

## RUNNING HEAD: PREFERENCE-MATCHING WORLDWIDE

Date: July 3, 2024

**A Worldwide Test of the Predictive Validity of  
Ideal Partner Preference-Matching**

Paul W. Eastwick<sup>1†</sup>, Jehan Sparks<sup>2</sup>, Eli J. Finkel<sup>3,4</sup>, Eva M. Meza<sup>1</sup>, Matúš Adamkovič<sup>5,6,7</sup>, Peter Adu<sup>8</sup>, Ting Ai<sup>9</sup>, Aderonke A. Akintola<sup>10</sup>, Laith Al-Shawaf<sup>11,12</sup>, Denisa Apriliawati<sup>13</sup>, Patrícia Arriaga<sup>14</sup>, Benjamin Aubert-Teillaud<sup>15</sup>, Gabriel Baník<sup>16</sup>, Krystian Barzykowski<sup>17</sup>, Carlota Batres<sup>18</sup>, Katherine J. Baucom<sup>19</sup>, Elizabeth Z. Beaulieu<sup>19</sup>, Maciej Behnke<sup>20,21</sup>, Natalie Butcher<sup>22</sup>, Deborah Y. Charles<sup>23</sup>, Jane M. Chen<sup>24</sup>, Jeong Eun Cheon<sup>25</sup>, Phakkanun Chittham<sup>26</sup>, Patrycja Chwiłkowska<sup>20</sup>, Chin Wen Cong<sup>27</sup>, Lee T. Copping<sup>22</sup>, Nadia S. Corral-Frias<sup>28</sup>, Vera Ćubela Adorić<sup>29</sup>, Mikaela Dizon<sup>30</sup>, Hongfei Du<sup>31</sup>, Michael I. Ehnmowo<sup>32</sup>, Daniela A. Escribano<sup>9</sup>, Natalia M. Espinosa<sup>33</sup>, Francisca Expósito<sup>34</sup>, Gilad Feldman<sup>35</sup>, Raquel Freitag<sup>36</sup>, Martha Frias Armenta<sup>28</sup>, Albina Gallyamova<sup>37</sup>, Omri Gillath<sup>9</sup>, Biljana Gjoneska<sup>38</sup>, Theofilos Gkinopoulos<sup>39</sup>, Franca Grafe<sup>40</sup>, Dmitry Grigoryev<sup>37</sup>, Agata Groycka-Bernard<sup>41</sup>, Gul Gunaydin<sup>42</sup>, Ruby Ilustrisimo<sup>43</sup>, Emily Impett<sup>44</sup>, Pavol Kačmár<sup>45</sup>, Young-Hoon Kim<sup>25</sup>, Mirosław Kocur<sup>46</sup>, Marta Kowal<sup>46</sup>, Maatangi Krishna<sup>47</sup>, Paul Danielle Labor<sup>48</sup>, Jackson G. Lu<sup>49</sup>, Marc Y. Lucas<sup>28</sup>, Wojciech Małecki<sup>50</sup>, Klara Malinakova<sup>51</sup>, Sofia Meißner<sup>40</sup>, Zdeněk Meier<sup>51</sup>, Michal Misiak<sup>46,52</sup>, Amy Muise<sup>53</sup>, Lukas Novak<sup>51</sup>, Jiaqing O<sup>54</sup>, Asil A. Özdoğru<sup>55,56</sup>, Haeyoung Gideon Park<sup>44</sup>, Mariola Paruzel-Czachura<sup>20</sup>, Zoran Pavlović<sup>57</sup>, Marcell Püske<sup>58</sup>, Gianni Ribeiro<sup>59,60</sup>, S. Craig Roberts<sup>61,46</sup>, Jan P. Röer<sup>40</sup>, Ivan Ropovik<sup>62,63</sup>, Robert M. Ross<sup>64</sup>, Ezgi Sakman<sup>65</sup>, Cristina E. Salvador<sup>33</sup>, Emre Selcuk<sup>42</sup>, Shayna Skakoon-Sparling<sup>66</sup>, Agnieszka Sorokowska<sup>41</sup>, Piotr Sorokowski<sup>46</sup>, Ognen Spasovski<sup>67</sup>, Sarah C. E. Stanton<sup>30</sup>, Suzanne L. K. Stewart<sup>68</sup>, Viren Swami<sup>69,70</sup>, Barnabas Szaszi<sup>58</sup>, Kaito Takashima<sup>71</sup>, Peter Tavel<sup>51</sup>, Julian Tejada<sup>36</sup>, Eric Tu<sup>53</sup>, Jarno Tuominen<sup>72</sup>, David Vaidis<sup>15</sup>, Zahir Vally<sup>73</sup>, Leigh Ann Vaughn<sup>74</sup>, Laura Villanueva-Moya<sup>34</sup>, Dian Wisnuwardhani<sup>75</sup>, Yuki Yamada<sup>71</sup>, Fumiya Yonemitsu<sup>76</sup>, Radka Žídková<sup>51</sup>, Kristýna Živná<sup>51</sup>, Nicholas A. Coles<sup>77</sup>

***Psychological Science Accelerator Affiliated Study: PSA-AF-001***

<sup>†</sup>Correspondence should be addressed to Paul W. Eastwick; E-mail: [eastwick@ucdavis.edu](mailto:eastwick@ucdavis.edu)

<sup>1</sup>University of California, Davis, Department of Psychology, USA, <sup>2</sup>UCLA Anderson School of Management, Behavioral Decision Making Group, USA, <sup>3</sup>Northwestern University, Department of Psychology and Morton O. Schapiro IPR Faculty Fellow, Institute for Policy Research, USA, <sup>4</sup>Kellogg School of Management, USA, <sup>5</sup>Slovak Academy of Sciences, Slovakia, <sup>6</sup>Charles University, Czech Republic, <sup>7</sup>University of Jyväskylä, Finland, <sup>8</sup>School of Health, Wellington Faculty of Health, Victoria University of Wellington, Wellington, New Zealand, <sup>9</sup>University of Kansas, Department of Social Psychology, USA, <sup>10</sup>Redeemer's University, Department of Behavioural Studies (Psy), Nigeria, <sup>11</sup>University of Colorado Colorado Springs, Dept of Psychology, USA, <sup>12</sup>Institute for Advanced Study in Toulouse (IAST), France, <sup>13</sup>UIN Sunan Kalijaga Yogyakarta, Indonesia, <sup>14</sup>Iscte-University Institute of Lisbon, CIS\_Iscte, Portugal, <sup>15</sup>Université Paris Cité, France, <sup>16</sup>University of Presov, Slovakia, <sup>17</sup>Jagiellonian University, Faculty of Philosophy, Institute of Psychology, Kraków, Poland, <sup>18</sup>Franklin and Marshall College, Department of Psychology, USA, <sup>19</sup>University of Utah, Department of Psychology, USA, <sup>20</sup>Adam Mickiewicz University, Department of Psychology and Cognitive Science, Poland, <sup>21</sup>Cognitive Neuroscience Center, Adam Mickiewicz University, Poznan, Poland, <sup>22</sup>Teesside University, Department of Psychology, UK, <sup>23</sup>Christ University, Department of Psychology, India, <sup>24</sup>Wellesley College, USA, <sup>25</sup>Yonsei University, Department of Psychology, South Korea, <sup>26</sup>Chulalongkorn University, Faculty of Psychology, Thailand, <sup>27</sup>Department of Social Science, Faculty of Social Science and Humanities, Tunku Abdul Rahman University of Management and Technology, Malaysia, <sup>28</sup>Universidad de Sonora, Mexico, <sup>29</sup>University of Zadar, Department of Psychology, Croatia, <sup>30</sup>University of Edinburgh, UK, <sup>31</sup>Beijing Normal University, Institute of Advanced Studies in Humanities and Social Sciences, Zhuhai, China, <sup>32</sup>University of Ibadan, Department of Psychology, Nigeria, <sup>33</sup>Duke University, Department of Psychology and Neuroscience, USA, <sup>34</sup>University of Granada, Department of Social Psychology, Spain, <sup>35</sup>University of Hong Kong, Hong Kong, <sup>36</sup>Universidade Federal de Sergipe, Brazil, <sup>37</sup>HSE University, Russia, <sup>38</sup>Macedonian Academy of Sciences and Arts, North Macedonia, <sup>39</sup>Institute of Psychology, Jagiellonian University in Krakow, Poland, <sup>40</sup>Witten/Herdecke University, Department of Psychology and Psychotherapy, Germany, <sup>41</sup>University of Wroclaw, Institute of Psychology, Poland, <sup>42</sup>Sabanci University, Faculty of Arts and Social Sciences, Turkey, <sup>43</sup>University of San Carlos, Philippines, <sup>44</sup>University of Toronto, Department of Psychology, Canada, <sup>45</sup>Pavol Jozef Šafárik University in Košice, Department of Psychology, Faculty of Arts, Slovakia, <sup>46</sup>University of Wroclaw, Institute of Psychology, IDN Being Human, Poland, <sup>47</sup>Independent Researcher, Bangalore, India, <sup>48</sup>University of the Philippines, Department of Psychology, Philippines, <sup>49</sup>Massachusetts Institute of Technology, MIT Sloan School of Management, USA, <sup>50</sup>University of Wroclaw, Institute of Polish Studies, IDN Being Human, Poland, <sup>51</sup>Palacký University Olomouc, Olomouc University Social Health Institute, Olomouc, Czech Republic, <sup>52</sup>University of Oxford, School of Anthropology & Museum Ethnography, UK, <sup>53</sup>York University, Toronto, ON, Canada, <sup>54</sup>Singapore Institute of Technology, Singapore, <sup>55</sup>Marmara University, Department of Psychology, Turkey, <sup>56</sup>Üsküdar University, Department of Psychology, Turkey, <sup>57</sup>University of Belgrade, Faculty of Philosophy, Department of Psychology, Serbia, <sup>58</sup>ELTE Eötvös Loránd University, Institute of Psychology, Budapest, Hungary, <sup>59</sup>The University of Queensland, School of Psychology, Australia, <sup>60</sup>University of Southern Queensland, School of Law and Justice, Australia, <sup>61</sup>University of Stirling, UK, <sup>62</sup>Charles University, Institute for Research and Development of Education, Czech Republic, <sup>63</sup>University of Presov, Faculty of Education, Slovakia, <sup>64</sup>Macquarie University, Department of Philosophy, Australia, <sup>65</sup>Bilkent University, Department of Psychology, Turkey, <sup>66</sup>Toronto Metropolitan University, Canada, <sup>67</sup>Ss. Cyril and Methodius University in Skopje, Faculty of Philosophy, North Macedonia, <sup>68</sup>University of Chester, School of Psychology, UK, <sup>69</sup>Anglia Ruskin University, School of Psychology, Sport, and Sensory Sciences, UK, <sup>70</sup>Perdana University, Centre for Psychological Medicine, Malaysia, <sup>71</sup>Kyushu University, Japan, <sup>72</sup>University of Turku, Department of Psychology and Speech-Language Pathology, Finland, <sup>73</sup>United Arab Emirates University, United Arab Emirates, <sup>74</sup>Ithaca College, Psychology Department, USA, <sup>75</sup>University of Indonesia, Indonesia, <sup>76</sup>Chuo University, Japan, <sup>77</sup>Stanford University, Center for the Study of Language and Information, USA

## Abstract

Ideal partner preferences (i.e., ratings of the desirability of attributes like attractiveness or intelligence) are the source of numerous foundational findings in the interdisciplinary literature on human mating. Recently, research on the predictive validity of ideal partner preference-matching (i.e., do people positively evaluate partners who match versus mismatch their ideals?) has become mired in several problems. First, articles exhibit discrepant analytic and reporting practices. Second, different findings emerge across laboratories worldwide, perhaps because they sample different relationship contexts and/or populations. This registered report—partnered with the Psychological Science Accelerator—uses a highly powered design ( $N=10,358$ ) across 43 countries and 22 languages to estimate preference-matching effect sizes. The most rigorous tests revealed significant preference-matching effects in the whole sample and for partnered and single participants separately. The “corrected pattern metric” that collapses across 35 traits revealed a zero-order effect of  $\beta=.19$  and an effect of  $\beta=.11$  when included alongside a normative preference-matching metric. Specific traits in the “level metric” (interaction) tests revealed very small (average  $\beta=.04$ ) effects. Effect sizes were similar for partnered participants who reported ideals before entering a relationship, and there was no consistent evidence that individual differences moderated any effects. Comparisons between stated and revealed preferences shed light on gender differences and similarities: For attractiveness, men’s and (especially) women’s stated preferences underestimated revealed preferences (i.e., they thought attractiveness was less important than it actually was). For earning potential, men’s stated preferences underestimated—and women’s stated preferences overestimated—revealed preferences. Implications for the literature on human mating are discussed.

Keywords: attraction, close relationships, human mating, ideals, matching hypothesis

## **A Worldwide Test of the Predictive Validity of Ideal Partner Preference-Matching**

The study of human mating is vast and interdisciplinary, spanning fields as diverse as economics (Hitsch et al., 2010), evolutionary psychology (Buss & Schmitt, 2019), family studies (Boxer et al., 2015), sociology (Lewis, 2016), and social/personality psychology (Fletcher et al., 2019). Despite the considerable depth and breadth of these fields, they share in common a key construct: *ideal partner preferences*. Ideal partner preferences are the attributes (e.g., attractiveness, intelligence, sense of humor) that people say they desire in a romantic partner and, for 80 years, scholars have been using this construct as the foundation for a variety of theories and models that explain how humans pursue and maintain mating relationships (Buss, 1989; Eagly & Wood, 1999; Fletcher, et al., 1999; Hill, 1945; see Eastwick et al., 2014, for a review).

