

Genome-wide Association Study identifies two novel loci for Gilles de la Tourette Syndrome

Fotis Tsetsos¹, Apostolia Topaloudi², Pritesh Jain², Zhiyu Yang², Dongmei Yu^{3,4}, Petros Kolovos¹, Zeynep Tumer^{5,6}, Renata Rizzo⁷, Andreas Hartmann⁸, Christel Depienne⁹, Yulia Worbe^{10,11}, Kirsten R. Müller-Vahl¹², Danielle C. Cath¹³, Dorret I. Boomsma^{14,15}, Tomasz Wolanczyk¹⁶, Cezary Zekanowski¹⁷, Csaba Barta¹⁸, Zsolia Nemoda¹⁸, Zsanett Tarnok¹⁹, Shanmukha S. Padmanabhuni²⁰, Joseph D. Buxbaum^{21,22,23,24,25,26}, Dorothy Grice^{21,25,26,27}, Jeffrey Glennon²⁸, Hreinn Stefansson²⁹, Bastian Hengerer³⁰, Evangelia Yannaki^{31,32}, John A. Stamatoyannopoulos^{33,34,35}, Noa Benaroya-Milshtein^{36,37}, Francesco Cardona³⁸, Tammy Hedderly³⁹, Isobel Heyman^{40,41}, Chaim Huyser^{42,43}, Pablo Mir^{44,45}, Astrid Morer^{46,47,48}, Norbert Mueller⁴⁹, Alexander Münchau⁵⁰, Kerstin J. Plessen^{51,52}, Cesare Porcelli⁵³, Veit Roessner⁵⁴, Susanne Walitza⁵⁵, Anette Schrag⁵⁶, Davide Martino⁵⁷, The TSAICG, The TSGeneSEE initiative, The EMTICS collaborative group, The TS-EUROTRAIN network, The TIC Genetics collaborative group, Jay A. Tischfield⁵⁸, Gary A. Heiman⁵⁸, A. Jeremy Willsey^{59,60}, Andrea Dietrich⁶¹, Lea K. Davis^{62,63}, James Crowley^{64,65,66}, Carol A. Mathews⁶⁷, Jeremiah M. Scharf^{9,4,68}, Marianthi Georgitsi^{1,69}, Pieter J. Hoekstra^{*61}, and Peristera Paschou^{*2}

¹Department of Molecular Biology and Genetics, Democritus University of Thrace, Alexandroupolis, Greece

²Department of Biological Sciences, Purdue University, West Lafayette, IN, USA

³Psychiatric and Neurodevelopmental Genetics Unit, Center for Genomic Medicine, Department of Psychiatry, Massachusetts General Hospital, Boston, MA, USA

⁴Stanley Center for Psychiatric Research, Broad Institute of MIT and Harvard, Cambridge, MA, USA

⁵Department of Clinical Genetics, Kennedy Center, Copenhagen University Hospital, Rigshospitalet, Denmark

⁶Department of Clinical Medicine, Faculty of Health and Medical Sciences, University of Copenhagen

⁷Child and Adolescent Neurology and Psychiatry, Department of Clinical and Experimental Medicine, University of Catania, Catania, Italy

⁸Department of Neurology, Hôpital de la Pitié-Salpêtrière, Paris, France

⁹Institute for Human Genetics, University Hospital Essen, Essen, Germany

¹⁰Assistance Publique Hôpitaux de Paris, Hôpital Saint Antoine, Paris France

¹¹French Reference Centre for Gilles de la Tourette Syndrome, Groupe Hospitalier Pitié-Salpêtrière, Paris, France

¹²Department of Psychiatry, Social psychiatry and Psychotherapy, Hannover Medical School, Hannover,

Germany

- ¹³ *Department of Psychiatry, University Medical Center Groningen & Rijksuniversiteit Groningen; MHC Drenthe Assen, the Netherlands*
- ¹⁴ *Institute for Anatomy and Cell Biology, Ulm University, Ulm, Germany*
- ¹⁵ *EMGO+ Institute for Health and Care Research, VU University Medical Centre, Amsterdam, Netherlands*
- ¹⁶ *Department of Child Psychiatry, Medical University of Warsaw, Warsaw, Poland*
- ¹⁷ *Laboratory of Neurogenetics, Department of Neurodegenerative Disorders, Mossakowski Medical Research Institute, Polish Academy of Sciences, Warsaw, Poland*
- ¹⁸ *Department of Molecular Biology, Institute of Biochemistry and Molecular Biology, Semmelweis University, Budapest, Hungary*
- ¹⁹ *Vadaskert Clinic for Child and Adolescent Psychiatry, Hungary*
- ²⁰ *Department of Management, Ball State University, Muncie, IN, USA*
- ²¹ *Department of Psychiatry, Icahn School of Medicine at Mount Sinai, USA*
- ²² *Seaver Autism Center for Research and Treatment, Icahn School of Medicine at Mount Sinai, USA*
- ²³ *Department of Genetics and Genomic Sciences, Icahn School of Medicine at Mount Sinai, USA*
- ²⁴ *Department of Neuroscience, Icahn School of Medicine at Mount Sinai, USA*
- ²⁵ *The Mindich Child Health and Development Institute, Icahn School of Medicine at Mount Sinai, USA*
- ²⁶ *Friedman Brain Institute, Icahn School of Medicine at Mount Sinai, USA*
- ²⁷ *Division of Tics, OCD, and Related Disorders, Icahn School of Medicine at Mount Sinai, USA*
- ²⁸ *Department of Cognitive Neuroscience, Donders Institute for Brain, Cognition and Behaviour, Radboud University Medical Center, Netherlands*
- ²⁹ *deCODE Genetics/Amgen, Iceland*
- ³⁰ *Boehringer Ingelheim Pharma GmbH & Co. KG, CNS Research, Germany*
- ³¹ *Hematology Department- Hematopoietic Cell Transplantation Unit, Gene and Cell Therapy Center, George Papanikolaou Hospital, Greece*
- ³² *Department of Medicine, University of Washington, WA, USA*
- ³³ *Altius Institute for Biomedical Sciences, WA, USA*
- ³⁴ *Department of Genome Sciences, University of Washington, WA, USA*
- ³⁵ *Department of Medicine, Division of Oncology, University of Washington, WA, USA*
- ³⁶ *Child and Adolescent Psychiatry Department, Schneider Children's Medical Centre of Israel, Petah-Tikva*
- ³⁷ *Sackler Faculty of Medicine, Tel Aviv University, Israel*
- ³⁸ *Department of Human Neurosciences, University La Sapienza of Rome, Rome, Italy*
- ³⁹ *Evelina London Children's Hospital GSTT, Kings Health Partners AHSC, London, UK*
- ⁴⁰ *UCL Great Ormond Street Institute of Child Health, University College London, London, UK*
- ⁴¹ *Psychological and Mental Health Services, Great Ormond Street Hospital for Children NHS Foundation Trust, London, UK*
- ⁴² *Levvel, Academic Center for Child and Adolescent Psychiatry, Amsterdam, The Netherlands*
- ⁴³ *Amsterdam UMC, Department of Child and Adolescent Psychiatry, Amsterdam, The Netherlands*
- ⁴⁴ *Unidad de Trastornos del Movimiento. Instituto de Biomedicina de Sevilla (IBiS). Hospital Universitario*

