in vital rates, as those 493 currently available do not allow for including interaction terms. This also limits the ability 494 to include size by climate interactions in a DLNM; although plant responses to both 495 fragmentation and climatic extremes can be size-specific (Bruna & Oli, 2005; Schwartz et 496 al., 2019). The ability to identify size-specific lagged responses may be especially 497 complicated given size and growth are rarely measured at the same time scale as SPEI and 498 other putative climatic drivers.

499 Finally, no analytical approach assessing the potential for demographic lags can

500 compensate for a lack of long-term data (Evers et al., 2021; Tenhumberg et al., 2018).

501 Unfortunately, long-term data monitoring the entire life-cycle of tropical taxa are rare, and

502 those doing so in fragmented landscapes are virtually nonexistent (Bruna & Ribeiro, 2005).

503 Without investing in collecting such data, generalizations regarding the demographic

504 consequences of climate change in these species rich and increasingly fragmented habitats

505 will continue to prove elusive. More generally, however, researchers need to consider how

506 delayed responses to climate could influence the interpretation of data in studies where the

507 organisms lifespan exceeds the study's duration.

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Data Availability Statement 515

516 Data used in this study are available at the Dryad Digital Repository; R code for analyses

517 and visualizations are archived at Zenodo [links included upon acceptance].



520 *Figure 1: (a-c) Time series of H. acuminata vital rates in 1-ha fragments (solid orange lines)* and continuous forest (dashed blue lines) and (d) drought occurrence in the study region. (a) 521 522 *Mean fold-change in plant plant size (log2(size_{t+1} / size_t)) varies by year and habitat. On* 523 average, plants grew in most years with the notable exception in 2003, in which on average 524 plants regressed in size in both habitats (i.e., fold-change < 0). Error bars represent the standard deviation. (b) The proportion of plants surviving from one transition year to the next 525 526 varied from 0.98 (CF in 1998-1999) to 0.91 (CF in 2003-2004). (c) The proportion of H. 527 acuminata above the size threshold for reproduction that flowered each year is on average 528 low but variable. The size threshold is determined by the upper 90th percentile size of 529 flowering plants across all years. (d) Monthly 3-month SPEI for our study region. Gray lines

- 530 represent values from different grid cells encompassing BDFFP; the dark line represents the
- 531 site mean. Colored stripes represent drought intensity: yellow = mild, orange = moderate, dark
- 532 *orange = severe, red = extreme.*



533

534 Figure 2: Survivorship curve for plants marked in the 1998 survey year; these plants comprise

- 535 49% of those in the complete demographic dataset. The percentage of these plants that were
- 536 still alive ten years later was 79.7% (1629/2055) in continuous forest vs. 72.4% (393/543) in
- 537 1-ha fragments.







Figure 4: Smooth effect of lagged SPEI on H. acuminata survival in (a) continuous forest, (b) 546 547 1-ha fragments, (c) and the difference between the two. Outlines show regions where the 548 effect of SPEI is significant, defined as those where the 95% confidence interval around the 549 response does not overlap the intercept. The bar on the bottom of each panel indicates the wet 550 seasons (black, November-May) and dry seasons (white, June-October). Surface is modeled as 551 a crossbasis function with cubic regression splines for each marginal basis. Model intercepts 552 were added to fitted values of the crossbasis function and back-transformed to the response 553 scale.





- 555 Figure 5: Smooth effect of lagged SPEI on H. acuminata growth in (a) continuous forest, (b) 1-
- 556 ha fragments, (c) and the difference between the two. Outlines show regions where the effect
- 557 of SPEI is significant, defined as those where the 95% confidence interval around the response
- 558 does not overlap the intercept. The bar on the bottom of each panel indicates the wet seasons
- 559 (black, November–May) and dry seasons (white, June–October).





