

RESEARCH ARTICLE

A fairness assessment of mobility-based COVID-19 case prediction models

Abdolmajid Erfani^{1*}, Vanessa Frias-Martinez^{2,3}

1 Department of Civil, Environmental, and Geospatial Engineering, Michigan Technological University, Houghton, MI, United States of America, **2** College of Information Studies, University of Maryland, College Park, MD, United States of America, **3** University of Maryland Institute for Advanced Computer Studies, University of Maryland, College Park, MD, United States of America

* aerfani@mtu.edu



OPEN ACCESS

Citation: Erfani A, Frias-Martinez V (2023) A fairness assessment of mobility-based COVID-19 case prediction models. PLoS ONE 18(10): e0292090. <https://doi.org/10.1371/journal.pone.0292090>

Editor: Emanuele Crisostomi, Università di Pisa, ITALY

Received: February 16, 2023

Accepted: September 12, 2023

Published: October 18, 2023

Peer Review History: PLOS recognizes the benefits of transparency in the peer review process; therefore, we enable the publication of all of the content of peer review and author responses alongside final, published articles. The editorial history of this article is available here: <https://doi.org/10.1371/journal.pone.0292090>

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Data Availability Statement: All data we use in this study are all publicly available. These datasets include the SafeGraph mobility data (<https://docs.safegraph.com/docs/social-distancing-metrics>), COVID-19 confirmed cases (<https://github.com/>

Abstract

In light of the outbreak of COVID-19, analyzing and measuring human mobility has become increasingly important. A wide range of studies have explored spatiotemporal trends over time, examined associations with other variables, evaluated non-pharmacologic interventions (NPIs), and predicted or simulated COVID-19 spread using mobility data. Despite the benefits of publicly available mobility data, a key question remains unanswered: are models using mobility data performing equitably across demographic groups? We hypothesize that bias in the mobility data used to train the predictive models might lead to unfairly less accurate predictions for certain demographic groups. To test our hypothesis, we applied two mobility-based COVID infection prediction models at the county level in the United States using SafeGraph data, and correlated model performance with sociodemographic traits. Findings revealed that there is a systematic bias in models' performance toward certain demographic characteristics. Specifically, the models tend to favor large, highly educated, wealthy, young, and urban counties. We hypothesize that the mobility data currently used by many predictive models tends to capture less information about older, poorer, less educated and people from rural regions, which in turn negatively impacts the accuracy of the COVID-19 prediction in these areas. Ultimately, this study points to the need of improved data collection and sampling approaches that allow for an accurate representation of the mobility patterns across demographic groups.

Introduction

The interactions between human mobility and epidemic spread have been studied unprecedentedly during the COVID-19 pandemic [1–8]. With these efforts, nonpharmaceutical interventions (such as national lockdowns) have been evaluated for their effectiveness and socio-economic impact on different groups [9–11], models have been developed to predict disease spatial diffusion [12, 13], and scenarios have been modeled to assess their outcomes [14–17]. Studies have demonstrated that mobility data are a meaningful proxy measure of social distancing [18], affect viral spreading [19, 20], and are useful for predicting the spread of COVID-19 [21–23].

CSSEGISandData/COVID-19), and sociodemographic information at US counties level (<https://www.census.gov/data/datasets.html>).

Funding: The authors acknowledge funding support from the National Science Foundation under Grant Numbers NSF 1750102 and NSF 2210572. The views, opinions, conclusions, or recommendations expressed in this research are those of the authors and do not necessarily reflect the view of the funding agencies.

Competing interests: The authors have declared that no competing interests exist.

In particular, to control the spread of new cases and plan efficiently for hospital needs and capacities during an epidemic, public health decision-makers require accurate predictions of future case numbers [7]. For example, a study by Ilin et al. (2021) showed that changes in mobility can be used to predict COVID-19 cases. Their study demonstrated that public mobility data can be used to develop reduced-form and simple models that mimic the behavior of more sophisticated epidemiological models for predicting COVID-19 cases on a 10-day basis [21]. Another study examined several state-of-the-art machine learning models and statistical methods and demonstrated how mobility data can improve prediction trends when used as exogenous information in models [22].

As discussed, mobility data from anonymized smartphones has been shown to improve COVID-19 case prediction models. However, mobility data bias has received little attention in this predictive context. There exist only just a handful of papers reporting demographic bias in mobility data due to differences in smartphone ownership and use [24–26]; and since data providers are not transparent about how mobility data is collected, or about the socio-economic and demographic groups represented in them, directly measuring and correcting bias in mobility data is difficult [27]. In this study, we hypothesize that the presence of socio-economic and demographic bias in the mobility data used to train the COVID-19 case predictive models, might result in unfairly less accurate predictions for particular socio-economic and demographic groups. Unfair predictions provided to decision makers e.g., predictions of COVID-19 cases for minority groups that are lower than reality, could in turn be used to unfairly assign more resources to population groups that do not necessarily need them.

To test our hypothesis, we evaluated the performance of two types of mobility-based COVID-19 case prediction models highly used by decision makers due to its interpretability: linear regressions and time series models. In contrast to more complex epidemiological models that are hard to tune due to its parametric nature, and deep learning models that are difficult to interpret, linear models and time series are easy to train and test [21, 28–30]. The models were trained using SafeGraph's mobility data, and performance was measured via predictive errors. To assess the fairness of the predictions, we analyzed the relationship between the model prediction errors and specific socio-economic and demographic features at the county level in the United States and across the two model types. Evaluating the performance of two diverse interpretable models allowed us to account for potential algorithmic bias i.e., bias introduced by the algorithm itself [31, 32]. If unfair predictions are pervasive across types of models trained and tested with the same data, we can partially attribute the unfairness to the mobility data itself.

Material and methods

In our study, we use mobility data from SafeGraph to build COVID-19 case prediction models; and we explore model performance across socio-economic and demographic features to potentially identify unfair results for specific groups i.e., differences in error distributions across social groups. We next describe these three types of datasets, with all being publicly available.

Human mobility

We used SafeGraph's publicly-available human mobility data at the county level in the US. SafeGraph uses location information extracted from smartphones to provide aggregate data characterizing mobility in terms of visit volumes to types of places and volumes of origin-destination (OD) flows [33]. For this study specifically, we used the data publicly available in the origin-destination-time (ODT) platform [34], that computes OD flows between counties as