Virgen del Rocío/CSIC/Universidad de Sevilla. Seville, Spain

⁴⁵ *Centro de Investigación Biomédica en Red sobre Enfermedades Neurodegenerativas (CIBERNED), Madrid, Spain*

⁴⁶ *Department of Child and Adolescent Psychiatry and Psychology, Institute of Neurosciences, Hospital Clinic Universitari, Barcelona, Spain*

⁴⁷ *Institut d'Investigacions Biomèdiques August Pi i Sunyer (IDIBAPS), Barcelona, Spain*

⁴⁸ *Centro de Investigación en Red de Salud Mental (CIBERSAM), Instituto Carlos III, Spain*

⁴⁹ *Department of Psychiatry and Psychotherapy, University Hospital, LMU Munich, Munich, Germany*

⁵⁰ *Institute of Systems Motor Science, University of Lübeck, Lübeck, Germany*

⁵¹ *Child and Adolescent Mental Health Centre, Mental Health Services, Capital Region of Denmark and University of Copenhagen, Copenhagen, Denmark*

⁵² *Division of Child and Adolescent Psychiatry, Department of Psychiatry, Lausanne University Hospital, Lausanne, Switzerland*

⁵³ *ASL BA, Maternal and Childhood Department; Adolescence and Childhood Neuropsychiatry Unit; Bari, Italy*

⁵⁴ *Department of Child and Adolescent Psychiatry, Faculty of Medicine, University Hospital Carl Gustav CarusTU Dresden, Dresden, Germany*

⁵⁵ *Department of Child and Adolescent Psychiatry and Psychotherapy, University of Zurich, Zurich, Switzerland*

⁵⁶ *Department of Clinical and Movement Neurosciences, UCL Queen Square, Institute of Neurology, University College London, London, UK*

⁵⁷ *Department of Clinical Neurosciences, Cumming School of Medicine & Hotchkiss Brain Institute, University of Calgary, Calgary, AB, Canada*

⁵⁸ *Department of Genetics and the Human Genetics Institute of New Jersey, Rutgers, the State University of New Jersey, Piscataway, NJ, USA*

⁵⁹ *Department of Psychiatry and Behavioral Sciences, UCSF Weill Institute for Neurosciences, University of California, San Francisco, San Francisco, CA, USA*

⁶⁰ *Quantitative Biosciences Institute, University of California, San Francisco, San Francisco, CA, USA*

⁶¹ *University of Groningen, University Medical Centre Groningen, Department of Child and Adolescent Psychiatry, Groningen, the Netherlands, Accare Child Study Center*

⁶² *Division of Genetic Medicine, Department of Medicine Vanderbilt University Medical Center Nashville, Nashville, TN, USA*

⁶³ *Vanderbilt Genetics Institute, Vanderbilt University Medical Center, Nashville, TN, USA*

⁶⁴ *Department of Genetics, University of North Carolina at Chapel Hill, Chapel Hill, NC, USA*

⁶⁵ *Department of Clinical Neuroscience, Karolinska Institutet, Stockholm, Sweden*

⁶⁶ *Department of Psychiatry, University of North Carolina at Chapel Hill, Chapel Hill, NC, USA*

⁶⁷ *Department of Psychiatry and Genetics Institute, University of Florida College of Medicine, USA*

⁶⁸ *Department of Neurology, Brigham and Women's Hospital, and the Department of Neurology, Massachusetts General Hospital, Boston, MA, USA*

⁶⁹ *1st Laboratory of Medical Biology-Genetics, School of Medicine, Aristotle University of Thessaloniki, Thessaloniki, Greece*

Abstract

Tourette Syndrome (TS) is a childhood-onset neurodevelopmental disorder of complex genetic architecture, characterized by multiple motor tics and at least one vocal tic persisting for more than one year. We performed a genome-wide meta-analysis integrating a novel TS cohort with previously published data, resulting in a sample size of 6,133 TS individuals and 13,565 ancestry-matched controls. We identified a genome-wide significant locus on chromosome 5q15 and one array-wide significant locus on chromosome 2q24.2. Integration of eQTL, Hi-C and GWAS data implicated the *NR2F1* gene and associated lncRNAs within the 5q15 locus, and the *RBMS1* gene within the 2q24.2 locus. Polygenic risk scoring using previous GWAS results demonstrated statistically significant ability to predict TS status in the novel cohort. Heritability partitioning identified statistically significant enrichment in brain tissue histone marks, while polygenic risk scoring on brain volume data identified statistically significant associations with right and left putamen volumes. Our work presents novel insights in the neurobiology of TS opening up new directions for future studies.

Introduction

Tourette Syndrome (TS) is a childhood-onset neurodevelopmental disorder characterized by multiple motor tics and at least one vocal tic persisting for more than one year [1]. The prevalence of TS is estimated in the range of 0.6-1% in school-aged children [2, 3]. It is a highly heritable disorder [4] with a population-based heritability estimated at 0.7 [5, 6] and SNP-based heritability estimates ranging from 0.21 [7] to 0.58 [4] of the total heritability. TS exhibits high polygenicity and its genetic background is influenced by both common and rare variants of small effect spread throughout the genome [4, 8, 9]. Two previously conducted genome-wide association studies (GWAS) [7, 10] have indicated enrichment of TS genetic susceptibility variants in tissues within the cortico-striatal and cortico-cerebellar circuits, and in particular, the dorsolateral prefrontal cortex [7, 10]. Furthermore, gene set analyses of GWAS data implicated ligand-gated ion channel signaling, lymphocytic, and cell adhesion and trans-synaptic signaling processes as potential biological underpinnings in the pathogenesis of TS [11]. Polygenic risk scores derived from the second TS GWAS can predict tic presence and severity at a statistically significant level [7, 12].

Here, present a genome-wide meta-analysis for TS integrating novel and previously published data resulting in a combined sample size of 6,133 TS individuals and 13,565 ancestry-matched controls. We identify a novel genome-wide significant locus in the novel (TS-EUROTRAIN) GWAS, and a second novel array-wide significant locus in the TS GWAS meta-analysis. These results provide further insight into the genetic basis of TS.

