

Additive genetic effects in interacting species jointly determine the outcome of caterpillar herbivory

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**1 Plant-insect interactions are common and important in basic and
2 applied biology. Trait and genetic variation can affect the outcome
3 and evolution of these interactions, but the relative contributions of
4 plant and insect genetic variation and how these interact remain un-
5 clear and are rarely subject to assessment in the same experimental
6 context. Here we address this knowledge gap using a recent host
7 range expansion onto alfalfa by the *Melissa* blue butterfly. Common
8 garden rearing experiments and genomic data show that caterpillar
9 performance depends on plant and insect genetic variation, with in-
10 sect genetics contributing to performance earlier in development and
11 plant genetics later. Our models of performance based on caterpillar
12 genetics retained predictive power when applied to a second com-
13 mon garden. Much of the plant genetic effect could be explained
14 by heritable variation in plant phytochemicals, especially saponins,
15 peptides, and phosphatidyl cholines, providing a possible mechanistic
16 understanding of variation in the species interaction. We find
17 evidence of polygenic, mostly additive effects within and between
18 species, with consistent effects of plant genotype on growth and de-
19 velopment across multiple butterfly species. Our results inform the-
20 ories of plant-insect coevolution and the evolution of diet breadth in
21 herbivorous insects and other host-specific parasites.**

Plant-insect interaction | Genomic prediction | Polygenic | Phytochemicals | Coevolution

1 A central challenge for the biological sciences is to under-
2 stand the causes and consequences of trait variation
3 within and among species. Experimental manipulations aimed
4 at understanding the molecular basis of organismal variation
5 have most often been done in settings stripped of all or most
6 ecological context. This approach can be fruitful for simple
7 traits, including some aspects of morphology (e.g., (1–4)), but
8 is lacking when it comes to interspecific interactions that in-
9 clude the evolution of crop pests, emerging infectious diseases,
10 and other host-parasite associations (5, 6).

11 Plants and herbivorous insects have contributed much to
12 our understanding of the formation and persistence of inter-
13 actions between hosts and parasites, in part because they
14 are experimentally tractable but also because insects are the
15 most diverse macroscopic organisms on the planet and their
16 specialized feeding habits play a role in their diversification
17 (7–11). Yet classic studies of the molecular basis of plant-insect
18 interactions have relied on candidate genes or targeted classes
19 of phytochemical compounds (e.g., (12–14)). More recently,
20 evolutionary geneticists have taken advantage of new technolo-

gies to explore the genetic basis of herbivory in a genomic context. With very few exceptions, these studies have focused on genetic variation in either herbivores or plants (15–19)(but see (20)), but rarely both in the same study and never to our knowledge paired with modern metabolomic approaches that allow for untargeted discovery of influential compounds (21). This leaves us with considerable uncertainty concerning the relative importance of heritable traits in herbivores and in plants for determining the outcome of plant-insect interactions. For example, particular genetic variants in an herbivore might be associated with increased feeding efficiency, but only when challenged with particular plant variants such as specific defensive metabolites or combinations of physical defenses (22). However, without an understanding of the genetic architecture of both the herbivore physiology and the plant traits, the evolutionary trajectory of the system cannot be understood in the context of available theoretical models or forecast with respect to the evolution of defense in the plant or increased performance in the herbivore. We address this need using a recent host range expansion onto alfalfa by the *Melissa* blue

Significance Statement

Studies of ecological interactions often ignore genetic variation and studies of coevolution have rarely assayed the genetics of hosts and parasites at the same time. We show that genetic differences among *Melissa* blue caterpillars and alfalfa plants account for 17–49% of the variability in caterpillar growth and survival. The genetic contribution includes heritable variation in defensive compounds including saponins. Our results suggest that the outcome of this plant-insect interaction is affected by many genes with mostly independent (additive) effects. Moreover, genetic differences among alfalfa plants have consistent effects on caterpillar growth in multiple butterfly populations and species. Our results thus advance understanding of the evolution of ecological interactions, including host-parasite interactions beyond herbivorous insects and plants.

ZG, CAB, JAF, CCN, SLL, LKL, and MLF designed the study. ZG, TS, CP, SAY, EP, MES, JGH, CAB, LKL and MLF conducted the experiments and generated the data. ZG, TS, CP, and MLF analyzed the data. ZG, CP and MLF drafted the manuscript. All authors helped edit and revise the manuscript.

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Summary of hypotheses and effects

Hypothesis	Plant genetics	Insect genetics	Combination
(i)	No	Yes	N/A
(ii)	Yes	No	N/A
(iii)	Yes	Yes	Additive
(iv)	Yes	Yes	Epistatic
Null	No	No	N/A



Fig. 1. Main hypotheses tested about the contribution of plant and insect genetics to caterpillar performance: (i) caterpillar performance is primarily affected by insect (*L. melissa*) genetics, (ii) caterpillar performance is primarily affected by plant (*M. sativa*) genetics, (iii) the genetics of the interacting species have similar effects on caterpillar performance and combine additively, (iv) the genetics of the interacting species have similar effects on caterpillar performance and combine epistatically, and (v) the null hypothesis that neither insect or plant genetic variation have an appreciable effect on caterpillar performance. The illustration (by R. Ribas) shows a *L. melissa* caterpillar feeding on alfalfa, while being tended by ants; additional biotic or abiotic factors, such as the presence of mutualistic ants, also affect caterpillar performance in the wild (25) but are not a component of this study.

butterfly, emphasizing the role of prediction when building an understanding of the functional genetic basis of a novel plant-insect interaction.

The Melissa blue butterfly (*Lycaeides melissa*) is widespread in western North America (23). It exists in isolated populations associated with larval host plants in the legume family, including many species of *Astragalus* and *Lupinus* (24, 25). The Melissa blue colonized alfalfa (*Medicago sativa*) after the plant was introduced to the western USA as a forage crop in the mid 1800s, and is now commonly found on naturalized (i.e., feral) alfalfa along roadsides and trails (24). Melissa blue butterflies show evidence of adaptation to alfalfa, but this host plant remains inferior to known native hosts in terms of caterpillar development with cascading life history effects (26–28). As insect growth and survival is often reduced on novel hosts, the lower quality of alfalfa for Melissa blue butterflies is likely typical of a general phenomenon (29). Alfalfa is phenotypically variable (30), and thus is not a homogeneous resource for Melissa blue butterflies. In particular, phenotypic variation among naturalized alfalfa populations, including phytochemical variation, affects Melissa blue caterpillar growth and host patch occupancy (25, 31, 32). However, it is unclear how much of this phenotypic variation has a genetic basis. Moreover, as is true for other plant-insect interactions, the relative contributions of plant (alfalfa) and insect (Melissa blue) genetic variation to the outcome of the interaction is unexplored, including whether growth and successful development from caterpillar to adult is influenced by additive or epistatic genetic variation in the interacting species.

Here, we use multiple common garden rearing experiments combined with multilocus genetic mapping and genomic prediction to build and test models that quantify the relative effects and interactions of alfalfa and Melissa blue genetic variation on caterpillar performance (i.e., growth and survival). We specifically test the following alternative hypotheses: (i) caterpillar performance is primarily affected by Melissa blue genetic variation and architecture, (ii) caterpillar performance is primarily affected by genetic variation and architecture in the host plant, (iii) the genetics of the interacting species have similar effects on caterpillar performance and combine additively, (iv) the genetics of the interacting species have similar effects on caterpillar performance and combine epistatically,

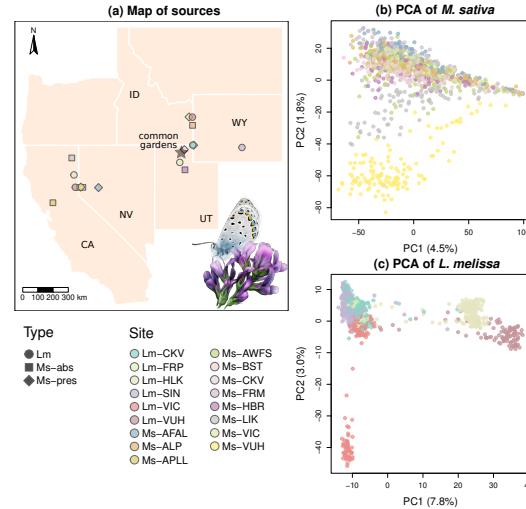


Fig. 2. (a) Map of plant (*M. sativa*) and insect (*L. melissa*) common garden source populations. Symbol shapes denote source type—Lm = *L. melissa*, Ms-abs = *M. sativa* site without *L. melissa* butterflies, and Ms-pres = *M. sativa* site with *L. melissa* butterflies, and are colored to indicate different populations within taxa. The inset illustration shows an adult *L. melissa* perched on *M. sativa* (illustration by R. Ribas). (b) Ordination of genetic variation via principal component analysis (PCA) for the *M. sativa* common garden plants. (c) Ordination of genetic variation via PCA for the *L. melissa* caterpillars from the rearing experiment. Points in (b) and (c) denote individual plants or caterpillars and are colored to match the map (a).

and (v) the null hypothesis that neither Melissa blue nor alfalfa genetic variation has an appreciable effect on caterpillar performance (Fig. 1). Genetic mapping of 1760 plant traits, including 1750 phytochemical metabolites, contributes to testing these hypotheses and also allows us to probe the functional basis of plant-genetic effects on caterpillar performance. Finally, we conduct complementary rearing experiments to test the consistency of plant genetic effects (i.e., their lack of interaction with herbivore genetics) across butterfly populations and species.