For many decades, scholars made the straightforward assumption that ideal partner preferences affected how positively people feel about their romantic partners—which is itself a key predictor of health and mortality (Robles et al., 2014). But only in the last 25 years have researchers begun to empirically examine the preference-matching question: That is, does a person positively evaluate a given romantic partner to the extent that the partner's attributes match the person's ideals? This matching hypothesis is the core novel prediction offered by the Ideal Standards Model—an influential model in the close-relationships tradition (Fletcher, et al., 2000; Fletcher et al., 1999; Simpson, et al., 2001)—and this hypothesis emerges in evolutionary psychological models as well (Buss, 1989; Conroy-Beam & Buss, 2016; Li & Meltzer, 2015; Shackelford & Buss, 1997; Sugiyama, 2005). Indeed, it is challenging to articulate what the ancestral, functional consequences of ideal partner preferences would be unless the match

between preferences and a partner's attributes had some meaningful association with romantic evaluations.

Does the empirical evidence support this matching hypothesis? In brief, the evidence is murky, and it has actually become murkier rather than clearer over time. Today, researchers can cite empirical papers supporting or refuting any point they wish to make about this matching hypothesis. This state of affairs is unfortunate, because precise effect size estimates for the matching hypothesis will have generative and theory-building implications no matter what they turn out to be. If the match between ideals and a partner's traits predicts romantic evaluations with (at least) modest effect sizes, then scholars should be able to assess participants' ideals and match them with new, compatible partners or determine whether their current relationships are likely to encounter difficulties. But if these effect sizes are small or near-zero, then explanations for the role of compatibility in human mating will need to become grounded in alternative theories that do not rely on attribute matching (e.g., the way two people co-construct their expectations, shared reality, or relationship narrative; Berscheid & Ammazzalorso, 2001, Eastwick, et al., 2023; Rossignac-Milon & Higgins, 2018). Inspired by other large collaborative replication efforts (Coles et al., 2020; Vohs et al., 2021), the current project aims to gather the *strongest possible evaluation* of the predictive validity of ideal partner preferences—perhaps the most interdisciplinary and theoretically central construct in research on human mating.

### **Ongoing Challenge #1: Lack of Standard Analytic Practices**

One reason that the predictive-validity evidence to date remains murky is differing analytic and reporting practices. There are many ways that the matching hypothesis has been operationalized—some more rigorous than others. Specifically, researchers have tested the predictive validity of ideal partner preferences in four primary ways: ideal-trait correlations, the

raw pattern metric, the corrected pattern metric, and the level metric. Our own systematic review yielded 35 published studies (Table S1) that have reported data that (a) examine participants' evaluations of a person they have met face-to-face (i.e., from speed-dating partners to established romantic partners), and (b) bear on at least one of these four approaches. The four approaches are illustrated with a mock dataset in Table S2.

First, scholars sometimes report *ideal-trait correlations*: for a particular trait, the researcher calculates the association between participants' ideals and the partners' traits (Example 1a in Table S2) in a sample that presumably involved some prior selection event (e.g., the partners are people whom the participants selected as a romantic partner). In other words, do people with a stronger preference for a trait end up with partners who are higher on the trait? However, the selection event is not used as a measured variable (i.e., there are no "unselected" partners)—so it cannot serve as a dependent measure—and no evaluative outcomes, such as attraction or relationship satisfaction, are collected (Conroy-Beam & Buss, 2016; Gerlach et al., 2019). Thus, these correlations are not rigorous tests of the matching hypothesis, as there are many alternative explanations for any such correlation (Eastwick et al., 2019; Fletcher et al., 2020). Indeed, the canonical papers using this approach (e.g., Fletcher et al., 1999; Murray et al., 1996) generally presumed that these correlations reflected a motivated reasoning process (e.g., people are motivated to believe that their current partner possesses the traits that they ideally want) rather than ideal partner preference-matching. These correlations are included in the analysis plan because they are available as a matter of course when conducting the more rigorous tests described next.

A second *pattern metric (raw)* approach uses the within-person correlation between (a) each participant's ideals and (b) a target partner's traits (usually rated by participants themselves)

across *all* available traits. Researchers subject this correlation to a Fisher  $z$ -transformation and then use it to predict an evaluative outcome (e.g., relationship satisfaction; Example 1b in Table S2). This approach typically reveals moderately sized associations ( $r = .20\text{-.40}$ ) with relationship satisfaction, which is consistent with the ideal partner preference-matching hypothesis. However, as methodologists have compellingly described (Wood & Furr, 2016; Rogers et al., 2018), this approach has a major shortcoming: The predictive power of the raw pattern metric is confounded with the normative desirability of the ideal traits and target partner traits that are used to calculate the within-person correlation. In other words, the raw pattern metric approach may have garnered support for the ideal partner preference-matching hypothesis because people tend to report positive evaluative outcomes when they think their partner has positive traits; thus, this approach does not uniquely test whether the *match* between ideals and partner traits has predictive effects. Approaches using Euclidean distance metrics share this shortcoming (e.g., Conroy-Beam et al., 2016; see Rogers et al., 2018).

A third *pattern metric (corrected)* approach follows Wood and Furr's (2016) recommendation to mean-center each ideal rating and partner trait rating ( $a$  and  $b$  in the paragraph above) prior to the calculation of the within-person correlation; just as with the raw pattern metric, this correlation can then be  $z$ -scored and used to predict an evaluative outcome (Example 1c in Table S2). This procedure removes the normative desirability confound and permits a clean test of the ideal partner preference-matching hypothesis, and published effect sizes range from near zero to  $r \sim .25$  (Eastwick et al., 2019; Fletcher et al., 2020; Lam et al., 2016).

A fourth *level metric* approach refers to the statistical interaction between the participant's ideal and the partner's trait (i.e., the ideal  $\times$  trait term) when predicting an

evaluative outcome (controlling for the main effects of the ideal and trait; example 1d in Table S2). For example, assume there is a positive association of (a) perceiving a partner to be funny with (b) attraction to that partner. The level metric tests whether this association is stronger (i.e., more positive) among participants who have high (rather than low) ideals for a funny partner—as if participants with high ideals are “weighting” the trait more positively in their evaluative judgments. This approach is designed to be implemented one-trait-at-a-time, which is critical when testing theories positing that ideals for specific attributes have functional outcomes (e.g., the hypothesis that heterosexual women have a stronger preference for *financial success* in a partner because they have historically needed to differentiate strong from weak providers more so than heterosexual men; Buss, 1989; Eastwick & Finkel, 2008; Eastwick et al., 2014; Li et al., 2013; Perusse, 1993). Significant effects emerge sporadically using this approach (e.g., Fletcher et al., 2020; Valentine et al., 2020), but high-powered level metric tests across multiple attributes are uncommon.

Critically, few papers report more than one of the four approaches (see Table S1), and researchers who draw conclusions from the weaker approaches (i.e., ideal-trait correlations, the raw pattern metric) are more likely to conclude support for the matching hypothesis than are researchers who use the stronger approaches (i.e., the corrected pattern metric, the level metric). The current registered report addressed the challenge of discrepant reporting practices by bringing together a diverse team of researchers who all committed to a preregistered analysis plan with all four analytic strategies described above.

### **Ongoing Challenge #2: Differences between Established Relationships and Initial Attraction**

A second reason that the state of the matching hypothesis is uncertain is that ideal partner preference-matching effects may depend on relationship context. The matching hypothesis has historically received support when participants evaluated a current romantic partner, as suggested by studies of established relationships (e.g., Fletcher et al., 1999, 2000, 2020; Zentner, 2005). But the hypothesis has not commonly been supported when participants evaluated a partner with whom they were not romantically involved, as suggested by studies of initial attraction (e.g., Eastwick & Finkel, 2008; Selerman & Gideon, 2022; Wu et al., 2018). However, direct comparisons of effect sizes for established relationship versus initial attraction partners remain elusive, as studies conducted in these two contexts typically differ from each other in innumerable ways.

To address context as a potentially critical moderator, the current project collected data on both established relationship and initial attraction partners using a method (adapted from Eastwick et al., 2011, and Sparks et al., 2020) that enables a clean comparison between these two contexts. Specifically, participants who were in an established relationship completed scales about their current romantic partner, and participants who were single completed the *identical scales* about the partner with whom they would most desire to have a romantic relationship. By using the same items and procedure in both relationship contexts, the two effect sizes can be compared to each other more straightforwardly than in prior studies.

Researchers have speculated that a difference between initial attraction and established relationship contexts could emerge because the ideal standards model primarily applies to long-term partnerships, and/or because participants only draw from their (abstract) ideal partner preferences once the relationship itself becomes an abstract entity with a hypothetical future (Eastwick et al., 2014; Meltzer et al., 2014). Nevertheless, there are two reasons for such a

difference that would be grounded in motivated perceptual processes rather than the ideal standards model *per se*. First, people may adjust their *perceptions of their partner's traits* to match their ideals (Murray et al., 1996), perhaps especially if they are happy in their current relationship. This interpretation is always plausible whenever participants provide their own ratings of a partner's traits—the most common method in this literature by far.<sup>1</sup> To examine this possibility, we also assessed each partner's level of formal education (e.g., high school, college degree)—a more objective measure that should be less subject to motivated re-interpretation than typical trait ratings of the partner. To the extent that preference-matching effects are a function of motivated perception of the partner's traits, the effect size for the level metric should be smaller for level of education.

Second, people may adjust their *ideals* to match their perceptions of their partner's traits (Gerlach et al., 2019; Neff & Karney, 2003), perhaps especially if they are happy in their current relationship. One way to address this alternative explanation is to collect participants' ideals before the relationship forms in the first place (Eastwick et al., 2011). To examine the possibility that people use their pre-relationship ideals when evaluating an ongoing relationship, we also recruited an additional sample through Cloud Research. These participants reported their ideals when single and then, after they started a new romantic relationship (several months later), they completed measures about their current romantic partner. To the extent that preference-matching

---

<sup>1</sup> To illustrate, 30 of the 35 studies in Table S1, or 86%, used this approach, whereas 23% asked partners to self-report their own traits, and 20% used some “objective” measure of the trait. (These numbers add to more than 100% because some studies employed multiple approaches.) The current study is primarily designed to establish robust effect size estimates for the (most common) participant-perception approach, which could then inform power analyses for future investigations of the other two (considerably more intensive, but usually less well-powered) approaches.

effects are a function of the motivated shifting of one's own ideals, the effect sizes in this “newly partnered” sample should be smaller.

### **The Current Research**

This collaborative effort produced the largest cross-national dataset of participants' evaluations and judgments about preferred-gender targets they know personally (e.g., romantic partners, friends, acquaintances). The specific research questions in the Primary Planned Analyses are outlined in Table 1. Research Questions (RQs) 1-4 rely on traditional null hypothesis significance testing; nevertheless, interpretations will focus primarily on effect size estimates vis-à-vis Cohen's (1992) small, medium, and large conventions. Effect sizes for the level metric (i.e., statistical interactions) will be interpreted as fractions of the attribute main effects. In tutorials of interaction statistical power (Baranger et al., 2023; Giner-Sorolla, 2018), a “knock out” interaction (i.e., interaction effect size beta = main effect size beta) is akin to a medium sized effect, and a “50% attention” interaction (i.e., interaction effect size beta = 50% of main effect size beta) is akin to a small effect. All four research questions were evaluated with all four analytic approaches described above.

Table 1 – Design Table: Primary Planned Analyses (Research Questions and Hypotheses)

	Research Question	Hypothesis	N
1	What is the (overall) effect size of ideal partner-preference matching?	a. Ideal-trait correlations ( $rs$ ) are greater than zero. b. The raw pattern metric ( $r$ ) is greater than zero. c. The corrected pattern metric ( $r$ ) is greater than zero. d. Level metric tests (interaction $\beta$ s) are greater than zero.	10,358 full sample
2	What is the effect size of ideal partner-preference matching in initial attraction contexts?	a. Ideal-trait correlations ( $rs$ ) are greater than zero. b. The raw pattern metric ( $r$ ) is greater than zero. c. The corrected pattern metric ( $r$ ) is greater than zero. d. Level metric tests (interaction $\beta$ s) are greater than zero.	4,152 subsample
3	What is the effect size of ideal partner-preference matching in established relationship contexts?	a. Ideal-trait correlations ( $rs$ ) are greater than zero. b. The raw pattern metric ( $r$ ) is greater than zero. c. The corrected pattern metric ( $r$ ) is greater than zero. d. Level metric tests (interaction $\beta$ s) are greater than zero.	5,544 subsample
4	Does the effect size of ideal partner-preference matching depend on initial attraction vs. established relationship context?	a. Ideal-trait correlations ( $rs$ ) are larger when reporting on current partners than desired partners. b. The raw pattern metric ( $r$ ) is larger when reporting on current partners than desired partners. c. The corrected pattern metric ( $r$ ) is larger when reporting on current partners than desired partners. d. Level metric tests (interaction $\beta$ s) are larger when reporting on current partners than desired partners.	4,152 vs. 5,544 subsamples

Note: All (a) ideal-trait correlations and (d) level metric tests involve 35 separate tests, one for each attribute in Tables 2 and 3. In these cases, we used a Holm-Bonferroni correction (Holm, 1979) to control the family-wise Type-I error rate, and we discuss possible power implications in the A Priori Power Analysis Plan section in the Supplemental Materials.