Methods and Materials

Datasets

The TS-EUROTRAIN dataset brings together three major TS cohorts, including 632 participants from the European Multicenter Tics in Children Study (EMTICS) [13], 763 participants from the TS-EUROTRAIN study [14], 238 participants from the TSGeneSEE study [15], and 52 participants from Sweden. These studies included participants from multiple European sites who were diagnosed using DSM-5 criteria for TS, consistent with previously published TS studies. In total, we collected samples from 1,685 individuals with TS (Supplementary Table 1). Additionally, 4,454 population control individuals were recruited from Ashkenazi Jewish, Greek, Hungarian, Polish, and Spanish sites. Ancestry-matched controls were also used from the following public datasets following appropriate approvals: British WTCCC2 1958 Birth Cohort samples (Study accession code: EGAS00000000028),

German control samples obtained from the POPGEN biobank [16], and French controls from the Three City Study [17], leading to a total of 8,558 general population controls (Supplementary Table 1). Written informed consent was obtained from all participants, as approved by the ethics committees of all participating institutions.

Genotyping, merging and imputation

Samples from the TS-EUROTRAIN dataset were genotyped on the Illumina HumanOmniExpress BeadChip. The control samples obtained from collaborators and public repositories were genotyped on multiple Illumina arrays; they were selected on the basis of maximizing marker overlap and compatibility (Supplementary Table 1). We applied standard GWAS quality control procedures to our data before and after the imputation, as described in previous GWAS performed by the Psychiatric Genomics Consortium (PGC). Quality control procedures included the removal of samples that fit any of the following criteria: call rate < 0.98 , absolute value of inbreeding coefficient < 0.2 , genomic sex discrepancy with reported sex, and formation of pairs with relatedness > 0.1875 . We applied variant quality control, excluding markers with call rate < 0.98 , differential missingness between cases and controls < 0.02 , Hardy-Weinberg equilibrium P value $< 1e-6$ in controls and $< 1e-10$ in cases.

Since we merged data from multiple sources, we performed the above quality control steps on each dataset separately. Imputation was performed on the Sanger Imputation Server using a reference panel of 64,940 European ancestry haplotypes (v1.1) from the Haplotype Reference Consortium (HRC) [18, 19]. We performed batch effect tests in samples of same status (case/control) between different sources, excluding markers that achieved a p-value $< 1e-5$. X chromosome data were absent from a significant fraction of the external datasets obtained; it was excluded from the final analysis. As a last step, in order to avoid ancestry bias, we matched the ancestry of the controls to TS individuals, at a three to one ratio, using the first five principal components as basis. Imputation and quality control resulted in a dataset of 1,438 individuals with TS and 4,356 controls on 2,949,675 markers.

Genome Wide Association Study

We used TS diagnosis as a categorical variable, and conducted a GWAS using an additive logistic regression model on the best guess genotypes produced by imputation. For the logistic regression we incorporated the principal components identified by Tracy-Widom statistics, as calculated by EIGENSOFT [20, 21], as well as sex and imputation batch. To control genomic inflation at SNPs with low minor allele frequency (MAF), we excluded SNPs when their MAF was $< 1\%$ and minor allele

count was <10 , in either cases or controls. We set the level of genome-wide significance at $P = 5e-8$. We used a basis of genomic distance $> 500kb$ and linkage disequilibrium (LD) to define independence between the associated loci.

We estimated confounding bias in our results by performing LD Score Regression with *ldsc* [22] and using the attenuation ratio, as well as the p-value of the intercept to evaluate our results. Results were plotted using *matplotlib* in Python. We produced the regional plots for the top loci using the Python package *region-plot* [23], which allows the use of local LD reference panels since our data were not imputed on the 1000 Genomes reference.

Meta-analysis

We conducted a meta-analysis with the results of the second TS GWAS study conducted by the TS Working Group of the Psychiatric Genomics Consortium (TSGWAS2) [7]. There was a known sample overlap between the TS-EUROTRAIN GWAS and TSGWAS2, and verified through genotypic identity-by-descent analysis, so the TS-EUROTRAIN GWAS was re-analyzed after excluding the overlapping samples, leading to a sample size of 1,314 cases and 4,077 ancestry-matched controls. The summary statistics of the re-analyzed TS-EUROTRAIN study and the TSGWAS2 meta-analysis were used as input to METASOFT [24]. METASOFT implements an array of methods for meta-analysis, especially in the case of heterogeneity in the results. In our study, we employed Han and Eskin's random effects model (RE2), which separates hypothesis testing from effect-size estimation, and is demonstrated to increase statistical power under heterogeneity compared to the conventional random effects model [25, 26]. We also employed METASOFT to produce m-values, that is, estimates of the posterior probability that an effect exists, with small values indicating absence of effect, large values indicating presence of effect, and intermediate values indicating ambiguity [24]. Results were plotted using *matplotlib* and *region-plot* [23] in Python.

Heritability and heritability partitioning

Heritability was estimated through LD Score Regression [22], after merging the alleles with the HapMap3 reference panel. We further investigated heritability partitioning into functional categories using stratified LD Score Regression [27]. TS heritability was partitioned into 53 functional categories as well as into 220 cell-type-specific and 10 cell-type-group-specific annotations produced on the data derived from the Roadmap Epigenomics Project [28]. The significance threshold for the heritability enrichments was defined at a Benjamini-Hochberg FDR < 0.05 . P-values were calculated

104 on the regression coefficient τ_c .

105 Genetic correlations

106 Bivariate LD score regression [22] was conducted to identify genetic correlations between the TS-
107 EUROTRAIN GWAS, the TS EUROTRAIN/GWAS2 meta-analysis, and TSGWAS2 [7]. We then exam-
108 ined each of these studies' cross-disorder correlations with obsessive-compulsive disorder (OCD)
109 [29], attention deficit hyperactivity disorder (ADHD) [30], major depressive disorder (MDD) [31],
110 autism spectrum disorders (ASD) [32], and anxiety [33]. To avoid confounding due to sample size,
111 we selected summary statistics from studies with more than 5,000 samples. For the correlation anal-
112 ysis we used the European LD scores and merged alleles based on the HapMap3 reference panel for
113 each trait, excluding markers residing in the Major Histocompatibility Complex region on chromo-
114 some 6. Significance threshold was defined by Benjamini-Hochberg false discovery rate as < 0.05 .
115 We visualized the results using the R packages *network-plot* and *ldsc-corr-plot* [34].

116 Polygenic Risk Scoring

117 We used PRSice-2 [35] for our Polygenic Risk Score (PRS) analysis. We performed a unilateral PRS
118 analysis between the TS-EUROTRAIN cohort and the TSGWAS2 [7] cohort, using the TSGWAS2 sum-
119 mary statistics as discovery and the TS-EUROTRAIN GWAS, after excluding the overlapping samples,
120 as the target dataset. The TSGWAS2 summary statistics were clumped on the LD information of the
121 TS-EUROTRAIN GWAS, using a window of 250kb and an r^2 threshold of 0.1. PRSice performed the
122 scoring on subsets of the dataset based on nine thresholds of p-value leniency (5E-08, 1E-04, 1E-
123 03, 0.01, 0.05, 0.1, 0.2, 0.5, 1). The resulting PRS was tested for association with TS, using logistic
124 regression with the previously mentioned ancestry components, sex, and imputation batch as co-
125 variates. The model fit for best p-value threshold was run using 10,000 permutations. Liability scale
126 was calculated on the variance explained by the PRS (R^2), using a TS population prevalence of 1%.

