

Joint Variational Autoencoders for Multimodal Imputation and Embedding

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Abstract. Single-cell multimodal datasets have measured various characteristics of individual cells, enabling a deep understanding of cellular and molecular mechanisms. However, multimodal data generation remains costly and challenging, and missing modalities happen frequently. Recently, machine learning approaches have been developed for data imputation but typically require fully matched multimodalities to learn common latent embeddings that potentially lack modality specificity. To address these issues, we developed an open-source machine learning model, Joint Variational Autoencoders for multimodal Imputation and Embedding (JAMIE). JAMIE takes single-cell multimodal data that can have partially matched samples across modalities. Variational autoencoders learn the latent embeddings of each modality. Then, embeddings from matched samples across modalities are aggregated to identify joint cross-modal latent embeddings before reconstruction. To perform cross-modal imputation, the latent embeddings of one modality can be used with the decoder of the other modality. For interpretability, Shapley values are used to prioritize input features for cross-modal imputation and known sample labels. We applied JAMIE to both simulation data and emerging single-cell multimodal data including gene expression, chromatin accessibility, and electrophysiology in human and mouse brains. JAMIE significantly outperforms existing state-of-the-art methods in general and prioritized multimodal features for imputation, providing potentially novel mechanistic insights at cellular resolution.

Keywords: Multimodal · Single-cell · Variational autoencoder · Deep learning · Data imputation · Data integration

1 Introduction

Understanding of molecular mechanisms at cellular resolution provides deeper insights into cellular function, development, and disease progression, but remains elusive. To this end, single-cell multimodal datasets have recently emerged by new sequencing technologies to measure various characteristics of single cells and identify cell functions (e.g., cell types). For instance, patch-seq [1] simultaneously measures gene expression and identifies electrophysiology, and morphological features (beyond omics), e.g., for the mouse visual cortex over 4,000 cells from several brain cell types [2]. [3] profiles single-cell gene expression and chromatin accessibility in the human developing brain. Thus, an integration of single-cell multimodal datasets can significantly aid in our understanding of biological mechanisms contributing to cell types and diseases through the automated discovery of inter-modal relationships.

Many methods have been developed to integrate multimodal datasets to improve prediction of cell types and cellular phenotypes (an outline of the process can be seen in Figure 1a). For instance, our recent DeepManReg [4] approach performs interpretable deep manifold learning which both improves phenotype prediction and is applicable to align single-cell multimodal data. Although these approaches primarily focus on prediction, the underlying idea can be extended to predict data present in separate modalities, referred to as cross-modal imputation going forward (a visual representation can be seen in Figure 1b). Cross-modal imputation is not new, but has been increasingly explored with the advent of deep learning. For instance, BABEL [5] focuses on conversion between RNA and ATAC-seq data matrices through the use of dual autoencoders, which have the particular advantage of being reusable. Moreover, autoencoders allow for flexibility

39 in latent space formulation. Polarbear [6] utilizes a similar approach, focusing on a conversion network be-
40 tween independently generated latent spaces for each modality. Both methods, however, focus specifically
41 on multi-omics data such as scRNA-seq and scATAC-seq and thus don't consider high nonlinearity from ad-
42 ditional single-cell modalities like electrophysiology. Moreover, they require fully correspondent multimodal
43 data (matched cells), limiting their capabilities with regards to data with missing modalities.

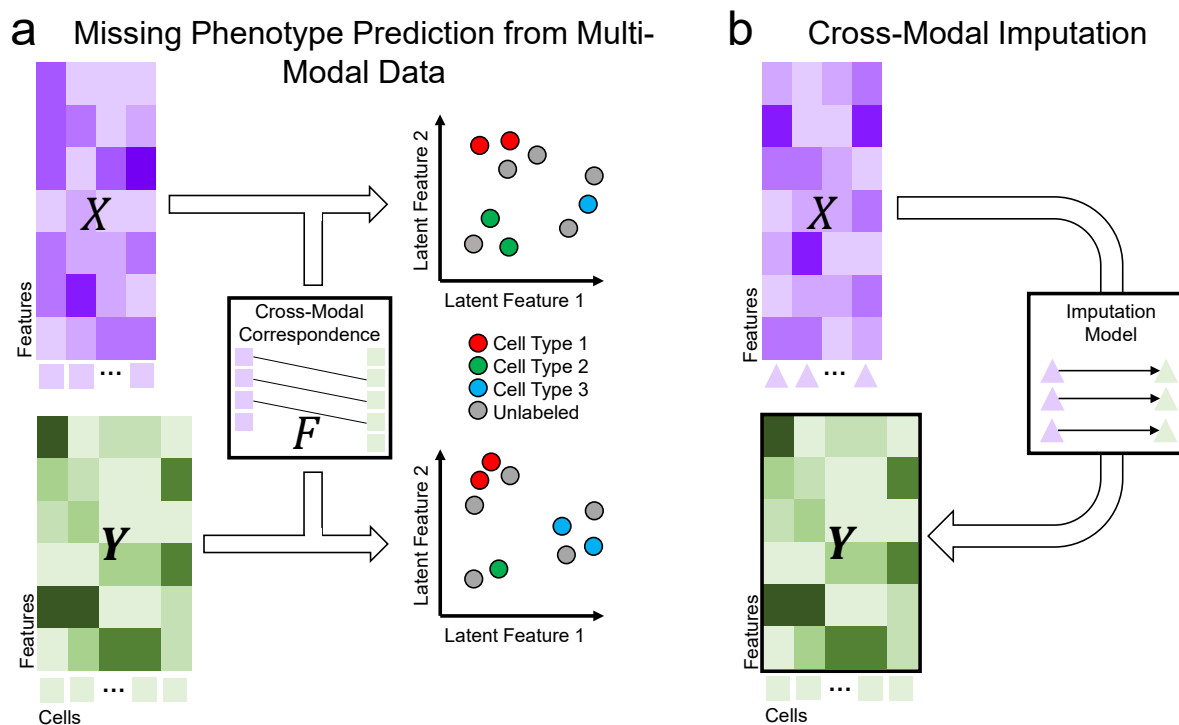


Figure 1. Challenges for multimodal data integration and imputation. (a) Using multimodal data to effectively predict missing phenotypes (e.g. cell types from multimodal single-cell data) is difficult due to heterogeneous features across modalities. Identifying similar latent spaces across modalities enables cross-modal comparison and, by extension, missing phenotype prediction. Machine learning (ML) can be used to discover similar cross-modal latent spaces and enable comparison and phenotype prediction. (b) Certain modalities are cost-prohibitive but may lend significant insight into biological mechanisms. As examples, scATAC-seq data for cell type epigenomics is expensive and electrophysiological data at single-cell resolution is difficult to produce. Imputing one modality from another using ML can alleviate these constraints.

44 Also, analyzing multimodal datasets incurs additional difficulties, including heterogeneous distributions,
45 multicollinearity, and varying reliability. Several approaches have been thus utilized in aligning multimodal
46 datasets (an outline of the process can be seen in Figure 1b). For instance, Unioncom [7] infers correspondence
47 information from each modality's distance matrices, then use a modified tSNE method for the final mapping.
48 MMD-MA [8] minimizes an objective function to maximize distribution similarity between datasets while
49 minimizing distortion. ScGLUE [9] uses an autoencoder model to map different modalities onto the same
50 latent space. ScDART [10] learns a latent space and cross-modal relationships simultaneously by predicting
51 gene activity across modalities before projecting to a single latent space. ScaEGAN [11] utilizes dual au-
52 toencoders in tandem with cycleGAN [12] to provide aligned latent representations for different modalities.
53 Certain methods rely on user-provided correspondence information to help inform alignment. For example,
54 CLUE [13] assumes completely aligned samples and introduces the impetus for aggregating correspondent
55 latent-space embeddings. ManiNetCluster [14] takes a user-provided correspondence matrix as input and
56 implements nonlinear manifold alignment (NLMA) by solving an eigenvalue problem. ManiNetCluster also
57 implements CCA [15] as a manifold alignment solution for multimodal data. If only partial correspondence
58 information is known, however, existing methods are limiting, and few are designed for such a use-case.

59 To address these issues, we developed Joint variational Autoencoders for Multimodal Imputation and
60 Embedding (JAMIE). An outline of JAMIE's capabilities can be found in Supplementary Table 1. JAMIE

61 trains a reusable joint variational autoencoder model to project available multimodalities onto similar latent
62 spaces (but still unique for each modality), allowing for enhanced inference of single-modality patterns [16].
63 To perform cross-modal imputation, data may be fed into an encoder, then the resultant latent space may
64 be processed by the decoder of the other modality. JAMIE is able to use partial correspondence information.
65 JAMIE combines the reusability and flexible latent space generation of autoencoders with the automated
66 correspondence estimation of alignment methods. We compared JAMIE to state-of-the-art methods on sim-
67 ulation and emerging single-cell multimodal data using gene expression, chromatin accessibility, and elec-
68 trophysiology in human and mouse brains. We found that JAMIE significantly outperforms other methods
69 (see evaluation in Section 2) and prioritized important multimodal features for multimodal imputation, also
70 providing potentially novel mechanistic insights at cellular resolution.

71 2 Results

72 As shown in Figure 2a, JAMIE utilizes a joint autoencoder model for data integration and imputation (see
73 details in Section 4. As input, JAMIE takes two data matrices \mathbf{X} and \mathbf{Y} of modalities X and Y as input.
74 Optionally, an additional correspondence matrix \mathbf{F} may be provided when the samples from two modalities
75 are partially corresponding (e.g. derived from the same single-cells). Encoders in JAMIE transform \mathbf{X} and
76 \mathbf{Y} into latent spaces which are aggregated using available corresponding information. Decoders in JAMIE
77 then predict reconstructions $\tilde{\mathbf{X}}$ and $\tilde{\mathbf{Y}}$ of the original modalities. Please see more details of JAMIE including
78 training, validation, and evaluation in Section 4.