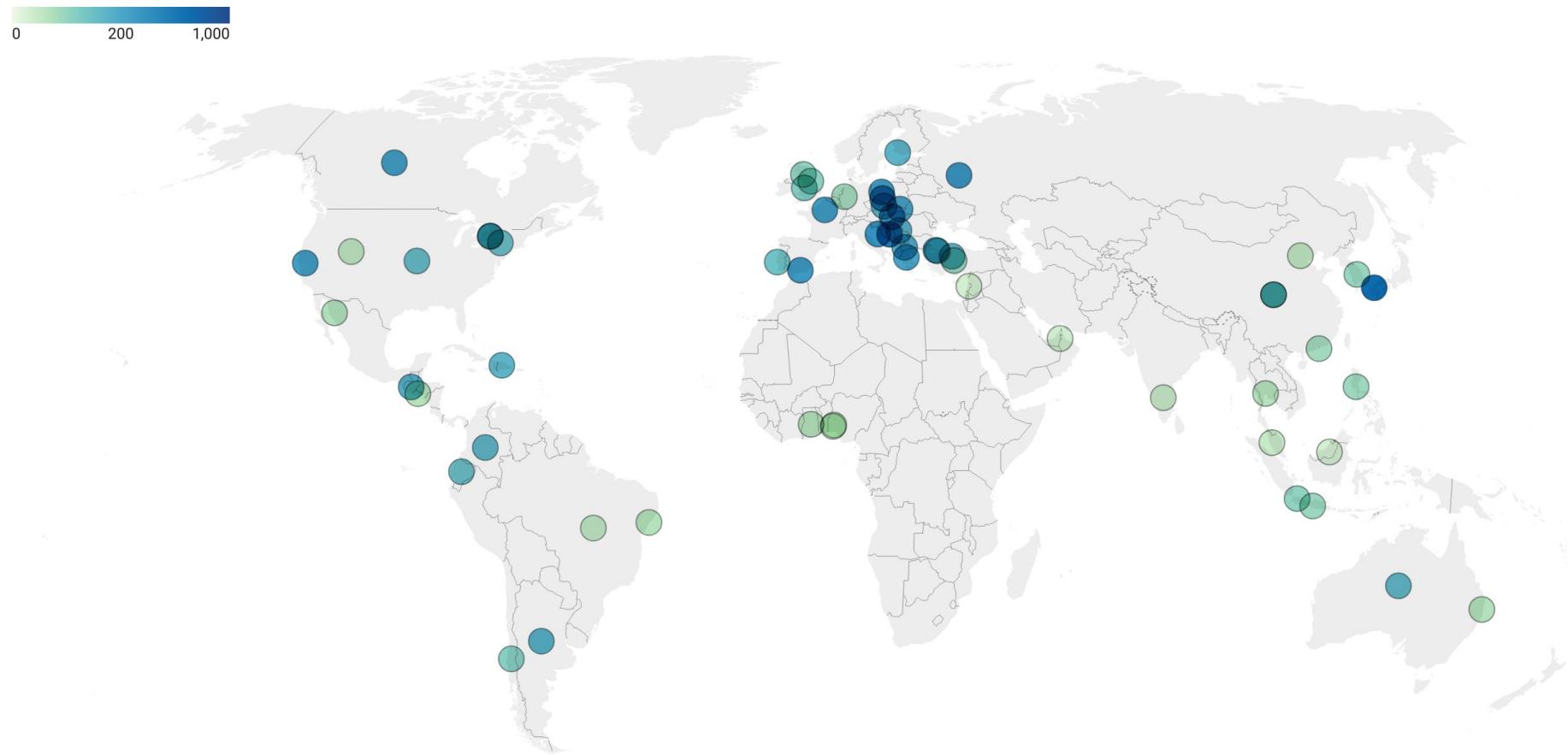
## Method

This study mimics the design of an influential, initial test of the predictive validity of ideal partner preference-matching (Fletcher et al., 1999, Study 6). Specifically, participants (a) provided their ideals on a variety of traits, (b) rated their current romantic partner on those same traits, and finally (c) reported an evaluation of their current partner as the outcome dependent measure. This procedure remains the gold-standard in this research space, but it was updated in three ways: (a) participants who were single were not excluded from participating, but were instead given the chance to evaluate the person with whom they most desire to have a romantic relationship (as in Eastwick et al., 2011, Study 3); (b) participants also evaluated three additional targets—peers of their preferred gender (as in Sparks et al., 2020)—to enable additional analytic tests (elaborated below); and (c) participants rated a larger set of traits (not just the traits highlighted in Fletcher et al., 1999, but also the Big Five personality traits; Goldberg, 1993).

## Ethics

Each research group ensured that they had approval from their institution's Ethics Committee or IRB to conduct the study, that the study was covered by the approved UC Davis IRB (exempt protocol 1898056-1 “The Preference Matching Project”), or that the study was exempt (see the Supplemental Materials for details).

**Figure 1 – 60 Samples Included in the Preference-Matching Project**



Note: Locations indicate the university where the data were collected or—in the cases of online community samples—the center of the relevant country. Map created with Datawrapper (Lorenz et al., 2012).

## **Participants**

Our final sample consisted of  $N = 10,358$  participants (after planned exclusions; see “Data Processing” for details) from 60 samples and 43 different countries (Table S3 and Figure 1). Some of the 60 samples assessed only student (undergraduate and graduate) participants ( $k = 22$  samples), some assessed only community participants ( $k = 8$ ), and some assessed a blend of student and community participants ( $k = 30$ ). Students typically received course credit, and community members were compensated in a manner determined appropriate for their local context (e.g., cash, electronic payments, gift cards, raffles, and some were not directly compensated).

Participants were  $M = 28.5$  years old ( $SD = 11.7$ ; we assumed that values less than 10 or greater than 100 were typos). In terms of gender,  $N = 6,833$  (66.0%) were women,  $N = 3,394$  (32.8%) were men,  $N = 127$  (1.2%) preferred to self-describe their gender, and  $N = 4$  provided no response. In terms of sexual orientation,  $N = 8,366$  (80.7%) were straight/heterosexual,  $N = 1,217$  (11.7%) were bisexual,  $N = 361$  (3.5%) preferred to self-describe,  $N = 202$  (2.0%) were gay,  $N = 162$  (1.6%) were lesbian, and  $N = 50$  (0.5%) either skipped this question or this question was intentionally omitted because queer identities were punishable in that context. In terms of education,  $N = 89$  (0.9%) reported “less than high school,”  $N = 3,601$  (34.8%) “high school,”  $N = 2,559$  (24.7%) “some college,”  $N = 2,556$  (24.7%) “four-year degree,”  $N = 1,370$  (13.2%) “Master’s degree,”  $N = 182$  (1.7%) “Doctorate or professional degree,” and  $N = 1$  provided no response.

## **Procedure**

The entire study consisted of a survey that could be completed on an electronic device. Data collection began on February 1, 2023 (after the stage 1 registered report was approved) and closed on November 10, 2023.

After providing consent and clicking a ReCAPTCHA button (to prevent bots from accessing the survey), participants completed two blocks of measures (in counterbalanced order). In the first block, they rated the desirability of 35 ideal partner preference attributes (as well as their ideal for a “high level of education,” to be used in a separate analysis), and they completed a brief set of demographic items and individual-difference measures.

The second block consisted of a set of items about specific partners. Using a procedure implemented successfully by Sparks et al. (2020, Study 2), participants were asked to provide the first name and last initial of four individuals whom they know personally.<sup>2</sup> They were instructed to choose individuals of their romantically preferred gender who are not related to them, who are around the same age as them (i.e., peers), and whom they have met in person. Participants who were in a romantic relationship were instructed to list their current romantic partner as the first of the four individuals; participants who were single were asked to list “the person with whom you would most desire to have a romantic relationship” (Eastwick et al., 2011, Study 3) as the first of the four individuals. Third, they rated the first of the four targets (i.e., the current partner or most desired partner) on the same set of 35 attributes. Fourth, they rated the first of the four targets on the romantic evaluation dependent measure. The third and fourth steps were counterbalanced. Fifth, they repeated the third and fourth steps (randomly counterbalanced for each target) for each of the remaining three targets (presented in a random order). The first target was completed

---

<sup>2</sup> This potentially identifying information was removed and replaced with numerical codes prior to the public posting of the dataset.

prior to the remaining three peer targets because the preregistered analysis plan focused on these targets in particular.

Data, analysis code, codebook, and preregistration (i.e., the Stage 1 manuscript) are available at [https://osf.io/b29vu/?view\\_only=35a15592f8b04cdfb9ab32f45c73f3c6](https://osf.io/b29vu/?view_only=35a15592f8b04cdfb9ab32f45c73f3c6).

## Materials

**Translation.** For surveys in languages other than English, participating laboratories translated the original English materials into the target language (see Table S3). All laboratories first used the translate feature in Qualtrics (which uses Google translate) to generate the initial translation, edited as necessary, and then had an independent researcher who was fluent in the target language read it over for comprehensibility. Then, consistent with translation best-practices (Benet-Martinez, 2007), one or more (different) researchers who were fluent in English and the target language back-translated, compared the back-translation against the original, and resolved discrepancies. Researchers at different universities who were administering surveys in the same language collaborated to ensure that their surveys were as similar as possible. In total, the surveys were administered in 22 different languages (see Table S3 for details).

**Ideal partner preferences.** Participants rated 35 attributes (Table S4) in an ideal romantic partner on a scale ranging from 1 (*not at all desirable*) to 11 (*highly desirable*). Scale derivation work by Fletcher et al. (1999) produced a popular measure of three factors: warmth/trustworthiness, vitality/attractiveness, and status/resources. We included five items assessing warmth/trustworthiness, five items assessing vitality/attractiveness, and four items assessing status/resources from this measure. We also included ten moderately-to-highly desirable traits that emerged in a more recent article using a similar scale-derivation procedure (Sparks et al., 2020), the Ten-Item Personality Inventory (i.e., a measure of the Big Five

personality traits; Gosling et al., 2003), and one trait with potentially crucial cross-cultural relevance (*smells good*; Roberts et al., 2020). The full collection of 35 attributes contained a mix of attributes that typically range from low to high levels of (self-reported) desirability in an ideal partner.

In addition, participants rated the extent to which “a high level of education is desirable” in their ideal romantic partner on a 1 (*not at all desirable*) to 11 (*highly desirable*) scale.

**Partner attributes.** Participants rated how the 35 attributes characterized each target on a scale from 1 (*not at all characteristic*) to 11 (*highly characteristic*). They also indicated the highest level of education that the partner had completed from a set of six options ranging from low to high (e.g., less than high school, high school, some college, four-year degree, Master’s degree, Doctorate or Professional degree). The wording of these categories was adapted to each countries’ educational context where needed; all adaptations contained six categories in ascending order. We decided a priori to treat this item separately from the other 35 attributes because it is distinct on both a conceptual (i.e., it is not really a psychological trait but rather an objective fact about a person) and measurement level.

**Romantic evaluation (dependent measure).** Participants reported their *romantic evaluation* of each of their four nominated targets on six items (“I am romantically interested in \_\_\_\_\_;” “\_\_\_\_\_ is the only person I want to be romantically involved with;” “\_\_\_\_\_ is very much my ideal romantic partner;” “It is important to me to see or talk with \_\_\_\_\_ regularly;” “\_\_\_\_\_ is the first person that I would turn to if I had a problem;” and “If I achieved something good, \_\_\_\_\_ is the person that I would tell first”) on a 1 (*strongly disagree*) to 11 (*strongly agree*) scale (see Table S5). Importantly, this measure was designed to be equally applicable to relationships with peers and with romantic partners (see Supplemental Material

below for scale-derivation details). Reliabilities were  $\alpha = .92$ ,  $\omega = .92$  on the full sample,  $\alpha = .91$ ,  $\omega = .91$  on the partnered sample, and  $\alpha = .85$ ,  $\omega = .85$  on the single sample.

**Individual-difference measures and demographic information.** Participants completed additional items including a 16-item measure of individualism/collectivism (e.g., “I’d rather depend on myself than others,” “Parents and children must stay together as much as possible,” Triandis & Gelfand, 1998), a 12-item measure of relational mobility (e.g., “They [the people around you] have many chances to get to know other people,” “It is easy for them to meet new people,” Thomson et al., 2018), and an item assessing relationship status (i.e., *yes* vs. *no* to “I am currently in a committed, romantic relationship”).

Participants also indicated the nature of their relationship with each of the four targets using the following (mutually exclusive) categories: spouse or fiancé, boyfriend/girlfriend/committed romantic partner, casual romantic/sexual partner, friend, colleague or co-worker, acquaintance, stranger. Additional individual differences and demographic information (beyond those referenced in the manuscript) are described in the Supplemental Materials.

**Attention checks.** In addition to the ReCAPTCHA button, there were two additional “directed query” attention checks (Abbey & Meloy, 2017). First, after the consent form, participants saw an item that lists the names of the seven continents and instructions that read: “If you are reading this query, please select ‘Other’ and type the word ‘nonsense’ in the blank to assure the researchers that you are reading the instructions.” Because some participants typed in a nonsense word into the blank space, we decided (before running any analyses) to use all participants who selected ‘Other’ and typed something in the space. Second, for the first target

only, the romantic evaluation items contained an additional item that stated: “Please select ‘3’ for this item to show that you are paying attention.”

**Relationship formation hypothesis.** As described above, one possible explanation for the stronger support for ideal partner preference-matching in established close relationship (vs. initial attraction) contexts is that people may be motivated to change their ideals to match their current partner’s attributes (Gerlach et al., 2019; Neff & Karney, 2003). To test this possibility, we collected a separate sample of  $N = 1,585$  participants (i.e., online workers from the “Cloud Research Approved List” on MTurk; Hauser et al., 2023) who completed two surveys at two points in time, about 3.5 months apart ( $M = 104$  days,  $SD = 12$ , range = 77-124). The sample consists of (a) participants who were in a relationship with the same partner at both time points ( $N = 709$ ), (b) participants who were single at both time points ( $N = 687$ ), and (c) participants who were single at the first time point and in a relationship with a new partner at the second time point ( $N = 189$ ).