Results

Overview of the primary common garden rearing experiment.

We planted a common garden comprising 1080 alfalfa (*M. sativa*) plants at the Greenville Experimental Farm near Logan, UT (41.765° N, 111.814° W) in 2018 (Fig. S1a). Seeds for this garden were collected from 11 naturalized (i.e., feral) *M. sativa* sites in the western USA, including five sites where *L. melissa* butterflies are found (Table S1, Fig. 2a). Caterpillars for the experiment were sourced from six sites by obtaining eggs from gravid *L. melissa* females in 2019. We detected substantial genetic variation and only subtle genetic differentiation among the source locations for alfalfa (161,008 SNPs, mean expected heterozygosity = 0.168, F_{ST} = 0.029) and for *L. melissa* (63,194 SNPs, mean expected heterozygosity = 0.065, F_{ST} = 0.045) (Figs. 2b, S2). Nearby SNPs (<100 bps) exhibited appreciable linkage disequilibrium in *M. sativa* (median r^2 = 0.050, 95th percentile = 0.862) and *L. melissa* (median r^2 = 0.002, 95th percentile = 0.052), but this decayed rapidly with physical distance with especially low levels of LD beyond 100 bps in *L. melissa* (Fig. S3).

The main rearing experiment was conducted in summer 2019. For this experiment, caterpillars were reared individually on each of the 1080 alfalfa plants. Rearing was done in a growth

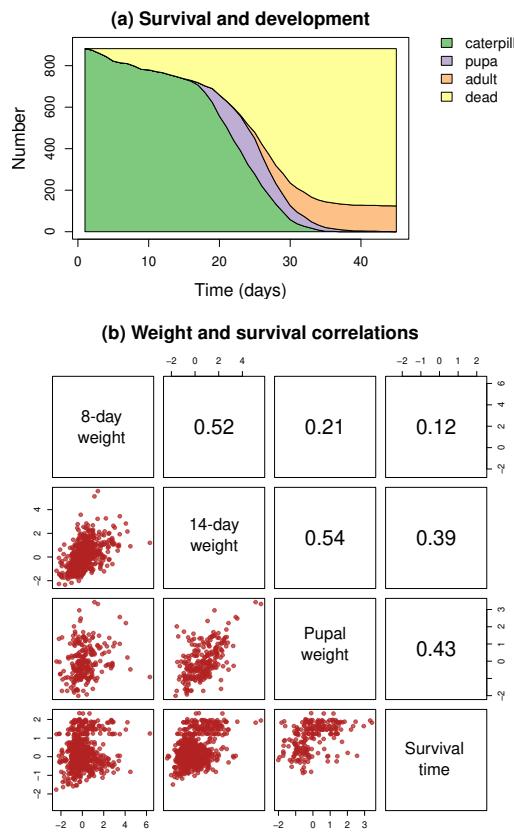


Fig. 3. (a) Plot shows survival and development of *L. melissa* over the course of the rearing experiment. Colored regions denote the number of individuals that were living caterpillar, pupa, adults or dead at each day post hatching. (b) Plots show pairwise correlations between *L. melissa* performance traits. Scatterplots are shown in the lower-triangle panels—each point denotes one individual—and Pearson correlations are reported in the corresponding upper triangle panels. Traits are given along the diagonal panels: 8-day weight, 14-day weight, pupal weight, and truncated survival time. Scatterplots and Pearson correlations are based on residuals after controlling for confounding environmental effects (see Methods for details).

performance in the common garden rearing experiment (Fig. 4a), consistent with our hypotheses (iii) and (iv) (Fig. 1). Specifically, *M. sativa* genetics (161,008 SNPs) explained between 2% (survival to 8 days) and 36% (14-day weight) of the variation in performance (mean across traits = 17%), and *L. melissa* genetics (63,194 SNPs) explained 5% (weight at pupation and survival to pupation) to 29% (8-day weight) of the variation in the same nine caterpillar performance measures (mean = 15%) (values denote point estimates of the percent variance explained, PVE; see Table S2 for credible intervals; cross-validation results are shown in the next section). Caterpillar genetics contributed more to performance metrics from early development (e.g., 8-day weight and survival to 8 and 14 days), whereas plant genetics mattered more for later development (e.g., 14-day weight, pupal weight, and survival to pupation and adult), resulting in a trend towards a negative relationship between caterpillar and plant genetic contributions across traits (Pearson $r = -0.52$, 95% CI = -0.88 to 0.22, $P = 0.15$). We detected mostly positive genetic correlations among performance traits (Fig. 4b), with similar but not identical genetic correlations calculated from *M. sativa* and *L. melissa* polygenic scores (Pearson correlation between *M. sativa* and *L. melissa* genetic correlations, $r = 0.80$, 95% CI = 0.63 to 0.89, $P = 5.923e^{-9}$). Polygenic scores in this context quantify the estimated effect of many plant or caterpillar genetic variants on a performance trait.

Mapping results suggested mostly a polygenic basis for the performance traits, with point estimates of > 10 loci affecting most traits (Tables S2, S3 and S4; Fig. 4c,d), but with more evidence of specific SNPs strongly associated with performance in *L. melissa*. This included ten SNPs with posterior probabilities of association (i.e., posterior inclusion probabilities) > 0.5 with at least one performance trait (Fig. 4d, Table S5). Some of these SNPs were in or near (< 20 kbps) genes with biologically plausible functions for affecting performance, such as *MSP-300*, *Lipase member H* and *Juvenile hormone acid O-methyltransferase*, all of which were associated with 8-day weight. For example, *MSP-300* affects muscle development and muscle-ectoderm attachment in *Drosophila* (33). Insect lipases metabolize fats, are expressed in gut tissue, and can affect survival and reproductive capacity in insects; *Lipase member H* in particular has further been associated with viral resistance in the moth *Bombyx mori* (34, 35). *Juvenile hormone acid O-methyltransferase* is involved in juvenile hormone biosynthesis and thus in the regulation of insect growth and development, especially metamorphosis (36, 37). A single *M. sativa* SNP was strongly associated with survival to pupation (posterior inclusion probability > 0.5 ; chromosome 1, position = 12,930,966 bp). This SNP was found in a gene encoding *TOM1-like protein 9* and was within 30 kbps of six additional genes, including two genes with known links to plant-insect interactions: *dentin sialophosphoprotein*, which is associated with soybean compensatory growth after cutworm herbivory (38), and *photosystem I reaction center subunit psaK*, which has been mechanistically linked to tolerance to aphids and aphid feeding preference in *Arabidopsis* (39) (Table S6). We obtained similar results with complementary genetic mapping analyses that included 20 genetic PCs as additional controls for population structure when estimating SNP-performance associations; this was true both in terms of the percentage of variation in performance explained (Pearson correlations $>$

chamber, with caterpillars fed fresh leaf tissue as needed. In this experiment, 26.1% of the caterpillars survived to pupation and 14.1% survived to eclose as adults (mean survival time = 21.8 days) (Fig. 3a). Mean *L. melissa* weights were 2.94 mg (SD = 2.13) at 8 days, 12.7 mg (SD = 7.71) at 14 days, and 20.0 mg (SD = 7.21) at pupation. Weight and survival were variable within and among groups of caterpillars from different source populations and within and among groups that consumed plants grown from different *M. sativa* source populations (Figs. S4 and S5). Weight and survival metrics of performance were positively correlated, including, 8-day weight vs. 14-day survival (Pearson $r = 0.0916$, 95% confidence interval [CI] = 0.0237–0.159), 14-day weight vs. survival to pupation ($r = 0.472$, 95% CI = 0.416–0.525), and pupal weight vs. survival to eclosion ($r = 0.449$, 95% CI = 0.342–0.545) (Fig. 3b). Past work has shown that weight and lifetime fecundity are highly correlated in *L. melissa* (26).