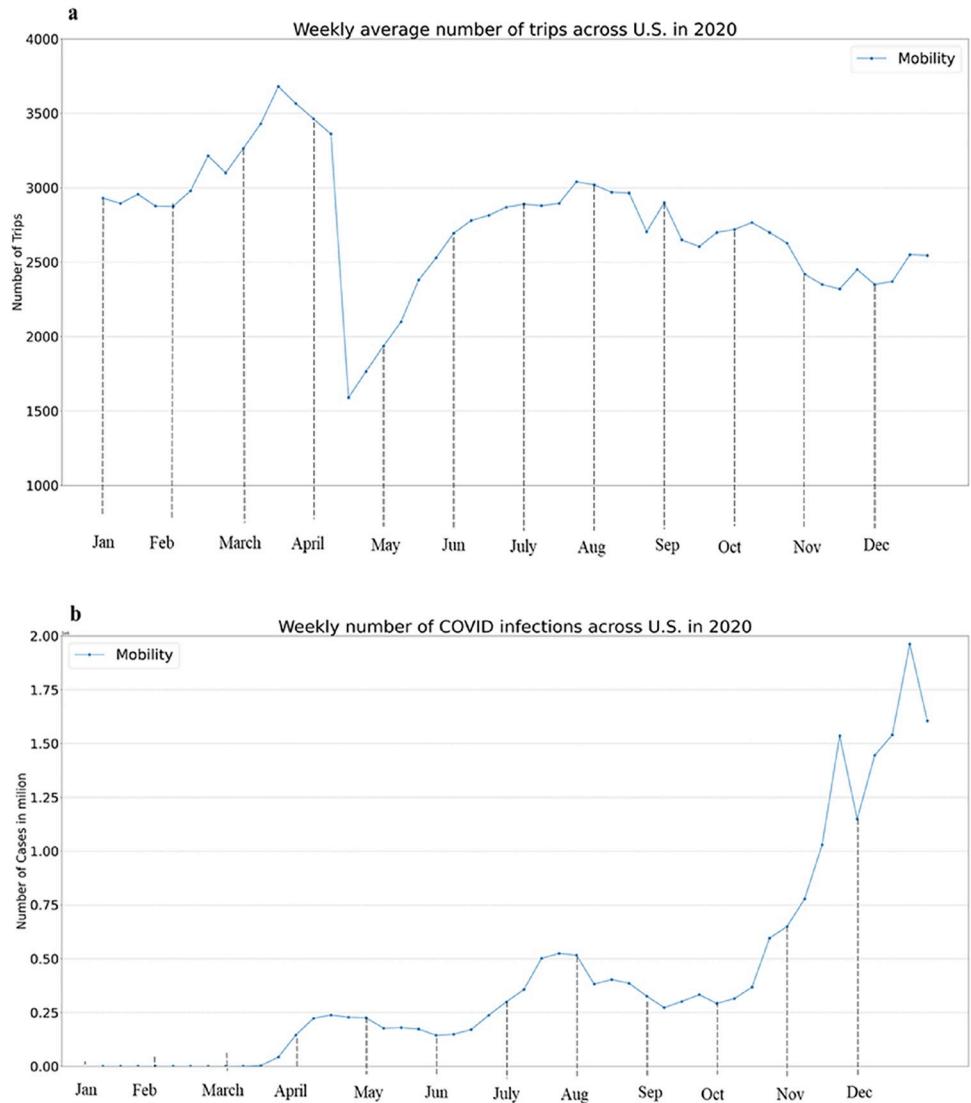


Fig 1. Data on mobility measures, COVID-19 infections. (a) Weekly average number of trips across the U.S. (b) Weekly new number of COVID infections across the U.S.

<https://doi.org/10.1371/journal.pone.0292090.g001>

the aggregation of trips that start at an individual's home county location (origin), with a destination defined as a stay location within a county for longer than a minute. OD flows between all counties in the US were collected throughout all days of the year 2020. Fig 1A illustrates how the average number of trips at county level across the US changed over the year 2020. According to various studies in the US using mobility data, the dataset collected in Fig 1A also shows similar trends of mobility change [35, 36].

COVID-19 cases

In order to obtain the cumulative and daily confirmed cases of COVID-19 for each county unit, we refer to the data repository compiled by the Johns Hopkins Center for Systems Science and Engineering (CSSE) [37]. As shown in Fig 1B, the weekly number of new COVID infections has been increasing over the year 2020.

Table 1. Summary statistics of input variables at the county level.

Variable	Description	Count	Mean	Min	Max
Population	Population	3,036	104,043.0	441	10,081,570
Income	Median household income (\$)	3,036	70,264.8	35,819	181,261
Education	Percentage of the population with a bachelor's degree and above (%)	3,036	22.0	3.2	77.6
Age	Median age	3,036	41.3	22.3	67.4
Smartphone	Percentage of the population who own smartphone (%)	3,036	72.7	25.0	92.4
NCHS	Urban-Rural Classification (1–6)	3,036	4.9	1.0	9.0

<https://doi.org/10.1371/journal.pone.0292090.t001>

Socio-economic and demographic data

Data on socio-economic and demographic characteristics at the county level was also collected from public databases (US American Community Survey census data, ACS 2020) [38]. Studies have shown that sociodemographic factors such as age, race, income, educational level, and area of residence can influence smartphone ownership and usage, which may have an impact on mobility data biases [39, 40]. Therefore, we collected a wide range of information, including the population, income, education, age, and ownership of smartphones at the county level. Also, we used the US National Center for Health Statistics (NCHS) Urban-Rural Classification Scheme for Counties [41], which assigns each county an ordinal code ranging from 1 (most urban) to 6 (most rural). Table 1 summarizes the demographic features of the study with some descriptive statistics.

Dataset preparation

We took the following steps to prepare the final dataset for modeling. Daily mobility data was collected from the origin-destination-time (ODT) platform [34] from April 14th to December 30th. The platform had daily mobility OD flows for 3,036 out of the 3,142 counties in the US. As a result, the total dataset size was of over 774,000 records. We used the Federal Information Processing Standard (FIPS) code, to match the daily number of COVID-19 cases per county with its corresponding mobility data. Therefore, the final dataset represented daily count of infections and mobility metrics per county in the US throughout the period of study. Following a similar procedure, we added socio-economic and demographic features at the county level to each data record using the FIPS code and the variables provided by the 5-year US ACS census from 2020 [38].

Methods

In this section we will describe (i) the two types of models used in the COVID-19 case prediction; (ii) the training and evaluation of these models; and (iii) the process proposed to evaluate the fairness of the predictions across socio-economic and demographic groups, as well as across models.

Models

Model 1: Linear regression (Ilin et al. [21], Wang et al. [42], Ayan et al. [43], and Sahin [44]). Several papers have suggested that linear regressions that combine mobility data with historical COVID-19 cases can successfully predict future cases [21, 42–44]. These models generally use different lags between mobility rates and COVID-19 cases to account for the infection period i.e., the period between the person's movement—and potentially interaction with others and infection—and the person testing positive for COVID-19. For this study we use Ilin

at al. linear model [21] because rather than picking one lag, they propose to consider multiple lags within the model encompassing the plethora of linear regressions that have been tested in the literature. Specifically, Ilin et al. (2021) use a distributed-lag model to estimate log confirmed infections as the dependent variable, with average mobility over lags 1–7, 8–14, and 15–21 days to predict the number of COVID-19 cases at a given day:

$$\log \frac{I_{it}}{I_{i,t-1}} = \beta_1 m_{1-7,it} + \beta_2 m_{8-14,it} + \beta_3 m_{15-21,it} + \epsilon_{it} \quad (1)$$

where i is the unit of analysis, $\log \frac{I_{it}}{I_{i,t-1}}$ is the first difference of log confirmed cases at time t , $m_{1-7,it}$, $m_{8-14,it}$, $m_{15-21,it}$ represents mobility measures averaged over lags 1–7, 8–14 and 15–21, respectively, and β_1 , β_2 , β_3 are model parameters to be estimated.

Model 2: Time series forecasting (Aji et al. [29], Zhao et al. [30], Zeng et al. [45], and Klein et al. [46]). The Autoregressive Integrated Moving Average (ARIMA) model is a statistical method that considers both past and present data for forecasting. An ARIMAX model, also known as ARIMA with multiple regressors, extends the basic ARIMA model to include other external variables for prediction. In the COVID-19 setting, mobility data and other sources of information have been used as regressors to potentially improve the predictive models [30, 45, 46]. For example, in their study Zhao et al. [30] conclude that with mobility data, time series forecasting provides accurate predictions with mobility data lags of between 8–10 days for dense or sparse populations respectively. In this study, we consider an ARIMAX (p, d, q) model that can be expressed as:

$$y_t = \beta_0 x_t + \sum_{j=1}^p \beta_j y_{t-j} + \epsilon_t + \sum_{j=1}^q \theta_j \epsilon_{t-j} \quad (2)$$

where y is the number of confirmed infections, x is the mobility change as exogenous variable lagged by 21 days (similar to Model 1's selection of lag), p is the Autoregressive (AR) parameter, q is the Moving Average (MA) parameter, d is the degree of first differencing to make data stationarity, ϵ is the error, and β_0 , β_j , θ_j are model parameters to be estimated. By using the Python package Auto Arima, we were able to generate the best p , d , and q values based on the data set, thus providing better forecasts [47]. To summarize, the lag of mobility, historic number of COVID cases can be used to predict future cases at unit of analysis.

Training and model evaluation

To train and evaluate the models, we used both historical COVID-19 data and mobility OD flows from mid-April to December 2020. Rather than using the raw mobility OD flows, we used a measure of mobility change over a baseline, which was calculated by dividing the daily mobility by the average daily mobility in February 2020, a non-holiday month before the COVID-19 pandemic. This is a common approach in prior COVID-19 case predictive models that use mobility data [21, 30].