The recruitment plan and demographics for this sample is described in detail in the Relationship Formation Hypothesis section of the Supplemental Materials; we preregistered that these participants would be analyzed separately from the main analyses that correspond to the Table S3 worldwide sample, given the procedural differences and the fact that these participants were all from the U.S.

These participants completed a subset of the measures reported above. Specifically, at time 1, they reported their ideal partner preferences and demographics in a 3-minute survey (for US\$1), and then at time 2, they completed the partner attribute and dependent measure items in a 10-minute survey (for US\$5). They completed the relationship status item on both surveys, but

the surveys did not include the additional individual differences and the three additional targets (i.e., these participants only completed items about the current partner or most desired partner).

## **Data Processing**

Once again, our final international sample consisted of  $N = 10,358$  participants. Not included in this value are the participants who were excluded from analyses because they (a) “straightlined” (i.e., give the same numerical response to) either the 35 ideal partner preference items or the 35 attribute ratings ( $N = 194$ ), (b) failed to pass both attention checks ( $N = 2,600$ ), or (c) failed to reach the debriefing screen ( $N = 6,932$ ; most of these participants stopped responding a short way into the survey).

Participants were included in the  $N = 10,358$  total and the overall analysis (i.e., research question 1) but excluded from the relationship status subgroup analyses if they (a) indicated that they were “single” but then categorized the first target they nominated as “spouse or fiancé” or “boyfriend/girlfriend/committed romantic partner,” or (b) indicated that they were “in a relationship” but then categorized the first target they nominated as anything other than “spouse or fiancé” or “boyfriend/girlfriend/committed romantic partner.”  $N = 662$  were included in the overall sample but excluded from the relationship status subgroup analyses for these reasons, which yielded  $N = 5,544$  participants in the “partnered” category (with an average relationship length of  $M = 6.3$  years and  $SD = 8.8$  years, assuming the  $N = 12$  values above 1,000 months were typos), and  $N = 4,152$  participants in the “single” category for analyses.

We did not anticipate, nor did we have, a high proportion of missing/incomplete data (less than 1% for all variables). Nevertheless, we also used predictive means matching using the mice package for R (van Buuren & Groothuis-Oudshoorn, 2011) to investigate the possible

consequences of missingness in a separate set of sensitivity analyses for Tables 2 and 3 (see Supplemental Materials).

## Results

### Primary Planned Analyses

As preregistered, these analyses pertained only to the first target that participants evaluated. All analyses were conducted as multilevel models that accounted for the nesting of participant within the  $k = 60$  samples (see Table 2 and 3 notes). Specifically, we included random intercept ( $u_0$ ) and slope ( $u_1$ ) terms in each analysis, and the random slope ( $u_1$ ) for sample was omitted when a given analysis did not converge. Overall, these random terms were fairly modest in magnitude: For the overall sample analyses reported in Table 2, random intercept ( $u_0$ ) terms accounted for 2.3% of the variance on average (i.e., 2.3% of the residual variance in the trait dependent measure was attributable to the sample) and random slope ( $u_1$ ) terms accounted for 0.3% of the variance. For the overall sample analyses reported in Table 3, random intercept ( $u_0$ ) terms accounted for 3.4% of the variance on average, and random slope ( $u_1$ ) terms accounted for 0.6% of the variance. In other words, the trait means (i.e., the DV in Table 2) and romantic evaluation DV (i.e., the DV in Table 3) showed some minor differences (about 3%) across samples. However, the association of ideals with traits (i.e., the associations in Table 2) and the association of traits with romantic evaluations (i.e., the associations in Table 3) differed very little (less than 1%) across samples.<sup>3</sup> All variables were standardized ( $M = 0$ ,  $SD = 1$ ) for each analysis.

As described above (and in Table S2), *ideal-trait correlations* refer to the between-persons association of the ideal rating and the partner attribute rating for a given attribute. One

---

<sup>3</sup> We calculated these % variance values using the r2mlm package in R (Shaw et al., 2020).

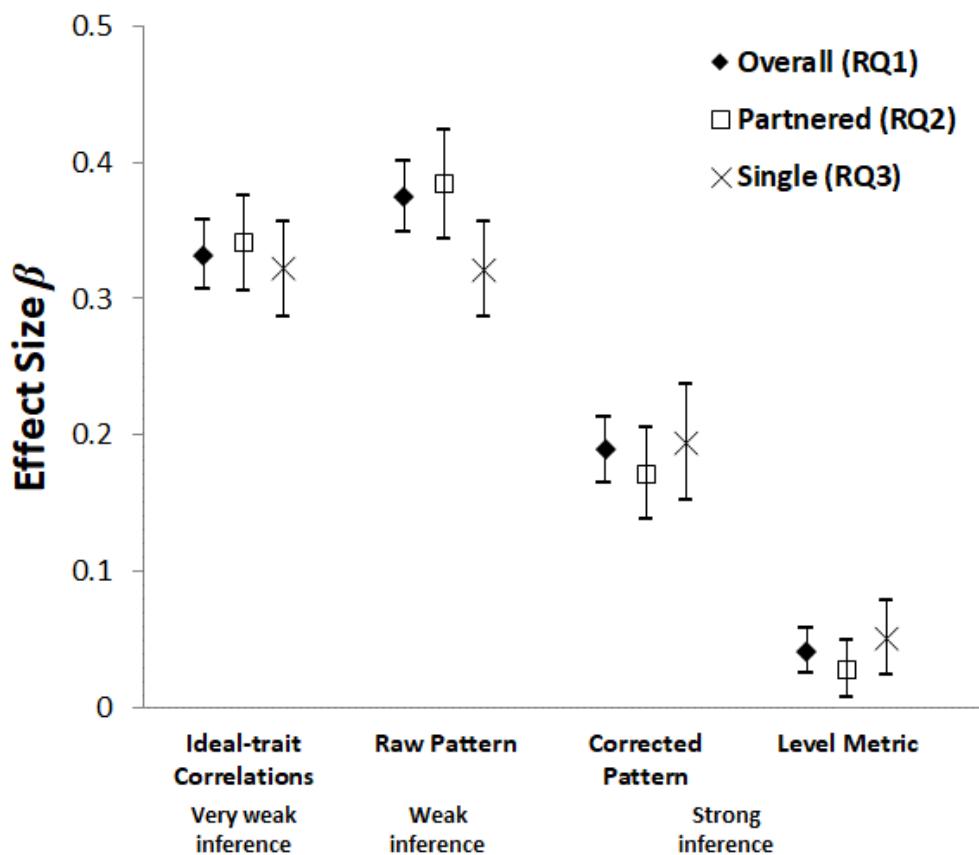
association is calculated for each attribute, and the dependent measure is not used in this calculation (Table 2). The *pattern metric (raw)* is the association between (a) a *Fisher-z* scored version of the within-person correlation between the 35 ideal ratings and the 35 partner-attribute ratings, and (b) the romantic evaluation measure. The *pattern metric (corrected)* is the association between (a) a *Fisher-z* scored version of the within-person correlation between the 35 ideal ratings and the 35 partner-attribute ratings after sample-mean centering all 70 items, and (b) the romantic evaluation measure. The *level metric* is the ideal  $\times$  attribute interaction predicting the romantic evaluation measure, controlling for the main effect of ideal and attribute (Table 3).

Given that we are assessing three constructs from Fletcher et al. (1999) and all five of the Big Five constructs (see Table 2), the ideal-trait correlations and level metric tests were conducted not only at the item level, but also at the construct level for the three Fletcher et al. (1999) constructs (i.e., warmth/trustworthiness, vitality/attractiveness, and status/resources) and the Big Five (i.e., Extraversion, Agreeableness, Conscientiousness, Emotional Stability, and Openness to Experience). The pattern metric analyses were calculated on the full set of 35 attributes, because such profile correlations require many items to assess reliably (Wood & Furr, 2016).

Given that the corrected pattern metric and the level metric provide the strongest tests of the ideal-partner preference matching hypothesis, our interpretations of the findings rely primarily on these effect sizes. We provide the ideal-trait correlations and raw pattern metric effect sizes for completeness and transparency. Importantly, the ideal-trait correlations and level metric analyses in Tables 2 and 3 require 35 statistical tests, one for each attribute. Therefore, we implemented a Holm-Bonferroni correction (Holm, 1979) for all instances where we conducted 35 statistical tests, and we only conclude support for attributes that pass this significance

threshold (i.e.,  $.05/35 = .0014$  = alpha for the lowest p-value of the 35;  $.05/34 = .0015$  = alpha for the second lowest p-value of the 35;  $.05/33 = .0015$  = alpha for the third lowest p-value, and so on). In all tables, the attributes are listed in the order that participants spontaneously nominated them in the classic Fletcher et al. (1999) paper (see Table S4). A summary of the central findings is depicted in Figure 2.

**Figure 2 – Results for Research Questions (RQs) 1-3**



Note: Values for ideal-trait correlations and level metric are averaged across the 35 traits. Bars depict upper and lower 95% CIs.

**Weak inference tests.** As anticipated, ideal-trait correlations (Table 2) were positive and significant across the board ( $\beta_I = .33$  across the 35 traits on average): Participants who had high

ideals for a trait tended to report that the target possessed higher amounts of that trait. These correlations trended higher for partnered ( $\beta_I = .34$  on average) than single ( $\beta_I = .32$  on average) participants, and 16 out of 35 of the partnered vs. single comparisons passed the Bonferroni-Holm correction. Nevertheless, these partnered vs. single differences tended to be very small.

Also as expected, the raw pattern metric (i.e., the within-person correlation between the 35 ideals and traits) predicted romantic interest strongly, with effect sizes in the medium-to-large range ( $\beta_I = .37$  in the full sample, see Table 3). As with the ideal-trait correlations, this association was slightly stronger for partnered ( $\beta_I = .38$ ) than single ( $\beta_I = .32$ ) participants.

In a non-preregistered analysis, we additionally examined whether a measure of Euclidean distance (i.e., the square root of the sum of the squared differences between ideals and traits; Rogers et al., 2018) predicted the romantic evaluation DV when used in place of the raw pattern metric. Results showed that this measure performed similarly: Larger Euclidean distances negatively predicted positive evaluations in the full sample ( $\beta_I = -.31, p < .001$ ), and for both partnered ( $\beta_I = -.31, p < .001$ ) and single ( $\beta_I = -.29, p < .001$ ) participants.

Table 2 –Ideal-Trait Correlations (Analysis Plan 1a through 4a)

	Attribute	Ideal-Trait Correlations				
		Overall	Partnered	Single	t for comparison	
1	Attractive	(V/A)	.29***	.28***	.31***	2.44
2	Intelligent		.35***	.38***	.31***	-4.69***
3	Humorous		.39***	.40***	.37***	-3.96***
4	Considerate	(W/T)	.31***	.30***	.30***	-1.04
5	Honest		.30***	.33***	.25***	-5.19***
6	Understanding	(W/T)	.31***	.31***	.29***	-1.19
7	Ambitious		.41***	.45***	.38***	-6.19***
8	Sporty and Athletic		.39***	.41***	.36***	-3.24**
9	Fun		.37***	.35***	.39***	-0.63
10	Sensitive	(W/T)	.36***	.36***	.36***	-1.79
11	A good lover	(V/A)	.34***	.33***	.32***	-4.24***
12	Nice body	(V/A)	.29***	.27***	.32***	2.69
13	Confident		.34***	.34***	.32***	-2.85

14	Sexy	(V/A)	.36***	.34***	.41***	1.91
15	Financially secure	(S/R)	.24***	.25***	.25***	-0.62
16	Supportive	(W/T)	.31***	.31***	.28***	-2.22
17	Dresses well	(S/R)	.33***	.34***	.31***	-2.78
18	A good listener	(W/T)	.28***	.26***	.29***	-2.50
19	Loyal		.27***	.33***	.20***	-8.43***
20	Successful	(S/R)	.29***	.30***	.28***	-3.61***
21	Adventurous	(V/A)	.38***	.39***	.38***	-4.09***
22	Good job	(S/R)	.28***	.30***	.27***	-2.22
23	Religious		.57***	.63***	.53***	-8.97***
24	Patient		.26***	.28***	.26***	-1.94
25	Extraverted, enthusiastic	(Ext)	.37***	.41***	.34***	-4.52***
26	Critical, quarrelsome	(Agr)	.39***	.39***	.42***	-0.10
27	Dependable, self-disciplined	(Con)	.31***	.33***	.29***	-3.61***
28	Anxious, easily upset	(Emo)	.27***	.28***	.27***	-0.83
29	Open to new experiences, complex	(Opn)	.36***	.37***	.34***	-3.15**
30	Reserved, quiet	(Ext)	.35***	.39***	.31***	-4.31***
31	Sympathetic, warm	(Agr)	.32***	.32***	.32***	-1.51
32	Disorganized, careless	(Con)	.25***	.24***	.26***	-0.29
33	Calm, emotionally stable	(Emo)	.27***	.29***	.25***	-3.82***
34	Conventional, uncreative	(Opn)	.34***	.35***	.32***	-3.09**
35	Smells good		.38***	.34***	.42***	-0.32
	W/T average		.41***	.40***	.39***	-2.95**
	V/A average		.40***	.37***	.43***	-0.74
	S/R average		.34***	.34***	.34***	-1.95
	Ext average		.36***	.42***	.31***	-5.76***
	Agr average		.37***	.36***	.38***	0.68
	Con average		.29***	.29***	.29***	-1.94
	Emo average		.27***	.27***	.26***	-1.79
	Opn average		.36***	.37***	.36***	-3.53***

Note: W/T: warmth/trustworthiness; V/A: vitality/attractiveness; S/R: status/resources; Ext: Extraversion; Agr: Agreeableness; Con: Conscientiousness; Emo: Emotional Stability; Opn: Openness to Experience. In the Big Five averages, Items 26, 28, 30, 32, and 34 were reverse scored. Values are the regression estimated betas ( $\beta_1$ 's) from the following equation: Partner attribute =  $\beta_0 + \beta_1$ Ideal +  $u_0 + u_1$ Ideal +  $\varepsilon$ . The random slope ( $u_1$ ) for sample is omitted when models do not converge.  $t$  for comparison refers to the  $\beta_3$  estimate in the following model, which tests the difference between the partnered and single columns: Partner attribute =  $\beta_0 + \beta_1$ Ideal +  $\beta_2$ RelStatus +  $\beta_3$ Ideal×RelStatus +  $u_0 + u_1$ Ideal +  $\varepsilon$  \*\*  $p < .01$ , \*\*\*  $p < .001$ . Asterisks are omitted for estimates that fail a Holm-Bonferroni test (Holm, 1979) within each column of 35 traits.