Plant and caterpillar genetic variation affect performance. Using multilocus genome-wide association methods (see Figs. S6, S7 and S8 for evidence of adequate Markov chain Monte Carlo performance), we found evidence that both *M. sativa* and *L. melissa* genetic variation contributed to caterpillar

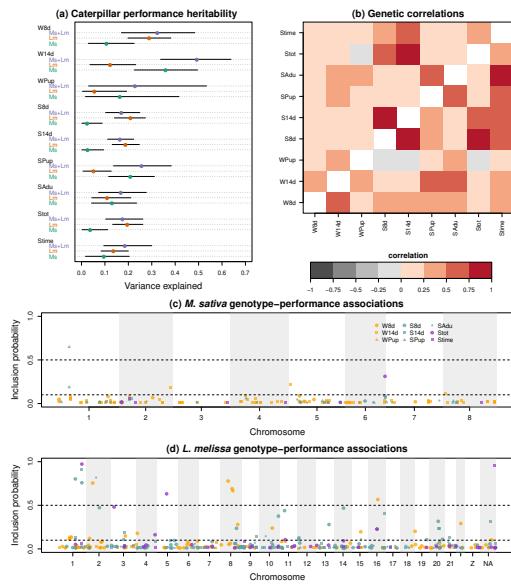


Fig. 4. Genetic mapping of caterpillar performance. (a) Dotchart shows Bayesian estimates of the proportion of trait variation explained by *M. sativa* genetics (Ms), *L. melissa* genetics (Lm), or both combined (Ms+Lm) for each caterpillar performance trait; W8d = 8-day weight, W14d = 14-day weight, Wpup = pupal weight, S8d = 8-day survival, S14d = 14-day survival, SPup = survival to pupation, SAdu = survival to adult, Stot = total survival time, and Stime = (truncated) survival time. Points and horizontal lines denote point estimates (posterior medians) and 95% equal-tail probability intervals, respectively. (b) Heatmap shows genetic correlations between pairs of caterpillar performance traits based on *M. sativa* genetics (lower triangle) or *L. melissa* genetics (upper triangle). Manhattan plots in (c) and (d) shown posterior inclusion probabilities (PIPs) for genotype-performance associations based on *M. sativa* and *L. melissa* SNPs, respectively. Points denote SNPs with different colors and symbols for different performance traits. Only SNPs with PIPs ≥ 0.01 are depicted. Horizontal lines at PIPs of 0.1 and 0.5 are included for reference.

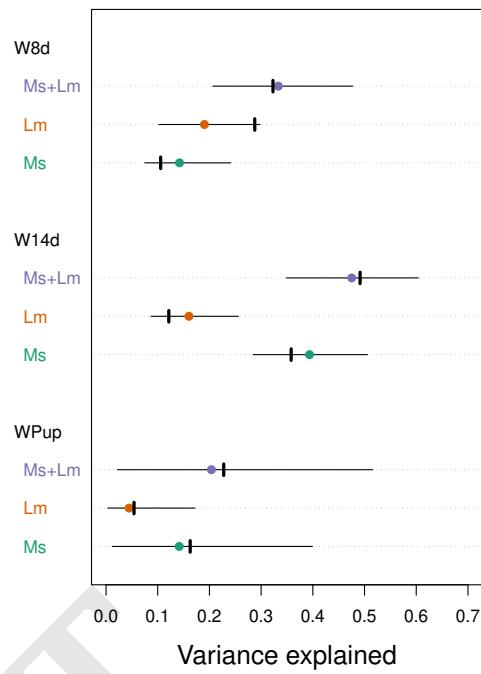


Fig. 5. Genetic mapping of caterpillar performance with epistasis. The dotchart shows Bayesian estimates of the proportion of trait variation explained by *M. sativa* genetics (Ms), *L. melissa* genetics (Lm), or both combined (Ms+Lm) for 8-day weight (W8d), 14-day weight (W14d), and pupal weight (Wpup). Points and horizontal lines denote point estimates (posterior medians) and 95% equal-tail probability intervals, respectively, for the proportion of trait variation explained by additive effects and pairwise epistatic effects. Vertical black lines denote point estimates (posterior medians) for the proportion of variation explained by additive genetic effects alone (as presented in Fig. 4a).

minimize the low power associated with testing all SNP-SNP interactions, we tested for marginal epistasis, that is for evidence of an epistatic interaction between each SNP and any of the other SNPs. We failed to find significant evidence of marginal epistasis among *M. sativa* SNPs, among *L. melissa* SNPs, or between *M. sativa* and *L. melissa* SNPs (i.e., no SNPs achieved genome-wide significance) (Figs. S12 and S13). Our failure to find epistasis in this manner could be driven in part by limited power to detect it. Thus, we next re-fit the multilocus genome-wide association models described above, but with additional terms for epistatic interactions. This allowed us to directly ask where including epistasis increases our ability to explain caterpillar performance. To do this in a statistically tractable way, we added pairwise interactions between the 150 SNPs with the most evidence of marginal epistasis (i.e., lowest P -values); this added an additional 11,175 terms to each model. Models including these epistatic effects failed to explain more of the variation in caterpillar performance than our purely additive models (Fig. 5). Thus, these direct tests of epistasis provide additional evidence against hypothesis (iv) and thus in favor of hypothesis (iii) (i.e., additivity within and between species) (Fig. 1). Consequently, we focus on the additive models in tests of predictive power below, before presenting additional tests of additivity versus epistasis in subsequent sections of this paper.

Predicting caterpillar performance from plant and caterpillar genotype. We next showed that our genotype-phenotype mod-

0.99, $P < 0.001$ for caterpillar and plant genetics) (Tables S7 and S8, Fig. S9), and in terms of specific SNP-performance associations (Pearson correlations for posterior inclusions probabilities, *M. sativa* $r = 0.76$, $P < 0.001$, *L. melissa* $r = 0.98$, $P < 0.001$) (Fig. S10).

We repeated the genetic mapping approach using a combined data set of both *M. sativa* and *L. melissa* genetic loci (i.e., the combined 224,202 SNPs) (genetic PCs were not included here or in subsequent analyses). The combined data set generally explained more of the variation in caterpillar performance, 17% to 49% (mean = 24%), than either *M. sativa* or *L. melissa* genetic loci alone. Moreover, the combined variation explained for each performance trait was well described by a model where the variance explained separately by plant and caterpillar genetics combined additively. Specifically, in a linear regression model, the percent variance in performance traits explained by plant and caterpillar genetics separately explained 97% of the variation in the estimates of the variance explained by the combined genetic data sets (linear regression, $\beta_{plant} = 1.17$, $P = 6.6e^{-6}$, $\beta_{caterpillar} = 0.80$, $P = 0.00037$, $r^2 = 0.97$) (Fig. S11), consistent with our hypothesis (iii) (Fig. 1).

Given the evidence of additivity of genetic effects between species presented above, we next turned to more direct tests of the hypothesis that epistatic interactions contribute to caterpillar performance, with a specific focus on caterpillar and pupal weight (see Methods for details and justification). To

els were moderately successful at predicting caterpillar performance. This is relevant both for validating these models and for demonstrating their potential utility and limitations in making predictions about effects and evolutionary trajectories in nature. Specifically, genomic predictions of performance from 10-fold cross validation exhibited statistically significant positive correlations with observed performance values for three out of ten performance traits for *M. sativa* genetics, five out of ten traits for *L. melissa* genetics, and six out of ten traits for *M. sativa* and *L. melissa* genetics combined (Fig. 6a). Especially pronounced positive correlations between observed and predicted performance were detected for 14-day weight based on *M. sativa* genetics and 8-day and 14-day survival based on *L. melissa* genetics. More generally, our ability to predict performance traits was well explained by our estimates of heritability (i.e., PVE): we calculated Pearson correlations of 0.89 (95% CI = 0.55 to 0.98, $P = 0.0013$) and 0.62 (95% CI = -0.073 to 0.91, $P = 0.074$) between PVE estimates and the correlation between observed and predicted traits for *M. sativa* and *L. melissa* genetics, respectively (Fig. 6b). In other words, we better predicted caterpillar performance for the performance traits that were more heritable.

Having demonstrated moderate predictive power within the main common garden, we next asked whether genotype-phenotype models estimated from this garden could successfully predict *L. melissa* performance for additional caterpillars fed *M. sativa* from a second, smaller common garden (the Gene Miller Life Science Garden; $N = 180$ plants) (Fig. S1). This second garden, planted in 2018 on the Utah State University campus ~ 2.5 km from the Greenville Experimental Farm garden, included plants from six of the 11 *M. sativa* source sites and caterpillars from each of the sites used in the main experiment. Survival rates for caterpillars reared on plants from this garden were similar to those reared on plants from the main garden (Fig. S14). Predictive performance for the second garden differed notably for *M. sativa* versus *L. melissa* genotype-phenotype models, with statistically significant positive correlations between observed and predicted trait values in the new garden for only one trait for *M. sativa* genetics versus six of the ten performance traits for *L. melissa* genetics (Fig. 6c). Predictions for the combined data set were similar to those based on *L. melissa* genetics alone. Consistent with these patterns, estimates of PVE from the main garden explained predictive power for *L. melissa* genetics (Pearson $r = 0.93$, 95% CI = 0.68 to 0.98), but not *M. sativa* genetics ($r = 0.17$, 95% CI = -0.56 to 0.75). Thus, unmeasured environmental differences likely limit our ability to predict performance from plant genetics across gardens to a much greater extent than for caterpillar genetics (plant growth environments differed, but caterpillar rearing environments did not) despite these gardens being separated by only ~ 2.5 km. Differences in the exact genetic composition of the two gardens could add to this effect.