79 After training a JAMIE model, its encoder for modality X and decoder for modality Y can be used
80 sequentially to impute from one modality to another (Figure 2b). Also, the latent spaces from the JAMIE
81 model can be used for phenotype prediction. Furthermore, the use of Shapley Additive Explanation values [17]
82 and similar importance evaluation methods then allows us to prioritize multi-modal features for imputation,
83 as in Figure 2c. These applications are further documented in Subsection 4.12.

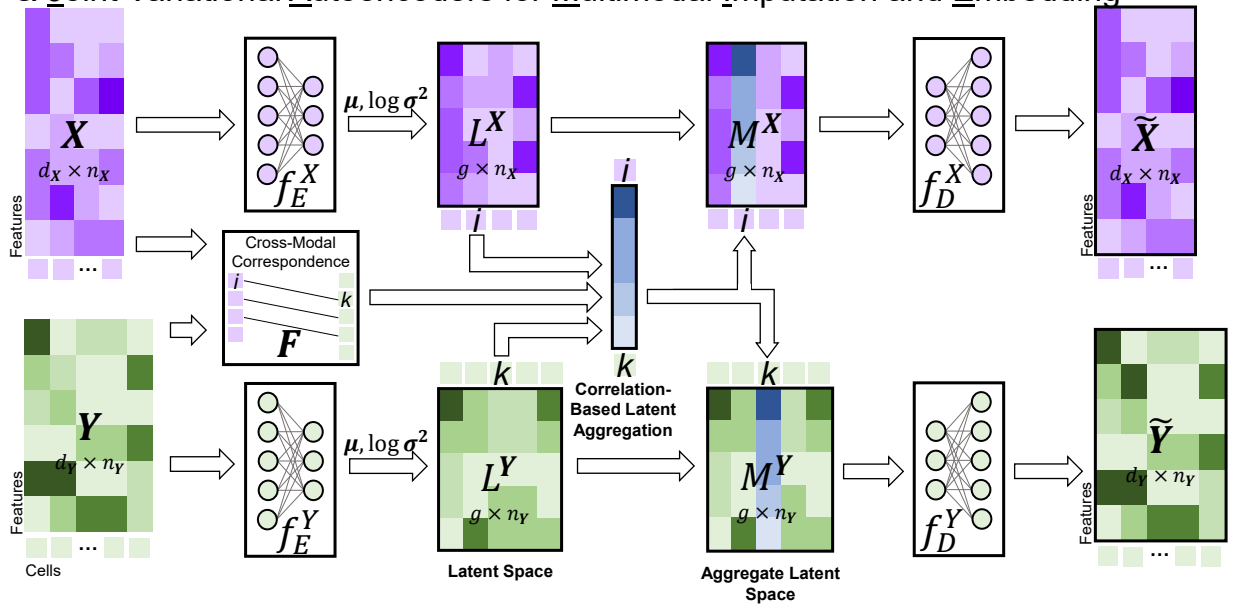
84 2.1 Simulated Multimodal Data

85 We first tested JAMIE on simulated single-cell multimodal data [8]. The simulation data was generated
86 by sampling from a Gaussian distribution on a branching manifold (Figure 3a). We found that the latent
87 embeddings of two modalities in JAMIE preserve the branching structure of the manifold while aligning
88 the cells of the same type in either modality and maintains cell type separation (Figure 3b). To quantify
89 the integration quality, we utilize two metrics: label transfer accuracy (LTA) [7,29], which measures cross-
90 modal phenotype separation, and fraction of samples closer to the true match (FOSCTTM), which measures
91 cross-modal alignment. More details can be found in Subsection 4.9. For separating cell types, JAMIE
92 (LTA= 0.976, FOSCTTM= 0.001) outperforms state of the art alignment methods NLMA (LTA= 0.970,
93 FOSCTTM= 0.001) in LTA and UnionCom (LTA= 0.947, FOSCTTM= 0.079) in both LTA and FOSCTTM
94 (Figure 3c).

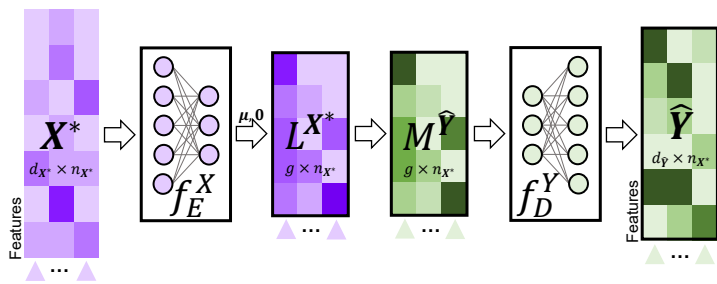
95 Also, we found that the imputed feature values by JAMIE are consistent with the measurements. For
96 instance, as shown in Figure 2d, the imputed values of Modality 1 features have high distribution similarity
97 with the measured values across cell types with average JS distance of 0.428 ± 0.097 (Figure 3d). Two
98 features with high similarity are highlighted with average JS distances 0.278 and 0.281 (Figure 3e), also
99 showing a preservation of expression changes across cell types (i.e., cell type 1 has a lower value than both
100 other cell types). The imputation performance is compared to a baseline method through correlation and
101 AUROC in subfigure f of each results figure. Each dot represents a cell and the axes are the performance
102 for each method. The red line is the space of equal performance between methods. For imputing the first
103 modality, we also see that JAMIE outperforms the baseline KNN using a two-tailed binomial test for 1,490
104 versus 466 features ($p < 1e-100$) in terms of AUROC and 1,147 versus 853 features ($p < 6e-11$) in terms
105 of correlation (Figure 3f). JAMIE exhibits similar outperformance for AUROC when imputing the second
106 modality (Supplementary Figure 1). We see that JAMIE is able to reliably predict features (See select cells
107 in Supplementary Figure 2).

108 We tested JAMIE on non-Gaussian simulation data [18] as well and found that JAMIE performed better
109 than the majority of state-of-the-art methods (Supplementary Figure 3). JAMIE achieved an LTA of 0.952

a Joint Variational Autoencoders for Multimodal Imputation and Embedding



b Cross-Modal Imputation



c Feature Interpretation

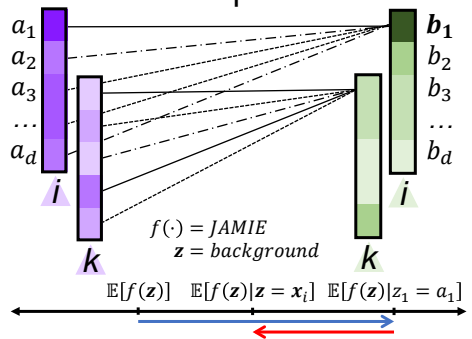


Figure 2. Joint variational Autoencoders for Multimodal Imputation and Embedding (JAMIE) uses variational autoencoders with a novel latent space aggregation technique in order to generate similar latent spaces for each modality. (a) As input, two data matrices \mathbf{X}, \mathbf{Y} are provided with optional *user-provided cross-modal correspondence* matrix \mathbf{F} . These data matrices are fed through encoding layers $f_E^X(\cdot), f_E^Y(\cdot)$ which provide μ_i^X, σ_i^X and μ_j^Y, σ_j^Y for all samples i in \mathbf{X} and j in \mathbf{Y} , respectively. Then, $L_i^X \sim \mathcal{N}(\mu_i^X, \sigma_i^X)$ and $L_j^Y \sim \mathcal{N}(\mu_j^Y, \sigma_j^Y)$ are sampled for all i and j to produce the latent spaces L^X, L^Y . Latent spaces are aggregated using \mathbf{F} to make M^X, M^Y . Finally, the decoding layers produce a reconstructed version of the original modalities $\tilde{\mathbf{X}} = f_D^X(M^X), \tilde{\mathbf{Y}} = f_D^Y(M^Y)$. The latent embeddings L^X, L^Y of the two modalities can be used in tandem for missing phenotype prediction. (b) The trained model can be reused for cross-modal imputation through coupling of encoders and decoders from different modalities. (c) The imputing function $f(\mathbf{x}) = f_D^Y(L^X) = \mathbf{y}, \mathbf{x} \in \mathbb{R}^{d_x}, \mathbf{y} \in \mathbb{R}^{d_y}$ is assessed using Shapley Additive Explanation values [17] which estimate contribution of each input feature by selectively masking the input feature vector with the background.

110 and FOSCTTM of < 0.001 . The second-best performing method (NLMA) had an LTA of 0.886 and FOS-
 111 CTM of 0.002, similar to the performance of JAMIE with only 75% correspondence information provided
 112 (LTA = 0.876, FOSCTTM = 0.016, [Supplementary Figure 4](#)). Further details on this dataset can be found in
 113 [Supplementary Section 2](#).

114 **2.2 Mouse Neuronal Gene Expression and Electrophysiology**

115 We applied JAMIE to infer cross-modal embeddings and impute gene expression and electrophysiological
 116 (ephys) data (inhibitory neuronal types: Lamp5, Serpinf1, Sst, Vip, Pvalb, Sncg) neuronal cells in the mouse
 117 visual cortex ([2], [Figure 4a](#)).