562 continuous forest, (b) 1-ha fragments, (c) and the difference between the two. Outlines show

regions where the effect of SPEI is significant, defined as those where the 95% confidence
interval around the response does not overlap the intercept. The bar on the bottom of each

565 panel indicates the wet seasons (black, November–May) and dry seasons (white, June–

566 October). Surface is modeled as a crossbasis function with cubic regression splines for each

567 marginal basis. Model intercepts were added to fitted values of the crossbasis function and

568 *back-transformed to the response scale.*

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1011	#>	Р	fs		1.5.0	2020-07-31	[?]
1012	#>	Р	generics		0.1.0	2020-10-31	[?]
1013	#>	Р	ggplot2	*	3.3.3	2020-12-30	[?]
1014	#>	Р	glue	*	1.4.2	2020-08-27	[?]
1015	#>	Р	goftest		1.2-2	2019-12-02	[?]
1016	#>	Р	gratia	*	0.6.0.9112	2021-06-11	[?]
1017	#>	Ρ	gridExtra		2.3	2017-09-09	[?]
1018	#>	Ρ	gtable		0.3.0	2019-03-25	[?]
1019	#>	P	haven		2.4.1	2021-04-23	[?]
1020	#>	P	here	*	1.0.1	2020-12-13	[?]
1021	#>	P	highr		0.9	2021-04-16	[?]
1022	#>	P	Hmisc	*	4.5-0	2021-02-28	[?]
1023	#>	P	hms		1.0.0	2021-01-13	[?]
1024	#>	P	htmlTable		2.1.0	2020-09-16	[?]
1025	#>	P	htmltools		0.5.1.1	2021-01-22	[?]
1026	#>	P	htmlwidgets		1.5.3	2020-12-10	[?]
1027	#>	P	httr		1.4.2	2020-07-20	[?]
1028	#>	P	igraph		1.2.6	2020-10-06	[?]
1029	#>	P	ianitor	*	2.1.0	2021-01-05	[?]
1030	#>	Ρ	ipeg		0.1-8.1	2019-10-24	[?]
1031	#>	P	isonlite		1.7.2	2020-12-09	[?]
1032	#>	P	knitr	*	1.33	2021-04-24	[?]
1033	#>	Ρ	labeling		0.4.2	2020-10-20	[?]
1034	#>	P	latex2exp	*	0.5.0	2021-03-18	[?]
1035	#>	Ρ	lattice	*	0.20-44	2021-05-02	[?]
1036	#>	P	latticeExtra		0.6-29	2019-12-19	[?]
1037	#>	P	lifecvcle		1.0.0	2021-02-15	[?]
1038	#>	P	lmomco	*	2.3.6	2020-03-14	[?]
1039	#>	P	Imoments		1.3-1	2019-03-15	[?]
1040	#>	P	lubridate	*	1.7.10	2023 03 25	[2]
1041	#>	P	magrittr		2.0.1	2020-11-17	[?]
1042	#>	P	MASS		7.3-54	2021-05-03	[2]
1043	#>	P	Matrix		1.3-3	2021-05-04	[?]
1044	#>	P	memoise		2.0.0	2021-01-26	[?]
1045	#>	P	mgcv	*	1.8-35	2021-04-18	[2]
1046	#>	P	modelr		0.1.8	2020-05-19	[?]
1047	#>	P	munsell		0.5.0	2018-06-12	[2]
1048	#>	P	mynfast		0.2.5 1	2020-10-14	[2]
1049	#>	P	mytnorm		1.1-1	2020-06-09	[2]
1050	#>	P	nlme	*	3,1-152	2021-02-04	[2]
1000	11.5				J.T T)2	2021 02-04	L •]