Genetic associations with plant traits explain the plant genetic contribution to caterpillar performance. Having shown that plant genetic variation affects caterpillar performance, we now focus on the Greenville Experimental Farm (see Fig. S1a) to identify possible components of the functional basis of the documented plant-genetic effects. This also allowed us to further test for additive versus epistatic interactions between plant and caterpillar genotypes on caterpillar per-

formance (see our hypotheses (iii) versus (iv) in Fig. 1). We first determined the extent to which genetic loci associated with caterpillar performance were also associated with other plant traits, including potential plant vigor or defense traits (17). Such genetic correlations could arise from pleiotropy, but also from linkage disequilibrium between distinct loci affecting the plant traits and caterpillar performance (i.e., genetic correlations do not demonstrate a causal genetic link between traits). Still, such an association would be consistent with the hypothesis that these traits, or other genetically correlated traits, constitute possible mechanisms by which plant genotype affects caterpillar performance. To do this, we measured and mapped 1760 plant traits in the Greenville Experimental Farm garden using the same multilocus mapping approach and *M. sativa* SNP data set described above. The traits included plant height, leaf length, leaf width, leaf area, leaf shape, leaf weight, specific leaf area, leaf toughness, trichome density, levels of herbivory on the plants in the field, and 1750 plant chemistry metabolites, which were quantified and characterized using liquid-chromatography combined with mass spectrometry (LC-MS; similar to (25, 32)).

We documented genetic variation affecting most of the plant traits, with mean PVEs of 20.5% for the non-chemical traits (minimum = 5.6%, maximum = 38.7%) and 10.9% (310 traits $> 20\%$ and 20 $> 50\%$) for the 1750 chemical traits (Table S9). Additionally, in the main Greenville Experimental Farm common garden, the distribution of PVE for the 1750 chemical traits differed markedly from that for 1750 matched, randomized traits, consistent with a clear genetic contribution to this variation in leaf metabolites (Fig. S15).

Multiple plant traits, including chemical and non-chemical traits, exhibited genetic correlations with each caterpillar performance trait; in other words, plant trait polygenic scores were correlated with caterpillar performance polygenic scores when inferred from plant genetics (Figs. 7a,b, S16). However, because of the large number of measured traits and genetic correlations among the plant traits (Fig. S17), many of the genetic correlations between plant traits and caterpillar performance were likely redundant. Thus, to identify the combined subset of traits most strongly predictive of caterpillar performance (and thus the best candidates for a mechanistic link to performance), we next fit a LASSO penalized regression model for the polygenic scores of each caterpillar performance trait (based on plant genetics) as a function of the polygenic scores for the 1760 plant traits. These models explained 41 to 80% of the variation in the caterpillar performance scores (mean = 69.2%, cross-validation predictive r^2 ranged from 0.39 to 0.76) (Table S10, Fig. 7c). On average 260 of the 1760 traits were retained in these models (i.e., given non-zero regression coefficients), with a range of 117 (survival time) to 347 (8-day survival) traits (Figs. 7d,e and S18). Both chemical and non-chemical traits were retained in the models. Non-chemical traits with the biggest effects included a positive effect of plant height on 14-day weight ($\beta = 0.037$), positive effects of trichome density ($\beta = 0.036$) and specific leaf area ($\beta = 0.031$) on survival to adulthood, and a negative effect of leaf toughness on survival to adulthood ($\beta = -0.34$). Consistent with a previous phenotypic assay of caterpillar performance and plant metabolomic variation in this system (32), top chemical traits included several saponins, including saponins (two distinct Medicagenic acids) associ-

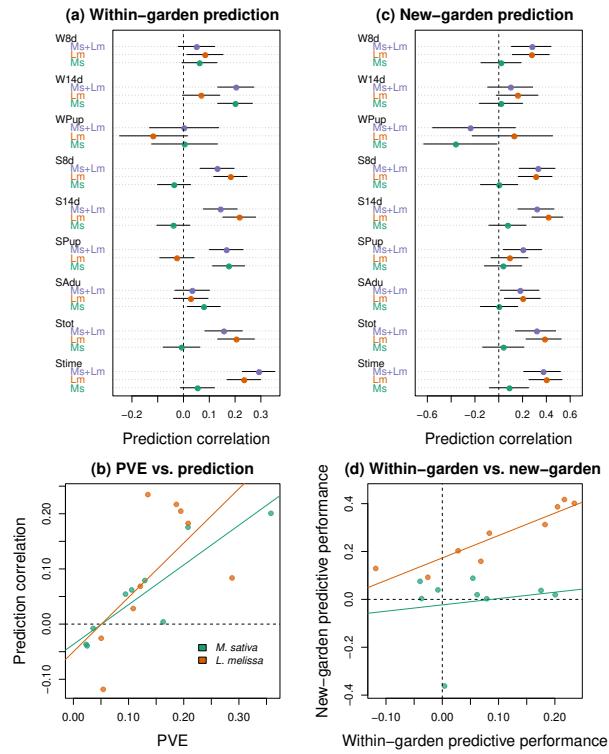


Fig. 6. Genomic prediction of caterpillar performance. (a) Dotchart shows Pearson correlations between cross-validation genomic predictions of phenotypes and the observed values based on *M. sativa* genetics (Ms), *L. melissa* genetics (Lm), or both combined (Ms+Lm) for each caterpillar performance trait; W8d = 8-day weight, W14d = 14-day weight, WPup = pupal weight, S8d = 8-day survival, S14d = 14-day survival, SPup = survival to pupation, SAdu = survival to adult, Stot = total survival time, and Stime = (truncated) survival time. Points and horizontal lines denote point estimates (posterior medians) and 95% equal-tail probability intervals, respectively. For example, a large value on the x axis indicates a high correlation between observed performance values and predictions from genotype based on cross validation. (b) Scatterplot of the proportion of variation explained by genetics (PVE) versus the Pearson correlation of genomic predictions from (a). Each point denotes a trait and is colored to indicate values from *M. sativa* or *L. melissa* genetics. Colored lines are best fits from ordinary linear regression, and a dashed line denotes the 0 value on the y-axis. (c) Dotchart similar to (a), but for genomic predictions of phenotypes in a second common garden (the Gene Miller Life Science Garden) based on the models fit from the main garden. (d) Scatterplot of correlations between observed caterpillar performance trait values and genomic predictions of these values using cross-validation within the main garden versus prediction for samples in the Gene Miller Life Science Garden based on the models fit for the main garden.

ated with effects on caterpillar weight and survival (Tables S12 and S13). The flavonoid glycoside Tricin 7-glucoside was associated with reduced survival, whereas several peptides (e.g., MESA.583 = C₁₃H₁₈O, a fragment of a N-acyl amine; MESA.615 = C₂₃H₄₃N₇O₇; MESA.849 = C₁₄H₁₉NO₃, a N-acyl amine) were associated with reduced weight or survival (Tables S12 and S13). Lastly, we fit LASSO regression models on the 1064 principal components (PCs) from an ordination of the plant trait and chemistry polygenic scores, which represent 1064 independent (orthogonal) variables. Our goal here was to provide additional evidence that multiple, distinct genetic factors contributed to explaining caterpillar performance polygenic scores. Models based on these predictors explained 27 to 76% of the variation in the caterpillar performance scores (mean = 56.6%, cross-validation predictive r^2 ranged from 0.25 to 0.72), with an average 180 of the 1064 PCs retained in the LASSO models (range = 52 to 337) (Fig. S19).