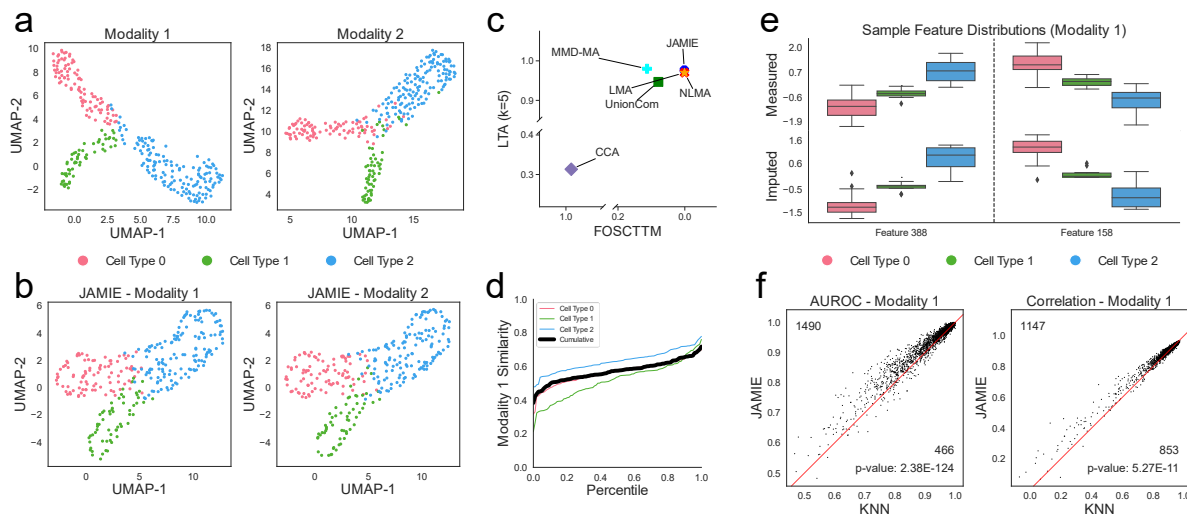


Figure 3. Simulated multimodal data [8]. (a) UMAP of original spaces colored by cell types. (b) UMAP of JAMIE latent spaces. (c) Fraction of samples closer to the true mean (x-axis) and label transfer accuracy (y-axis) of JAMIE and state-of-the-art methods for cell type separation using all correspondence information available. CCA: canonical correlation analysis [15,14], LMA: linear manifold alignment [14], MMD-MA [8], NLMA: nonlinear manifold alignment [14], UnionCom [7]. (d) Cumulative distributions of similarity (1 – JS distance) between measured and imputed feature values in Modality 1. The black line indicates the average similarity across cell types while the colored lines each correspond to a single cell type. (e) Measured (top) and imputed (bottom) values of two select features in Modality 1 across cell types ($n = 300$). Boxes span from the upper to lower quartiles and a line indicates the median. Whiskers extend to extremes up to 1.5 times the interquartile range from the upper and lower quartiles. Any outliers beyond this range are plotted individually. (f) Imputation performance for Modality 1 of JAMIE versus a baseline KNN by AUROC and correlation. We utilize a two-tailed binomial test to generate p-values.

118 We found that both gene and ephys embeddings identified by JAMIE can effectively separate those cell
 119 types (Figure 4b). Further, JAMIE (LTA= 0.944, FOSCTTM= 0.002) outperforms other alignment methods
 120 in both LTA and FOSCTTM, e.g., LMA (LTA= 0.907, FOSCTTM= 0.072) and UnionCom (LTA= 0.887,
 121 FOSCTTM= 0.124) (Figure 4c).

122 Also, JAMIE performs imputation consistently across cell types and generally maintains ephys changes
 123 across cell types with an average JS distance of 0.537 ± 0.115 (Figure 4d). Two ephys features with high
 124 similarity demonstrate this preservation with average JS distances 0.314 and 0.316 (Figure 4e). Finally, ephys
 125 features are imputed significantly better than the baseline KNN, with 33 of 39 features ($p < 2e-5$) performing
 126 better on JAMIE for both AUROC and correlation (Figure 4f). The imputation performance is visualized
 127 for select cells in Supplementary Figure 2.

128 JAMIE was able to generalize to phenotypes beyond cell type, achieving a relatively high LTA (0.650)
 129 when predicting cortical layers compared to other state-of-the-art methods (Supplementary Figure 5). For
 130 reference, NLMA achieved an LTA of 0.663 while LMA, the next best, only achieved an LTA of 0.515.

131 In addition to the mouse visual cortex, we also tested JAMIE on gene expression and ephys features in the
 132 mouse motor cortex. JAMIE maintains cell type separation in the latent space (Supplementary Figure 6b).
 133 JAMIE (LTA= 0.899, FOSCTTM= 0.002) outperforms several methods in both LTA and FOSCTTM, e.g.,
 134 LMA (LTA= 0.897, FOSCTTM= 0.044) (Supplementary Figure 6c). Unioncom (LTA= 0.248, FOSCTTM=
 135 0.445), CCA (LTA= 0.273, FOSCTTM= 0.366), and MMD-MA (LTA= 0.246, FOSCTTM= 0.478) are
 136 unable to align the datasets.

137 JAMIE also preserves ephys changes across cell types, with average JS distance 0.497 ± 0.244 (Supplemen-
 138 tary Figure 6d). Two ephys features with high similarity demonstrate the preserving nature of JAMIE with
 139 JS distances 0.380 and 0.405 (Supplementary Figure 6e). Finally, JAMIE significantly outperforms the base-
 140 line KNN using a two-tailed binomial test with 22 of 29 features for correlation ($p < 9e-3$) (Supplementary
 141 Figure 6f).

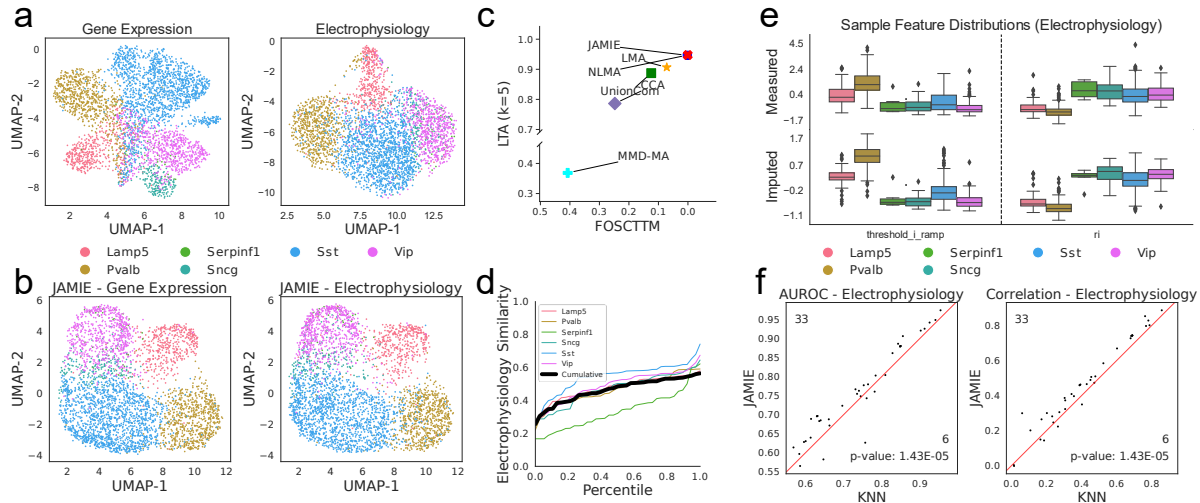


Figure 4. Gene expression and electrophysiological features in the mouse visual cortex [2]. (a) UMAP of single cells by gene expression data (left) and electrophysiological features (right), colored by cell types (inhibitory neuronal types). (b) JAMIE latent spaces of (a). (c) Fraction of samples closer to the true mean (x-axis) and label transfer accuracy (y-axis) of JAMIE and state-of-the-art methods for cell type separation using all correspondence information available. CCA: canonical correlation analysis [15,14], LMA: linear manifold alignment [14], MMD-MA [8], NLMA: nonlinear manifold alignment [14], UnionCom [7]. (d) Cumulative distributions of similarity (1 – JS distance) between measured and imputed electrophysiological features by gene expression. The black line indicates the average similarity across cell types while the colored lines each correspond to a single cell type. (e) Measured (top) and imputed (bottom) values of two electrophysiological features across cell types ($n = 3,654$). Boxes span from the upper to lower quartiles and a line indicates the median. Whiskers extend to extremes up to 1.5 times the interquartile range from the upper and lower quartiles. Any outliers beyond this range are plotted individually. (f) Performance for imputing electrophysiological features of JAMIE versus a baseline KNN by AUROC and correlation. We utilize a two-tailed binomial test to generate p-values.

142 2.3 Gene Expression and Chromatin Accessibility in Human Brain

143 To further investigate the emerging single-cell multiomics data on gene regulation, we apply JAMIE to the
 144 gene expression and chromatin accessibility data of the developing human cerebral cortex (scRNA-seq and
 145 scATAC-seq by 10x Multiome) [3] (Figure 5a). The chromatin accessibility data measures the accessibility
 146 of open chromatin regions (OCRs) by peak signals. OCRs play key epigenomic roles in regulating gene
 147 expression.

148 JAMIE embeddings separate cell types, in contrast to UMAP alone (Figure 5a-b). In comparison of cell
 149 type separations, JAMIE (LTA= 0.959, FOSCTTM< 0.001) also outperforms all other methods in both
 150 LTA and FOSCTTM compared with NLMA (LTA= 0.767, FOSCTTM= 0.002) and LMA (LTA= 0.775,
 151 FOSCTTM= 0.002) (Figure 5c). CCA (LTA= 0.544, FOSCTTM= 0.930) and UnionCom (LTA= 0.458,
 152 FOSCTTM= 0.494) failed to align due to the complexity of the data. We note that JAMIE with provided
 153 75% correspondence information (LTA= 0.951, FOSCTTM= 0.047) and 50% correspondence information
 154 (LTA= 0.936, FOSCTTM= 0.106) also outperform all other methods in terms of LTA (Supplementary
 155 Figure 4).