1051	#>	Ρ	nnet		7.3-16	2021-05-03	[?]
1052	#>	Ρ	numDeriv		2016.8-1.1	2019-06-06	[?]
1053	#>	Ρ	patchwork	*	1.1.1	2020-12-17	[?]
1054	#>	Ρ	pillar		1.6.0	2021-04-13	[?]
1055	#>	Ρ	pkgbuild		1.2.0	2020-12-15	[?]
1056	#>	Ρ	pkgconfig		2.0.3	2019-09-22	[?]
1057	#>	Ρ	pkgload		1.2.1	2021-04-06	[?]
1058	#>	Ρ	png		0.1-7	2013-12-03	[?]
1059	#>	Ρ	prettyunits		1.1.1	2020-01-24	[?]
1060	#>	Ρ	processx		3.5.2	2021-04-30	[?]
1061	#>	Ρ	ps		1.6.0	2021-02-28	[?]
1062	#>	Ρ	purrr	*	0.3.4	2020-04-17	[?]
1063	#>	Ρ	qqplotr	*	0.0.5	2021-04-23	[?]
1064	#>	Ρ	R6		2.5.0	2020-10-28	[?]
1065	#>	Ρ	RColorBrewer		1.1-2	2014-12-07	[?]
1066	#>	Ρ	Rcpp	*	1.0.6	2021-01-15	[?]
1067	#>	Ρ	readr	*	1.4.0	2020-10-05	[?]
1068	#>	Ρ	readxl	*	1.3.1	2019-03-13	[?]
1069	#>	Ρ	remotes		2.3.0	2021-04-01	[?]
1070	#>	Ρ	renv		0.13.2	2021-03-30	[?]
1071	#>	Ρ	reprex		2.0.0	2021-04-02	[?]
1072	#>	Ρ	rlang		0.4.11	2021-04-30	[?]
1073	#>	Ρ	rmarkdown	*	2.7	2021-02-19	[?]
1074	#>	Ρ	robustbase		0.93-7	2021-01-04	[?]
1075	#>	Ρ	rpart		4.1-15	2019-04-12	[?]
1076	#>	Ρ	rprojroot		2.0.2	2020-11-15	[?]
1077	#>	Ρ	rstudioapi		0.13	2020-11-12	[?]
1078	#>	Ρ	rvest		1.0.0	2021-03-09	[?]
1079	#>	Ρ	scales		1.1.1	2020-05-11	[?]
1080	#>	Ρ	sessioninfo		1.1.1	2018-11-05	[?]
1081	#>	Ρ	snakecase		0.11.0	2019-05-25	[?]
1082	#>	Ρ	SPEI	*	1.7	2017-06-07	[?]
1083	#>	Ρ	statmod	*	1.4.35	2020-10-19	[?]
1084	#>	Ρ	stringi		1.5.3	2020-09-09	[?]
1085	#>	Ρ	stringr	*	1.4.0	2019-02-10	[?]
1086	#>	Ρ	survival	*	3.2-11	2021-04-26	[?]
1087	#>	Ρ	tarchetypes	*	0.2.0	2021-05-11	[?]
1088	#>	Ρ	targets	*	0.4.2	2021-04-30	[?]
1089	#>	Ρ	testthat		3.0.2	2021-02-14	[?]
1090	#>	Ρ	tibble	*	3.1.1	2021-04-18	[?]
1091	#>	Ρ	tidyr	*	1.1.3	2021-03-03	[?]
1092	#>	Ρ	tidyselect		1.1.1	2021-04-30	[?]
1093	#>	Ρ	tidyverse	*	1.3.1	2021-04-15	[?]
1094	#>	Ρ	tsibble	*	1.0.1	2021-04-12	[?]
1095	#>	Ρ	tsModel	*	0.6	2013-06-24	[?]
1096	#>	Ρ	usethis		2.0.1	2021-02-10	[?]
1097	#>	Ρ	utf8		1.2.1	2021-03-12	[?]
1098	#>	Ρ	vctrs		0.3.8	2021-04-29	[?]
1099	#>	Ρ	withr		2.4.2	2021-04-18	[?]
1100	#>	Ρ	xfun		0.22	2021-03-11	[?]