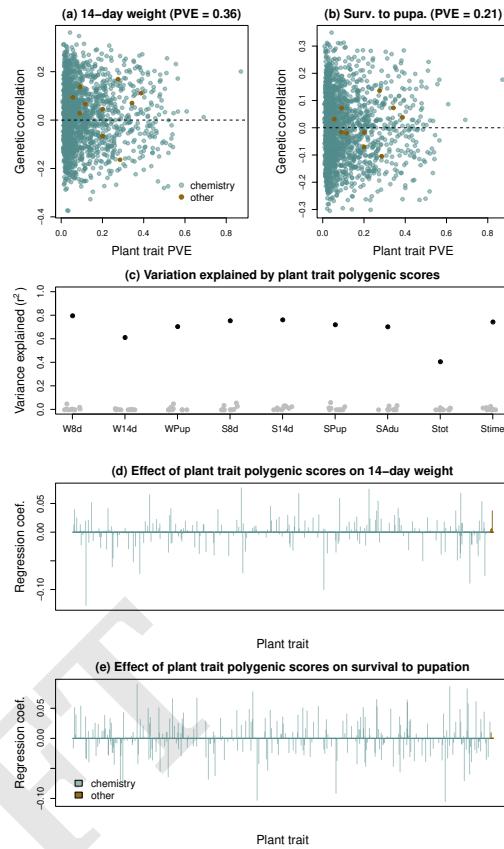
Compared to predicting polygenic scores for caterpillar performance, our ability to predict caterpillar performance at the phenotypic level from plant-trait polygenic scores was notably reduced (Table S10, Figs. S20 and S21). This was expected as plant genetics only explained a modest proportion of the variation in performance and thus the ability to explain variation in these traits (not just polygenic scores) was necessarily capped by performance-trait heritabilities. Still, when considering all performance traits together, plant trait polygenic scores explained more of the trait variation than expected by chance (Fisher combined test, $\chi^2 = 34.42$, df = 18, $P = 0.011$). This signal was driven primarily by association of plant traits with 8 and 14 day weight and survival to pupation and eclosion.

Lastly, we determined the extent to which the association of plant trait polygenic scores with caterpillar performance polygenic scores (both inferred from plant genetics) was affected by *L. melissa* genotype. Such an interaction would suggest

caterpillar performance is affected by epistatic interactions between *M. sativa* and *L. melissa* genotypes, as predicted by our hypothesis (iv) (Fig. 1). We used principal component (PC) scores from the first four principal components of the *L. melissa* genotype matrix, which together accounted for ~15% of the *L. melissa* genetic variation, as summaries of *L. melissa* genotype. We then fit LASSO penalized regression models for caterpillar performance polygenic scores as a function of these PC scores, plant trait polygenic scores, and interactions between each plant trait polygenic score and each of the four PCs. This allowed us to test for epistasis at the level of plant morphology and phytochemistry polygenic scores from *M. sativa* and four axes of *L. melissa* genetic background and thereby avoid the lack of power that would be associated with exhaustively testing SNP-SNP interactions (nonetheless, these models still included $4 \times 1760 = 7040$ possible interaction terms). We found no evidence of epistasis between *M. sativa* and *L. melissa* affecting caterpillar performance. Specifically, including these interaction terms in the models actually reduced the variance explained by the LASSO models (Table S11) and the interaction terms were retained less frequently in the models than the non-interaction terms (Figs. S22 and S23). We obtained similar results when fitting models for caterpillar performance trait values rather than polygenic scores, with a smaller proportion of interaction terms retained in the model for most traits (Figs. S24 and S25) and no overall increase in variance explained by models with versus without interactions (i.e., the variance explained in 14-day weight doubled, but the variance explained in 8-day weight was halved, and there was no detectable general increase in variance explained across traits) (Fig. S26). Thus, these results support our hypothesis (iii) with additive contributions of plant and caterpillar genetics (Fig. 1).

442 **Plant and caterpillar genetics have consistent effects on performance.** We conducted two additional experiments to determine the extent to which genetic differences among *M. sativa* plants or populations had consistent effects on caterpillar performance for different butterfly populations and species. This constitutes another test of additivity versus epistasis for plant and insect genotypes (our hypotheses (iii) versus (iv) in Fig. 443 1) and of the potential for our findings to provide general predictions beyond our main study populations. In the first of 444 these experiments, *L. melissa* (Lycaenidae) caterpillars from 445 four populations were reared on greenhouse-grown *M. sativa* sourced 446 from six sites (Table S14). Two additional butterfly species, *Colias eurytheme* (a legume specialist) (Pieridae) 447 and *Vanessa cardui* (a generalist that rarely feeds on alfalfa) 448 (Nymphalidae), were reared on these same plants. Whereas 449 only modest genetic differences exist among the *L. melissa* 450 populations (Fig. 2) (23, 24), these three butterfly species are 451 deeply divergent (\sim 100 million years) creating substantial 452 opportunities for the effect of *M. sativa* genotype and phenotype 453 to interact with genetic differences among the butterfly taxa (40). 454 Caterpillars were fed leaf tissue from multiple individual plants, but 455 each caterpillar was given plants from a single source population and 456 leaves from each plant were fed to all three butterfly species. 457 Survival rates were highest for *C. eurytheme*, followed by *L. melissa* and lastly *V. cardui* 458 (Fig. S27). Plant population (here used as a proxy for plant 459 genotype) explained \sim 3–10% of the variation in 8-day weight 460 for each butterfly species, and 9–14% of the variation in 14-day 461 weight, with larger effects in the butterfly species less-well 462 adapted to *M. sativa* (Table S15). Caterpillar population 463 explained a small but non-zero proportion of the variation in 464 8-day weight in *L. melissa* (this could not be assessed in the 465 other species), but not a significant amount of variation in 466 14-day weight. Thus, consistent with our main results above, 467 genetic differences among plant and caterpillar populations 468 (caterpillar populations for *L. melissa* only) explained variation 469 in caterpillar performance, with plant genetics mattering more 470 for 14-day weight than 8-day weight and caterpillar genetics 471 mattering more for 8-day weight than 14-day weight. Plant 472 population and plant maternal family also explained variation 473 in plant growth and development traits, consistent with our 474 common garden results above (Table S9). Importantly, the 475 effect of each plant population on caterpillar performance was 476 remarkably consistent across *L. melissa* populations and even 477 across different species, with moderate to large positive 478 correlations (though not always significantly so) in the effect of 479 each plant population on 8 and 14-day weight across all pairs 480 of population and species (Fig. S28).

490 The final complementary experiment used the same three 491 butterfly species: *L. melissa*, *C. eurytheme*, and *V. cardui*, 492 but instead involved feeding each caterpillar leaf tissue from 493 a single *M. sativa* plant from a third common garden near 494 the University of Nevada (UNR Main Station in Reno, NV; 495 Fig. S1). We used these data to ask whether the effect of 496 plant genotype (here, individual plant) on caterpillar weight 497 was consistent across species. We detected modest, positive 498 pairwise correlations between the three species of caterpillars, 499 suggesting a degree of similarity of plant genotypes that affect 500 performance of these different herbivorous species (Fig. S29). 501 Specifically, the correlations were as follows: *V. cardui* vs. *L. 502 melissa* $r = 0.33$ ($P = 0.015$, $t = 2.52$, $df = 52$); *C. eurytheme*



503 **Fig. 7.** Associations between plant trait polygenic scores and caterpillar performance 504 polygenic scores. Scatterplots show genetic correlations between plant chemistry 505 and other plant traits and 14-day caterpillar weight (a) and survival to pupation (b) 506 inferred from plant genetics as a function of the proportion of plant trait variation 507 explained by genetics (PVE). A dashed horizontal line denotes a genetic correlation of 508 zero. Panel (c) shows the variance explained by lasso regression models of caterpillar 509 performance polygenic scores estimated from plant genetics as a function of polygenic 510 scores for 1750 plant chemistry traits and 10 non-chemistry traits. Black dots denote 511 inferred values of r^2 and gray dots show similar estimates using randomized plant trait 512 polygenic scores (10 random data sets each). Panels (d) and (e) show standardized 513 regression coefficients from the lasso models for 14-day weight (d) and survival to 514 pupation (e).

515 vs. *L. melissa* $r = 0.43$ ($P = 0.0010$, $t = 3.48$, $df = 52$); *C. 516 eurytheme* vs. *V. cardui* $r = 0.15$ ($P = 0.28$, $t = 1.08$, $df = 52$). 517 Thus, these two experiments combined with our main results 518 show that genetic variation within *M. sativa* affects caterpillar 519 performance across populations and species of butterflies in a 520 remarkably consistent manner, consistent with the additivity 521 hypothesis (hypothesis (iii) in Fig. 1).