156 Moreover, we used JAMIE to impute gene expression from OCRs (peaks) and vice versa. Imputed gene
 157 expression values are preserved across cell types with average JS distance 0.484 and standard deviation
 158 0.245 (Figure 5d). We observe wider distributions in the measured data when compared with the imputed
 159 data, likely leading to relatively high JS distance with a large number of cell types. We highlight two genes
 160 with specifically high similarity, having JS distances 0.260 and 0.332 (Figure 5e). Finally, JAMIE significantly
 161 outperforms the imputation baseline using a two-tailed binomial test, better imputing 18,478 of 19,071 genes
 162 for AUROC ($p < 1e-100$) and 20,140 of 22,592 genes for correlation ($p < 1e-100$) from OCRs (Figure 5f).
 163 JAMIE exhibits similar imputation performance when imputing ATAC features from RNA (Supplementary
 164 Figure 7). Imputation of OCRs from gene expression also outperforms the KNN baseline using the same
 165 statistical measure ($p < 1e-100$) (Supplementary Figure 7).

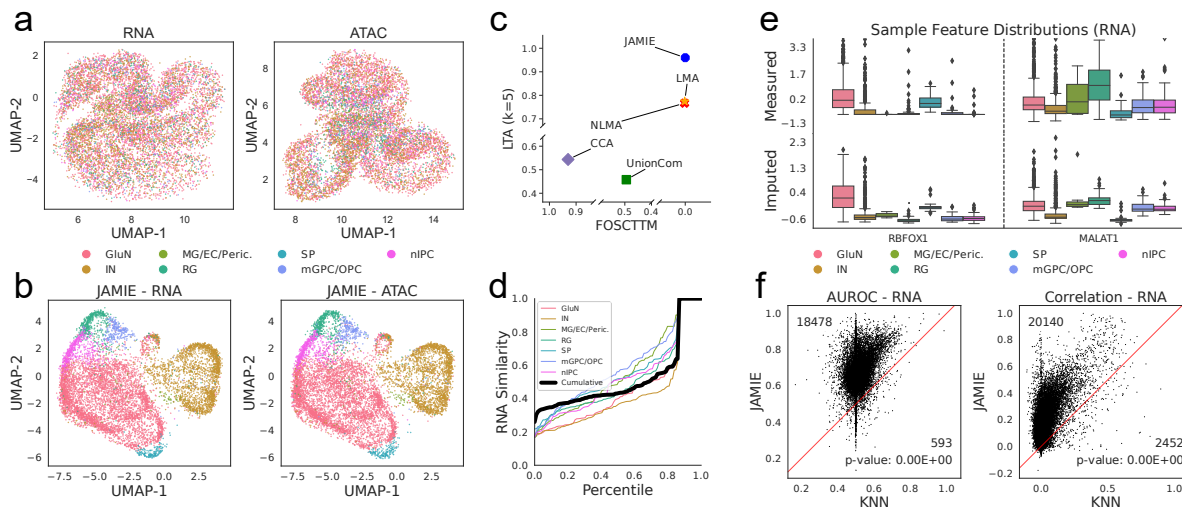


Figure 5. Gene expression and chromatin accessibility of single cells in the developing brain at 21 postconceptional weeks [3]. (a) UMAP of single cells in the human developing brain by gene expression (left) and the accessibility activities of open chromatin regions (right), colored by cell type. (b) JAMIE latent space of a. (c) Fraction of samples closer to the true mean (x-axis) and label transfer accuracy (y-axis) of JAMIE and state-of-the-art methods for cell type separation using all correspondence information available. CCA: canonical correlation analysis [15,14], LMA: linear manifold alignment [14], NLMA: nonlinear manifold alignment [14], UnionCom [7]. (d) Cumulative distributions of similarity (1 - JS distance) between measured and imputed gene expression by chromatin accessibility. The black line indicates the average similarity across cell types while the colored lines each correspond to a single cell type. (e) Measured (top) and imputed (bottom) values of two genes across cell types ($n = 8,981$). Boxes span from the upper to lower quartiles and a line indicates the median. Whiskers extend to extremes up to 1.5 times the interquartile range from the upper and lower quartiles. Any outliers beyond this range are plotted individually. (f) Performance for imputing genes of JAMIE versus a baseline KNN by AUROC and correlation. We utilize a two-tailed binomial test to generate p-values.

166 A strong integration method we mentioned in Section 1 is scGLUE [9], which utilizes a combined au-
 167 toencoder and graph model. To facilitate a comparison between scGLUE and JAMIE, we also ran both
 168 methods on [19], which was utilized in [9]. We found that JAMIE (LTA= 0.859 and FOSCTTM< 0.001)
 169 outperformed scGLUE (LTA= 0.854 and FOSCTTM= 0.038) in both LTA and FOSCTTM (Supplementary
 170 Figure 8). NLMA (LTA= 0.259 and FOSCTTM= 0.035) was unable to achieve cell type separation.

171 2.4 Gene Expression and Chromatin Accessibility in Colon Cancer

172 To further test imputation performance, we compared JAMIE to BABEL [5] using scRNA-seq and scATAC-
 173 seq data in colon adenocarcinoma COLO-320DM cells in Supplementary Figure 9 [20]. JAMIE significantly
 174 outperforms BABEL in imputing both modalities by a two-tailed binomial test. In particular, JAMIE was
 175 better on 12,309 ($p < 1e-100$) and 13,334 ($p < 1e-100$) genes by AUROC and correlation, respectively,
 176 for gene expression imputation from chromatin accessibility (scATAC-seq, OCR peaks) (Supplementary
 177 Figure 9a). For imputing OCRs from gene expression, JAMIE was better on 28,120 ($p < 1e-100$) and 49,936
 178 ($p < 1e-100$) features by AUROC and correlation, respectively (Supplementary Figure 9b).

179 2.5 Biological Interpretability for Cross-Modal Imputation

180 To avoid the black-box nature of many deep learning models, we applied SHAP [17] to prioritize features
 181 for cross-modal imputation (Section 4.12). Specifically, this analysis gives top features in one modality for
 182 imputing a given feature in another modality.

183 As shown in Figure 6a, JAMIE prioritized open chromatin regions (OCRs) for imputing the gene,
 184 DENND1B, which is a gastric cancer related gene on chromosome 1 [21]. Further, keeping OCRs closer
 185 to the location of DENND1B generally results in better imputation performance. For instance, removing

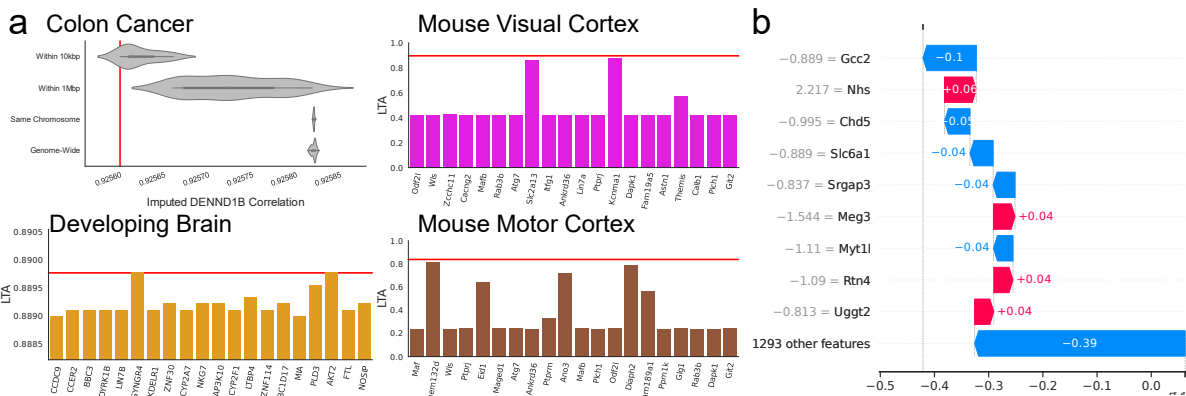


Figure 6. Feature prioritization for cross-modal imputation and embedding. (a) Top left - JAMIE imputation of DENND1B gene with select open chromatin regions removed in colon cancer data [20]. The x-axis is the correlations of the imputed gene expression data with and without select open chromatin regions removed. Aside from the *same chromosome* and *genome-wide* results, all distributions have significantly different means using a one-tailed t-test (All $p < 2e-10$). Top right - LTA for JAMIE with select genes removed in the mouse visual cortex [2]. Bottom left - LTA for JAMIE with select genes removed in the developing brain [3]. Bottom right - LTA for JAMIE with select genes removed in the mouse motor cortex [30]. For each, 14 of the most significant values are displayed along with 6 randomly sampled genes for background. Red Line: Baseline value with no chromatin regions or genes removed. (b) Waterfall plot of select important genes for imputing electrophysiological feature *fast_trough_t_long_square* in mouse visual cortex patch-seq dataset [2]. X-axis: SHAP importance values, Red: Positive SHAP value, Blue: Negative SHAP value.

186 OCRs within 10kb produces lower correlation than elsewhere within the chromosome, suggesting the possi-
 187 ble capability of JAMIE to reveal the importance of genomic proximity from chromatin accessibility to gene
 188 expression.

189 For the developing brain, JAMIE identifies several important genes, including MIA and BBC3 (Both
 190 LTA= 0.889) for contributing JAMIE embeddings to separate the cell types. MIA has been linked to increased
 191 risk for neurodevelopmental disorders [22] and BBC3 has been linked to cell death in the adult brain [23].
 192 Additionally, JAMIE identifies the gene SST (LTA= 0.423) as an important gene in the mouse visual cortex.
 193 SST is known to be directly related to visual discrimination [24]. Many cell type marker genes were found
 194 within the top 200 prioritized genes, with all possible found within the top 400.

195 3 Discussion

196 JAMIE is a novel deep neural network model for cross-modal estimation. It works for complex, mixed, or
 197 partial correspondence multimodal data facilitated by a novel latent embedding aggregation methodology
 198 reliant on a joint variational autoencoder structure. In addition to its outperformances as above, JAMIE
 199 is also computationally efficient and requires little memory usage (Supplementary Table 2). Moreover, the
 200 pretrained model along with learnt cross-modal latent embeddings may be reused for downstream analyses.