Table 3 – Effect Sizes for Tests of Ideal Partner Preference-Matching (Analysis Plan 2b-4b, 2c-4c, 2d-4d)

Analysis		Overall	Partnered	Single	t for comparison
<b>Pattern metric</b>					
Raw		.37***	.38***	.32***	3.06**
Corrected		.19***	.17***	.19***	3.27**
<b>Level Metric</b>					
1 Attractive	(V/A)	.02**	.00	.05***	3.14**
2 Intelligent		.03***	.00	.03	3.46***
3 Humorous		.04***	.01	.06***	4.30***
4 Considerate	(W/T)	.00	-.04***	.04***	5.79***
5 Honest		.02	-.01	.02	2.72
6 Understanding	(W/T)	.02	-.01	.04***	4.72***
7 Ambitious		.07***	.05***	.08***	3.69***
8 Sporty and Athletic		.07***	.06***	.08***	2.41
9 Fun		.02	-.03**	.05***	6.45***
10 Sensitive	(W/T)	.06***	.07***	.06***	0.80
11 A good lover	(V/A)	.04***	.02	.06***	1.72
12 Nice body	(V/A)	.02	.01	.06***	3.66***
13 Confident		.04***	.01	.04***	3.70***
14 Sexy	(V/A)	.02**	.02	.04***	2.74
15 Financially secure	(S/R)	.04***	.04***	.06***	2.51
16 Supportive	(W/T)	.01	-.01	.02	3.59***
17 Dresses well	(S/R)	.03***	.03	.04***	2.14
18 A good listener	(W/T)	.01	-.02	.05***	5.62***
19 Loyal		.03***	.03**	.02	0.46
20 Successful	(S/R)	.05***	.03	.06***	3.97***
21 Adventurous	(V/A)	.05***	.07***	.07***	3.36***
22 Good job	(S/R)	.05***	.05***	.06***	2.16
23 Religious		.13***	.10***	.07***	-0.21
24 Patient		.01	-.02	.04	4.01***
25 Extraverted, enthusiastic	(Ext)	.07***	.09***	.03	-1.63
26 Critical, quarrelsome	(Agr)	.08***	.10***	.08***	1.36
27 Dependable, self-disciplined	(Con)	.03***	-.01	.06***	5.27***
28 Anxious, easily upset	(Emo)	.07***	.05***	.08***	3.25**
29 Open to new experiences, complex	(Opn)	.05***	.05***	.06***	3.78***
30 Reserved, quiet	(Ext)	.09***	.09***	.07***	0.28
31 Sympathetic, warm	(Agr)	.02	-.01	.04	4.38***
32 Disorganized, careless	(Con)	.04***	.05***	.05***	1.73
33 Calm, emotionally stable	(Emo)	.03***	.02	.04***	2.10
34 Conventional, uncreative	(Opn)	.07***	.09***	.05***	-0.04
35 Smells good		.01	.03	.02	1.06
W/T average		.00	-.03***	.02	4.50***
V/A average		.01	-.02*	.05***	5.30***
S/R average		.03***	.03**	.07***	4.30***
Ext average		.07***	.08***	.04**	-1.43

Agr average	.03**	.04***	.05***	1.22
Con average	.03**	.01	.06***	4.73***
Emo average	.05***	.01	.07***	3.82***
Opn average	.05***	.05***	.05**	2.41*

Note: W/T: warmth/trustworthiness; V/A: vitality/attractiveness; S/R: status/resources. Ext: Extraversion; Agr: Agreeableness; Con: Conscientiousness; Emo: Emotional Stability; Opn: Openness to Experience. In the Big Five averages, Items 26, 28, 30, 32, and 34 were reverse scored. Values for pattern metric (raw) and pattern metric (corrected) are the regression estimated beta ( $\beta_1$ ) from the following equation: Romantic evaluation =  $\beta_0 + \beta_1$ PatternMetric +  $u_0 + u_1$ PatternMetric +  $\varepsilon$ . Values for the level metric are the ideal  $\times$  trait interaction estimated beta's ( $\beta_3$ 's) from the following equation: Romantic evaluation =  $\beta_0 + \beta_1$ Ideal +  $\beta_2$ PartnerAttribute +  $\beta_3$ Ideal $\times$ PartnerAttribute +  $u_0 + u_1$ PartnerAttribute +  $\varepsilon$ . In all cases, the random slope ( $u_1$ ) for sample is omitted when models do not converge. “*t* for comparison” for the pattern metric tests refers to the  $\beta_3$  estimate in the following model: Romantic evaluation =  $\beta_0 + \beta_1$ PatternMetric +  $\beta_2$ RelStatus +  $\beta_3$ PatternMetric $\times$ RelStatus +  $u_0 + u_1$ PatternMetric +  $\varepsilon$ . “*t* for comparison” for the level metric tests refers to the  $\beta_7$  estimate in the following model: Romantic evaluation =  $\beta_0 + \beta_1$ Ideal +  $\beta_2$ PartnerAttribute +  $\beta_3$ Ideal $\times$ PartnerAttribute +  $\beta_4$ RelStatus +  $\beta_5$ Ideal $\times$ RelStatus +  $\beta_6$ PartnerAttribute $\times$ RelStatus +  $\beta_7$ Ideal $\times$ PartnerAttribute $\times$ RelStatus +  $u_0 + u_1$ PartnerAttribute +  $\varepsilon$ . \*  $p < .05$ , \*\*  $p < .01$ , \*\*\*  $p < .001$ . Asterisks are omitted for estimates that fail a Holm-Bonferroni test (Holm, 1979) within each column of 35 traits.

**Strong inference tests.** The corrected pattern metric successfully predicted the romantic evaluation ( $\beta_1 = .19$  in the full sample, see Table 3). In other words, a pure measure of preference-matching across 35 different traits predicted the evaluative dependent measure with a small-to-medium effect size. The association was actually larger in the single ( $\beta_1 = .19$ ) than the partnered ( $\beta_1 = .17$ ) subsample, but the difference was quite small.<sup>4</sup>

The level metric results were more modest, although many were significantly different from zero (Table 3). As with the corrected pattern metric, these effects tended to be larger for single than partnered participants, although again, such differences were very small. Intriguingly, level metric interaction effects tended to be larger for traits that are not as commonly assessed in this research space, like religiosity and extraversion. The level metric interaction effects were

<sup>4</sup> Some perspectives (e.g., Biesanz, 2010; Fletcher et al., 2020; Wood et al., 2019) add a measure of normative matching alongside “distinctiveness” metrics like these. Using this approach, effect sizes are about half as large as those reported here, but still significant; see the “Normative Preference-Matching” section below.

quite small for traits that are normatively very desirable and commonly studied, like warmth/trustworthiness and vitality/attractiveness traits.

Overall, the level metric effect sizes illustrate why such interactions have been hard to detect in prior studies: The average interaction  $\beta_3 = .04$  is a 15% attenuation interaction given the average  $\beta_2 = .27$ . To put the effect size challenges in context, we used the Shiny App InteractionPoweR (Baranger et al., 2023; Finsaas et al., 2021) and the average values across all the 35 level metric tests:  $\beta_1 = .02$ ,  $\beta_2 = .27$ ,  $\beta_3 = .04$  (see equation in note of Table 3), and the average ideal-trait correlation  $\beta = .33$  from Table 2. Using these values, achieving 80% power to detect an interaction effect of  $\beta_3 = .04$  would require  $N = 4,475$  participants.<sup>5</sup> (The largest level metric effect—religiosity—would still require  $N = 470$  to achieve 80% power.) In summary, the current data suggest that level metric effects do exist, but such interactions will require substantial, if not enormous, resources to detect.

**Level of education level metric analysis.** It was also possible to test the level metric interaction for level of education using the same multilevel analyses described in the Table 3 note (Romantic evaluation =  $\beta_0 + \beta_1$ Ideal +  $\beta_2$ PartnerAttribute +  $\beta_3$ Ideal×PartnerAttribute +  $u_0 + u_1$ PartnerAttribute +  $\varepsilon$ ) using the ideal “level of education” item and the partner’s actual level of education (coded on a 6-point continuous scale). We calculated this estimate for the overall sample, single participants, and partnered participants, and we also tested the difference between single and partnered participants. For the overall sample, this interaction was  $\beta_3 = .06$ ,  $t(1991.84)$

---

<sup>5</sup> The power to detect a standardized interaction effect  $\beta$  is very close to the power to detect a correlation of size  $\beta$ , with two caveats: (a) larger main effects of the two interacting variables (in this case, ideals and attribute perceptions) increase power, and (b) a larger correlation between the two main effects can increase or decrease power, depending on the size of the main effects (Baranger et al., 2023). These mitigating forces are not especially large in these analyses, and so the  $N$  required to achieve 80% power to detect  $\beta_3 = .04$  (4,475) is only slightly smaller than the  $N$  required to achieve 80% power to detect  $r = .04$  (4,900).

$= 6.33, p < .001$ . For single participants, this interaction was  $\beta_3 = .03, t(1421.33) = 1.95, p = .051$ ; for partnered participants, this interaction was  $\beta_3 = .03, t(5522.28) = 2.70, p = .007$ ; and the difference between single and partnered participants was not significant,  $t(9020.83) = 0.50, p = .618$ .

**Relationship formation hypothesis.** This hypothesis pertains to the separate sample of CloudResearch participants who completed the surveys at two time points, about 3.5 months apart. We conducted the raw pattern metric, corrected pattern metric, and level metric analyses on these three samples (see Table 4). Some of the findings echoed the Table 3 results for the full international sample. For example, for both the steadily partnered and the newly partnered sample, the raw pattern metric was considerably larger than the corrected pattern metric (especially in the steadily partnered sample), but the corrected pattern metric was still significant and of a meaningful effect size ( $\beta = .24$ ). Estimates for the steadily partnered and newly partnered sample were similar, suggesting that both sets of participants maintained their ideals over the intervening months and drew from them when evaluating their partners, regardless of whether or not they were dating that partner when they reported their ideals at time 1. Intriguingly, for the single participants, the corrected pattern metric was essentially zero: Unlike the participants in the international single sample in Table 3, ideal partner preference-matching seemed to have no bearing on the evaluations of these single participants—a finding we revisit in the Discussion. Once again, level metric findings were erratic and small on average (the smaller sample size here yielded a larger range of negative and positive values, relative to Table 3); preferences for religiosity and extraversion perhaps deserve additional study going forward nonetheless.

With respect to level of education: For the steadily partnered sample, the level metric interaction was  $\beta_3 = .05$ ,  $t(705) = 1.46$ ,  $p = .144$ ; for steadily single participants, this interaction was  $\beta_3 = .03$ ,  $t(683) = 0.77$ ,  $p = .444$ ; for newly partnered participants, this interaction was  $\beta_3 = .05$ ,  $t(185) = 0.71$ ,  $p = .480$ . The difference between these three samples was not significant,  $F(2,1573) = 0.12$ ,  $p = .890$ .