522 **Discussion.** From an ecological perspective, the greatest 523 diversity of life is not counted in the number of species or other 524 taxonomic units, but in the diversity of inter-specific interactions 525 (41). The ubiquity of plant-feeding insects has made 526 them a focal point for understanding the evolution, persistence, 527 and variability of interactions (9, 42, 43). The outcomes 528 of plant-insect interactions (e.g., caterpillar survival) might 529 depend on genetic variation within each species and these 530 genetic effects could compound additively or non-additively. 531 Taken all together, our results support the hypothesis that 532 both plant (alfalfa) and insect (Melissa blue butterfly) 533 genotype matter for caterpillar growth and survival, and that these 534

522 contributions are mostly additive (our hypothesis (iii) in Fig. 523 1). These results are qualitatively similar to those reported 524 in another study (20), which identified individual plant (*Ara- 525 bidopsis thaliana*) and caterpillar (*Pieris rapae*) genes affecting 526 caterpillar performance. The advance over previous work that 527 we offer here is in quantitative, genomic prediction of cater- 528 pillar performance, which, in contrast to the identification of 529 specific genes, provides a formal connection from trait genet- 530 ics to models of evolution for quantitative traits (44). We 531 specifically demonstrated that the combined effects of plant 532 and insect genotype explain a substantial proportion of vari- 533 ation in caterpillar growth and survival (17–49%), and that 534 these mostly-additive effects can predict performance from 535 genotypes in cross-validation analyses. Moreover, models that 536 included pairwise epistatic effects failed to explain caterpillar 537 performance better than the additive-only models. We were 538 able to identify specific traits and phytochemicals associated 539 with the plant contribution to performance, most notably plant 540 size, and several saponins, peptides, and phosphatidyl cholines. 541 Whereas some of these classes of chemicals (e.g., saponins) 542 are best known as insect toxins or feeding-deterrants (e.g., 543 (45–47)), our results suggest these classes include molecules 544 with positive and negative effects on performance, consistent 545 with other recent metabolomic work (25, 32). We also found 546 evidence that plant genotype had consistent effects on perfor- 547 mance in multiple butterfly populations and distantly related 548 species, including a second legume specialist (*C. eurytheme*) 549 and a generalist (*V. cardui*). This too is consistent with results 550 from the only other similar study (20), which documented con- 551 served changes in gene expression in response to herbivores 552 across multiple plant and butterfly species. This consistency 553 is relevant to the predictability and nature of the evolution of 554 plant-insect interactions, as we discuss more below.

555 Our results have clear implications for the study of coevolution, 556 which takes many forms and pertains to the formation of 557 new species and new interactions (43). Quantitative theories of 558 coevolution have historically been dominated by gene-for-gene 559 models, in which the fitness of a particular genetic variant 560 in (for example) a parasite is conditioned on the presence 561 of a specific gene in the host (22). Evidence in support of 562 gene-for-gene models has come mostly from plant-pathogen 563 systems (22) (but see (48)). In contrast, diffuse models of 564 coevolution relax some of the expectations for gene-by-gene 565 interactions, and have been favored by researchers working 566 with more macroscopic parasites, including herbivorous insects 567 (49). However, relevant investigations in plants and insects 568 have mostly relied on experiments that contrast categories of 569 individuals (strains or biotypes) rather than more comprehen- 570 sive or continuous variation in genetically variable populations 571 (reviewed in (12)), which has left the field with uncertainty 572 regarding the most relevant theoretical context for the diver- 573 sity of evolving plant-insect interactions. The results that 574 we report are not consistent with the gene-for-gene model of 575 coevolution, as the performance of our focal herbivore was 576 both highly polygenic and successfully predicted without in- 577 teractions between caterpillar and plant genotypes. Instead, 578 our results suggest that genetic differences in plant quality 579 and defense have similar effects regardless of insect genotype 580 or even species.

581 Our results also shed light on the evolution of diet breadth 582 and host use in herbivorous insects. Specifically, the finding

583 of substantial heritable variation in the Melissa blue butterfly 584 for growth and survival suggests that ongoing adaptation to 585 alfalfa, which at present is a marginal host (26), is not con- 586 strained by a lack of genetic variation. This is consistent with 587 earlier work on this system (28). Likewise, alfalfa appears 588 to harbor genetic variation to evolve traits that reduce the 589 success of the Melissa blue even further, and this inference 590 likely extends to other herbivores given the consistent effects 591 of plant variation on other butterfly species reported here and 592 on other herbivores in an observational study (25). While the 593 persistence of plant genetic variation affecting herbivores might 594 be attributable to the age of these interactions (since most 595 herbivores of alfalfa in North America are recent colonists), 596 we suspect other factors are more important. First, the asym- 597 metry in our predictions, with consistent caterpillar-genetic 598 effects but not plant-genetic effects on performance between 599 common gardens, suggests a major role for plasticity in the 600 effect of plant genotype on caterpillar performance. This is not 601 surprising given considerable evidence that biotic and abiotic 602 environmental factors affect plant quality and plant defenses 603 in alfalfa (31) and other plants (50), but does mean genetic 604 variation in performance measured in the lab and common 605 garden might not strongly predict effects in specific natural 606 populations (51). Moreover, other biotic and abiotic factors 607 could contribute more to caterpillar growth and survival in 608 the wild, and some of these could interact with plant genotype. 609 For example, recent work has shown that the abundance of 610 ants, which tend Melissa blue caterpillars and thereby reduce 611 the threat from enemies (see image in Fig. 1), greatly increases 612 caterpillar survival and population persistence on alfalfa, with 613 ant abundance indirectly affected by alfalfa phytochemistry 614 (25, 52). In contrast to the complexity of plant effects, the 615 more consistent effects of caterpillar genetic variants raises the 616 possibility that the ability of herbivores to successfully utilize 617 plants might more readily evolve, while the ability of plants to 618 evolve defenses will be more contingent (on local environments, 619 etc.). This again supports a diffuse model of coevolution (49) 620 and could eventually help us understand the accumulation of 621 host-specific herbivores on plants through evolutionary time.

622 Beyond issues specific to herbivorous insects and their host 623 plants, genetic variation within species is important for host- 624 parasite interactions (53), including for example susceptibility 625 to parasitic diseases in humans and other animals being a 626 function of both genetic variation in the hosts and among 627 pathogen strains (54). However, as is the case for plant-insect 628 interactions, genomic investigations of other pairwise inter- 629 actions have rarely considered both species simultaneously, 630 but have focused on either the host or parasite. If epistatic, 631 among-species interactions were common (as assumed by the 632 gene-for-gene model of coevolution), the piecewise approach 633 (focusing on one interacting species rather than the pair) might 634 be a major roadblock to progress in understanding the evolu- 635 tion of these systems. However, if additivity and consistency 636 of polygenic effects hold generally, as documented in the plant 637 and herbivores studied here, a focus on one species in an inter- 638 action might not be misleading, and might inform predictive 639 models, but this hypothesis remains to be tested with other 640 interacting species.

Materials and Methods

642 **Establishing the primary common garden.** We planted a common
643 garden comprising 1080 alfalfa plants at the Greenville Experimental
644 Farm near Logan, Utah (41.765° N, 111.814° W) in 2018 (Fig. S1a).
645 Seeds for this garden were collected from 11 naturalized alfalfa sites
646 in the western USA, including five sites where *L. melissa* are found,
647 and six sites lacking *L. melissa* butterflies (Table S1). An average
648 of 4.9 seeds were planted from each of 220 maternal plants (with
649 an average of 97.6 seeds planted from each site, $SD = 8.6$, range
650 = 77 to 105) (Table S1). See “Establishing the primary common
651 garden” in the SI for additional details.

652 **Caterpillar husbandry and performance assays.** We obtained *L.*
653 *melissa* eggs from gravid females collected from six sites between
654 June 16th and July 4th 2019 (Table S1). As in past work, gravid
655 females were caged with a few sprigs of host plant (*M. sativa*) and
656 allowed to lay eggs (17, 26, 28). Eggs were kept in a Percival in-
657 cubator (model no. 136VL) at 27°C with 14 hours light:10 hours
658 dark. Upon hatching, caterpillars were assigned randomly to feed
659 on a specific *M. sativa* plant. Each neonate caterpillar was carefully
660 transferred to a Petri dish with a sprig of fresh plant material (a
661 few leaflets) with the stem of the plant tissue wrapped in a damp
662 Kimwipe. We verified each caterpillar was alive and uninjured after
663 transfer. The Petri dish containing the caterpillar was then returned
664 to the incubator. Caterpillars were given fresh leaf tissue *ad libitum*
665 and were checked daily for survival, pupation and eclosion as adults.

666 As metrics of performance, we measured 8-day and 14-day cater-
667 pillar weight, and weight at pupation using a Mettler Toledo XPE105
668 analytical microbalance (Mettler Toledo). Weights were recorded
669 to the nearest 0.01 mg, and we took the mean of two indepen-
670 dent weight measurements. *Lycaeides melissa* caterpillars generally
671 spend 20 to 30 days as larvae (17), and weight and lifetime fecundity
672 are highly correlated in *L. melissa* (26). We then considered the
673 following nine performances metrics: 8-day caterpillar weight (mg),
674 14-day caterpillar weight (mg), weight at pupation (mg), survival to
675 8 days (binary), survival to 14 days (binary), survival to pupation
676 (binary), survival to adult (binary), total survival time (integer val-
677 ued), and truncated survival time (integer valued). For truncated
678 survival time, we truncated survival at the maximum number of
679 days required for any of the caterpillars to reach eclosion; this avoids
680 caterpillars that developed slowly but never pupated or eclosed from
681 having the longer survival times than caterpillar that successfully
682 eclosed as adults.