127 Biological annotation of results

128 We applied FUMA [36] to perform gene-based and gene-set analyses on the results from the TS-
129 EUROTRAIN GWAS and the subsequent TS-EUROTRAIN/GWAS2 meta-analysis. The genetic vari-
130 ants were assigned to protein-coding genes based on their GRCh37 build genomic position, using
131 a ± 20 kb window size. After quality control, 18,089 genes contained at least one variant and as such
132 were used for the gene-based analysis, thus setting the Bonferroni threshold at $p < 2.764e - 6$. The

gene-based association results were subsequently used for gene-set enrichment analysis under a competitive model. Tissue Expression Analysis was conducted on the GTEx v8 expression data [37, 38]. We investigated chromatin contact points through Capture Hi-C data available from the 3D Genome Browser [39], using promoter-centered long-range chromatin interaction data derived from human dorsolateral prefrontal cortex tissues [40].

We performed a set-based association analysis using PLINK [41, 42] on the gene-sets that were previously identified as significantly associated with TS [11, 43]. We used logistic regression as the association model on the genotypes and principal components that were identified by Tracy-Widom statistics in the GWAS. Another repetition of this step was performed with the χ^2 association test, to test for this method's robustness to population structure. We proceeded to run the analysis on all TS-EUROTRAIN samples, using a 10kb genomic window size and a million permutations. Since the permutations were performed on the phenotypic status of the samples, and only served as a method of association of the trait with the gene sets, we also corrected the results by defining the significance threshold through Bonferroni correction.

PRs for brain anatomy

Using the TS-EUROTRAIN/GWAS2 meta-analysis summary statistics as base, we computed TS polygenic risk scores (PRS_{TS}) for individuals in the UK Biobank [44] using PRSice-2 [35] and subsequently evaluated the association between risk scores and subjects' 14 subcortical volumes (FIRST). After quality controls, 29,829 samples with brain MRI phenotypes available were included for the analysis. We assessed the association with PRS_{TS} computed using independent SNPs with meta-analyzed p-value under various thresholds (0.05, 0.001, 1e-5). For each threshold, clumping of base summary statistics was carried out under $r^2 = 0.1$ with a ± 500 kb window. Regressions between PRS_{TS} and brain volume measurements were evaluated controlling for age, sex, and the top five PCs as covariates. Bonferroni significance threshold for FIRST measurements was $p < 1.19e - 3$ ($0.05/(14 \times 3)$).

Results

Mega-analysis of TS-EUROTRAIN GWAS

GWAS analysis was performed as a mega-analysis, on the combined genetic data of all collected samples, using a logistic regression model on the best-guess genotypes (genotype probability > 0.9) with INFO score > 0.9 and MAF > 0.01 . As covariates, we included the ancestry components 1,2,4,

162 and 5 to account for population stratification as identified by ANOVA statistics (Supplementary Table
163 2), sex as a binary index to account for sex stratification, and imputation batch to account for bias
164 due to array and imputation batch effects.

165 The TS-EUROTRAIN GWAS identified three novel highly-correlated ($r^2 > 0.8$) genome-wide sig-
166 nificant SNPs, located near the *NR2F1 Antisense RNA 1* long non-coding RNA (*NR2F1-AS1* lncRNA)
167 locus (Supplementary Figure 1, Figure 1a). The strongest signal was found for rs2453763 (chr5:92376460:T/A,
168 $OR = 0.7512$, $P = 2.62e - 8$, $MAF=0.3581$), a variant 350kb upstream of *NR2F1-AS1*, and asso-
169 ciated with decreased risk for TS. The imputation statistics for this SNP indicate high imputation
170 quality ($MAF=0.3581$, $INFO=0.99$). The proximal SNPs were rs2009416 (chr5:92415111:T/C, $OR =$
171 0.7532 , $P = 3.31e - 8$, $MAF=0.3562$) and rs1496337 (chr5:92411293:T/C, $OR = 0.7534$, $P = 3.33e - 8$,
172 $MAF=0.3563$). Conditional analysis using the lead SNP as covariate showed no secondary signals
173 in the region. The top ($P < 1e - 5$) loci detected in the novel GWAS are reported in Table 1. LD
174 Score Regression analysis of the summary statistics did not provide evidence for genomic inflation
175 ($\lambda_{GC} = 1.07$, intercept=1.0061, intercept p-value=0.28, attenuation ratio=0.0887).

176 **Meta-analysis with TSGWAS2**

177 The TS-EUROTRAIN GWAS was then meta-analyzed with summary statistics results from the previ-
178 ous largest meta-analysis of TS to date (TSGWAS2) [7] using Han and Eskin's random effects model
179 [24, 25]. Since there was a known sample overlap between the two studies, the TS-EUROTRAIN
180 GWAS was re-analyzed after excluding the overlapping samples (1,314 cases and 4,077 ancestry-
181 matched controls), with the results being very similar to the full dataset TS-EUROTRAIN GWAS (Sup-
182 plementary Figure 2). Only variants overlapping in both studies were included, leading to a total of
183 6,133 cases, 13,565 controls and 1,955,677 variants in the final meta-analysis.

184 The TS-EUROTRAIN/GWAS2 meta-analysis (Figure 2) identified the three genome-wide signif-
185 icant SNPs of the TS-EUROTRAIN GWAS, with rs2453763 again being the top hit (chr5:92376460:T/A,
186 random effects $P = 4.05e - 08$) along with an array-wide significant SNP, rs10209244 (chr2:161561898:A/G,
187 random effects $P = 6.16e - 08$, $MAF = 0.01$) that resides 200kb downstream of *RBMS1* (Figure 1b).
188 The top loci detected from the meta-analysis are reported in Table 2. LD Score Regression analysis of
189 the summary statistics did not provide evidence for genomic inflation ($\lambda_{GC} = 1.16$, intercept=1.016,
190 intercept p-value=0.11, attenuation ratio=0.0869).

191 Genetic relationship between the TS-EUROTRAIN GWAS and TSGWAS2

192 The SNP with the strongest signal in the previously published TSGWAS2 study was absent from the
193 TS-EUROTRAIN dataset due to the differences in reference panels used (1000 Genomes for TSG-
194 WAS2 and HRC for the novel study) and stringent batch effect quality control performed on the novel
195 dataset. We observed no genome-wide significant heterogeneity (Cochran's Q-test $p\text{-value} < 5e-8$)
196 in the meta-analysis.