The two models were trained at the county level on a daily basis using both COVID-19 case numbers and changes in mobility OD flows as independent variables to predict future cases. Socio-economic and demographic data were not used to train the models. That information was exclusively used during the fairness evaluation. For the linear regression model (Model 1), 21 days of mobility and past COVID-19 case data were used at a time for the training, and the trained model was used to test 1-day and 7-day predictions. We implemented a 1-day sliding window to replicate this train-test approach throughout the time period of analysis and reported average daily prediction error rates. Similarly, the time series model (Model 2), was

also trained using a typical training-testing window approach for time series predictions [48], with a 90-day training dataset. Using a 1-day sliding window on the training dataset, this approach resulted in predictions available from early August to the end of the year. Different training lengths were evaluated for both models, and the ones with the best accuracies were selected. In this process, thousands of regressions and ARIMAX models are trained at the county level on a daily basis to be able to predict COVID-19 cases. Once trained, each model was used to predict the number of COVID-19 cases for two lookaheads: 1-day (next day) and 7-days time (week) intervals at the county level, as predictions on a daily and weekly basis are a common theme in previous studies [21, 29, 30, 42–45].

Finally, the model performance was evaluated via the error rate, which was calculated on a daily basis based on the difference between the actual number of COVID cases and predictions as Eq 3. A mean absolute percentage error rate (MAPE) is calculated by averaging the error rates for specific counties over a given time period.

$$\text{Error rate}_t = \left| \frac{\text{Prediction value}_t - \text{Actual number of COVID cases}_t}{\text{Actual number of COVID cases}_t} \right| \quad (3)$$

$$\text{MAPE} = \frac{100\%}{n} \sum_{t=1}^n \text{Error rate}_t \quad (4)$$

Fairness analysis

We analyzed the fairness of the predictions for each model by computing the weekly MAPE per lookahead (1-day and 7-day) at the county level, followed by a spearman rank-order correlation analysis between the average weekly error rate across counties in the US and their socio-economic and demographic characteristics presented in the data section: household income (average household income), smartphone ownership (percentage of households owning smartphones), population, education level (bachelor's degree), urbanity-rurality level (NCHS classification), and age (median age). A spearman correlation provides an opportunity to investigate the monotonic relationship between two continuous variables of demographic features and model accuracy. A monotonic relationship occurs when the variables change together, but not necessarily at the same rate [49, 50]. Using the P-value to evaluate the correlation analysis significance, we can assess whether performance is similar (fair) or not (unfair) between social groups. To discuss correlation strength, and based on prior work, we will use 0.3 as the correlation coefficient threshold between a high and low correlation [51], or a weak and a moderate correlation [52].

Results

Model performance

First, we discuss the COVID-19 case prediction performance of the two models presented (see Fig 2).

As we can observe, both models predict the number of next day cases (1-day) with an average weekly error rate of 10–20%, and the number of cases in a week (7-days) with an average weekly error rate of 30–40%. Models' performance is in a comparable range to previous studies [21, 29, 30, 42–45], but with the difference that we reported the results for the entire year of 2020 and the US, not specific regions or COVID waves.

Fairness performance

For Model 1, we observe a negative and statistically significant spearman rank correlation between the prediction error rate and income (R (1-day) = -0.13, R (7-days) = -0.08, p -

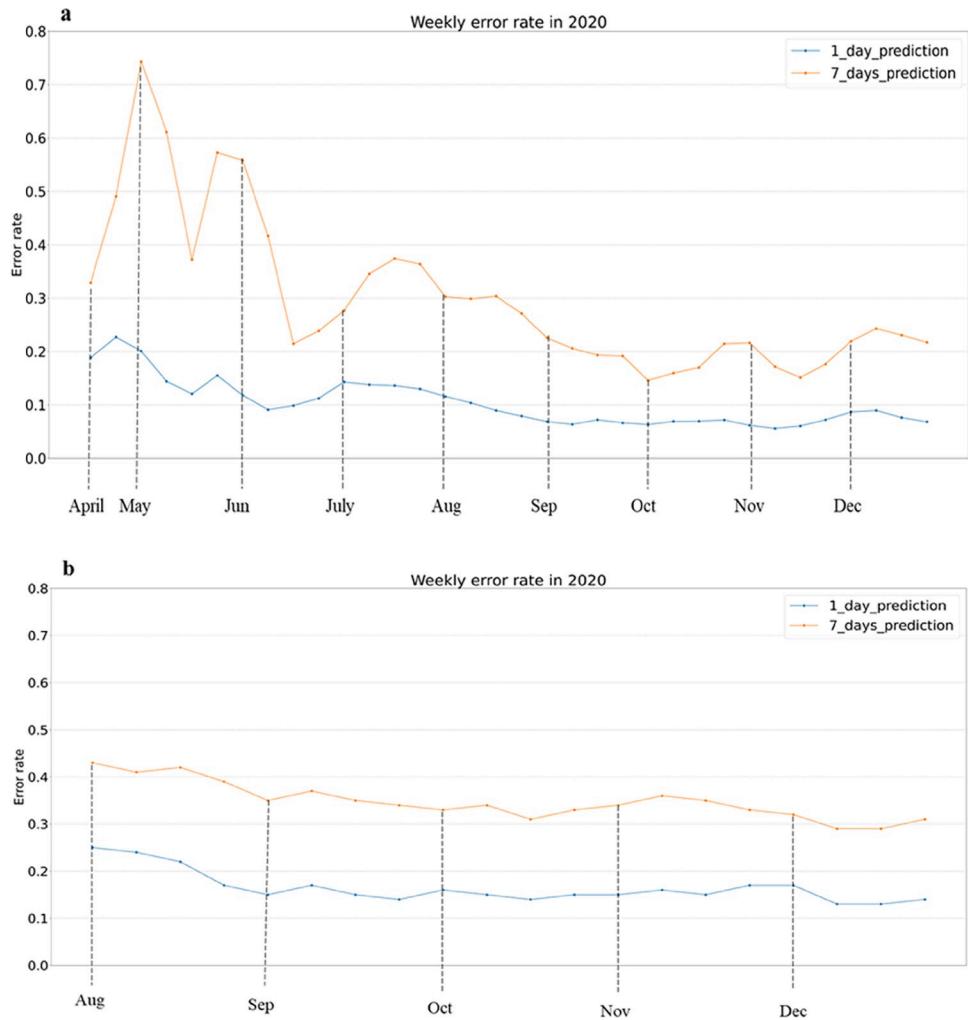


Fig 2. Prediction error rate on a weekly basis. (a) Regression model (b) Time series model.

<https://doi.org/10.1371/journal.pone.0292090.g002>

value < 0.001), smartphone ownership (R (1-day) = -0.14, R (7-days) = -0.09, p-value < 0.001), population (R (1-day) = -0.11, R (7-days) = -0.07, p-value < 0.001), bachelor degree (R (1-day) = -0.13, R (7-days) = -0.09, p-value < 0.001). The results suggest that Model 1 – a regression model of COVID-19 cases with mobility – performs better (has fewer errors) in counties with higher incomes, higher smartphone ownership, larger populations, and higher educational levels. On the other hand, correlation analysis indicates a weak and positive relationship between NCHS code and error rate (R (1-day) = 0.21, R (7-days) = 0.15, p-value < 0.001) and median age (R (1-day) = 0.12, R (7-days) = 0.09, p-value < 0.01). Therefore, as rurality, and age increased, the model's error rate increased, suggesting it performs worse in rural areas and among older communities (Fig 3 represents the weekly correlations for some of these features).