Table 4 – Relationship Formation Hypothesis

Analysis		Steadily Partnered	Steadily Single	Newly Partnered	F for comparison
<b>Pattern Metric</b>					
Raw		.50***	.18***	.39***	20.86***
Corrected		.24***	.01	.24***	10.78***
<b>Level Metric</b>					
1 Attractive	(V/A)	.00	.02	.07	0.56
2 Intelligent		-.05	-.03	.08	1.93
3 Humorous		.03	.02	.14***	3.44
4 Considerate	(W/T)	.02	-.03	-.09	1.28
5 Honest		-.06	-.02	.08	1.75
6 Understanding	(W/T)	.00	-.02	-.07	0.43
7 Ambitious		.12***	.02	.16	3.68
8 Sporty and Athletic		.06	.05	.10	0.28
9 Fun		-.03	-.01	.08	0.95
10 Sensitive	(W/T)	.03	-.01	.06	0.53
11 A good lover	(V/A)	-.07	-.03	.07	2.43
12 Nice body	(V/A)	-.04	.01	.06	1.12
13 Confident		.03	.07	.15	1.16
14 Sexy	(V/A)	.04	-.01	.01	0.67
15 Financially secure	(S/R)	.10	-.02	.19	5.33
16 Supportive	(W/T)	.06	-.04	.04	2.80
17 Dresses well	(S/R)	.03	.02	.26***	6.54
18 A good listener	(W/T)	-.01	-.08	.02	1.77
19 Loyal		-.01	-.05	.01	0.72
20 Successful	(S/R)	.03	-.03	.18	3.62
21 Adventurous	(V/A)	.10	.02	.21**	3.87
22 Good job	(S/R)	.14***	.02	.15	4.03
23 Religious		.24***	-.01	.37***	13.51***
24 Patient		.09	.01	-.07	2.96
25 Extraverted, enthusiastic	(Ext)	.09	-.01	.08	2.14
26 Critical, quarrelsome	(Agr)	.15***	.05	.10	2.25
27 Dependable, self-disciplined	(Con)	.05	-.03	.19**	4.94
28 Anxious, easily upset	(Emo)	.06	.02	.05	0.36
29 Open to new experiences, complex	(Opn)	.01	.09	.13	2.18

30	Reserved, quiet	(Ext)	.14***	.02	.16	4.12
31	Sympathetic, warm	(Agr)	.07	.04	-.06	1.69
32	Disorganized, careless	(Con)	.01	-.01	.02	0.16
33	Calm, emotionally stable	(Emo)	.10	.02	.07	1.34
34	Conventional, uncreative	(Opn)	.03	.04	-.04	0.66
35	Smells good		.03	-.04	.10	2.50
	W/T average		-.04	-.05	-.05	0.06
	V/A average		-.06	.00	.04	1.55
	S/R average		.07*	.00	.23***	4.54*
	Ext average		.13***	.03	.07	2.42
	Agr average		.10**	.08*	-.03	1.41
	Con average		.05	-.04	.11	2.58
	Emo average		.08*	.00	.07	1.36
	Opn average		-.03	.07	.00	2.24

Note: W/T: warmth/trustworthiness; V/A: vitality/attractiveness; S/R: status/resources. Ext: Extraversion; Agr: Agreeableness; Con: Conscientiousness; Emo: Emotional Stability; Opn: Openness to Experience. In the Big Five averages, Items 26, 28, 30, 32, and 34 were reverse scored. Note that in these analyses, there is no within-sample dependency. Values for pattern metric (raw) and pattern metric (corrected) are the regression estimated betas ( $\beta_1$ ) from the following equation: Romantic evaluation =  $\beta_0 + \beta_1$ PatternMetric +  $\varepsilon$ . Values for the level metric are the ideal  $\times$  trait interaction estimated beta's ( $\beta_3$ 's) from the following equation: Romantic evaluation =  $\beta_0 + \beta_1$ Ideal +  $\beta_2$ PartnerAttribute +  $\beta_3$ Ideal $\times$ PartnerAttribute +  $\varepsilon$ . RelStatus is a 3-level categorical variable, so "F for comparison" for the pattern metric tests refers to the omnibus test of the two  $\beta_3$  estimates in the following model: Romantic evaluation =  $\beta_0 + \beta_1$ PatternMetric +  $\beta_2$ RelStatus +  $\beta_3$ PatternMetric $\times$ RelStatus +  $\varepsilon$ . "F for comparison" for the level metric tests refers to the omnibus test of the two  $\beta_7$  estimates in the following model: Romantic evaluation =  $\beta_0 + \beta_1$ Ideal +  $\beta_2$ PartnerAttribute +  $\beta_3$ Ideal $\times$ PartnerAttribute +  $\beta_4$ RelStatus +  $\beta_5$ Ideal $\times$ RelStatus +  $\beta_6$ PartnerAttribute $\times$ RelStatus +  $\beta_7$ Ideal $\times$ PartnerAttribute $\times$ RelStatus +  $\varepsilon$ . \*  $p < .05$ , \*\*  $p < .01$ , \*\*\*  $p < .001$ . Asterisks are omitted for estimates that fail a Holm-Bonferroni test (Holm, 1979) within each column of 35 traits.

## Exploratory Descriptive Analyses

Table 5 presents descriptive analyses of the average preferences of participants in the dataset, both stated (i.e., rated ideals) and revealed (i.e., the association between the attribute and the evaluative dependent measure; Wood & Brumbaugh, 2009). Colloquially speaking, the ideal partner preference ratings (i.e., the means for each attribute) capture the extent to which people generally say that each attribute is important in an ideal partner, whereas the revealed preferences (i.e., the slopes for each attribute predicting the dependent variable) capture the

extent to which each attribute *actually predicts* people's romantic evaluations of partners.<sup>6</sup> This table also includes the rank ordering of both sets of 35 preferences.

On the whole, stated and revealed preferences aligned in terms of ranking, although some intriguing differences did emerge. For example, the attributes "confident," "a good listener," "patient," and "calm, emotionally stable" ranked considerably more highly as stated preferences than as revealed preferences. In contrast, the attributes "attractive," "a good lover," "nice body," "sexy," and "smells good" ranked considerably more highly as revealed preferences than as stated preferences. In fact, "a good lover" was the #1 largest revealed preference but actually ranked 12<sup>th</sup> in terms of stated preferences. (We also conducted separate analyses on the partnered and single subsamples, revealing identical conclusions; see Tables S10 and S11 in the Supplemental Materials.)

Table 5 also calculates gender differences in the preference for attractiveness (i.e., the average of the items "attractive," "nice body," and "sexy") and earning potential (i.e., the average of the items "ambitious," "financially secure," and "good job"). Some theoretical perspectives anticipate that men will place greater weight on attractiveness, and women will place greater weight on earning potential (Buss, 1989). These gender differences indeed emerged when participants reported their stated preferences. Nevertheless, consistent with past meta-analytic work (Eastwick et al., 2014) and the very small level metric analyses documented in Table 3, these gender differences did not emerge in participants' revealed preferences.

We can also use the Table 5 ranking approach to illuminate *why* a gender difference incongruity emerges between stated and revealed preferences. Men's stated preferences tended to

---

<sup>6</sup> This analysis applies at the level of the entire dataset on the primary target only; we calculate a related form of revealed preference (which we call a "functional preference"; Ledgerwood et al., 2018) that makes use of all four targets in a later section.

underestimate the value they actually placed on “attractive,” “nice body,” and “sexy” by about 6 ranks (out of 35; 1 = highest ranked, 35 = lowest ranked) on average (see Table S12 in the Supplemental Materials for details). That is, their stated preferences for these three traits ranked 9, 18, and 17 (respectively) but their revealed preferences for these three traits ranked 7, 13, and 6. However, women underestimated the value they placed on these three traits by a full 13 ranks (out of 35): Their stated preferences for these three traits ranked 18, 28, and 23 (respectively) but their revealed preferences for these three traits ranked 8, 17, and 5 (i.e., about the same as men). As for “ambitious,” “financially secure,” and “good job,” men’s stated preferences underestimated their value by about 4 ranks: Their stated preferences for these three traits ranked 25, 25 (tied), and 27 (respectively) but their revealed preferences for these three traits ranked 22, 24, and 20. In contrast, women’s stated preferences *overestimated* their value by about 4 ranks: Their stated preferences for these three traits ranked 22, 17, and 18 (respectively) but their revealed preferences for these three traits ranked 24, 25, and 21 (i.e., again, about the same as men). In summary, both men’s and women’s stated preferences appeared to underestimate the weight they place on attractiveness, but this underestimation effect was more pronounced for women than for men. In contrast, men’s stated preferences slightly underestimated the weight they placed on earning potential, and women’s stated preferences slightly overestimated the weight they placed on earning potential.

Table 5 – Descriptive Statistics for Stated and Revealed Preferences

Attribute	N	Stated Preferences			Revealed Preferences	
		M	SD	Rank	$\beta$	Rank
Attractive	10343	8.86	1.89	16	.42***	8
Intelligent	10348	9.39	1.65	9	.38***	12
Humorous	10345	9.34	1.78	11	.36***	13
Considerate	10343	9.59	1.60	7	.40***	10
Honest	10347	10.08	1.38	2	.43***	5
Understanding	10346	9.84	1.46	4	.42***	7
Ambitious	10344	8.13	2.34	24	.22***	24
Sporty and Athletic	10347	7.16	2.43	29	.10***	29
Fun	10351	9.43	1.66	8	.38***	11
Sensitive	10340	8.10	2.35	25	.28***	19
A good lover	10338	9.26	1.99	12	.56***	1
Nice body	10348	8.02	2.15	26	.32***	16
Confident	10347	8.77	1.87	17	.18***	26
Sexy	10342	8.39	2.20	19	.42***	6
Financially secure	10342	8.38	2.19	20	.20***	25
Supportive	10346	9.92	1.52	3	.49***	3
Dresses well	10344	8.19	2.18	23	.24***	22
A good listener	10346	9.69	1.60	5	.35***	14
Loyal	10345	10.10	1.53	1	.51***	2
Successful	10344	8.22	2.22	22	.29***	17
Adventurous	10338	7.89	2.36	27	.16***	27
Good job	10342	8.30	2.17	21	.24***	21
Religious	10340	4.83	3.33	31	.04**	31
Patient	10342	9.35	1.70	10	.29***	18
Extraverted, enthusiastic	10343	7.70	2.19	28	.13***	28
Critical, quarrelsome	10339	3.41	2.65	33	-.04**	33
Dependable, self-disciplined	10348	9.26	1.81	12	.33***	15
Anxious, easily upset	10341	3.10	2.24	34	.02	32
Open to new experiences, complex	10346	8.64	2.10	18	.23***	23
Reserved, quiet	10338	5.53	2.66	30	.07***	30
Sympathetic, warm	10345	9.61	1.59	6	.40***	9
Disorganized, careless	10340	2.80	2.17	35	-.05***	34
Calm, emotionally stable	10344	9.26	1.75	12	.26***	20
Conventional, uncreative	10343	4.02	2.56	32	-.07***	35
Smells good	10346	9.10	1.98	15	.45***	4
W/T average	10356	9.43	1.29		.48***	
V/A average	10353	8.48	1.56		.50***	
S/R average	10353	8.27	1.81		.31***	
Ext average	10344	7.08	1.78		.04**	
Agr average	10348	9.10	1.61		.25***	
Con average	10348	9.23	1.56		.21***	
Emo average	10347	9.08	1.56		.14***	
Opn average	10350	8.31	1.75		.19***	

Attribute	N	Stated Preferences		Revealed Preferences				
		M	SD	$\beta$	Gender Diff.			
Attractiveness Composite					Gender Diff.			
Heterosexual Men	2935	8.73	1.70	$t = 13.10^{***}$	$d = 0.22$	.46***	0.19	0.02
Heterosexual Women	5408	8.35	1.80			.45***		
Earning Potential Composite								
Heterosexual Men	2933	7.50	1.85	$t = 27.51^{***}$	$d = 0.71$	.27***	0.78	0.00
Heterosexual Women	5410	8.74	1.63			.28***		

Note: Effect sizes  $d$  and  $q$  are coded such that positive effect sizes are in the predicted direction. Gender differences were only calculated for participants who identified as a man or a woman and who selected the option “straight/heterosexual” for their sexuality. Stated preferences are means. Revealed preferences are  $\beta_1$  terms in the equation: Romantic evaluation =  $\beta_0 + \beta_1$ PartnerAttribute +  $u_0 + u_1$ PartnerAttribute +  $\varepsilon$ . In all cases, the random slope ( $u_1$ ) for sample is omitted when models do not converge.

## Secondary Planned Analyses

**Normative preference-matching.** A difference between the effect sizes associated with the raw pattern metric and the corrected pattern metric implies—but does not directly test—the idea that participants positively evaluate partners to the extent that they perceive those partners to have consensually desirable traits (Fletcher et al., 2020). The direct test of this idea entails calculating a normative pattern metric: the association between (a) a Fisher  $z$ -scored version of the within-person correlation between the *sample average* of the 35 ideal ratings (not the participant’s own rating) and (the participants’ own ratings of) the 35 partner-attribute ratings, and (b) the romantic evaluation measure.

Using the multilevel analyses described in the Table 3 note, we calculated this estimate for the overall sample, single participants, and partnered participants, and we also tested the difference between single and partnered participants. For the overall sample, this effect was  $\beta_1 = .37$ ,  $t(39.35) = 29.38$ ,  $p < .001$ . For single participants, this effect was  $\beta_1 = .32$ ,  $t(33.68) = 18.14$ ,  $p < .001$ ; for partnered participants, this effect was  $\beta_1 = .39$ ,  $t(56.99) = 19.84$ ,  $p < .001$ ; and the difference between single and partnered participants was significant,  $t(6765.48) = 2.17$ ,  $p = .030$ .

These effect sizes suggest that, when participants perceived that partners had normatively “ideal” traits, they evaluated those partners very positively, regardless of their own idiosyncratic ideal partner preferences.

In some research areas that examine analogous forms of multivariate matching (e.g., Biesanz, 2010; Fletcher et al., 2020; Wood et al., 2019), it is common practice to predict a dependent measure from both the normative and distinctive metrics simultaneously. Similarly, we can predict the romantic evaluation DV using the following equation:

$$\text{Romantic evaluation} = \beta_0 + \beta_1 \text{NormativePatternMetric} + \beta_2 \text{CorrectedPatternMetric} + u_0 + u_1 \text{NormativePatternMetric} + u_2 \text{CorrectedPatternMetric} + \varepsilon \quad (\text{Eq. 1})$$

Using this approach, the normative preference-matching effects closely approximate the effect sizes when included in the equation alone: in the full sample,  $\beta_1 = .34$ ,  $t(37.85) = 26.59$ ,  $p < .001$ ; in the single subsample,  $\beta_1 = .29$ ,  $t(32.15) = 15.76$ ,  $p < .001$ ; in the partnered subsample  $\beta_1 = .37$ ,  $t(59.33) = 20.21$ ,  $p < .001$ . However, the corrected pattern metric effect sizes were approximately half the size of what they were when included in the equation alone: in the full sample,  $\beta_2 = .11$ ,  $t(31.42) = 10.71$ ,  $p < .001$ ; in the single subsample,  $\beta_2 = .13$ ,  $t(41.70) = 6.69$ ,  $p < .001$ ; in the partnered subsample,  $\beta_2 = .09$ ,  $t(47.09) = 5.66$ ,  $p < .001$ . In other words, idiosyncratic preference-matching offers a small ( $\beta = .09\text{--}.13$ ), yet significant, boost above and beyond normative preference-matching, and normative preference-matching is approximately 3 times as large.