201 We evaluated JAMIE using the evaluation metrics that were designed for the alignment methods such
 202 as FOSCTTM and LTA. Despite JAMIE not being explicitly designed to provide common latent spaces
 203 in each modality like alignment, JAMIE’s FOSCTTM remains comparable to, and in many applications
 204 better than, existing methods in this paper. In addition, LTA appears to be better under complex use cases
 205 featuring unfiltered and noisy data. This suggests that the use of modality-specific and aggregated latent
 206 spaces in JAMIE allows for flexibility in forming the latent embeddings for better separate cell types from
 207 noisy samples.

208 In cases with low user-provided correspondence, we found that the alignment loss becomes more impor-
 209 tant. As seen in Supplementary Figure 10a, imputation performance is aided through the inclusion of align-
 210 ment loss with low (50%) user-provided correspondence. In highly-correspondent applications, the alignment
 211 loss may be disregarded to speed up computation by eliminating the computation time of \tilde{F} (Supplementary
 212 Table 2).

213 We have demonstrated the capability of JAMIE to impute between two modalities bidirectionally. In
214 practice, imputing a biological outcome from the source modality is likely to be more useful than imputing
215 a source modality from a biological outcome. Examples include imputing electrophysiological features from
216 gene expression and gene expression from epigenomic regulation (e.g., chromatin accessibility). However,
217 our analyses for partial correspondence provide another imputation use case. As stated in our introduction,
218 missing modalities occur frequently when measuring single-cell multimodal data due to its process complexity.
219 This generates data with partial correspondence, i.e., with unmatched cells between modalities. In this case,
220 we showed that JAMIE with partial correspondence is able to outperform a baseline KNN imputation method
221 with full correspondence (using all samples). For instance, as shown in [Supplementary Figure 10b](#), JAMIE
222 imputation using 50% of user-provided correspondence significantly outperforms the baseline by a two-tailed
223 binomial test using all user-provided correspondence while imputing electrophysiological features from gene
224 expression in the mouse visual cortex ($p < 2e-3$).

225 JAMIE started as a vanilla deterministic autoencoder but was later changed to be variational because of
226 the superior performance of VAEs regarding interpolation in the latent space [25]. Intuitively, the continuous
227 latent space allows for easy sampling and interpolation [26]. Additionally, the proposed aggregation method
228 in [Subsection 4.2](#) relies on the interpolation of latent representations and is improved after switching to a
229 continuous latent space. Though VAEs perform random sampling in the latent space, the process does not
230 add any noticeable randomness to the prediction. Given the same input, the encoder indeed can generate
231 different latent variables due to random sampling during training, but our decoder still reproduces very
232 similar reconstructions. This is because the decoder is trained to map latent samples referring to the same
233 input to very close reconstructions [26]. On the topic of random sampling, the dropout layers do not add
234 randomness to the prediction. They only serve to generalize the model and prevent overfitting by adding
235 redundancy. We chose a default value of $p = 0.6$ to maximize generalizability without adversely affecting
236 performance with default hyperparameters. From [Supplementary Figure 11](#) we see that dropout values of
237 0.4 and 0.6 have roughly equivalent mean performance while 0.8 has slightly degraded results.

238 We can split JAMIE into *Preparation* and *Training* phases. The former is preprocessing, PCA, and cal-
239 culating the inferred correspondence matrix while the latter is the joint VAE model detailed in [Figure 2](#).
240 [Supplementary Table 2](#) shows the consumed time and memory of each phase in seven datasets. For larger
241 datasets, most computation time is spent on the *Preparation* phase. We see that *Training* time scales primar-
242 ily with cell count while peak memory usage generally scales with both cell count and number of features.

243 Over the course of runtime, we also found that the KL-divergence dictated the pace of training, as the
244 other losses converge quickly alone ([Supplementary Figure 12](#)). Visually, we also see that, as κ ([Subsec-
245 tion 4.3](#)) increases, an effect can be seen in the alignment and combination losses while the reconstruction
246 loss steadily declines, indicating that the reconstruction and KL losses dominate the training. Without KL-
247 annealing, we found that other losses were unable to converge, even with the KL-loss having a relatively low
248 weight. This was computed on our non-Gaussian simulation data [18].

249 Training variational autoencoders is time consuming for larger datasets. Thus, prior feature selection
250 such as the automated PCA in JAMIE helps alleviate time requirements. Data preprocessing is also crucial
251 to avoid large or repeated features that can disproportionately shape the features of the low-dimensional
252 embedding, due to the use of reconstruction loss. For cross-modal imputation specifically, the diversity of
253 the training dataset has to be carefully considered to avoid biasing the final model, and negatively affecting
254 the generalizability. JAMIE could also potentially be extended to align datasets from different sources rather
255 than different modalities, such as gene expression measured under different conditions. Because the novel
256 contributions of JAMIE are mainly in the processing of the latent space, many changes in structure and
257 datatype are straightforward to achieve.

258 4 Methods

259 [Figure 2a](#) shows the training process of JAMIE. Two numeric data matrices of modalities X and Y are
260 used as input. $\mathbf{X} \in \mathbb{R}^{d_{\mathbf{X}} \times n_{\mathbf{X}}}$ is a data matrix of modality X with $d_{\mathbf{X}}$ features and $n_{\mathbf{X}}$ samples. The i th
261 column of \mathbf{X} , i.e. $\mathbf{x}_i \in \mathbb{R}^{d_{\mathbf{X}}}$, is the feature vector of the i th sample of modality X . $\mathbf{Y} \in \mathbb{R}^{d_{\mathbf{Y}} \times n_{\mathbf{Y}}}$ is a data
262 matrix of modality Y with $d_{\mathbf{Y}}$ features and $n_{\mathbf{Y}}$ samples. Similarly, the j th column of \mathbf{Y} , i.e. $\mathbf{y}_j \in \mathbb{R}^{d_{\mathbf{Y}}}$, is
263 the feature vector of the j th sample of modality Y . Each row of \mathbf{X} and \mathbf{Y} is standardized, e.g., $mean = 0$
264 and $standard\ deviation = 1$. If available, a user-provided correspondence matrix $\mathbf{F} \in \mathbb{R}_{\geq 0}^{n_{\mathbf{X}} \times n_{\mathbf{Y}}}$ is used to aid

265 construction of the similar latent spaces during training, where $F_{ij} = 1$ implies one-to-one correspondence
266 (i.e. same cells different modalities) between cell i in modality X and cell j in modality Y , $F_{ij} = 0$ implies
267 no known correspondences, and $0 < F_{ij} < 1$ implies partial confidence of correspondence.

268 Jamie utilizes joint variational autoencoders to learn similar latent spaces for each of the two modalities,
269 $L^X \in \mathbb{R}^{g \times n_X}$ and $L^Y \in \mathbb{R}^{g \times n_Y}$, where g is the selected dimensionality of the latent space. Higher g will
270 lead to better reconstruction loss, but the features will not be as dense information-wise. Lower g will
271 condense information more efficiently, but at the cost of reconstruction and imputation performance. We
272 experimentally found $g = 32$ to be robust across multiple datasets, and use this as the default for JAMIE.
273 During training, we utilize correlation-based latent aggregation. In order to learn these latent spaces, JAMIE
274 minimizes the following losses:

$$275 \ell_{\text{total}} = \ell_{\text{kl}} + \alpha \ell_{\text{combination}} + \beta \ell_{\text{reconstruction}} + \gamma \ell_{\text{alignment}}. \quad (1)$$

276 ℓ_{kl} takes the Kullback-Leibler (KL) divergence between the inferred distribution of our variational au-
277 toencoders (VAEs) and the multivariate standard normal, aiding with the continuity of the latent space.
278 $\ell_{\text{combination}}$ enforces similarity of correspondent samples and takes the weight α . $\ell_{\text{reconstruction}}$ is the summed
279 mean squared error between the reconstructions and original data matrices and is assigned the weight β .
280 $\ell_{\text{alignment}}$ uses inferred cross-modal correspondence to shape the resultant latent space and is assigned the
281 weight γ . The weights are user-defined and their default values are detailed in [Subsection 4.8](#). An overview
282 of the JAMIE algorithm can be seen in [Figure 2](#).

283 The default hyperparameters work for a wide variety of cases, and are used for our applications unless
284 otherwise mentioned, but may be customized to the user's preference.

285 ℓ_{kl} affects the continuity of the latent space and aids in interpolation and stability. This is particularly
286 important for imputation applications. Although this effect is not obvious from the integrated space itself
287 ([Supplementary Figure 13a](#)), the imputation performance is significantly improved with this loss ([Supple-
288 mentary Figure 13b](#)).

289 $\ell_{\text{combination}}$ controls the focus on similarity between correspondent features across modalities. A higher
290 weight will result in closer cross-modal cell representations at the risk of losing per-modality detail. $\ell_{\text{reconstruction}}$
291 controls the ability of the latent space to encode raw information from each modality. A larger weight will
292 make cell representations more adherent to their single-modal representations at the cost of cross-modal rep-
293 resentation similarity of correspondent samples. The effects of these losses are very pronounced when removed
294 from the training process, and result in lower integrated space reliability ([Supplementary Figure 13a](#)).

295 Finally, $\ell_{\text{alignment}}$ is primarily helpful for low-correspondence applications where the shape of the latent
296 space is not clear from context. A higher weight will increase the similarity of cross-modal correspondent
297 cell representations by changing the shape of the resultant latent space. This can come at the cost of cell
298 type mixing and cluster accuracy but can result in better imputation performance in low-correspondence
299 applications ([Supplementary Figure 10a](#)).

300 The latent dimension g can be lowered for more compressed representations, potentially at a loss in
301 accuracy ([Supplementary Figure 14a](#)). This can be useful in applications which present overfitting, normally
302 when the dataset has very few features or samples. The number of principal components used in the JAMIE
303 model can be adjusted to change model size, which can speed up calculation at the cost of performance
304 ([Supplementary Figure 14b](#)). This may also be lowered to avoid overfitting.