For Model 2, we observe a negative and statistically significant spearman rank correlation between the prediction error rate and the income (R (1-day) = -0.13, R (7-days) = -0.09, p-value < 0.001), smartphone ownership (R (1-day) = -0.13, R (7-days) = -0.10, p-value < 0.001), population (R (1-day) = -0.11, R (7-days) = -0.11, p-value < 0.001), bachelor degree (R (1-day) = -0.13, R (7-days) = -0.08, p-value < 0.001). These results show that Model

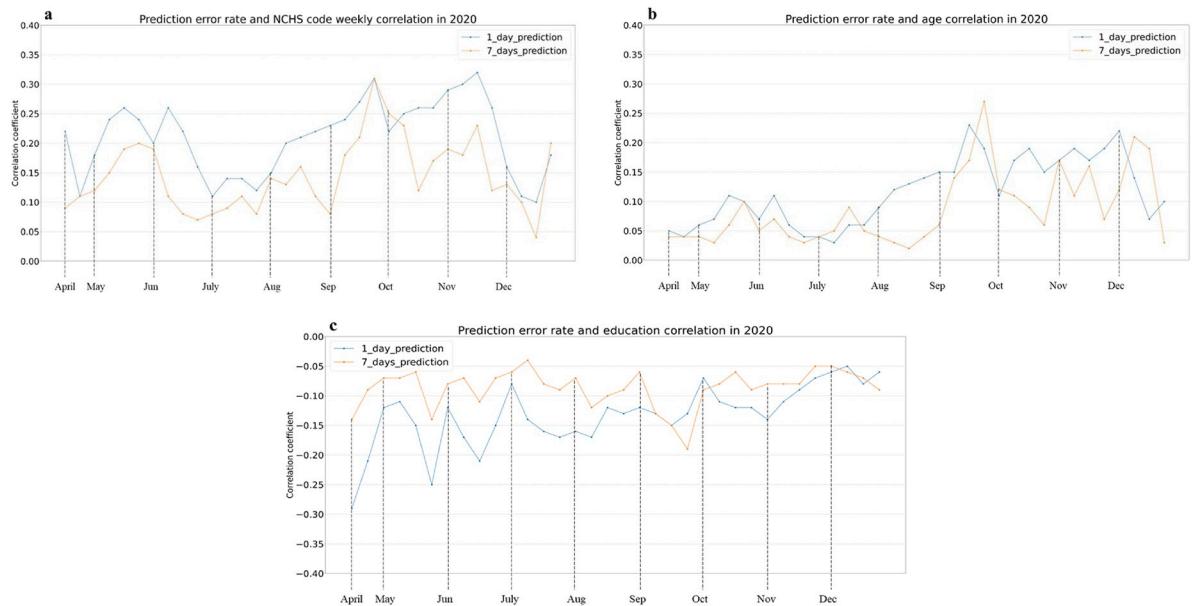


Fig 3. Correlation analysis with selected factors on a weekly basis for Model 1. (a) NCHS code (b) Age, and (c) education.

<https://doi.org/10.1371/journal.pone.0292090.g003>

2—an ARIMAX with mobility data as an exogenous variable—performs better (i.e., with lower errors) in counties whose income, smartphone ownership, population, and educational levels are higher. Fig 4 shows weekly correlations for some of these features. On the other hand, the correlation analysis also reveals a weak and positive relationship between the error rate and the NCHS code (R (1-day) = 0.20, R (7-days) = 0.21, p -value < 0.001) and median age (R (1-day) = 0.08, R (7-days) = 0.09, p -value < 0.01). In other words, the model's error rate increased as

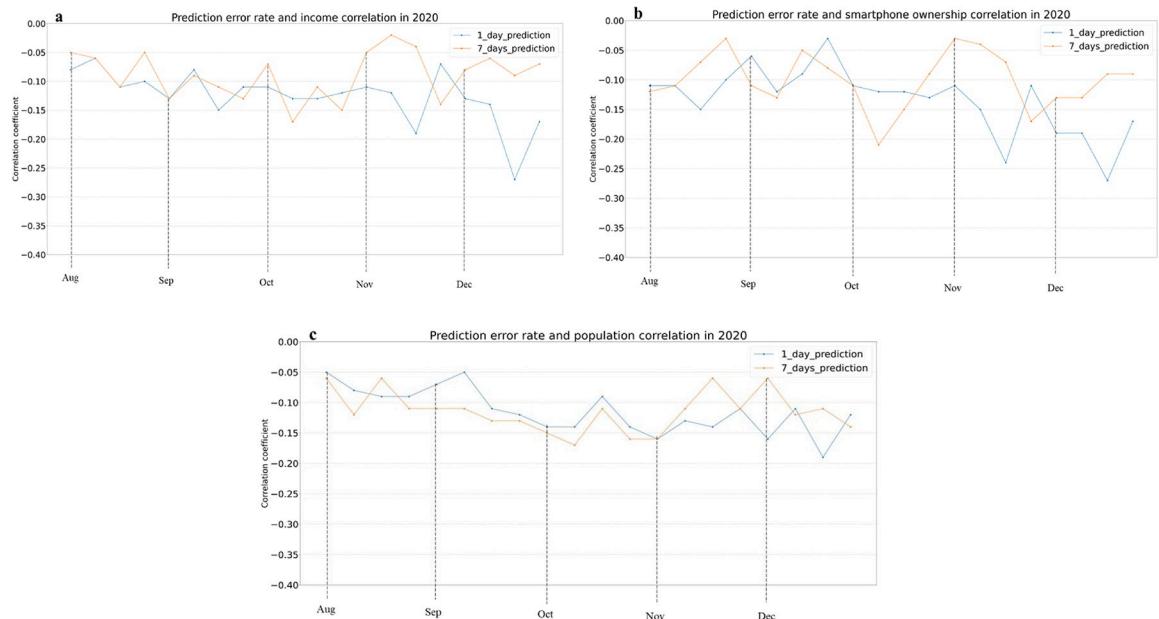


Fig 4. Correlation analysis with selected factors on a weekly basis for Model 2. (a) Income (b) Smartphone ownership, and (c) Population.

<https://doi.org/10.1371/journal.pone.0292090.g004>

Table 2. County-level correlations between Model 1 error rate and sociodemographic features. (Note: Statistical significance: *** p_value < 0.001, ** p_value < 0.01, * p_value < 0.05).

	Income		Smartphone		Population		Education		NCHS		Age	
	1 day	7 day	1 day	7 day	1 day	7 day	1 day	7 day	1 day	7 day	1 day	7 day
April	-0.20***	-0.09***	-0.19***	-0.09***	-0.14***	-0.07***	-0.25***	-0.12***	0.17***	0.11***	0.05*	0.04*
May	-0.14***	-0.06***	-0.16***	-0.07***	-0.14***	-0.09***	-0.16***	-0.09***	0.23***	0.17***	0.08***	0.03**
Jun	-0.14***	-0.09***	-0.14***	-0.07***	-0.09***	-0.05***	-0.16***	-0.08***	0.21***	0.11***	0.07***	0.05***
July	-0.14***	-0.07***	-0.08***	-0.06***	-0.07***	-0.05***	-0.14***	-0.07***	0.13***	0.09***	0.05*	0.06***
August	-0.15***	-0.07***	-0.15***	-0.08***	-0.11***	-0.05***	-0.15***	-0.10***	0.20***	0.14***	0.12***	0.03*
September	-0.12***	-0.12***	-0.19***	-0.17***	-0.13***	-0.10***	-0.13***	-0.13***	0.26***	0.20***	0.18***	0.16***
October	-0.10***	-0.11***	-0.16***	-0.14***	-0.14***	-0.11***	-0.11***	-0.10***	0.26***	0.22***	0.16***	0.13***
November	-0.11***	-0.07***	-0.17***	-0.09***	-0.11***	-0.07***	-0.10***	-0.07***	0.29***	0.18***	0.18***	0.13***
December	-0.09***	-0.06***	-0.09***	-0.09***	-0.07***	-0.06***	-0.06***	-0.07***	0.14***	0.12***	0.13***	0.14***

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rurality and age increased, revealing a model that performs worse in rural environments, and among older populations. Due to the replication of these findings in models 1 and 2, which controls for algorithmic bias, we posit that this model is unfair in part because it uses biased mobility data, although bias in the way COVID-19 case data is gathered (e.g., under-reporting) could also influence its outcome.