197 To explore the relationship between the TS-EUROTRAIN GWAS and TSGWAS2, we used LD Score
198 Regression [22] to compute their genetic correlation, after excluding the overlapping samples. LD
199 Score Regression identified a strong genetic correlation between the two studies ($r_g = 0.95$, $p =$
200 $6e-8$), and provided evidence of consistency across them (Figure 3). Investigation of the gene sets
201 found previously associated with TS [11, 43] also successfully replicated the associations for the lym-
202 phocytic, the ligand-gated ion channel signaling, the cell adhesion and trans-synaptic signaling, as
203 well as the astrocyte-neuron metabolic coupling gene sets (Supplementary Table 3).

204 PRS analysis displayed consistency between the two studies. PRS were computed using the sum-
205 mary statistics of TSGWAS2 as a training dataset and the TS-EUROTRAIN raw genotypes as discovery
206 in PRSice [45]. The best fit p -value threshold was determined at $p = 0.1182$ (model fit $p = 1.26e-28$)
207 (Figure 4). Maximum variance explained at the best fit model was estimated by Nagelkerke's R^2 at
208 3.3%.

209 Cross-disorder analysis

210 Pairwise genetic correlations were computed between our results and five traits that have been found
211 to be correlated in previous studies with the TSGWAS2 results [46, 47] using LD Score Regression
212 [22]. Benjamini-Hochberg FDR correction with an $\alpha = 0.05$ was used to correct for multiple testing.
213 After correction, the TS-EUROTRAIN GWAS was significantly correlated with ADHD ($r_g = 0.28$, $p =$
214 $7.7e-3$), and the TS-EUROTRAIN/GWAS2 meta-analysis was significantly correlated with ADHD
215 ($r_g = 0.19$, $p = 5.1e-3$) and OCD ($r_g = 0.42$, $p = 1e-4$) (Figure 3).

216 Heritability estimation and partitioning

217 We used LD Score Regression [22] to estimate the SNP-based heritability (h_{SNP}^2) using the summary
218 statistics of the novel GWAS and the meta-analysis. The summary statistics were merged with the
219 HapMap3 marker panel provided by the authors. For the TS-EUROTRAIN GWAS, analysis yielded an
220 observed scale h_{SNP}^2 estimate of 0.4385($SE : 0.1167$) on the observed scale, and 0.2736($SE : 0.0728$)

221 assuming a TS prevalence of 0.01 on the liability scale, while the LD score regression analysis inter-
222 cept was computed at 1.0157($SE : 0.013$) (p-value 0.028) and the ratio of stratification to polygenic-
223 ity was estimated at 0.0863($SE : 0.0711$). For the meta-analysis the h^2_{SNP} estimate was 0.3504($SE :$
224 0.0439) and 0.2184($SE : 0.0269$) on the liability scale.

225 We proceeded to partition the heritability by functional genomic categories using stratified LD
226 Score Regression [27] on the full baseline model and a model based on the Roadmap epigenomics
227 data, as provided by the authors [27]. The full baseline model included 24 main overlapping func-
228 tional categories and identified statistically significant enrichment in two categories, after Benjamini-
229 Hochberg FDR correction at an $\alpha = 0.05$. The H3K4me1 sites category (enrichment value 1.61,
230 $P = 9.5e-4$) was the top significant signal in the analysis, with the conserved elements category (en-
231 richment value 2.05, $P = 3.8e-3$) being the second significant signal. The Roadmap model includes
232 epigenomic mapping data from 395 tissues [28] and when applied to our data for heritability parti-
233 tioning, yielded 13 statistically significant modifications after Benjamini-Hochberg FDR correction
234 at an $\alpha = 0.05$. These 13 signals marked the enrichment of the histone marks H3K27ac, H3K4me1,
235 and H3K9ac on five brain tissues. H3K27ac was identified on the angular gyrus, the cingulate gyrus,
236 the dorsolateral prefrontal cortex, and the inferior temporal lobe; H3K4me1 on the angular gyrus,
237 the cingulate gyrus, the dorsolateral prefrontal cortex, the inferior temporal lobe, and the substan-
238 tia nigra; H3K9ac on the angular gyrus, the anterior caudate, the dorsolateral prefrontal cortex, and
239 the inferior temporal lobe. The results for the full baseline model and the Roadmap model can be
240 further explored in Supplementary Tables 4 and 5.

241 **Biological annotation and enrichment analysis**

242 Functional mapping, annotation, and gene set enrichment using the FUMA pipeline did not pro-
243 duce significant results. The identified top signals from the TS-EUROTRAIN GWAS and the TS-
244 EUROTRAIN/GWAS2 meta-analysis reside in large intergenic regions whose distance from their cor-
245 related genes (Supplementary Figure 5) exceeded the distance limits set by the software, and were
246 thus excluded from the annotation step of the pipeline. The top signal of the gene-based analysis
247 was *RANGAP1* ($P = 3.36e-6$) on chromosome 22; it did not meet the genome-wide significance
248 threshold ($P = 2.8e-6$) (Supplementary Table 6). MAGMA tissue expression analysis using FUMA
249 did not produce any statistically significant results. MAGMA tissue expression analysis on the TS-
250 EUROTRAIN GWAS using the 53 GTEx tissue sample set indicated putative enrichment in various
251 brain tissues, with the top signals in the hypothalamus, putamen, and nucleus accumbens. Analy-

sis using the 30 GTEx tissue sample set indicated potential enrichment in the brain, followed by the colon, adrenal gland, and pituitary (Supplementary Figure 3). MAGMA tissue expression analysis of the meta-analysis, using the 53 tissue sample set from GTEx, indicated stronger putative enrichment in various brain tissues, with the top signals in the cerebellar hemisphere, cerebellum, and frontal cortex (area BA9). Using the 30 tissue sample set from GTEx, stronger evidence of potential enrichment could be identified in the brain, followed by the pituitary and the ovary (Supplementary Figure 4).

Brain volumes associated through PRS

Association between the meta-analysis PRS_{TS} and brain volume measurements highlighted the previously described relationship [48] between genetic risk for TS and bilateral putamen volumes under multiple PRS p-value thresholds. The strongest (and only statistically significant) associations were found between PRS_{TS} computed using 116 independent SNPs with $p < 0.001$ and right putamen volume (PRS-p = $3.70E-05$, $r^2 = 0.264$), as well as left putamen volume (PRS-p = $1.25E-04$, $r^2 = 0.257$), both being negatively associated with TS genetic risk (Supplementary Table 7).

Discussion

We report results from a novel study, as well as the largest GWAS meta-analysis on TS to date, including 6,133 TS individuals of European ancestry and 13,844 matched controls. We report two novel independent genome-wide significant loci associated with TS: one on chromosome 5q15 upstream of *NR2F1* in TS-EUROTRAIN GWAS (1,438 TS cases and 4,356 controls) and another on chromosome 2q24.2 downstream of *RBMS1* in the final meta-analysis. These results confirm the value of collaborative efforts towards expanding sample size and datasets available for analysis (such as the TS-EUROTRAIN, EMTICS, TSGeneSEE, and PGC initiatives). However, this study still only captures a small fraction of the risk for TS attributable to common variants. Larger studies are necessary and warranted.