305 4.1 Distribution of Latent Space Features

306 Intermediate feed-forward neural networks are utilized to generate distributions for each feature in the
307 g -dimensional latent space for each modality. These will be henceforth referred to as encoders, denoted
308 $f_E^X(\cdot)$, $f_E^Y(\cdot)$ for each modality. The variational encoders have four hidden layers. The encoder of modality
309 X transforms the feature dimension as follows: $n_X \rightarrow 2n_X \rightarrow n_X \rightarrow g \rightarrow 2g$. For modality X , the top
310 half of the fourth layer output corresponds to means and the bottom half corresponds to log variance per
311 sample and feature as in [Equation 2a](#). The encoder is similar for Y . Batch normalization and leaky ReLU
312 activations are included wherever possible, and dropout layers are added ($p = 0.6$) for imputation models.

313 For each sample \mathbf{x}_i and \mathbf{y}_j , we sample the multivariate Gaussian

$$314 \quad [f_E^X(\mathbf{X})]_i = \begin{bmatrix} \boldsymbol{\mu}_i^X \\ \boldsymbol{\sigma}_i^X \end{bmatrix} \quad \mathbf{L}_i^X \sim \mathcal{N}(\boldsymbol{\mu}_i^X, \text{diag}(\boldsymbol{\sigma}_i^X)) \quad \boldsymbol{\mu}_i^X, \mathbf{L}_i^X \in \mathbb{R}^g \quad \boldsymbol{\sigma}_i^X \in \mathbb{R}_{>0}^g \quad i \in \{1, 2, \dots, n_X\}, \quad (2a)$$

$$315 \quad [f_E^Y(\mathbf{Y})]_j = \begin{bmatrix} \boldsymbol{\mu}_j^Y \\ \boldsymbol{\sigma}_j^Y \end{bmatrix} \quad \mathbf{L}_j^Y \sim \mathcal{N}(\boldsymbol{\mu}_j^Y, \text{diag}(\boldsymbol{\sigma}_j^Y)) \quad \boldsymbol{\mu}_j^Y, \mathbf{L}_j^Y \in \mathbb{R}^g \quad \boldsymbol{\sigma}_j^Y \in \mathbb{R}_{>0}^g \quad j \in \{1, 2, \dots, n_Y\}, \quad (2b)$$

317 where $[f_E^X(\mathbf{X})]_i$ and $[f_E^Y(\mathbf{Y})]_j$ are the encoded results of \mathbf{x}_i and \mathbf{y}_j , respectively, and $\text{diag}(\cdot)$ is a diagonal
 318 matrix generated from a vector. \mathbf{L}_i^X and \mathbf{L}_j^Y are sampled from the multivariate Gaussian distributions
 319 indicated above. We define $\mathbf{L}_i^X \in \mathbb{R}^g$ as the i th column of \mathbf{L}^X , representing the feature vector of the i th
 320 sample. $\mathbf{L}_j^Y \in \mathbb{R}^g$ is defined similarly.

321 The VAE is able to approximately map any non-Gaussian input data onto the multivariate Gaussian
 322 latent space (using the encoder) and reconstruct the input data from the space using the decoder [26].
 323 Therefore, the multimodal input to JAMIE is not restricted to a Gaussian distribution.

324 4.2 Aggregation of the Latent Spaces

325 Using our user-provided correspondence matrix $\mathbf{F} \in \mathbb{R}^{n_X \times n_Y}$, we can perform an aggregate calculation to
 326 combine the latent space vectors of known aligned points. In particular,

$$327 \quad \mathbf{M}^X = (\mathbf{L}^X + \delta \mathbf{L}^Y \mathbf{F}^T) (\text{diag}(1 + \delta \mathbf{F} \mathbf{1}_{n_Y}))^{-1}, \quad (3a)$$

$$328 \quad \mathbf{M}^Y = (\mathbf{L}^Y + \delta^{-1} \mathbf{L}^X \mathbf{F}) (\text{diag}(1 + \delta^{-1} \mathbf{F}^T \mathbf{1}_{n_X}))^{-1}, \quad (3b)$$

330 where the aggregate feature vector $\mathbf{M}^X \in \mathbb{R}^{g \times n_X}$ has g features and n_X samples and $\mathbf{M}^Y \in \mathbb{R}^{g \times n_Y}$ has g
 331 features and n_Y samples. $\mathbf{1}_{n_x} \in \mathbb{R}^{n_x}$ and $\mathbf{1}_{n_y} \in \mathbb{R}^{n_y}$ are vectors of all ones. δ refers to the relative weighting
 332 between modalities X and Y in the aggregation function. We learn δ over the course of training. A value
 333 $\delta < 1$ implies that \mathbf{X} is weighted more than \mathbf{Y} during aggregation, while $\delta > 1$ implies the opposite. The
 334 construction of $\mathbf{M}^X, \mathbf{M}^Y$ is shown visually in Figure 2a in the formulation of the blue vector. Figure 2a
 335 shows the case $\mathbf{F}_{ik} = 1$ with all other entries in row i and column k being 0, which results in simply averaging
 336 the latent feature vectors of \mathbf{x}_i and \mathbf{y}_k .

337 The linear aggregation technique is backed by the nonlinear transformation of the encoder, which is able
 338 to account for varying timings, magnitude changes, and distributions due to its deep structure. In Figure 2,
 339 this is the blue latent space vector combination of correspondent samples. We term the technique *correlation-*
 340 *based latent aggregation*. This moves correspondent latent embeddings in similar directions over the course
 341 of training and is key in the formation of similar latent spaces. Correlation-based latent aggregation is the
 342 primary motivation for the use of the variational autoencoder framework, as it allows for this intuitive
 343 representation of partially aligned datasets. We can now adjust our low-dimensional data embeddings by
 344 using aggregates wherever possible to produce our final aggregate latent spaces \mathbf{M}^X and \mathbf{M}^Y .

345 4.3 Continuity of the Aggregate Latent Spaces

346 Continuous \mathbf{L}^X and \mathbf{L}^Y facilitate imputation, thus, we adopt the VAE architecture [25] and the correspond-
 347 ing KL loss:

$$348 \quad \ell_{\text{kl}} = \frac{\kappa}{g} \sum_{k=1}^g \mathbf{d}_k, \quad (4)$$

349 where \mathbf{d}_k takes the KL divergence of the k th feature between the distribution of our latent space estimate
 350 and a multivariate standard normal as the targeted distribution:

$$351 \quad \mathbf{d}_k = -\frac{1}{2n_X} \sum_{i=1}^{n_X} (1 + 2 \log([\boldsymbol{\sigma}_i^X]_k) - ([\boldsymbol{\mu}_i^X]_k)^2 - ([\boldsymbol{\sigma}_i^X]_k)^2) \\ - \frac{1}{2n_Y} \sum_{j=1}^{n_Y} (1 + 2 \log([\boldsymbol{\sigma}_j^Y]_k) - ([\boldsymbol{\mu}_j^Y]_k)^2 - ([\boldsymbol{\sigma}_j^Y]_k)^2). \quad (5)$$

352 The parameter κ is slowly increased from 0 to 1 over the course of training to accelerate convergence [27],
 353 further detailed in [Subsection 4.7](#). This process is further referred to as *KL-annealing*. KL divergence is
 354 described in [Subsection 4.7](#).

355 ℓ_{kl} will lower to zero as the latent space’s distribution approaches the standard multivariate Gaussian.
 356 This encourages a continuous latent space through the nonzero standard deviations. The mean approaching
 357 zero encourages a small, clustered latent space. This will negatively affect metrics such as silhouette distance,
 358 but can aid in imputing outlier or unseen single-cell data.

359 4.4 Similarity of the Original and Aggregate Latent Spaces

360 The structure alone with JAMIE already produces shared latent spaces M^X and M^Y . To balance L^X and
 361 L^Y and keep both latent spaces the same scale, we implement a loss proportional to the square norm of the
 362 difference between each latent space and its aggregated matrix M :

$$363 \ell_{\text{combination}} = \frac{1}{gn_X} \|L^X - M^X\|_F^2 + \frac{1}{gn_Y} \|L^Y - M^Y\|_F^2. \quad (6)$$

364 The combination loss makes the aggregate matrices as similar as possible to the originals while still main-
 365 taining cross-modal corresponding cell representation similarity. In practice, as long as the two modalities
 366 contain unique information and are non-trivially corresponding, the combination and reconstruction losses
 367 will balance each-other out, as the reconstruction loss ensures that information unique to each modality is
 368 encoded in the latent representations. This balance allows JAMIE to form the latent space flexibly while
 369 preserving as much single-modal detail as possible.

370 4.5 Similarity of the Original and Reconstructed Data Matrices

371 Decoders are structured in the reverse manner as encoders. It uses fully connected layers to transform the
 372 feature dimension as follows: $g \rightarrow n_X \rightarrow 2n_X \rightarrow n_X$. The structure is similar for Y . Batch normalization and
 373 leaky ReLU are included where possible, and dropout layers are added ($p = 0.6$) for imputation models. We
 374 denote the decoders $f_D^X(\cdot)$, $f_D^Y(\cdot)$ for each modality and use each to get an estimate of the original datasets
 375 $\tilde{X} = f_D^X(M^X) \in \mathbb{R}^{d_X \times n_X}$ and $\tilde{Y} = f_D^Y(M^Y) \in \mathbb{R}^{d_Y \times n_Y}$. The quality of the reconstruction, and thereby
 376 the information retained in the latent space, can be evaluated by:

$$377 \ell_{\text{reconstruction}} = \frac{1}{d_X n_X} \|X - \tilde{X}\|_F^2 + \frac{1}{d_Y n_Y} \|Y - \tilde{Y}\|_F^2. \quad (7)$$

378 4.6 Correspondence of the Aggregate Latent Spaces

379 We find the optimal inferred correspondence matrix $\tilde{F} \in \mathbb{R}_{\geq 0}^{n_X \times n_Y}$ by solving the following optimization
 380 problem:

$$381 \begin{aligned} & \min_{\alpha, \tilde{F}} \|\alpha K_X - \tilde{F} K_Y \tilde{F}^T\|_F^2 \\ & \text{s.t. } \tilde{F} \geq 0, \quad \alpha \geq 0, \\ & \tilde{F} \mathbf{1}_{n_Y} = \mathbf{1}_{n_X}, \quad \tilde{F}^T \mathbf{1}_{n_X} \leq \mathbf{1}_{n_Y}, \end{aligned} \quad (8)$$

382 where $[K_X, K_Y] = [dist(X), dist(Y)]$ are the intra-dataset distance. This generally uses geodesic distance
 383 and is similar to the quadratic assignment problem.