**Individual difference moderation.** It is plausible that ideal partner preference-matching effects vary across studies in the existing literature due to individual differences across participant populations. A study by Lam et al. (2016) points to the intriguing possibility that there are important cross-cultural factors at play. In this reasonably large ( $N = 472$ ) study, these scholars found that the corrected pattern metric had a significant predictive association with relationship evaluations in Taiwan ( $r = .22$ ) but not in the U.S. ( $r = .05$ ), and the difference

between these two correlations was significant. Reasons for a Taiwan-U.S. difference remain somewhat speculative, but one relevant distinction between these two cultures is relational mobility—that is, the ability to meet new people and select into (and out of) relationships on the basis of personal desires (Kito et al., 2017; Thomson et al., 2018; Yuki & Schug, 2012). Americans, by virtue of their higher relational mobility, might be more likely than Taiwanese to “try out” relationships that mismatch their ideals, perhaps especially if they presume that they could later end the relationship with minimal consequences. Then, if people are motivated on average to feel positively about their partners after investing time and energy into the relationship (Joel & MacDonald, 2021), high relational mobility populations may include a larger proportion of people with ideal-mismatching partners who nevertheless report high satisfaction. A second potentially relevant distinction is individualism-collectivism (Triandis & Gelfand, 1998), as individuals in collectivistic cultures may be especially likely to adopt the ideal partner preferences of their parents (Locke et al., 2020). If the attributes of one’s romantic partner implicate family members in collectivistic societies, this fact may motivate collectivistic (but not individualistic) individuals to remain attuned to the extent to which the partner mismatches their ideals.

To test whether relational mobility (i.e., the average of the 12 items; Thomson et al., 2018), individualism (i.e., either the 4-item horizontal individualism or 4-item vertical individualism subscales; Triandis & Gelfand, 1998), or collectivism (i.e., either the 4-item horizontal collectivism or 4-item vertical collectivism subscales; Triandis & Gelfand, 1988) affect ideal partner preference-matching, we examined whether these five individual difference measures moderated all the analyses reported in Table 3 that pertained to research questions 1-3 (i.e., effect sizes associated with the overall sample, single participants and partnered

participants). Again, we used Bonferroni-Holm correlations for each set of 35 tests. Table 6 uses “+” signs to indicate positive, significant interaction terms (i.e., ideal preference-matching is stronger among participants who are *higher* in relational mobility/individualism/collectivism), and “-” signs to indicate negative significant interaction terms (i.e., ideal preference-matching is stronger among participants who are *lower* in relational mobility/individualism/collectivism).

The predicted direction of moderation is depicted in a row at the top of Table 6. Reliabilities for relational mobility were  $\alpha = .82$  ( $\omega = .81$ ) on the full sample,  $\alpha = .82$  ( $\omega = .81$ ) on the partnered sample, and  $\alpha = .82$  ( $\omega = .80$ ) on the single sample; reliabilities for horizontal individualism were  $\alpha = .71$  ( $\omega = .72$ ) on the full sample,  $\alpha = .69$  ( $\omega = .70$ ) on the partnered sample, and  $\alpha = .73$  ( $\omega = .74$ ) on the single sample; reliabilities for vertical individualism were  $\alpha = .67$  ( $\omega = .68$ ) on the full sample,  $\alpha = .67$  ( $\omega = .67$ ) on the partnered sample, and  $\alpha = .69$  ( $\omega = .69$ ) on the single sample; reliabilities for horizontal collectivism were  $\alpha = .74$  ( $\omega = .74$ ) on the full sample,  $\alpha = .74$  ( $\omega = .74$ ) on the partnered sample, and  $\alpha = .73$  ( $\omega = .73$ ) on the single sample; reliabilities for vertical collectivism were  $\alpha = .69$  ( $\omega = .72$ ) on the full sample,  $\alpha = .68$  ( $\omega = .71$ ) on the partnered sample, and  $\alpha = .69$  ( $\omega = .72$ ) on the single sample.

Very few of these interactions were statistically significant. And, crucially, in the full table, 21 interactions were in the predicted direction of moderation (i.e., no shading in Table 6), and 23 interactions were in the opposite of the predicted direction (i.e., grey shading). For example, when interactions emerged for the corrected pattern metric, they tended to be positive interactions (8 out of 9 times), regardless of whether the prior literature anticipated that these interactions would be negative (relational mobility, individualism) or positive (collectivism). Given the ambiguity of these results and related concerns about moderation with measured variables (Rohrer et al., 2022), we hesitate before interpreting them any more deeply.

Table 6 – Secondary Planned Analyses

Analysis	Overall			Partnered			Single			Functional Prefs	
	R	Ind	Col	R	Ind	Col	R	Ind	Col	H	V
	H	V	H	V	H	V	H	V	H	V	
<b>Predicted Direction of Moderation</b>											
<b>Pattern Metric</b>											
Raw				-	-	-	-	-	-	-	2.61%
Corrected	+	+		+			+	+	+	+	6.87%
<b>Level Metric</b>											
1 Attractive			(V/A)						+		1.21%
2 Intelligent											0.53%
3 Humorous											2.02%
4 Considerate			(W/T)				+				1.47%
5 Honest							+				0.98%
6 Understanding			(W/T)								2.04%
7 Ambitious											4.87%
8 Sporty and Athletic											4.53%
9 Fun							+				1.68%
10 Sensitive			(W/T)								4.15%
11 A good lover			(V/A)								4.12%
12 Nice body			(V/A)				+				3.82%
13 Confident							+	+			3.51%
14 Sexy			(V/A)		+				+		3.68%
15 Financially secure			(S/R)								4.11%
16 Supportive			(W/T)								1.34%
17 Dresses well			(S/R)								3.32%
18 A good listener			(W/T)								2.29%
19 Loyal											1.49%
20 Successful			(S/R)								2.71%
21 Adventurous			(V/A)								4.75%
22 Good job			(S/R)								4.13%
23 Religious							-	-	-		4.77%
24 Patient											4.59%
25 Extraverted, enthusiastic			(Ext)				-				4.98%
26 Critical, quarrelsome			(Agr)								3.60%
27 Dependable, self-disciplined			(Con)								1.67%
28 Anxious, easily upset			(Emo)								4.94%
29 Open to new experiences, complex			(Opn)								3.96%
30 Reserved, quiet			(Ext)								3.51%
31 Sympathetic, warm			(Agr)								2.36%
32 Disorganized, careless			(Con)								3.75%
33 Calm, emotionally stable			(Emo)								3.11%
34 Conventional, uncreative			(Opn)								2.97%
35 Smells good											3.02%
W/T average				-			-				2.13%

V/A average	+	+	+	+	+	2.01%
S/R average						3.36%
Ext average						2.91%
Agr average					+	2.10%
Con average		+				2.61%
Emo average	+		+			4.01%
Opn average				-		2.75%

Note: W/T: warmth/trustworthiness; V/A: vitality/attractiveness; S/R: status/resources. Ext: Extraversion; Agr: Agreeableness; Con: Conscientiousness; Emo: Emotional Stability; Opn: Openness to Experience; R: Relational mobility moderation; Ind: Individualism moderation (H = Horizontal, V = Vertical); Col: Collectivism moderation (H = Horizontal, V = Vertical). In the Big Five averages, Items 26, 28, 30, 32, and 34 were reverse scored. Individual-difference moderation values for pattern metric (raw) and pattern metric (corrected) derive from the interaction beta ( $\beta_3$ ) from the following equation: Romantic evaluation =  $\beta_0 + \beta_1$ PatternMetric +  $\beta_2$ IndividualDifference +  $\beta_3$  PatternMetric  $\times$  IndividualDifference +  $u_0 + u_1$ PatternMetric +  $\varepsilon$ . Values for the level metric derive from the interaction beta's ( $\beta_7$ 's) from the following equation: Romantic evaluation =  $\beta_0 + \beta_1$ Ideal +  $\beta_2$ PartnerAttribute +  $\beta_3$ Ideal  $\times$  PartnerAttribute +  $\beta_4$ IndividualDifference +  $\beta_5$ Ideal  $\times$  IndividualDifference +  $\beta_6$ PartnerAttribute  $\times$  IndividualDifference +  $\beta_7$ Ideal  $\times$  PartnerAttribute  $\times$  IndividualDifference +  $u_0 + u_1$ PartnerAttribute +  $\varepsilon$ . In all cases, the random slope ( $u_1$ ) for sample is omitted when models do not converge. The “+” indicates significant positive moderation; the “-” indicates significant negative moderation; the predicted pattern of moderation is depicted in the first five rows. The “+” and “-” signs were omitted for estimates that failed a Holm-Bonferroni test (Holm, 1979) within each column of 35 traits. Shaded “+” and “-” signs are in the opposite of the predicted direction. Functional preferences refer to the  $R^2_{t(v)}$  variance estimate from Rights & Sterba (2019) that captures the percentage of variance (out of 100%) accounted for by individual differences in the association of the attribute/pattern metric with the romantic evaluation dependent measure.

**Functional preferences.** Given that participants rated four total targets in the primary sample, it was possible to calculate each participant's *functional preference* for each attribute (Ledgerwood et al., 2018). A functional preference (also called a “driver of liking,” Lawless & Heymann, 2010) is the strength with which an attribute (e.g., attractiveness) predicts *a given person's romantic evaluations* across a series of targets—how much the attribute “matters” for a given participant. In this case, each participant's functional preference can be measured as the association of an attribute with the dependent measure across the four targets. Functional preferences in this context are very similar to the revealed preferences described above. The distinction is that a functional preference (typically) refers to a preference that has been

measured separately for each participant, and this requires that the participant rates multiple targets. The descriptive analyses in Table 5 only used the first (primary) target that participants evaluated.

A new approach by Rights and Sterba (2019) permits the calculation of the extent to which these functional preferences exhibit stable individual differences across targets. Specifically, the *R* package *r2mlm* (Shaw et al., 2020) provides the percentage of variance accounted for by the random effects component (i.e., “slope variation” or  $R^2_{t(v)}$ ) for a particular attribute as a fraction of the total variance.<sup>7</sup>

We calculated these values for all 35 attributes, the 3 Fletcher et al. (1999) constructs (both jointly and separately), the 5 Big Five traits, and the 2 pattern metric analyses (Table 6). The  $R^2_{t(v)}$  variance estimates for the 35 attributes, the 3 Fletcher et al. (1999) constructs, and the 5 Big Five constructs essentially denote the extent to which there are stable individual differences in the tendency for some people to exhibit stronger functional preferences than other people for a given attribute (Eastwick, Finkel, et al., 2023).

These values tended to be larger than zero, but they were fairly modest: The average of the 35 traits was  $R^2_{t(v)} = 3.1\%$ , and no trait exceeded 5%. In other words, individual differences in the way that participants weigh a given trait accounts for about 3% of the variance in romantic evaluations. We also calculated the  $R^2_{t(v)}$  variance estimates for the two pattern metric analyses; these values denote the extent to which there are stable individual differences in the tendency to desire a partner who matches (vs. mismatches) one’s ideals across all attributes. For example, the results for the corrected pattern metric indicated that individual differences in the way that

---

<sup>7</sup> Unlike the analyses above, this analysis ignores nesting within sample, and we conduct the analysis on a dataset that contains four rows per participant (one for each target). Now, the random slope effect captures variability across *participants* (not samples) in the extent to which the attribute predicts the dependent measure.

people weigh the match between ideals and traits across all traits accounts for about 7% of the variance in romantic evaluations.

## Discussion

### Central Takeaways

This is the first report from the Preference-Matching Project: the largest examination of ideal partner preference-matching to date ( $N = 10,358$  participants). In brief, ideal partner preference-matching predicted romantic evaluations—when collapsing across a large array of traits. That is, the effect size for the corrected pattern metric was modest but meaningful ( $\beta = .19$ ), and it did not differ appreciably between the partnered ( $\beta = .17$ ) vs. single ( $\beta = .19$ ) subsamples. Normative desirability proved to be an important consideration, too: Participants who perceived that partners matched the normative (i.e., sample-wide) ideal partner strongly desired those partners ( $\beta = .37$ ). When included together (as recommended by Beisanz, 2010; Fletcher et al., 2020), normative preference-matching remained strong ( $\beta = .34$ ), while the corrected pattern metric was cut to a small (but still significant) effect size ( $\beta = .11$ ). Approaches like the raw pattern metric and Euclidean distance revealed medium-to-large effects, likely because they blend normative and distinctive matching together into a single measurement mixture (Rogers et al., 2018; Wood & Furr, 2016).

The level metric (i.e., ideal  $\times$  trait interaction) tests were also highly informative. These effects were quite small on average ( $\beta = .04$ ), which may be why they have rarely been significant in prior studies ( $\beta = .04$  would require a sample size of  $N = 4,475$  to detect with 80% power; no prior study was even close to achieving such a large sample size, see Table S1). Also notable is that the level metric tests for the commonly studied, highly desirable attributes in this literature (e.g., traits in the warmth/trustworthiness and vitality/attractiveness categories) did not

even differ from zero in the full sample ( $\beta = .00-01$ ). Alternatively, traits that are rarely studied in this literature and that received moderate ideal ratings on average showed much larger level metric effects, like extraversion ( $\beta = .07$ ) and religious ( $\beta = .13$ ). It appears that the predictive validity of specific traits is more likely to be detectable for traits that land in a middling range of desirability (i.e., what could be called “horizontal” attributes; Hitsch et al., 2010), rather than traits that are highly normatively desirable (i.e., “vertical” attributes).