683 **Generating the genetic data.** We extracted DNA from 1236 *M. sativa*
684 plants and 1079 *L. melissa* caterpillars, pupae or adults reared on
685 these plants. We then generated partial genome sequences for each
686 organism using our genotyping-by-sequencing approach (23, 55);
687 this produced \sim 2.5 billion reads for *M. sativa* and \sim 2.5 billion
688 reads for *L. melissa* (see “DNA extraction and sequencing” in the
689 SI). We then aligned the DNA sequences to the *M. sativa* or *L.*
690 *melissa* genome and identified SNPs using **samtools** (versions 1.10),
691 **bcftools** (version 1.9) and **GATK** (version 4.1) (56, 57) (see “DNA
692 sequence alignment and variant calling” in the SI). After filtering, we
693 retained 161,008 SNPs for *M. sativa* and 63,194 SNPs for *L. melissa*.
694 We then estimated genotypes using the Bayesian (ad)mixture model
695 implemented in **entropy** (version 2.0) (23, 58) (see “Inference of
696 genotypes and genetic variation” in the SI). Patterns of genetic
697 variation were then summarized with principal component analysis
698 (PCA), and by calculating measures of linkage disequilibrium (LD)
699 and genetic differentiation among samples from different source
700 populations (i.e., F_{ST}) (see “Inference of genotypes and genetic
701 variation” in the SI).

702 **Preparing the caterpillar performance data for genetic mapping.** We
703 removed potential confounding variation from the caterpillar per-
704 formance data prior to analyzing genotype-performance associations.
705 First, we regressed each of the nine caterpillar performance met-
706 rics on caterpillar hatch date (to control for temporal effects) and
707 source population (to control for potential non-genetic, e.g., mater-
708 nal environment, effects). This was done with the **lm** function in R.
709 Next, we used distance-based Moran’s eigenvector maps to remove
710 possible effects of space (location) within the common garden. This
711 procedure involves creating spatial variables based on a PCA of
712 a truncated (nearest neighbors) Euclidean distance matrix (i.e., a
713 principal coordinates analysis), where distance was defined from the

714 spatial layout of the common garden (59). We then used forward
715 selection of variables following (60) to select spatial variables (eigen-
716 vectors) that explained the variation in each trait. Specifically, we
717 first tested for a significant (at $P < 0.05$) fit of a model with all of
718 the spatial variables. If and only if this full model was significant,
719 we began adding spatial variables to a null model one at a time
720 based on the extent to which they increased the total model r^2 .
721 This procedure continued until either: (i) the P -value for the most
722 recently added variable was > 0.05 , (ii) the total r^2 exceeded the
723 original r^2 from the full model with all variables, (iii) adding the
724 new variable did not increase the model r^2 , or (iv) 200 spatial
725 covariates had been added. The final models explained 18 to 51%
726 of the variation in plant traits (mean = 35%) with 22 to 77 covariates
727 retained; however, a model with no spatial covariates was selected
728 for most caterpillar performance traits with 14-day weight being the
729 sole exception (20 covariates explaining 14% of the trait variation).
730 Scaled residuals from the final model for each trait were then used
731 for genetic mapping.

732 **Multilocus genetic mapping of caterpillar performance.** We tested for
733 associations between (i) *M. sativa* SNPs (161,008 SNPs), (ii) *L.*
734 *melissa* SNPs (63,194 SNPs), and (iii) SNPs from both species
735 combined (224,202 SNPs), and each of the nine caterpillar per-
736 formance metrics (i.e., the residuals from the models described in
737 the previous paragraph). We performed these analyses using Bayesian
738 sparse linear mixed models (BSLMMs), which we fit with **gemma**
739 (version 0.95alpha) (61). A key advantage of this approach for
740 genotype-phenotype association analyses is that, unlike traditional
741 genome-wide association (GWA) mapping methods that test each
742 genetic marker separately, the BSLMM approach fits all SNPs in a
743 single model and thus mostly avoids issues related to testing large
744 numbers of null hypotheses. The BSLMM method assumes that
745 trait values are determined by a polygenic term and a vector of the
746 (possible) measurable effects of each SNP on the trait (β) (61).
747 Bayesian Markov chain Monte Carlo (MCMC) with variable selec-
748 tion is used to infer the posterior inclusion probability (PIP) for each
749 SNP, that is, the probability that each SNP has a non-zero effect or
750 association, and the effect size conditional on it being non-zero (62).
751 The polygenic term denotes each individual’s expected deviation
752 from the mean phenotype based on all of the SNPs. This term ac-
753 counts for phenotypic covariances among individuals caused by their
754 relatedness or overall genetic similarity (61). The kinship matrix
755 also serves to control for population structure and relatedness when
756 estimating effects of individual SNPs (β) along with their PIPs.
757 Similarly, SNPs in linkage disequilibrium (LD) with the same causal
758 variant effectively account for each other, such that only one or the
759 other is needed in the model, and this redundancy is captured by
760 the posterior inclusion probabilities. Moreover, in the context of our
761 study, mapping with plants grown from seed in a common garden
762 and caterpillars reared from eggs in growth chambers substantially
763 reduces some issues related to the confounding effects of population
764 structure, such as genotype-environment correlations, that com-
765 monly cause problems in human association-mapping studies (63)
766 and more generally in observational studies of human genetics (64).
767

768 The hierarchical structure of the model makes it possible to estimate
769 additional parameters that describe aspects of a trait’s genetic
770 architecture (17, 61, 62, 65). These include the percentage of
771 the phenotypic variance explained (PVE) by additive genetic effects
772 (which includes β and the polygenic term, and should approach the
773 narrow-sense heritability), the percentage of the PVE due to SNPs
774 with measurable effects or associations (PGE, the percentage of the
775 phenotypic variance explained by genic effects, which is based only
776 on β), and the number of SNPs with measurable associations ($n\gamma$).
777 All of these metrics use MCMC to integrate over uncertainty in the
778 effects of individual SNPs, including whether these are non-zero.
779 Lastly, using this BSLMM approach, it is also possible to obtain
780 genomic-estimated breeding values (GEBVs) or polygenic scores,
781 that is, the expected trait value for an individual from the additive
782 effects of their genes, as captured by both β and the polygenic term
783 (17, 65).

784 For each of the nine caterpillar performance metrics and three
785 genetic data sets, we conducted 10 MCMC runs with **gemma**, each
786 comprising 1 million iterations and a 200,000 iteration burn-in.
787 Every 10th MCMC sample was retained to form the posterior
788 distribution. Polygenic scores (i.e., genomic-estimated breeding
789