384 The procedure originates from an extension on [28], which optimized a similar expression but assumed
 385 one-to-one correspondence on inter-dataset cell pairs. [7] extended this to the concept of one-to-many corre-
 386 spondence between modalities and created the implementation in [Equation 8](#).

387 Once solved, \tilde{F}_{ij} denotes the probability that sample x_i is matched with sample y_j . The objective of \tilde{F}
 388 is to provide complete correspondence information. \tilde{F} will generally contain only one-to-one relationships,
 389 while true inter-modal relationships are often layered and complex. In practice, we may consider combining
 390 F and \tilde{F} by a weighted average if the user-provided correspondence is especially low (see [Algorithm 1](#)).

391 To ensure that correspondent samples have similar aggregate latent representations we introduce $\ell_{\text{alignment}}$.
 392 This metric relies on our inferred correspondence \tilde{F} rather than the user-provided correspondence matrix

393 $\tilde{\mathbf{F}}$. Unlike \mathbf{F} , our inferred correspondence $\tilde{\mathbf{F}}$ contains information about samples which are likely to not
394 correspond. We can utilize this dense information to shape our latent space such that contributing samples
395 are clustered:

$$396 \ell_{\text{alignment}} = \frac{1}{gn_{\mathbf{X}}} \|\mathbf{M}^{\mathbf{X}} - \mathbf{M}^{\mathbf{Y}} \tilde{\mathbf{F}}^T\|_{\mathbf{F}}^2. \quad (9)$$

397 The alignment loss is primarily effective when the user-provided correspondence is weak (containing few
398 highly-correspondent pairs). This loss shapes the integrated latent spaces and aids in obtaining cross-modal
399 latent similarity by inferring correspondence using unsupervised manifold alignment. Any unsupervised cor-
400 respondence method may be substituted for $\tilde{\mathbf{F}}$.

401 4.7 Early Stopping and KL Annealing

402 Early stopping is performed by keeping track of the best ℓ_{total} over the course of training. JAMIE is given
403 a minimum number of epochs (default, 2,500) over which time no early stopping can take place. Over this
404 period, κ is gradually increased from 0 to 1. The concept originates from [27]. Once this period is over, the
405 training continues until the algorithm does not produce a new best loss for a set number of epochs (default,
406 500) or until the maximum number of epochs is achieved (default, 10,000). This process is further detailed
407 in [Algorithm 1](#).

408 4.8 JAMIE Algorithm

409 In [Algorithm 1](#), datasets \mathbf{X} and \mathbf{Y} are provided pre-standardized along with an optional correspondence
410 matrix \mathbf{F} as described in [Section 4](#). The user-given variables are τ_{min} (default, 2,500) τ_{max} (default, 10,000),
411 which are the minimum and maximum number of epochs for training. b_s (default, 512) is the batch size.
412 (α, β, γ) are the weights for the model with defaults (1000, 31.25, 31.25). $\tau_{\text{max_lapses}}$ (default, 500) $\ell_{\text{min_change}}$
413 (default, 1e-8) control the behavior of the early-stopping algorithm.

414 ρ is an experimental feature which controls how much the inferred correspondence matrix is utilized
415 over the user-provided correspondence. The weighted average $\rho\mathbf{F} + (1 - \rho)\tilde{\mathbf{F}}$ is used in [Equation 3a](#) and
416 [Equation 3b](#). By default, $\rho = 1$. This is not explored in-depth within the paper, but is a useful feature for
417 low-correspondence applications.

418 4.9 Performance Evaluation

419 **Label Transfer Accuracy (LTA):** To assess phenotype (e.g., cell type, cortical layer) separation, we
420 use label transfer accuracy (LTA) [7,29]. LTA measures the phenotype-identification accuracy of a nearest
421 neighbor classifier trained on the second modality and evaluated on the first. Higher values generally indicate
422 better phenotype separation.

423 **Fraction of Samples Closer to the True Match (FOSCTTM):** We use this metric to evaluate the
424 average fraction of samples closer to the truly aligned sample (in both modalities) than its actual pairing.
425 This provides an estimate of how closely the two modalities are aligned. Lower values generally indicate
426 closer cross-modal alignment.

427 **Pearson Correlation:** We use Pearson correlation mainly to compare imputed and measured feature
428 values, providing an estimate of how well the imputation method captures the variance of the measured
429 feature values. For the purposes of imputation, higher values usually indicate better performance.

430 **Area under the ROC curve (AUROC):** We use AUROC to evaluate imputed features for prediction
431 of high or low measured expression values. To test this, we use AUROC on the median-binarized measured
432 data, that is, each feature value above the feature median is treated as 1, and all others are treated as 0.

433 **Jensen-Shannon Distance (JS distance):** It is also important that an imputation method preserves
434 individual cell type distributions. We use JS distance, which is the square root of the Jensen-Shannon
435 divergence. Jensen-Shannon divergence is the average KL divergence between two distributions and their
436 average (a symmetrized version of KL divergence). KL divergence measures the difference between two
437 probability distributions, and increases proportionally. As a baseline for imputation performance, a KNN
438 model ($k = 5$) is used to generate new multimodal data based on the average of nearest neighbors in the
439 training set.

Algorithm 1 JAMIE

Input: $\mathbf{X}, \mathbf{Y}, \mathbf{F}$
Params: $\tau_{\min}, \tau_{\max}, b_s, (\alpha, \beta, \gamma), \rho, \tau_{\max_lapses}, \ell_{\min_change}$
Output: $f_E^X(\cdot), f_E^Y(\cdot), f_D^X(\cdot), f_D^Y(\cdot), \mathbf{L}^X, \mathbf{L}^Y$

- 1: $[\mathbf{K}_X, \mathbf{K}_Y] = [dist(\mathbf{X}), dist(\mathbf{Y})]$
- 2: $\tilde{\mathbf{F}} = \arg \min_{\alpha, \tilde{\mathbf{F}}} \|\alpha \mathbf{K}_X - \tilde{\mathbf{F}} \mathbf{K}_Y \tilde{\mathbf{F}}^T\|_F^2$
- 3: $\mathbf{F} \leftarrow \rho \mathbf{F} + (1 - \rho) \tilde{\mathbf{F}}$
- 4: $\ell_{\text{best}} \leftarrow \infty, \tau_{\text{lapses}} \leftarrow 0$
- 5: $\tau_{\text{current}} \leftarrow 0$
- 6: Initialize $f_E^X, f_E^Y, f_D^X, f_D^Y$ with random weights
- 7: **while** $\tau_{\text{current}} < \tau_{\max}$ and $\tau_{\text{lapses}} < \tau_{\max_lapses}$ **do**
- 8: Randomly sample $\mathbf{X}_s \in \mathbb{R}^{d_X \times b_s}$ from $\mathbf{X} \in \mathbb{R}^{d_X \times n_X}$ and $\mathbf{Y}_s \in \mathbb{R}^{d_Y \times b_s}$ from $\mathbf{Y} \in \mathbb{R}^{d_Y \times n_Y}$
- 9: Extract submatrix $\mathbf{F}_s \in \mathbb{R}^{b_s \times b_s}$ from $\mathbf{F} \in \mathbb{R}^{n_X \times n_Y}$ and $\tilde{\mathbf{F}}_s \in \mathbb{R}^{b_s \times b_s}$ from $\tilde{\mathbf{F}} \in \mathbb{R}^{n_X \times n_Y}$
- 10: Optionally transform $\mathbf{X}_s, \mathbf{Y}_s$ using PCA trained on \mathbf{X}, \mathbf{Y}
- 11: Normalize columns of \mathbf{F}_s and $\tilde{\mathbf{F}}_s$ to summation 1
- 12: Randomly sample $L_i^{X_s}$ from distributions $f_E^X(\mathbf{X}_s)$ and $L_j^{Y_s}$ from distributions $f_E^Y(\mathbf{Y}_s)$ for $i, j \in \{1, 2, \dots, b_s\}$
- 13: $\mathbf{M}^{X_s} \leftarrow (\mathbf{L}^{X_s} + \delta \mathbf{L}^{Y_s} \mathbf{F}_s^T) (\text{diag}(1 + \delta \mathbf{F}_s \mathbf{1}_{b_s}))^{-1}$
- 14: $\mathbf{M}^{Y_s} \leftarrow (\mathbf{L}^{Y_s} + \delta^{-1} \mathbf{L}^{X_s} \mathbf{F}_s) (\text{diag}(1 + \delta^{-1} \mathbf{F}_s^T \mathbf{1}_{b_s}))^{-1}$
- 15: $\tilde{\mathbf{X}}_s \leftarrow f_D^X(\mathbf{M}^{X_s})$ and $\tilde{\mathbf{Y}}_s \leftarrow f_D^Y(\mathbf{M}^{Y_s})$
- 16: If $\mathbf{X}_s, \mathbf{Y}_s$ were transformed using PCA, perform inverse transform on $\tilde{\mathbf{X}}_s, \tilde{\mathbf{Y}}_s$
- 17: $\kappa \leftarrow \left(1 + \exp\left(\frac{-5(\tau_{\text{current}} - 0.5\tau_{\min})}{0.5\tau_{\min}}\right)\right)^{-1}$
- 18: Calculate ℓ_{kl} via [Equation 4](#)
- 19: Calculate $\ell_{\text{combination}}$ via [Equation 6](#)
- 20: Calculate $\ell_{\text{reconstruction}}$ via [Equation 7](#)
- 21: Calculate $\ell_{\text{alignment}}$ via [Equation 9](#)
- 22: $\ell_{\text{total}} \leftarrow \ell_{\text{kl}} + \alpha \ell_{\text{combination}} + \beta \ell_{\text{reconstruction}} + \gamma \ell_{\text{alignment}}$
- 23: **if** $\tau_{\text{current}} \geq \tau_{\min}$ **then**
- 24: **if** $\ell_{\text{total}} < \ell_{\text{best}} - \ell_{\min_change}$ **then**
- 25: $\ell_{\text{best}} \leftarrow \ell_{\text{total}}$
- 26: $\tau_{\text{lapses}} \leftarrow 0$
- 27: **else**
- 28: $\tau_{\text{lapses}} \leftarrow \tau_{\text{lapses}} + 1$
- 29: **end if**
- 30: **end if**
- 31: Update $f_E^X, f_E^Y, f_D^X, f_D^Y$ using gradient descent on ℓ_{total}
- 32: $\tau_{\text{current}} \leftarrow \tau_{\text{current}} + 1$
- 33: **end while**
- 34: $\mathbf{L}^X \leftarrow f_E^X(\mathbf{X}), \mathbf{L}^Y \leftarrow f_E^Y(\mathbf{Y})$