To summarize the fairness analysis across models, Tables 2 and 3 provide the monthly correlation averages between the sociodemographic factors at the county level and the error rates for Models 1 and 2 1-day and 7-day predictions, respectively. As discussed, due to the diverse size of the optimal training windows, Model 1 predictions run from April till December, while Model 2 predictions are produced from August till December. With a few fluctuations, and as discussed in the weekly analyses in Figs 3 and 4, both models show the same pattern of results throughout 2020: lower prediction errors in large, highly educated, wealthy, young, and urban counties. Given that the strength of the correlations found is weak, we posit that all socio-economic and demographic features are related to significant, albeit weak, bias. We do not observe statistically significant differences in the strengths across features. We do however see that correlation coefficients are smaller for 7-day predictions than 1-day predictions, which might point to more negligible bias for these models. Nevertheless, the statistical results do not support any feature being more biased than other. In addition, it is important to highlight that for certain socio-economic and demographic features in Tables 2 and 3, we observe some variance in the correlation coefficients across months, with earlier pandemic months showing higher coefficients. We posit that these might be due to mobility behaviors being more entropic later during the pandemic, which might make it harder to find associations. As prior work has shown, uniform mobility behaviors made mobility data more useful in predictive models at the onset of the pandemic than in later periods [20, 53].

Table 3. County-level correlations between Model 2 error rate and sociodemographic features. (Note: Statistical significance: *** p_value < 0.001, ** p_value < 0.01, * p_value < 0.05).

	Income		Smartphone		Population		Education		NCHS		Age	
	1 day	7 day	1 day	7 day	1 day	7 day	1 day	7 day	1 day	7 day	1 day	7 day
August	-0.09*	-0.07*	-0.12**	-0.08*	-0.08*	-0.09**	-0.10**	-0.06	0.14**	0.15***	0.08*	0.08**
September	-0.12***	-0.12**	-0.08*	-0.09*	-0.09**	-0.12**	-0.14**	-0.09**	0.15**	0.16***	0.05*	0.05*
October	-0.12**	-0.13***	-0.12***	-0.14***	-0.13**	-0.15***	-0.11***	-0.09**	0.19***	0.23***	0.08*	0.08**
November	-0.12***	-0.06*	-0.15***	-0.08*	-0.14***	-0.11***	-0.14***	-0.07*	0.26***	0.31***	0.11*	0.16**
December	-0.18***	-0.08**	-0.21***	-0.11**	-0.15***	-0.11***	-0.17***	-0.08***	0.27***	0.22***	0.09**	0.09**

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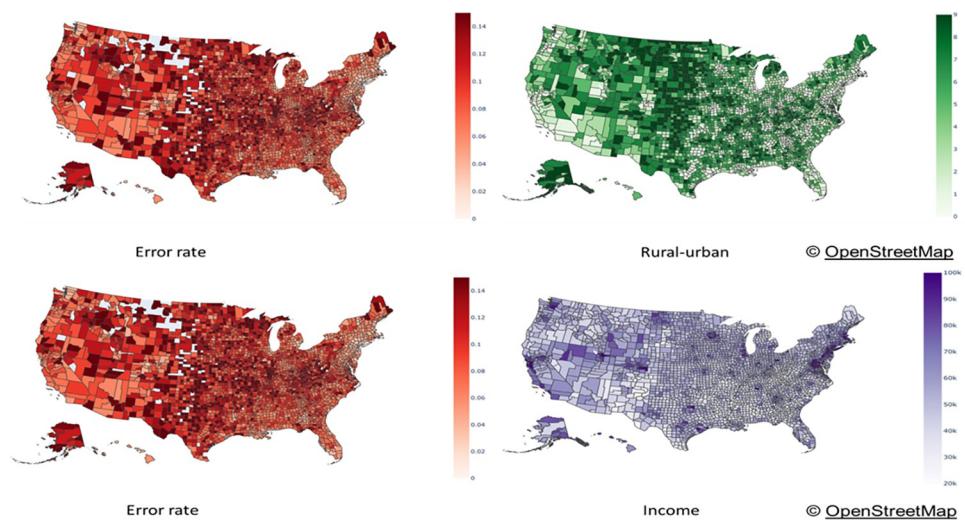


Fig 5. Spatial comparison of error rate and demographic features for Model 1. (These plots have been generated with [Plotly open source graphing libraries](#) using base maps from [OpenStreetMap](#). OpenStreetMap is open data, licensed under the Open Data Commons [Open Database License \(ODbL\)](#) by the [OpenStreetMap Foundation \(OSMF\)](#)).

<https://doi.org/10.1371/journal.pone.0292090.g005>

Fig 5 shows two comparative visualizations between the average 1-day prediction error rate for the regression model (Model 1) and two demographic features namely urbanity level and household income. Visualizations for the other demographic features are shown in the supplemental materials, S1 Fig in [S1 File](#). The visualizations show trends in line with the quantitative results discussed before i.e., that areas with a higher rurality (dark green) and areas with a lower income rate (white) have a higher error rate (dark red). We observe two interesting patterns. First, the error rates are much higher across the eastern states of the Great Plains (vertical line from North Dakota to Texas) which represent some of the highest categories of rurality and some of the lowest income rates (outside of metropolitan areas in the region). Second, the error rates are higher in the Appalachian region (from southern New York to northern Mississippi) which is associated to some of the lowest income ratios in the country. Similar visualizations were observed for 1-day predictions for Model 2. Visualizations for 7-day predictions did not show clear spatial trends, possibly due to the weaker correlations reported. Map figures are generated using the Plotly package, an open source Python package.

Discussion

To combat the COVID-19 pandemic, governments and private companies around the world were promoting the use of digital public health technologies for data collection and processing [54–58]. Through the use of GPS, cellular networks and Wi-Fi, smartphones can collect and aggregate location data in real-time to monitor population flows, identify transmission hot-spots, determine the effectiveness of non-pharmacologic interventions [59], and predict future COVID-19 cases [7, 20, 25].

Using SafeGraph's mobility data, we examined whether two popular predictive models that use mobility data to predict COVID-19 cases over time, performed fairly across social groups. Our findings revealed a correlation between a county's socio-economic and demographic characteristics and the models' error rates. In particular, we observed that the prediction errors were lower in large, highly educated, wealthy, young, and urban counties. Given that the findings were similar across models, thus controlling for algorithmic bias, we posit that the

presence of bias in the mobility data negatively impacts the model predictions by unfairly outputting case numbers with higher errors for specific social groups. Furthermore, our results show that mobility data appears to be less likely to capture older, poorer, and less educated users. Thus, allocating public health resources based on such mobility data could disproportionately harm seniors and minorities at high risk.

For both predictive models, we observed higher biases in 1-day predictions compared to 7-day predictions. We posit this is possibly due to the already reported difficulty in predicting COVID-19 cases for higher lookaheads [60], thus resulting in noisier predictions which might in turn generate lower correlations between model performance and sociodemographic characteristics. When comparing the regression and time series models in terms of their biases, we did not observe any statistically significant difference between the reported correlation coefficients. A Wilcoxon rank sum test, also known as a Mann-Whitney U test [61], was implemented to compare two independent groups of coefficients on a 1-day prediction period (Tables 2 and 3). We chose this test because it is a non-parametric statistical test that is not based on assumptions of normality or equal variance. With a P-value of 0.289, the null hypothesis was accepted pointing to no significant difference between bias coefficients between the two models.

To generalize smartphone-derived insights over a population, the mobility data must reveal information about the population without bias i.e., information that is representative across socio-economic and demographic groups. However, due to the lack of ground truth data about the socio-economic and demographic characteristics of the population whose mobility data is collected, this study has also shown that investigating performance fairness can provide valuable insights into potential mobility data biases.