The genome-wide significant locus identified in the TS-EUROTRAIN/GWAS2 meta-analysis resides in the 5q15 chromosomal region (SNP rs2453763). The associated SNP is located within *CTD-2091N23.1*, a gene encoding a long non-coding RNA that has yet to be functionally characterized. It is found upstream of a gene cluster that harbors the genes *NR2F1-AS*, *NR2F1*, and *lnc-NR2F1*. The GTEx portal [38] reports SNP rs2453763 to be significantly associated as a splicing quantitative trait locus (sQTL) for *CTD-2091N23.1* in the tibial nerve, and as an eQTL for *NR2F1* and *NR2F1-AS* in the

esophagus smooth muscles and for *CTD-2091N23.1* in cultured fibroblasts (Table 3). Capture Hi-C [39] showed strong evidence for the SNP being related to the regulation of *NR2F1* (Supplementary Figure 5).

Moreover, the 5q15 region has previously been implicated in neurodevelopmental phenotypes [49–52]. The 5q14-5q15 regions have been reported to contain fragile sites that are associated with genomic and epigenomic instability as well as linked to schizophrenia and autism [53]. The exact genes are yet to be identified, with recent evidence suggesting a role for *NR2F1*-related genes, and more intriguingly, the *Inc-NR2F1* gene. *Inc-NR2F1* is a long non-coding RNA locus discovered to be recurrently mutated in individuals with autism spectrum disorders and intellectual disability, with translocations in this locus reported to show patterns of Mendelian inheritance [54]. A functional study of *Inc-NR2F1* identified its role in neuronal maturation *in vitro* through expression regulation of a network of genes that have been linked to autism [54]. Functional studies of the *NR2F1* gene also have indicated its critical role for neurodevelopment through investigations into human and mouse haploinsufficiency [55], insertion of point mutations in mouse models that lead to excitatory/inhibitory neuronal imbalance [56], and the study of knock-out mouse models [57]. Notably, in the absence of *NR2F1*, an imbalance between oligodendrocytes and astrocytes develops, leading to postnatal hypomyelination and astrogliosis [55]. *NR2F1* is a highly conserved orphan nuclear receptor which is a regulator of transcription. It belongs to the steroid/thyroid hormone nuclear receptor superfamily, involved in a wide range of roles, including cell differentiation, cancer progression, and central and peripheral neurogenesis [58]. A multitude of pathogenic variants have been identified in *NR2F1*, leading to Bosch-Boonstra-Schaaf optic atrophy syndrome, and autosomal dominant neurodevelopmental disorder (OMIM: 615722) [59]. *NR2F1* is also known by its historical name, *COUP-TF1*; it is a target of the androgen receptor (AR), along with *SOX9* and *OCT4* [60]. *NR2F1* interacts with *SOX9* [60, 61] and represses a host of targets in multiple tissues, including *CYP17A1*, Oxytocin gene *OXT*, and *OCT4* [62]. Especially in the case of *CYP17A1*, encoding for a key enzyme of steroid biosynthesis, *NR2F1* and SF-1 exert opposing effects, as repressor and activator, respectively [63].

The second locus (revealed by the meta-analysis) was located in the 2q24.2 chromosomal region (SNP rs10209244), downstream of the *RBMS1* gene. Capture Hi-C [39] showed strong evidence for the SNP being related to the regulation of *RBMS1* (Supplementary Figure 5). In previous GWAS, *RBMS1* has been associated with lymphocyte, neutrophil, and white blood cell count [64, 65], educational attainment and mathematical ability [66], and addiction to tobacco and alcohol [67]. *RBMS1* has been implicated in estradiol production on granulosa cells, through inactivation of c-Myc, which is its downstream target [68, 69]. Specifically, *RBMS1* mRNA stability is controlled through miR-383,

315 which is in turn positively regulated by SF-1 [68, 69]. These results hint that steroid regulatory path-
316 ways may be involved in TS pathogenesis. Steroid hormones have been proposed to play a funda-
317 mental role in TS and sexual dimorphism of the central nervous system; androgenic hormones, in
318 particular, are likely to exacerbate symptoms of the disorder [70].

319 We sought to validate our results through means of heritability correlation patterns and poly-
320 genic risk scoring. Heritability analysis in the TS-EUROTRAIN dataset indicated that a large pro-
321 portion of TS SNP-based heritability can be attributed to common variants ($h^2_{SNP} = 0.4385$), in con-
322 cordance with the estimate ranges in previous investigations [4, 7]. The attenuation ratio was suffi-
323 ciently low at 0.0887, attributing inflation to polygenicity in the samples. After exclusion of the over-
324 lapping samples, heritability correlation between the previous TS GWAS and the TS-EUROTRAIN
325 dataset indicated an almost complete correlation between the two studies. Polygenic prediction in
326 the TS-EUROTRAIN cohort using the TSGWAS2 results as discovery achieved significant predictive
327 levels, on par with the inter-cohort predictive Nagelkerke's R^2 levels in the previous TS GWAS [7] and
328 substantially increased by more than an order of magnitude compared to tic prediction in a general
329 population cohort [12].

330 Investigations into the genetic basis of TS have long been hampered by heterogeneity between
331 studies. We addressed this by employing a random-effects meta-analytic method, Han and Eskin's
332 random effects model [25], a model shown to perform particularly well in detecting true positives in
333 the presence of heterogeneity in complex disorders [26].

334 The neuroendocrine system has long been hypothesized to be involved in the neurobiology of TS
335 [71]. The hypothalamus-pituitary-gonadal (HPG) axis has been hypothesized to be implicated in TS
336 tic exacerbation [71], while a series of investigations have been launched into deciphering the role of
337 stress through the hypothalamus-pituitary-adrenal (HPA) axis [72, 73]. The FUMA eQTL analysis did
338 not provide statistically significant results, however, it demonstrated distinct patterns of enrichment
339 towards brain tissues, the pituitary, and the ovary when performed on the meta-analysis results, and
340 towards the adrenal glands, the brain, and the pituitary when performed on the TS-EUROTRAIN
341 GWAS results. While future analyses will be needed to replicate these findings, these results might
342 suggest the potential involvement of the HPG and the HPA axes in TS pathogenesis.

343 In summary, our GWAS meta-analysis of 6,133 cases and 13,565 controls identified a genome-
344 wide significant locus on chromosome 5q15 and one array-wide significant locus on chromosome
345 2q24.2. Integration of eQTL and GWAS data implicate the *NR2F1* gene and associated lncRNAs
346 within the 5q15 locus, and Hi-C data provides some evidence for the RBMS1 gene within the 2q24.2
347 locus. While this study is, to our knowledge, the largest TS GWAS meta-analysis to date, when com-

pared to GWAS in other neuropsychiatric disorders, it is clear that even larger studies are warranted. Increased statistical power will further enable the identification of more leads for potential future interventions. Future plans of collaborative efforts will focus on vastly reinforcing sample sizes, increasing power to adequately perform fine-mapping and eQTL analyses with the aim to move towards the elucidation of the underlying biology of TS.