1101	#>	P xml2	1.3.2	2020-04-23 [?]
1102	#>	P yaml	2.2.1	2020-02-01 [?]
1103	#>	source		
1104	#>	CRAN (R 4.0.2)		
1105	#>	CRAN (R 4.0.2)		
1106	#>	CRAN (R 4.0.2)		
1107	#>	CRAN (R 4.0.2)		
1108	#>	CRAN (R 4.0.2)		
1109	#>	CRAN (R 4.0.2)		
1110	#>	CRAN (R 4.0.2)		
1111	#>	CRAN (R 4.0.2)		
1112	#>	CRAN (R 4.0.2)		
1113	#>	CRAN (R 4.0.2)		
1114	#>	(R 4.0.2)		
1115	#>	(R 4.0.2)		
1116	#>	(R 4.0.2)		
1117	#>	(R 4.0.2)		
1118	#>	(R 4.0.2)		
1110	#\>	(R 4 0 2)		
1120	#\\$	(R 4 0 2)		
1120	#>	(R 4.0.2)		
1121	#\\	(R 4 0 2)		
1122	#\\$	(R 4 0 2)		
1123	#\\$	(R 4 0 2)		
1125	#\	(R 4.0.2)		
1125	π/ #\	(R 4.0.2)		
1120	т/ #\	(R 4.0.2)		
1127	π/ #\	(R 4.0.2)		
1120	#/ #\	(R 4.0.2)		
1120	#/ #\	(R 4.0.2)		
1121	#/ #\	(R 4.0.2)		
1122	#/ #\	(R 4.0.2)		
1132	#/ #\	(R 4.0.2)		
1133	#/ #\	(R 4.0.2)		
1125	#/ #\	(R 4.0.1)		
1135	#/ #\	(R 4.0.2)		
1127	#/ #\	(R 4.0.2)		
1120	#> #\	(R 4.0.2)		
1120	#> #\	(R 4.0.2)		
1139	#> #\	(R 4.0.2)		
1140	#> #\	(R 4.0.2)		
1141	#> #\	(R 4.0.2)		
1142	#>	(R 4.0.2)		
1143	#> #`	(R 4.0.2)		
1144	#>	CRAN (R 4.0.2)		
1145	#>	CRAN (R 4.0.2)		00250-2-4
1146	#>	Github (gavinsim	oson/gratia	@9359C3Q)
114/	#>	CRAN (R 4.0.0)		
1148	#>	CRAN (R 4.0.2)		
1149	#>	CRAN (R 4.0.2)		
1150	#>	CRAN (R 4.0.2)		

1151	#>	CRAN	(R 4.0.2)
1152	#>	CRAN	(R 4.0.2)
1153	#>	CRAN	(R 4.0.2)
1151	#\	CDAN	(R 1.0.2)
	#2	CRAN	(R 4.0.2)
1155	#>	CRAN	(R 4.0.2)
1156	#>	CRAN	(R 4.0.2)
1157	#>	CRAN	(R 4.0.2)
1158	#>	CRAN	(R 4.0.2)
1159	#>	CRAN	(R 4.0.2)
1160	#\	CRAN	$(R \land 0 \ 0)$
1161	#\ #\	CRAN	(R 4 0 2)
1162	#\ #\	CRAN	(R 4 0 2)
1162	#>	CRAN	(R 4 0 2)
1164	π/ #\		(R + 0.2)
1165	#/ #\		(R 4.0.2)
1166	#/ #\		(R 4.0.2)
1167	#> #\		(R 4.0.2)
1160	#> #\		(R 4.0.2)
1160	#> #\		(R 4.0.2)
1109	#> #\		(R 4.0.2)
1170	#> #\		(R 4.0.2)
11/1	#> #\		(R 4.0.2)
1172	#> #\	CRAN	(R 4.0.2)
11/3	#>	CRAN	(R 4.0.2)
11/4	#>	CRAN	(R 4.0.2)
11/5	#>	CRAN	(R 4.0.2)
11/6	#>	CRAN	(R 4.0.2)
11//	#>	CRAN	(R 4.0.2)
1178	#>	CRAN	(R 4.0.2)
1179	#>	CRAN	(R 4.0.2)
1180	#>	CRAN	(R 4.0.2)
1181	#>	CRAN	(R 4.0.2)
1182	#>	CRAN	(R 4.0.2)
1183	#>	CRAN	(R 4.0.2)
1184	#>	CRAN	(R 4.0.2)
1185	#\	CRAN	(R / 0.2)
1105	π/ #\	CDAN	(R + 0.2)
1100	#2	CRAN	(R 4.0.2)
110/	#>	CRAN	(R 4.0.2)
1188	#>	CRAN	(R 4.0.0)
1189	#>	CRAN	(R 4.0.2)
1190	#>	CRAN	(R 4.0.2)
1191	#>	CRAN	(R 4.0.2)
1192	#>	CRAN	(R 4.0.2)
1193	#>	CRAN	(R 4.0.2)
1194	#>	CRAN	(R 4.0.2)
1195	#>	CRAN	(R 4.0.2)
1196	#>	CRAN	(R 4.0.2)
1197	#>	CRAN	(R 4.0.2)
1198	#>	CRAN	(R 4.0.2)
1199	#>	CRAN	(R 4.0.2)
1200	#>	CRAN	(R 4.0.2)