Finally, as the research community moves forward with the use of mobility data in COVID-19 case prediction models, we think it is important to consider the following set of recommendations. First, and whenever possible, we strongly suggest applying sampling bias mitigation approaches to correct for under-represented groups in the data, as prior work has successfully done [48, 62, 63]. Second, mitigation approaches might not always be possible, due to the lack of demographic information about the individuals whose cell phone data is being collected. For that reason, we encourage the research community working with mobility data to report fairness analyses together with the performance of the predictive models proposed. We hope that these practices will enhance pandemic management via case prediction models that are more transparent and fairer, and that will allow for more equitable decision making.

Limitations

While this study addressed potential biases in mobility data currently used by two types of predictive models, there are a number of limitations related to modeling and dataset biases that require clarification.

We chose a baseline period of one pre-covid month in 2020 to model ‘normal behavior’. This choice was determined by the limited availability of “free” mobility data. Although, ideally, mobility baselines should be from a pre-covid period e.g., 2019, we were limited by the availability of free SafeGraph data, which started in 2020. Testing different baselines is an important research question, but that would require having access to additional mobility data that was not available.

Other limitations include the decision to add mobility data in the predictive models using mobility changes from up to 21 days prior to the prediction date. Although earlier mobility periods could be considered, the probability that mobility patterns prior to 21 days might translate into a COVID-19 infection is extremely low given that the incubation periods known

for COVID-19 and its variants can be up to 14 days [64, 65]. Prediction results are only reported for 1 and 7 days ahead despite the fact that different testing lookaheads might provide diverse outcomes.

We have reported fairness analysis results in terms of correlation coefficients between performance and socio-economic variables. Nevertheless, statistically significant correlations reflect the probability of such a correlation occurring rather than its strength. Correlation coefficient strengths can be interpreted differently across scientific fields, and authors should avoid overinterpreting associations [51, 66, 67]. Based on prior work utilizing spearman rank correlation in the context of medicine and big data analysis, we have selected a correlation coefficient of 0.3 as the threshold between high and low correlation [51], or weak and moderate correlation [52]. As a final point, in addition to mobility data, there are other sources of data that might include biases and have an effect on prediction performance, such as diverse COVID-19 case reporting methodologies and US census tract data, both of which are used in this paper. It is important to note that these are potential biases in our study, and future work should look into their effect on the fairness analysis presented in this paper.

Finally, we made several conscious modeling choices. First, we focused on linear regressions and time series models due to its interpretability. In contrast with more complex epidemiological models that are hard to tune due to its parametric nature, and deep learning models that are difficult to interpret, linear models and time series are easy to train and test [21, 28–30]. Second, we have avoided incorporating socio-economic and demographic variables as input to the linear regression and time series prediction models. This choice was based on prior work showing that the addition of demographic features as input to predictive models is not only controversial but also potentially harmful [68]. In fact, it has been argued that using socio-economic or demographic data as predictor may instead reinforce bias and generate predictions based primarily on demographic variables rather than on more actionable parameters, thus perpetuating inequalities [69]. Future studies could consider incorporating demographic features as inputs to the predictive models to replicate the fairness analysis presented in this paper. Third, we trained individual prediction models per county. Future work should explore a unified model that learns COVID-19 trends for all counties. This approach would allow for inclusion of county-level variables indicative of population vulnerability directly in the model, potentially yielding more accurate results.

Supporting information

S1 File.

(DOCX)

Author Contributions

Conceptualization: Abdolmajid Erfani, Vanessa Frias-Martinez.

Data curation: Abdolmajid Erfani.

Formal analysis: Abdolmajid Erfani.

Methodology: Abdolmajid Erfani, Vanessa Frias-Martinez.

Project administration: Vanessa Frias-Martinez.

Resources: Vanessa Frias-Martinez.

Supervision: Vanessa Frias-Martinez.

Validation: Vanessa Frias-Martinez.

Visualization: Abdolmajid Erfani.

Writing – original draft: Abdolmajid Erfani.

Writing – review & editing: Abdolmajid Erfani, Vanessa Frias-Martinez.

References

1. Alessandretti L. What human mobility data tell us about COVID-19 spread. *Nature Reviews Physics*. 2022 Jan; 4(1):12–3. <https://doi.org/10.1038/s42254-021-00407-1> PMID: 34877474
2. Rutten P, Lees MH, Klous S, Heesterbeek H, Sloot PM. Modelling the dynamic relationship between spread of infection and observed crowd movement patterns at large scale events. *Scientific Reports*. 2022 Sep 1; 12(1):14825. <https://doi.org/10.1038/s41598-022-19081-z> PMID: 36050348
3. Hu S, Xiong C, Yang M, Younes H, Luo W, Zhang L. A big-data driven approach to analyzing and modeling human mobility trend under non-pharmaceutical interventions during COVID-19 pandemic. *Transportation Research Part C: Emerging Technologies*. 2021 Mar 1; 124:102955. <https://doi.org/10.1016/j.trc.2020.102955> PMID: 33456212
4. Hu T, Wang S, She B, Zhang M, Huang X, Cui Y, et al. Human mobility data in the COVID-19 pandemic: characteristics, applications, and challenges. *International Journal of Digital Earth*. 2021 Sep 2; 14(9):1126–47. <https://doi.org/10.1080/17538947.2021.1952324>
5. Nouvellet P, Bhatia S, Cori A, Ainslie KE, Baguelin M, Bhatt S, et al. Reduction in mobility and COVID-19 transmission. *Nature communications*. 2021 Feb 17; 12(1):1090. <https://doi.org/10.1038/s41467-021-21358-2> PMID: 33597546
6. Kartal MT, Depren Ö, Depren SK. The relationship between mobility and COVID-19 pandemic: Daily evidence from an emerging country by causality analysis. *Transportation Research Interdisciplinary Perspectives*. 2021 Jun 1; 10:100366. <https://doi.org/10.1016/j.trip.2021.100366> PMID: 36844006
7. Wellenius GA, Vispute S, Espinosa V, Fabrikant A, Tsai TC, Hennessy J, et al. Impacts of social distancing policies on mobility and COVID-19 case growth in the US. *Nature communications*. 2021 May 25; 12(1):3118. <https://doi.org/10.1038/s41467-021-23404-5> PMID: 34035295
8. Gutiérrez-Jara JP, Vogt-Geisse K, Cabrera M, Córdoba-Lepe F, Muñoz-Quezada MT. Effects of human mobility and behavior on disease transmission in a COVID-19 mathematical model. *Scientific Reports*. 2022 Jun 27; 12(1):10840. <https://doi.org/10.1038/s41598-022-14155-4> PMID: 35760930
9. Coleman N, Gao X, DeLeon J, Mostafavi A. Human activity and mobility data reveal disparities in exposure risk reduction indicators among socially vulnerable populations during COVID-19 for five US metropolitan cities. *Scientific Reports*. 2022 Sep 22; 12(1):15814. <https://doi.org/10.1038/s41598-022-18857-7>
10. Gozzi N, Tizzoni M, Chinazzi M, Ferres L, Vespignani A, Perra N. Estimating the effect of social inequalities on the mitigation of COVID-19 across communities in Santiago de Chile. *Nature communications*. 2021 Apr 23; 12(1):2429. <https://doi.org/10.1038/s41467-021-22601-6> PMID: 33893279
11. Chang S, Pierson E, Koh PW, Gerardin J, Redbird B, Grusky D, et al. Mobility network models of COVID-19 explain inequities and inform reopening. *Nature*. 2021 Jan; 589(7840):82–7. <https://doi.org/10.1038/s41586-020-2923-3> PMID: 33171481
12. Canino MP, Cesario E, Vinci A, Zarin S. Epidemic forecasting based on mobility patterns: an approach and experimental evaluation on COVID-19 Data. *Social Network Analysis and Mining*. 2022 Dec; 12(1):116. <https://doi.org/10.1007/s13278-022-00932-6> PMID: 35996384
13. Chinazzi M, Davis JT, Ajelli M, Gioannini C, Litvinova M, Merler S, et al. The effect of travel restrictions on the spread of the 2019 novel coronavirus (COVID-19) outbreak. *Science*. 2020 Apr 24; 368(6489):395–400. <https://doi.org/10.1126/science.aba9757> PMID: 32144116
14. Hu S, Luo W, Darzi A, Pan Y, Zhao G, Liu Y, et al. Do racial and ethnic disparities in following stay-at-home orders influence COVID-19 health outcomes? A mediation analysis approach. *PloS one*. 2021 Nov 11; 16(11): e0259803. <https://doi.org/10.1371/journal.pone.0259803> PMID: 34762685
15. Mahmoudi J, Xiong C. How social distancing, mobility, and preventive policies affect COVID-19 outcomes: Big data-driven evidence from the District of Columbia–Maryland–Virginia (DMV) megaregion. *PloS one*. 2022 Feb 17; 17(2): e0263820. <https://doi.org/10.1371/journal.pone.0263820> PMID: 35176031
16. Aleta A, Martin-Corral D, Pastore y Piontti A, Ajelli M, Litvinova M, Chinazzi M, et al. Modelling the impact of testing, contact tracing and household quarantine on second waves of COVID-19. *Nature Human Behaviour*. 2020 Sep; 4(9):964–71. <https://doi.org/10.1038/s41562-020-0931-9> PMID: 32759985