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Tables

Table 1: Top regions ($p < 1E-5$) uncovered in the TS-EUROTRAIN GWAS (1,438 cases and 4,356 controls on 2,949,675 variants). One variant was identified as genome-wide significant ($p < 5e-8$). Chromosome and region (based on hg19) are shown for index SNPs ($LD - r^2 > 0.1$), as well as number of LD-associated markers in proximity (N). A1 refers to associated allele. The odds ratio (OR) and standard error (SE) are shown for the association between A1 and TS. MAF indicates the allelic frequency of allele 1 in the dataset. The reported nearest genes were determined by genomic location (± 500 kb). The analysis was restricted to variants with $MAF \geq 0.01$ and information quality (INFO) score ≥ 0.9 . Chromosome X was not analyzed, since it was absent from a significant portion of the acquired datasets.

SNP ID	Chr	P-value	A1	OR	SE	N	Location	KB	Genes
rs2453763	5	$2.623e-08$	T	0.7512		25	chr5:92322427..92559372	236.946	NR2F1-AS1
rs2197383	3	$4.681e-07$	A	0.5948		165	chr3:79889497..80380401	490.905	ROBO1
rs3773057	3	$1.388e-06$	T	0.1981		23	chr3:29563164..29627731	64.568	RBMS3, RBMS3-AS3
rs9382365	6	$1.829e-06$	G	0.6769		77	chr6:54418351..54531232	112.882	FAM83B, TINAG
rs152061	5	$2.085e-06$	T	1.262		196	chr5:64778944..64989139	210.196	ADAMTS6, CENPK, ERBB2IP, NLN, PPWD1, SGTB, TRAPPC13, TRIM23
rs2278796	1	$6.882e-06$	T	1.261		9	chr1:204951209..204971553	20.345	CNTN2, DSTYK, NFASC, RBBP5, TMCC2, TMEM81
rs34940828	3	$7.531e-06$	C	2.104		64	chr3:123213895..123398466	184.572	ADCY5, CCDC14, MYLK, MYLK-AS1, PTPLB, SEC22A
rs2076218	1	$9.746e-06$	A	1.25		3	chr1:209745395..209768699	23.305	C1orf74, CAMK1G, DIEXF, G0S2, HSD11B1, IRF6, LAMB3, MIR205, MIR205HG, MIR4260, TRAF3IP3

Table 2: Top regions ($p < 1E-5$) uncovered by GWAS meta-analysis (TS-EUROTRAIN and TSGWAS2 [7] - 6,133 cases, 13,565 controls on 1,955,677 variants) using Han and Eskin's random effects model [25]. One variant was identified as genome-wide significant ($p < 5e-8$) and one as array-wide significant ($p < 1e-7$). Chromosome and region (based on hg19) are shown for index SNPs ($LD-r^2 > 0.1$), as well as number of LD-associated markers in proximity (N). The reported nearest genes were determined by genomic location (± 500 kb). MAF indicates the allelic frequency of the minor allele in the dataset.

SNP ID	Chr	P-value	N	Location	KB	Genes
rs2453763	5	$4.054e-08$	25	chr5:92322427..92559372	236.946	NR2F1-AS1
rs10209244	2	$6.156e-08$	29	chr2:161422880..161676570	253.691	MIR4785, RBMS1
rs13401916	2	$2.441e-07$	13	chr2:161945103..162055548	110.446	LOC100996579, LOC101929512, PSMD14, TANK, TBR1
rs139469	22	$9.997e-07$	33	chr22:41451185..41627527	176.343	ACO2, CHADL, DNAJB7, EP300, EP300-AS1, L3MBTL2, MIR1281, MIR4766, MIR6889, PHF5A, RANGAP1, RBX1, SLC25A17, ST13, TEE, TOB2, XPNPEP3, ZC3H7B

Table 3: Significant SNP-gene pairings identified through GTEx eQTL and sQTL data [37, 38]. Four significant associations were identified for SNP rs2453763, while no significant associations were identified for rs10209244. rs2453763 is an eQTL for three genes on two tissues, and an sQTL for one gene on one tissue. Reported are the symbol of the associated gene, the respective associated tissue, and the normalized effect size (NES). a) GTEx eQTL associations for the top variant in *NR2F1* (rs2453763). b) GTEx sQTL associations for the top variant in *NR2F1* (rs2453763)

Gene Symbol	SNP Id	P-Value	NES	Tissue
<i>NR2F1</i>	rs2453763	$5.7E-10$	0.25	Esophagus - Muscularis
<i>NR2F1-AS1</i>	rs2453763	$1.9E-7$	0.22	Esophagus - Muscularis
<i>CTD-2091N23.1</i>	rs2453763	$2.3E-4$	-0.19	Cells - Cultured fibroblasts

Gene Symbol	SNP Id	P-Value	NES	Tissue	Intron Id
<i>CTD-2091N23.1</i>	rs2453763	$4.6E-09$	0.39	Nerve - Tibial	93051776:93075658:clu_40848

Figures

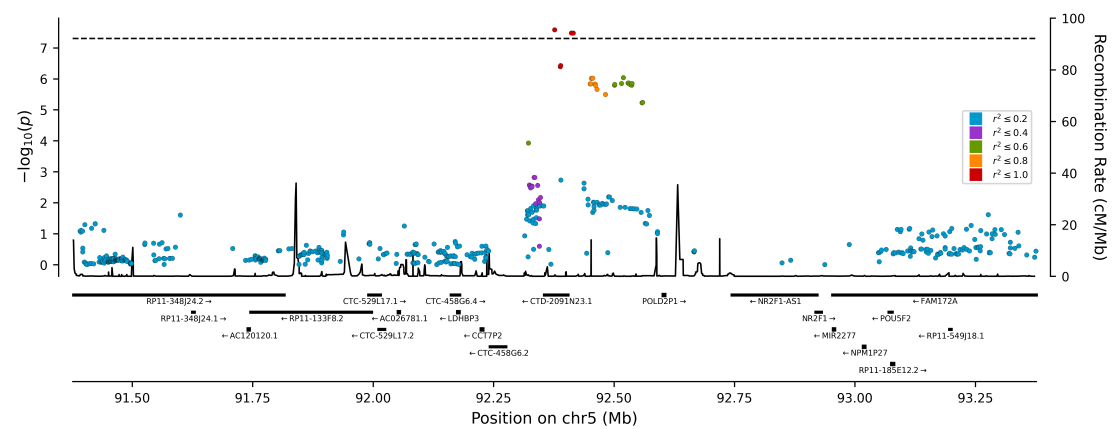


Figure 1: Regional plots using the TS-EUROTRAIN dataset as the base for LD calculations. In red are the markers that are in high LD with the lead marker ($0.8 \leq r^2 \leq 1$). a) Regional plot of the top TS-EUROTRAIN GWAS locus (NR2F1). b) Regional plot of the second top meta-analysis locus (RBMS1)