values) were then calculated from the genetic data sets and model- averaged effect estimates for each SNP locus; these incorporate the polygenic term as is standard in genomic prediction methods (e.g., (66, 67)). Genetic covariance matrixes were computed from the estimated polygenic scores.	862 863
As noted above, the kinship matrix and multilocus approach of the BSLMM in gemma control for confounding effects of population structure and relatedness when testing for individual SNP- phenotype association, but nonetheless, this methods could fail to fully capture complex patterns of structure (see, e.g., (63)). Thus, to verify the robustness of our results, we fit additional models using the BSLMM approach in gemma where we included the first 20 genetic PCs as potential covariates to further account for population structure. This was done as described above, except that the analysis was only conducted for <i>M. sativa</i> and <i>L. melissa</i> SNPs separately. We compared this to our main results both in terms of the percentage of variation in performance explained by genetics (PVE) and specific SNP-performance trait associations.	864 865 866 867 868 869 870 871 872 873 874 875 876 877 878 879
Direct tests of epistatic genetic effects on caterpillar performance. We tested for epistatic interactions affecting caterpillar performance among (i) the 161,008 <i>M. sativa</i> SNPs, (ii) the 63,194 <i>L. melissa</i> SNPs, and (iii) the 224,202 SNPs from both species (this includes within and between species epistatic interactions). We conducted these tests with MAPIT (https://github.com/loriananthony/MAPIT) (68). Exhaustive testing of all pairwise SNP-SNP interactions suffers from low statistical power because of the large number of tests involved. The statistical method in MAPIT overcomes the problem of low power by instead testing for marginal epistatic effects, that is testing the null hypothesis that a given SNP does not interact with any of the other SNPs (i.e., that the variance component for epistatic effects is 0) (68). This is done without trying to identify the specific SNPs a focal SNP interacts with. We computed <i>P</i> -values for tests of marginal epistasis using the recommended hybrid method that first implements a <i>z</i> -test to compute a <i>P</i> -value and then re-computes the <i>P</i> -value with the Davies method if the initial values is less than 0.05 (as in (18)).	880 881 882 883 884 885 886 887 888 889 890 891 892 893 894 895 896 897
For many of the survival traits, we observed an unexpected excess of very low <i>P</i> -values, especially for <i>L. melissa</i> SNPs and for 8 and 14-day survival (Fig. S30). We strongly suspect this is a statistical artefact, especially as these measures constitute residuals from integer or binary traits and the control kinship matrix consists of relatedness based on plant and insect genetics, a combination of complications that could be problematic for this method and inflate type-I errors (note that this differs from the BSLMM in gemma where the multilocus approach allows SNPs to serve as controls for each other). Given our concern that these results are not biologically meaningful, we conservatively focus only on the weight measurements when presenting these tests of epistasis, as these do not appear to suffer from the same issue (see Figs. S12 and S13).	898 899 900 901 902 903 904 905 906 907 908
Even with the MAPIT method, a potential exists for tests of epistasis to be underpowered, especially in terms of achieving strict, genome-wide significance. Thus, we conducted additional analyses using the BSLMM approach from gemma to test for associations between <i>M. sativa</i> and <i>L. melissa</i> genetics and caterpillar performance but where we included pairwise epistatic effects among SNPs with the most evidence of marginal epistasis from the MAPIT analyses (similar to (18)). Our goal was to ask whether including these additional epistatic terms improved the explanatory power of the model. In these analyses, we considered only the caterpillar weight traits (for the reasons noted above). We included either (i) the top 150 SNPs with the lowest <i>P</i> -values for marginal epistasis within species (for analyses with only <i>M. sativa</i> or <i>L. melissa</i> SNPs) or (ii) the top 75 SNPs from each species with the lowest <i>P</i> -values for marginal epistasis in the combined species analysis. We then created new genetic covariates for all pairwise interactions between pairs of the 150 SNPs ($\frac{150 \times 149}{2} = 11,175$ potential epistatic effects). We did this by taking the product of the centered and standardized genotypes for each pair of loci. These were then included in the BSLMM model for gemma (though not in the construction of the kinship matrix, which was solely based on additive effects). We fit these models as described above, except we increased the number of MCMC iterations and burnin to 2 million and 400,000, respectively. We then determined the total PVE in weight explained by the models with additive and epistatic effects for <i>M. sativa</i> genetics, <i>L. melissa</i> genetics, and both <i>M. sativa</i> and <i>L. melissa</i> genetics combined.	909 910 911 912 913 914 915 916 917 918 919 920 921 922
Within-garden cross-validation and genomic prediction. We used 10- fold cross-validation to assess our ability to predict performance traits from <i>M. sativa</i> genetic data, <i>L. melissa</i> genetic data, and the combined genetic data from <i>M. sativa</i> and <i>L. melissa</i> . To do this, we first randomly assigned each observation to one of ten test data sets. Then, for each test data set, we estimated genotype- phenotype associations using gemma as described above, but based only on the 90% of individuals not in that test data set. For this, we used a single MCMC run comprising 1 million iterations, a 200,000 iteration burn-in, and a thinning interval of 10. We then used gemma to predict the phenotypes of the 10% of individuals held back for the test set (these individuals were not used to fit the model); this was done with the predict option in gemma . We then quantified predictive performance using the Pearson correlation between the genomic predictions of each performance metric and the observed values.	864 865 866 867 868 869 870 871 872 873 874 875 876 877 878 879
Gene Miller Life Science Garden set up and genomic prediction. We further tested our ability to predict caterpillar performance trait values from genotypes by generating genomic predictions of performance for caterpillars reared on <i>M. sativa</i> from a second, smaller common garden comprising 180 <i>M. sativa</i> —The Gene Miller Life Science Garden (see “Establishing the Gene Miller Life Science Garden” in the SI for details). We used leaf tissue from these plants for rearing <i>L. melissa</i> caterpillars in the summer of 2019 exactly as described for the main common garden at the Greenville Experimental Farm (see ‘Caterpillar husbandry and performance assays’ above for details). This parallel experiment was conducted at the same time as the main experiment. Plant and caterpillar samples from this parallel experiment were sequenced along with the samples from the Greenville Experimental Farm experiment. We successfully obtained genetic data from 172 <i>M. sativa</i> and 156 caterpillars of the 180 involved in this experiment. These genetic data were processed along with those from the main garden (see ‘DNA sequence alignment and variant calling’ above for details).	880 881 882 883 884 885 886 887 888 889 890 891 892 893 894 895 896 897
We then used the estimated, model-averaged effects from the BSLMM fits in gemma from the main garden to predict performance traits based on plant, caterpillar, or plant and caterpillar genotypes for these individuals. We compared these genomic predictions (i.e., polygenic scores computed from the main-garden models) to the observed performance trait values for these caterpillars. This was done using residuals after removing effects of hatch date and block (i.e., plot) within the USU garden. As with the within- garden cross-validation analyses described in the previous section, predictive power was measured by the Pearson correlation between the predicted and observed performance trait values.	898 899 900 901 902 903 904 905 906 907 908
Plant trait measurements and phytochemical analysis. We measured a series of morphological traits potentially associated with plant vigor or resistance to insects (e.g., putative structural plant defenses) (17, 69, 70) for each of the 1080 <i>M. sativa</i> plants in the Greenville Experimental Farm common garden: plant height, leaf length, leaf width, leaf area, leaf shape, leaf weight, specific leaf area, leaf toughness, trichome density, levels of herbivory on the plants in the field, and 1750 plant chemistry metabolites, which were quantified and characterized using liquid-chromatography combined with mass spectrometry (LC-MS). See “Plant trait measurements” and “Sample extraction and phytochemical analysis” in the SI for details. We further annotated the 20 phytochemicals that were most strongly associated with caterpillar performance (see “Structural annotations of phytochemicals” in the SI).	909 910 911 912 913 914 915 916 917 918 919 920 921 922
Multilocus genetic mapping of plant traits. We tested for associations between the <i>M. sativa</i> SNPs (161,008 SNPs) and 1760 plant traits: leaf length, leaf width, leaf area, leaf shape, leaf weight, SLA, trichome density, leaf toughness, plant height, field herbivory and 1750 metabolomic chemical features (see the previous two sections for details). This was done using the 1080 <i>M. sativa</i> plants from the main common garden at the Greenville Experimental Farm in Logan, UT. We first removed possible effects of spatial location within the garden as captured by distance-based Moran’s eigenvector maps using forward selection of variables (60), exactly as described for the	923 924 925 926 927 928 929 930 931 932

933 caterpillar performance traits above (see ‘Preparing the caterpillar
 934 performance data for genetic mapping’). As with the caterpillar
 935 performance traits, genotype-plant trait associations were estimated
 936 by fitting BSMMs with **gemma** (version 0.95alpha) (61). For each
 937 of the 1760 plant traits, we conducted 10 MCMC runs with **gemma**,
 938 each comprising 1 million iterations and a 200,000 iteration burn-
 939 in. Every 10th MCMC sample was retained to form the posterior
 940 distribution. Polygenic scores were then calculated from the genetic
 941 data sets and model-averaged effect estimates for each SNP locus.
 942 Genetic covariance matrixes were computed from the estimated
 943 polygenic scores. The model-fitting procedure was repeated with
 944 1760 randomized plant trait data sets (i.e., values of each of the
 945 original traits were permuted among plants) to verify that the
 946 distribution genotype-phenotype associations from the real data set
 947 differed from null expectations.

948 **LASSO regression models.** We used least absolute shrinkage and
 949 selection operator (LASSO) regression to (i) identify the subset of
 950 plant traits with polygenic scores that best predicted caterpillar-
 951 performance polygenic scores and (ii) estimate the direction and
 952 magnitude of these associations (see “LASSO regression models”
 953 in the SI). We fit additional LASSO models (i) using PCs of the
 954 1760 plant trait polygenic scores as covariates and (ii) to deter-
 955 mine whether plant-trait polygenic scores could explain and predict
 956 caterpillar performance at the phenotypic level. Lastly, we fit an
 957 additional models to evaluate the extent to which plant-genetic
 958 effects interacted with caterpillar genetics to affect performance (see
 959 “LASSO regression models” in the SI).

960 **Complementary USU greenhouse and Nevada common garden rear- 961 ing experiments.** An additional rearing experiment was conducted
 962 using *M. sativa* grown in a USU greenhouse to (i) replicate the general
 963 effect of *M. sativa* genotype on caterpillar performance and (ii)
 964 determine whether different plant genotypes had consistent effects
 965 of caterpillar performance across different butterfly populations
 966 and species (i.e., *Colias eurytheme* and *Vanessa cardui*). We per-
 967 formed yet another rearing experiment with the same three species
 968 of caterpillars using an experimental *M. sativa* garden experimental
 969 garden at the University of Nevada, Reno (Fig. S1). Together, these
 970 experiments provide additional tests of additivity versus epistasis
 971 with respect to genetic differences among butterfly populations and
 972 among deeply divergent species. See “Complementary USU green-
 973 house experiment” and “Complementary Nevada common garden
 974 rearing experiment” in the SI for details.

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