17. Beigi P, Haque M, Rajabi MS, Hamdar S. Bike Share's Impact on COVID-19 Transmission and Bike Share's Responses to COVID-19: A case study of Washington DC. arXiv preprint arXiv:2205.05011. 2022 May 10. <https://doi.org/10.138/s41598-020-77751-2>
18. Pan Y, Darzi A, Kabiri A, Zhao G, Luo W, Xiong C, et al. Quantifying human mobility behaviour changes during the COVID-19 outbreak in the United States. *Scientific Reports*. 2020 Nov 26; 10(1):20742. <https://doi.org/10.1038/s41598-020-77751-2> PMID: 33244071
19. Levin R, Chao DL, Wenger EA, Proctor JL. Insights into population behavior during the COVID-19 pandemic from cell phone mobility data and manifold learning. *Nature Computational Science*. 2021 Sep; 1(9):588–97. <https://doi.org/10.1038/s43588-021-00125-9>
20. Badr HS, Gardner LM. Limitations of using mobile phone data to model COVID-19 transmission in the USA. *The Lancet Infectious Diseases*. 2021 May 1; 21(5): e113. [https://doi.org/10.1016/S1473-3099\(20\)30861-6](https://doi.org/10.1016/S1473-3099(20)30861-6) PMID: 33152270
21. Ilin C, Annan-Phan S, Tai XH, Mehra S, Hsiang S, Blumenstock JE. Public mobility data enables COVID-19 forecasting and management at local and global scales. *Scientific reports*. 2021 Jun 29; 11(1):1–1. <https://doi.org/10.1038/s41598-021-92892-8>
22. García-Cremades S, Morales-García J, Hernández-Sanjaime R, Martínez-España R, Bueno-Crespo A, Hernández-Orallo E, et al. Improving prediction of COVID-19 evolution by fusing epidemiological and mobility data. *Scientific Reports*. 2021 Jul 26; 11(1):1–6. <https://doi.org/10.1038/s41598-021-94696-2>
23. Alali Y, Harrou F, Sun Y. A proficient approach to forecast COVID-19 spread via optimized dynamic machine learning models. *Scientific Reports*. 2022 Feb 14; 12(1):1–20. <https://doi.org/10.1038/s41598-022-06218-3>
24. Coston A, Guha N, Ouyang D, Lu L, Chouldechova A, Ho DE. Leveraging administrative data for bias audits: assessing disparate coverage with mobility data for COVID-19 policy. In *Proceedings of the 2021 ACM Conference on Fairness, Accountability, and Transparency* 2021 Mar 3 (pp. 173–184). <https://doi.org/10.1145/3442188.3445881>
25. Milusheva S, Bjorkgren D, Viotti L. Assessing Bias in Smartphone Mobility Estimates in Low Income Countries. In *ACM SIGCAS Conference on Computing and Sustainable Societies* 2021 Jun 28 (pp. 364–378). <https://doi.org/10.1145/3460112.3471968>
26. Schlosser F, Sekara V, Brockmann D, Garcia-Herranz M. Biases in human mobility data impact epidemic modeling. arXiv preprint arXiv:2112.12521. 2021 Dec 23.
27. Grantz K.H., Meredith H.R., Cummings D.A., Metcalf C.J.E., Grenfell B.T., Giles J.R., et al., The use of mobile phone data to inform analysis of COVID-19 pandemic epidemiology. *Nature communications*, 2020 Sep 30; 11(1), p.4961. <https://doi.org/10.1038/s41467-020-18190-5> PMID: 32999287
28. Khan FM, Gupta R. ARIMA and NAR based prediction model for time series analysis of COVID-19 cases in India. *Journal of Safety Science and Resilience*. 2020 Sep 1; 1(1):12–8. <https://doi.org/10.1016/j.jnlssr.2020.06.007>
29. Aji BS, Rohmawati AA. Forecasting number of COVID-19 cases in Indonesia with ARIMA and ARIMAX models. In *2021 9th International Conference on Information and Communication Technology (ICoICT)* 2021 Aug 3 (pp. 71–75). IEEE. <https://doi.org/10.1109/ICoICT52021.2021.9527453>
30. Zhao J, Han M, Wang Z, Wan B. Autoregressive count data modeling on mobility patterns to predict cases of COVID-19 infection. *Stochastic environmental research and risk assessment*. 2022 Dec; 36(12):4185–200. <https://doi.org/10.1007/s00477-022-02255-6> PMID: 35765667
31. Kordzadeh N, Ghasemaghaei M. Algorithmic bias: review, synthesis, and future research directions. *European Journal of Information Systems*. 2022 May 4; 31(3):388–409. <https://doi.org/10.1080/0960085X.2021.1927212>
32. Johnson I, McMahon C, Schöning J, Hecht B. The effect of population and "structural" biases on social media-based algorithms: A case study in geolocation inference across the urban-rural spectrum. In *Proceedings of the 2017 CHI conference on Human Factors in Computing Systems* 2017 May 2 (pp. 1167–1178). <https://doi.org/10.1145/3025453.3026015>
33. SafeGraph. Social Distancing Metrics. <https://docs.safegraph.com/docs/social-distancing-metrics> (2020).
34. Li Z, Huang X, Hu T, Ning H, Ye X, Huang B, et al. ODT FLOW: Extracting, analyzing, and sharing multi-source multi-scale human mobility. *Plos one*. 2021 Aug 5; 16(8): e0255259. <https://doi.org/10.1371/journal.pone.0255259> PMID: 34351973
35. Xiong C., Hu S., Yang M., Luo W., & Zhang L. Mobile device data reveal the dynamics in a positive relationship between human mobility and COVID-19 infections. *Proceedings of the National Academy of Sciences*, 117(44), 27087–27089 (2020).