<u>~</u>`

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1201	#> CRAN (R 4.0.2)
1202	#> CRAN (R 4.0.2)
1203	#> CRAN (R 4.0.2)
1204	#> CRAN (R 4.0.2)
1205	#> CRAN (R 4.0.2)
1206	#> CRAN (R 4.0.2)
1207	#> CRAN (R 4.0.2)
1208	#> CRAN (R 4.0.2)
1209	#> CRAN (R 4.0.2)
1210	#> CRAN (R 4.0.2)
1211	# CRAN (R 4.0.2)
1212	# CRAN (R 4.0.2)
1213	H^{2} CRAN (R 4.0.2)
1210	H^{2} CRAN (R 4.0.2)
1211	$H_{2} = CRAN (R 4.0.2)$
1216	$H_{2} = CRAN (R 4.0.2)$
1210	$H_{2} = CRAN (R 4.0.2)$
1217	H^{2} CRAN (R 4 0 2)
1210	H^{*} CRAN (R 4 0 2)
1220	H^{2} CRAN (R 4 0 2)
1220	H^{2} CRAN (R 4 0 2)
1221	$H_{2} = CRAN (R 4.0.2)$
1222	H^{2} CRAN (R 4 0 2)
1223	H^{2} CRAN (R 4 0 2)
1221	H^{*} CRAN (R 4 0 2)
1225	# CRAN (R 4 0 2)
1220	# CRAN (R 4 0 2)
1227	# CRAN (R 4.0.2)
1220	H^{2} CRAN (R 4 0 2)
1230	H^{2} CRAN (R 4 0 2)
1230	# (RAN (R 4.0.2)
1231	$\#_{2} = CRAN (R + 0.2)$
1232	#> Chan (h 4:0:2)
1233	<pre>#/ #/ [1] /Users/scottericr/Documents/HeliconiaDemography/reny/library/R-</pre>
1234	π [1] / 0.5et 3/ 3cotter 1ct / bocuments/ herroniabellogi aphy/renv/ribrary/ κ^{-}
1235	4.07.800_04-appie-dal wini7.0
1230	<pre>m/ [2] /nrivate/var/folders/h /2vfnvvls5vs401tmhhh3wadh0000gn/T/RtmnKX4NCE/renv-</pre>
1237	system_library
1230	+> [2]
1235	#/ [J] /nrivate/var/folders/h /2vfnyyls5vs401tmbhb3wgdb0000gn/T/PtmnyEg51N/renv-
1240	system_libnary
1241	
1242	#> P — Loaded and on-disk path mismatch.
1244	The current Git commit details are:
1245 1246	<pre>#> Local: emilio /Users/scottericr/Documents/HeliconiaDemography #> Remote: emilio @ origin</pre>

```
1247
```

(https://github.com/BrunaLab/HeliconiaDemography.git)
#> Head: [5e4dd7c] 2021-06-22: minor edits to results and discussion