Figure 2: The Manhattan plot for the genome-wide association meta-analysis of Tourette Syndrome with the TS-EUROTRAIN and the TSGWAS2 datasets (6,133 TS cases and 1,3565 controls of European descent on 1,955,677 variants) using Han and Eskin's random effects model [25]. The $-\log_{10}(p)$ values for the association tests (two-tailed) are shown on the y axis and the chromosomes are ordered on the x axis. One genetic locus on chromosome 5 surpassed the genome-wide significance threshold ($p < 5e-8$; indicated by the red line). One genetic locus surpassed the array-wide significance threshold ($p < 1e-7$). Gray and black differentiate adjacent chromosomes.

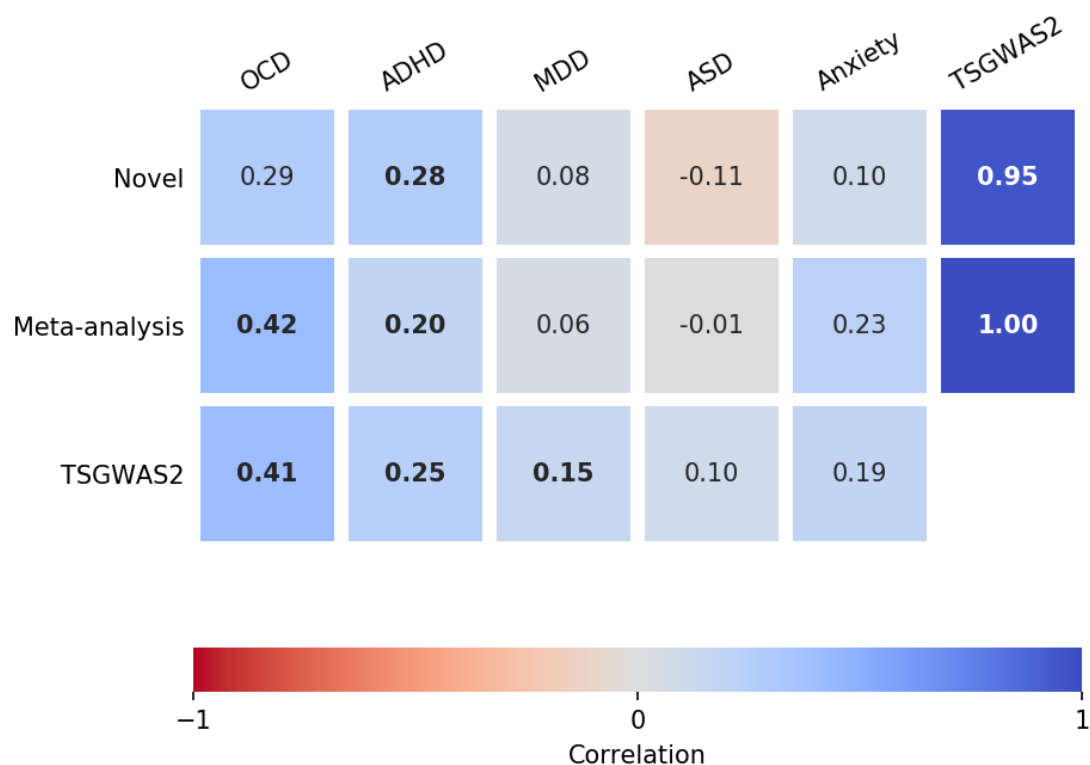


Figure 3: Genetic correlations with Tourette Syndrome. The genetic correlations were estimated with bivariate LD score regression [22]. We showcase the correlations between three TS studies (TS-EUROTRAIN, TS-EUROTRAIN/TSGWAS2 meta-analysis, and TSGWAS2) and five neuropsychiatric disorders (OCD, ADHD, MDD, ASD, and anxiety) previously correlated with TS. The number in each square denotes the correlation r_g . In bold are the correlations that were identified as statistically significant after Benjamini-Hochberg FDR correction ($\alpha=0.05$).

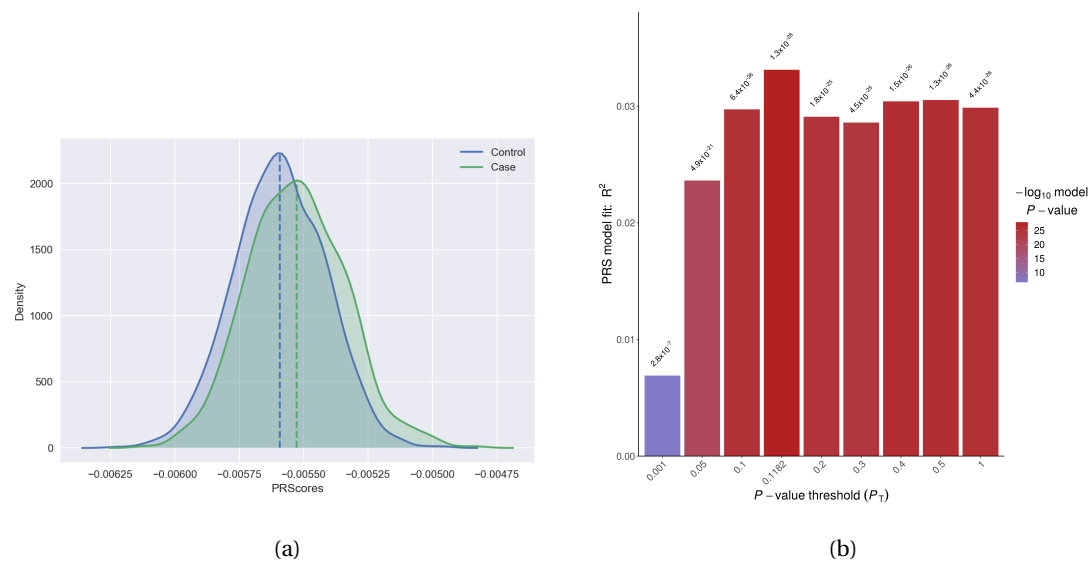


Figure 4: Polygenic Risk Scoring analysis using the TSGWAS2 dataset [7] as discovery and the TS-EUROTRAIN dataset as target. Best fit p-value threshold was determined at $p=0.1182$ (model fit $p=1.26e-28$). Maximum variance explained at the best fit model was estimated by Nagelkerke's R^2 at 3.3%. a) PRS distribution comparison between cases and controls for the best fit model. b) PRS histogram for each p-value bin, including the best fit bin.