36. Lee M. et al. Human mobility trends during the early stage of the COVID-19 pandemic in the United States. *PLoS One*, 15(11), e0241468 (2020). <https://doi.org/10.1371/journal.pone.0241468> PMID: 33166301
37. COVID-19 Data Repository by the Center for Systems Science and Engineering (CSSE). Johns Hopkins University (2020). <https://github.com/CSSEGISandData/COVID-19>
38. U.S. Census Bureau. Annual Estimates of the Resident Population for Selected Age Groups by Sex for the United States, States, Counties and Puerto Rico Commonwealth and Municipios. (2019) <https://www.census.gov/data/datasets.html>.
39. Kim Y, Briley DA, Ocepek MG. Differential innovation of smartphone and application use by sociodemographics and personality. *Computers in Human Behavior*. 2015 Mar 1; 44:141–7. <https://doi.org/10.1016/j.chb.2014.11.059>
40. Rahmati A, Tossell C, Shepard C, Kortum P, Zhong L. Exploring iPhone usage: the influence of socio-economic differences on smartphone adoption, usage and usability. In Proceedings of the 14th international conference on Human-computer interaction with mobile devices and services 2012 Sep 21 (pp. 11–20). <https://doi.org/10.1145/2371574.2371577>
41. Ingram D. D., & Franco S. J. 2013 NCHS urban-rural classification scheme for counties. US Department of Health and Human Services, Centers for Disease Control and Prevention, National Center for Health Statistics (2014).
42. Wang R, Ji C, Jiang Z, Wu Y, Yin L, Li Y. A short-term prediction model at the early stage of the COVID-19 pandemic based on multisource urban data. *IEEE Transactions on Computational Social Systems*. 2021 Mar 5; 8(4):938–45. <https://doi.org/10.1109/TCSS.2021.3060952> PMID: 35582632
43. Ayan N, Chaskar S, Seetharam A, Ramesh A, Antonio AD. Mobility-aware COVID-19 Case Prediction using Cellular Network Logs. In 2021 IEEE 46th Conference on Local Computer Networks (LCN) 2021 Oct 4 (pp. 479–486). IEEE. <https://doi.org/10.1109/LCN52139.2021.9525023>
44. Şahin M. Forecasting COVID-19 cases based on mobility. *MANAS Journal of Engineering*. 2020; 8 (2):144–50.
45. Zeng C, Zhang J, Li Z, Sun X, Olatosi B, Weissman S, et al. Spatial-temporal relationship between population mobility and COVID-19 outbreaks in South Carolina: time series forecasting analysis. *Journal of medical Internet research*. 2021 Apr 13; 23(4):e27045. <https://doi.org/10.2196/27045> PMID: 33784239
46. Klein B. et al. Forecasting hospital-level COVID-19 admissions using real-time mobility data. *medRxiv* (2022).
47. Januschowski T., Gasthaus J., & Wang Y. Open-Source Forecasting Tools in Python. *Foresight: The International Journal of Applied Forecasting*, 2019, 55.
48. Ulyah S. M., & Mardianto M. F. F. (2019, December). Comparing the performance of seasonal arimax model and nonparametric regression model in predicting claim reserve of education insurance. In *Journal of Physics: Conference Series* (Vol. 1397, No. 1, p. 012074). IOP Publishing.
49. Puth MT, Neuhäuser M, Ruxton GD. Effective use of Spearman's and Kendall's correlation coefficients for association between two measured traits. *Animal Behaviour*. 2015 Apr 1; 102:77–84. <https://doi.org/10.1016/j.anbehav.2015.01.010>
50. Arik SÖ, Shor J, Sinha R, Yoon J, Ledsam JR, Le LT, et al. A prospective evaluation of AI-augmented epidemiology to forecast COVID-19 in the USA and Japan. *NPJ digital medicine*. 2021 Oct 8; 4(1):146. <https://doi.org/10.1038/s41746-021-00511-7> PMID: 34625656
51. Akoglu H. User's guide to correlation coefficients. *Turkish journal of emergency medicine*. 2018 Sep 1; 18(3):91–3. <https://doi.org/10.1016/j.tjem.2018.08.001> PMID: 30191186
52. Xiao C., Jiaqi Y., Rui M., and Chunming R., Using Spearman's correlation coefficients for exploratory data analysis on big dataset. *Concurrency and Computation: Practice and Experience* 28, no. 14, 2016: 3866–3878.
53. Gatalo O, Tseng K, Hamilton A, Lin G, Klein E. Associations between phone mobility data and COVID-19 cases. *The Lancet Infectious Diseases*. 2021 May 1; 21(5):e111. [https://doi.org/10.1016/S1473-3099\(20\)30725-8](https://doi.org/10.1016/S1473-3099(20)30725-8) PMID: 32946835
54. Gasser U, Ienca M, Scheibner J, Sleigh J, Vayena E. Digital tools against COVID-19: taxonomy, ethical challenges, and navigation aid. *The lancet digital health*. 2020 Aug 1; 2(8): e425–34. [https://doi.org/10.1016/S2589-7500\(20\)30137-0](https://doi.org/10.1016/S2589-7500(20)30137-0) PMID: 32835200
55. Li L, Erfani A, Wang Y, Cui Q. Anatomy into the battle of supporting or opposing reopening amid the COVID-19 pandemic on Twitter: A temporal and spatial analysis. *Plos one*. 2021 Jul 13; 16(7): e0254359. <https://doi.org/10.1371/journal.pone.0254359> PMID: 34255783
56. Nia ZM, Ahmadi A, Bragazzi NL, Woldegerima WA, Mellado B, Wu J, et al. A cross-country analysis of macroeconomic responses to COVID-19 pandemic using Twitter sentiments. *Plos one*. 2022 Aug 24; 17(8): e0272208. <https://doi.org/10.1371/journal.pone.0272208> PMID: 36001531

57. Whitelaw S, Mamas MA, Topol E, Van Spall HG. Applications of digital technology in COVID-19 pandemic planning and response. *The Lancet Digital Health*. 2020 Aug 1; 2(8): e435–40. [https://doi.org/10.1016/S2589-7500\(20\)30142-4](https://doi.org/10.1016/S2589-7500(20)30142-4) PMID: 32835201
58. Hickey PJ, Erfani A, Cui Q. Use of LinkedIn Data and Machine Learning to Analyze Gender Differences in Construction Career Paths. *Journal of Management in Engineering*. 2022 Nov 1; 38(6):04022060.
59. Budd J, Miller BS, Manning EM, Lampos V, Zhuang M, Edelstein M, et al. Digital technologies in the public-health response to COVID-19. *Nature medicine*. 2020 Aug 26; (8):1183–92. <https://doi.org/10.1038/s41591-020-1011-4> PMID: 32770165
60. Kumaresan V, Balachandar N, Poole SF, Myers LJ, Varghese P, Washington V, et al. Fitting and validation of an agent-based model for COVID-19 case forecasting in workplaces and universities. *Plos one*. 2023 Mar 23; 18(3):e0283517. <https://doi.org/10.1371/journal.pone.0283517> PMID: 36952500
61. Natarajan S, Lipsitz SR, Fitzmaurice GM, Sinha D, Ibrahim JG, Haas J, et al. An extension of the Wilcoxon rank sum test for complex sample survey data. *Journal of the Royal Statistical Society: Series C (Applied Statistics)*. 2012 Aug; 61(4):653–64. <https://doi.org/10.1111/j.1467-9876.2011.01028.x> PMID: 23913985
62. Griffin G.P., Mulhall M., Simek C. and Riggs W.W., Mitigating bias in big data for transportation. *Journal of Big Data Analytics in Transportation*, 2020 Jan 27; 2(1), pp.49–59. <https://doi.org/10.1007/s42421-020-00013-0>
63. Garber M.D., Labgold K. and Kramer M.R., On selection bias in comparison measures of smartphone-generated population mobility: an illustration of no-bias conditions with a commercial data source. *Annals of Epidemiology*, 2022 March 12; 70, pp.16–22. <https://doi.org/10.1016/j.annepidem.2022.03.003> PMID: 35288279
64. Collins S., Starkman E., Coronavirus Incubation Period, 2022 Dec 30, <https://www.webmd.com/covid/coronavirus-incubation-period>
65. Helmer J., Why the COVID-19 Incubation Period Changes and How That Can Affect Us, 2023 May 19.
66. Pomyen Y, Segura M, Ebbels TM, Keun HC. Over-representation of correlation analysis (ORCA): a method for identifying associations between variable sets. *Bioinformatics*. 2015 Jan 1; 31(1):102–8. <https://doi.org/10.1093/bioinformatics/btu589> PMID: 25183485
67. Erfani A, Cui Q. Predictive risk modeling for major transportation projects using historical data. *Automation in Construction*. 2022 Jul 1; 139:104301.
68. Baker R. S., Karumbaiah S., Esbenshade L., & Vitale J. M. Using Demographic Data as Predictor Variables: a Questionable Choice, 2022 Dec 19; <https://doi.org/10.35542/osf.io/y4wv>
69. Paquette L., Ocumpaugh J., Li Z., Andres A., & Baker R. Who's Learning? Using Demographics in EDM Research. *Journal of Educational Data Mining*, 2020, 12(3), 1–30.