440 **p-Values:** When we compare two methods in terms of imputation performance, we utilize the two-tailed
441 binomial test with the null hypothesis that both methods have an equal probability of imputing better on
442 any given feature.

443 **4.10 Hyperparameters and Validation**

444 Principal Component Analysis (PCA) is used to lower the dimensionality of data with many features. The
445 default value for this preprocessing is 512 features, and is automatically inverted when applying imputation.
446 The hyperparameters used for all experiments in the paper are the default losses detailed in [Section 4](#), the
447 minimal and maximal iterations detailed in [Subsection 4.7](#), and a batch size of 512.

448 When performing integration in [Section 2](#), no data is withheld, hyperparameters are left at default, and
449 dropout is set to $p = 0$ to facilitate maximal learning of single-modal nuances, regardless of dataset specificity.
450 This is done because many alignment methods are unable to be reused, and perform integration only once
451 before having to be recalculated. For imputation, hyperparameters are left unchanged but dropout is set
452 at $p = 0.6$ to facilitate generalization. 20 percent of cells within each dataset are withheld at random. Any
453 imputation performance calculations, such as those in [Figure 3f](#), are performed on this withheld data.

4.11 Datasets and Preprocessing

We applied JAMIE to four main single-cell multimodal datasets: (1) simulated multimodal data generated by sampling from a Gaussian distribution on a branching manifold in MMD-MA (N=300, 3 cell types) [8]; (2) Patch-seq gene expression and electrophysiological features of single neuronal cells in the mouse visual cortex (N=3,654, 6 cell types) [2] and mouse motor cortex (N=1,208, 9 cell types) [30]; (3) 10x Multiome gene expression and chromatin accessibility data of 8,981 cells in the human developing brain (21 postconceptional weeks, covering 7 major cell types in human cerebral cortex) [3]; (4) scRNA-seq gene expression and scATAC-seq chromatin accessibility data of 4,301 cells from the COLO-320DM colon adenocarcinoma cell line (which was also analyzed by BABEL [5]) [20]. The input datasets to JAMIE were standardized (mean = 0, standard deviation = 1) across features and large numbers of input features were reduced dimensionality by PCA to 512 PCs as needed for efficient computation.

4.12 Applications

Phenotype Prediction from Multimodal Data: Integration on multimodal datasets can improve classification, knowledge of phenotype, and understanding of complex biological mechanisms. Given two datasets \mathbf{X}, \mathbf{Y} and correspondence \mathbf{F} , we generate $\mathbf{L}^{\mathbf{X}}, \mathbf{L}^{\mathbf{Y}}$ from the trained encoders $f_E^{\mathbf{X}}(\cdot), f_E^{\mathbf{Y}}(\cdot)$ as in Subsection 4.1. Clustering can be performed or classification methods may be directly built atop the reusable latent space. Clustering on these latent spaces confers several advantages, the main of which being the incorporation of both modalities in the process of feature generation. We can then actively predict which samples correspond (if not already known) and perform related tasks such as cell type prediction. On partially-annotated datasets, cells in similar clusters should be of similar cell type. There is generally no need to use a complicated clustering or classification algorithm, as JAMIE should do most of the separation as a part of the latent space generation. The networks may be further analyzed to provide a clearer picture of the relationship between data features and phenotype. To visualize this cell type clustering, UMAP [31] is performed on high-dimensional data for visualization.

Cross-Modal Imputation: There are several methods to perform cross-modal imputation, but many methods are not able to prove that they have learned underlying biological mechanisms for the purpose of imputation. When utilized for cross-modal imputation, we can predict our missing data with a more rigorous foundation than that offered by a feed-forward network or linear regression. Given training data \mathbf{X}, \mathbf{Y} , we can train the model. With new data $\mathbf{X}^* \in \mathbb{R}^{d_{\mathbf{X}} \times n_{\mathbf{X}^*}}$, we can predict its correspondent latent embedding $\hat{\mathbf{Y}} = f_D^{\mathbf{Y}}(\mathbf{L}^{\mathbf{X}^*}) \in \mathbb{R}^{d_{\mathbf{Y}} \times n_{\mathbf{X}^*}}$ which has true values \mathbf{Y}^* . Then, we have predicted correspondent cells using a latent space which likely confers understanding of the relationship between data features and phenotype.

This has motivation within the given losses. The alignment loss incentivizes the low-dimensional embeddings for each modality to be on the same latent space. Additionally, since we are predicting correspondent cells, $\mathbf{F} = \mathbb{I}$. Ideally then, $\mathbf{L}^{\mathbf{X}^*} = \mathbf{M}^{\mathbf{X}^*} = \mathbf{M}^{\mathbf{Y}^*} = \mathbf{L}^{\mathbf{Y}^*}$. So, if we assume perfect and total alignment, $f_D^{\mathbf{Y}}(\mathbf{L}^{\mathbf{X}^*}) = f_D^{\mathbf{Y}}(\mathbf{L}^{\mathbf{Y}^*})$. This prediction is shown in Figure 2b. Of course, this ideal scenario is not always the case. The efficacy of this cross-modal imputation method is explored within the results section of this paper, where 20 percent of the data is taken as validation, and the model is trained on the remainder.

Latent and Imputed Feature Interpretation: For interpretation of our trained model, we employ [17], henceforth referred to as SHAP. SHAP looks at individual predictions produced by a model and will assess importance of various input features through sample modulation. This can be used for a variety of interesting applications. If a target variable is easily separable by phenotype, SHAP can determine relevant features for further study. Additionally, if we perform imputation, SHAP can expose inter-modal linkages learned by the model. Given a model f and sample $\mathbf{x} \in \mathbb{R}^n$, SHAP values $\phi_0, \dots, \phi_{n-1}$ are learned such that $\mathbb{E}[f(\mathbf{z})] + \sum_{i \in \{j, \dots, m\}} \phi_i = \mathbb{E}[f(\mathbf{z}) | \mathbf{z}_j = x_j, \dots, \mathbf{z}_m = x_m]$ for a background feature vector $\mathbf{z} \in \mathbb{R}^n$. If $\{j, \dots, m\} = \{0, 1, \dots, n-1\}$, then the sum of SHAP values and the background output will equal $f(\mathbf{x})$, with each ϕ_i proportional to the impact of x_i on the model output.

Another useful technique involves choosing a key metric for classification (e.g., LTA) or imputation (e.g., Correspondence between imputed and measured features) and evaluating the metric with each feature sequentially removed (replaced with its background) from the model. Then, if the key metric becomes worse, it suggests that the removed feature is more important to the outcome of the model.

504 **5 Data Availability**

505 The MMD-MA simulation dataset can be downloaded from <https://noble.gs.washington.edu/proj/mmd-ma/>. Our simulation data may be downloaded from <https://github.com/daifengwanglab/JAMIE>.
506
507 Processed Patch-seq gene expression and electrophysiological features for the mouse visual and motor
508 cortices are available at <https://github.com/daifengwanglab/scMNC>. Raw Patch-seq datasets are avail-
509 able at [30,2]. Single-cell RNA-seq and ATAC-seq data on the human developing brain can be down-
510 loaded at <https://github.com/GreenleafLab/brainchromatin/blob/main/links.txt> under the head-
511 ing *Multiome*. Single-cell RNA-seq and ATAC-seq of colon adenocarcinoma data can be found at <https://github.com/wukevin/babel>. Processed datasets for SNARE-seq adult mouse cortex data can be down-
512 loaded from <https://scglue.readthedocs.io/en/latest/data.html>.
513

514 **6 Code Availability**

515 All code was implemented in Python using PyTorch and the source code is publicly available at <https://github.com/daifengwanglab/JAMIE> [32]. Since Code Ocean provides an interactive platform for compu-
516 tational reproducibility [33], we have also provided an interactive version of our code for reproducing results
517 and figures at [34].
518

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525 **8 Author Contributions Statement**

526 D.W. conceived and supervised the study. N.C.K. developed and implemented the methodology. X.H. and
527 D.W. verified the methods. N.C.K. performed visualization and analysis. N.C.K, X.H., and D.W. edited and
528 wrote the manuscript. All authors read and approved the final manuscript.

529 **9 Competing Interests Statement**

530 The authors declare no competing interests.

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