In two cases, expected moderation effects failed to emerge. First, we did not find much evidence that certain people or certain populations were especially likely to rely on their ideals. Our preregistered tests of potentially relevant individual differences (Table 6) revealed no interpretable pattern. Indeed, the multi-level modeling approach of Rights and Sterba (2019) suggested that the slope random effects corresponding to the sample ( $u_1$ ) in Table 3 explained less than 1% of the variance. In other words, the average association between an attribute and the romantic evaluation dependent measure tended not to vary reliably depending on which of the 60 samples generated it. Slightly larger (but still modest) amounts of variability emerged for (a) the tendency for some people to desire particular traits more than others across four different partners (3.1%; Table 6), (b) the tendency for mean levels of the traits to vary across samples (2.3%; cultures vary in the extent to which participants view partners as “humorous” or “ambitious”), and (c) the tendency for mean levels of the dependent measure to vary across samples (3.4%; cultures vary in the extent to which participants are happy in their relationships). These latter three types of effects might be more promising candidates for tests of moderation.

Second, for the most part, effect sizes in the partnered and single samples were similar. Many scholars (including several in the project coordinator group of the current project) once believed that ideal partner preference-matching was more likely to predict outcomes in

established relationships rather than initial attraction contexts (Eastwick et al., 2014). It is possible that the earlier literature suggested this pattern because studies of ongoing relationships classically gravitated toward the uncorrected pattern metric (which reveals medium-to-large effects; Fletcher et al., 1999, 2000), whereas initial attraction studies were inspired by perspectives on gender differences for specific traits in isolation (e.g., attractiveness, earning potential) and therefore tended to rely on level metric tests (which reveal very small effects; Eastwick & Finkel, 2008).

Nevertheless, one curious data point remains: Why did the single participants in the relationship formation subsample show no effects whatsoever? These participants first reported their ideals in isolation while they were single. Then, about 3.5 months later, these (still single) participants completed the rest of the procedure. The corrected pattern and level metric tests suggested that these participants were not drawing from their previously reported ideals at all (Table 4). And yet, this separation of 3.5 months seemed to matter very little for participants who were partnered at both time points, or participants who were single at time 1 and partnered at time 2. There are perhaps two ways of explaining these data. First, perhaps the people who were single at both time points had several rejection experiences in the interim, and their ideals changed more than the single participants who had the acceptance experience of becoming partnered during this time frame (Charlot et al., 2019). Second, perhaps single people who are very attracted to a particular partner are motivated to interpret the partner's traits in line with their ideals, but only if they have recently been reminded of their ideals. Researchers in this area should keep a keen eye on whether single participants are reporting their ideals and measures about a potential partner at the same or a different moment in time. (Most speed-dating studies, for example, ask participants to report their ideals on an intake form, and then participants

evaluate potential partners several days later.) This seemingly incidental methodological feature may matter a great deal for reasons that are not yet clear.

Finally, we presented a new approach that allows researchers to explore the distinction between stated preferences (i.e., preferences for traits as rated on scales) and revealed preferences (i.e., preferences as captured by the strength of the association between the trait and the DV). When the 35 attributes were ranked in the whole sample, it was possible to document cases where stated preference judgments (relatively) overestimated revealed preference judgments: Participants actually liked attributes like “confident,” “a good listener,” “patient,” and “calm, emotionally stable” less than they thought they did. In other cases, participants’ stated preferences were underestimates, as in the case of “attractive,” “a good lover,” “nice body,” “sexy,” and “smells good.” This approach was also able to illuminate why gender differences emerge for stated (but not revealed) preferences for attractiveness and earning potential attributes (Table 5). Specifically, for attractiveness, both men’s and women’s stated preferences underestimated their revealed preferences, but women’s tendency to underestimate proved far stronger than men’s. For earning potential, a “mirror image” pattern emerged such that men’s stated preferences underestimated their revealed preferences but women’s stated preferences overestimated their revealed preferences. Moving forward, this approach could be used to examine other research questions on accuracy and bias using various measures of preferences.

### **Strengths and Limitations**

This study has a number of strengths. Our partnership with the Psychological Science Accelerator (Moshontz et al., 2018) meant that the data were collected across 43 countries using a questionnaire that had been translated (and back-translated) into 22 different languages. Critically, our highly powered design meant that the estimates of effect sizes throughout this

paper are far more precise than is typical in most studies in this research area. Also, this paper was approved as a registered report, which meant that the design and analytic approach were reviewed before the data were collected.

This study also makes several important theoretical contributions. The pattern of effect sizes suggests that studies are far more likely to find empirical support to the extent that they focus on matching across many variables simultaneously rather than single attributes in isolation (e.g., gender differences in specific attributes; Eastwick & Finkel, 2008; a “top-3 most important” attributes approach, Sparks et al., 2020). Furthermore, the fact that effect sizes tended to be about three times larger for normative matching rather than the corrected pattern metric sheds new light on the intuitive idea that “people know what they want in a partner.” Yes, people’s stated preferences capture the attributes that are *generally* desirable in partners, but a given person’s *distinctive* preferences only modestly (but still significantly) capture the attributes that they find especially desirable. These estimates also help clarify theories about the origin and nature of relationship variance (i.e., compatibility), as they represent one of the strongest attempts to use attribute-matching to explain why people are more likely to experience attraction and romantic contentment with some partners rather than others. The current data suggest that the corrected pattern metric across 35 traits may be able to explain 2-4% of relationship variance. But of course, SRM approaches suggest that romantic evaluative measures are mostly (i.e., > 50%) comprised of relationship variance (Kenny, 2019). The lion’s share of human romantic compatibility remains unaccounted for, and we may have to stretch beyond attribute-matching concepts like similarity and preference-matching to explain it (Eastwick et al., 2023).

This study also has some limitations. This study only used measured variables, and experimental approaches will be required to understand the causal consequences of ideals

(Eastwick, Smith, et al., 2019; Rohrer et al., 2022). Furthermore, the participants' partners did not actually take part in this study, and effect sizes will likely decline across the board if the partner's (rather than the participant's) reports of the partner's traits are used instead (Hromatko et al., 2015). If one conservatively estimated that the zero-order corrected pattern metric would decline to (say)  $r = .10$ , a sample size of  $N = 779$  participants would be necessary to achieve 80% power—a challenging but not impossible task. Also, the 35 attributes that we assessed here are certainly not exhaustive, and our results suggest the wisdom of testing the predictive validity of other traits that (a) receive middling (i.e., not especially high) normative desirability ratings, or (b) are prioritized in some cultures more than others. Finally, even though we sampled participants from all over the world, most of them had at least a high school level of education, and many of them likely live in situations where they have substantial freedom of choice over who they could select as a romantic partner. Future research would need to examine how mate evaluations take place in contexts where people themselves have limited input over whom they are expected to court or marry.

## Conclusion

The current study partnered with the Psychological Science Accelerator to test the predictive validity of ideal partner preferences across 43 different countries. Results revealed that ideals did indeed have predictive power, although results were highly dependent on whether preference-matching was conceptualized as a normative match ( $\beta$ s ranging from .30-.40), an idiosyncratic or distinctive match ( $\beta$ s .10-.20), or as the level of specific traits (average  $\beta = .04$ ). These data—especially given the size and breadth of the dataset—should be able to provide effect size benchmarks for future studies of human mate preferences, regardless of whether

researchers are interested in stated preferences, revealed preferences, or preference-matching effects.

### **Data and Code Availability**

Data, codebook, and analysis scripts, are openly available through the Open Science Framework: [https://osf.io/b29vu/?view\\_only=35a15592f8b04cdfb9ab32f45c73f3c6](https://osf.io/b29vu/?view_only=35a15592f8b04cdfb9ab32f45c73f3c6).

## References

Abbey, J. D., & Meloy, M. G. (2017). Attention by design: Using attention checks to detect inattentive respondents and improve data quality. *Journal of Operations Management*, 53, 63-70.

Arciniega, G. M., Anderson, T. C., Tovar-Blank, Z. G., & Tracey, T. J. (2008). Toward a fuller conception of Machismo: Development of a traditional Machismo and Caballerismo Scale. *Journal of Counseling Psychology*, 55, 19-33.

Bahns, A. J., Crandall, C. S., Gillath, O., & Preacher, K. J. (2017). Similarity in relationships as niche construction: Choice, stability, and influence within dyads in a free choice environment. *Journal of Personality and Social Psychology*, 112, 329-355.

Balzarini, R. N., Muise, A., Dobson, K., Kohut, T., Raposo, S., & Campbell, L. (2021). The detriments of unmet sexual ideals and buffering effect of sexual communal strength. *Journal of Personality and Social Psychology*, 120, 1521-1550.

Baranger, D. A., Finsaas, M. C., Goldstein, B. L., Vize, C. E., Lynam, D. R., & Olino, T. M. (2023). Tutorial: Power analyses for interaction effects in cross-sectional regressions. *Advances in Methods and Practices in Psychological Science*, 6(3), 25152459231187531.

Benet-Martinez, V. (2007). Cross-cultural personality research: Conceptual and methodological issues. In Robins, R. W., Fraley, C. R., Krueger, R. F. (Eds) *Handbook of Research Methods in Personality Psychology* (pp. 170-89). New York: Guilford Press.

Berscheid, E., & Ammazzalorso, H. (2001). Emotional experience in close relationships. In G. J. O. Fletcher & M. S. Clark (Eds.), *Blackwell handbook of social psychology: Interpersonal processes* (pp. 308-330). Malden, MA: Blackwell.

Biesanz, J. C. (2010). The social accuracy model of interpersonal perception: Assessing individual differences in perceptive and expressive accuracy. *Multivariate Behavioral Research*, 45(5), 853-885.

Botwin, M. D., Buss, D. M., & Shackelford, T. K. (1997). Personality and mate preferences: Five factors in mate selection and marital satisfaction. *Journal of Personality*, 65, 107-136.

Boxer, C. F., Noonan, M. C., & Whelan, C. B. (2015). Measuring mate preferences: A replication and extension. *Journal of Family Issues*, 36, 163-187.

Buyukcan-Tetik, A., Campbell, L., Finkenauer, C., Karremans, J. C., & Kappen, G. (2017). Ideal standards, acceptance, and relationship satisfaction: Latitudes of differential effects. *Frontiers in Psychology*, 8, 1691.

Burri, R. P., Welling, L. L. M., & Puts, D. A. (2011). Mate-preference drives mate-choice: Men's self-rated masculinity predicts their female partner's preference for masculinity. *Personality and Individual Differences*, 51, 1023-1027.

Buss, D. M. (1989). Sex differences in human mate preferences: Evolutionary hypotheses tested in 37 cultures. *Behavioral and Brain Sciences*, 12, 1-49.

Buss, D. M., & Schmitt, D. P. (2019). Mate preferences and their behavioral manifestations. *Annual Review of Psychology*, 70, 77-110.

Campbell, L., Chin, K., & Stanton, S. C. (2016). Initial evidence that individuals form new relationships with partners that more closely match their ideal preferences. *Collabra: Psychology*, 2, 2.

Campbell, L., Overall, N. C., Rubin, H., & Lackenbauer, S. D. (2013). Inferring a partner's ideal discrepancies: Accuracy, projection, and the communicative role of interpersonal behavior. *Journal of Personality and Social Psychology*, 105, 217-233.

Campbell, L., Simpson, J., Kashy, D., & Fletcher, G. (2001). Ideal standards, the self, and flexibility of ideals in close relationships. *Personality and Social Psychology Bulletin, 27*, 447-462.

Cheung, I., Campbell, L., LeBel, E. P., Ackerman, R. A., Aykutoğlu, B., Bahník, Š., ... & Yong, J. C. (2016). Registered replication report: Study 1 from Finkel, Rusbult, Kumashiro, & Hannon (2002). *Perspectives on Psychological Science, 11*, 750-764.

Cohen, J. (1992). A power primer. *Psychological Bulletin, 112*, 155.

Coles, N. A., March, D. S., Marmolejo-Ramos, F., Banaruee, H., Butcher, N., Cavallet, M., ... & Marozzi, M. (2020). The Many Smiles Collaboration: A Multi-Lab Foundational Test of the Facial Feedback Hypothesis. *Nature Human Behaviour*.

Conroy-Beam, D., & Buss, D. M. (2016). Do mate preferences influence actual mating decisions? Evidence from computer simulations and three studies of mated couples. *Journal of Personality and Social Psychology, 111*, 53-66.

Conroy-Beam, D., Goetz, C. D., & Buss, D. M. (2016). What predicts romantic relationship satisfaction and mate retention intensity: Mate preference fulfillment or mate value discrepancies? *Evolution and Human Behavior, 37*, 440-448.

Driebe, J. C., Stern, J., Penke, L., & Gerlach, T. M. (in press). Probing the predictive validity of ideal partner preferences for future partner traits and relationship outcomes across 13 years. *European Journal of Personality*.

Eagly, A. H., & Wood, W. (1999). The origins of sex differences in human behavior: Evolved dispositions versus social roles. *American Psychologist, 54*, 408-423.

Eastwick, P. W. (2009). Sex differences in mate preferences revisited: The role of ideals in romantic partner selection and retention. (*unpublished doctoral dissertation*). Evanston, IL: Northwestern University.

Eastwick, P. W., Eagly, A. H., Finkel, E. J., & Johnson, S. E. (2011). Implicit and explicit preferences for physical attractiveness in a romantic partner: A double dissociation in predictive validity. *Journal of Personality and Social Psychology, 101*, 993-1011.

Eastwick, P. W., & Finkel, E. J. (2008). Sex differences in mate preferences revisited: Do people know what they initially desire in a romantic partner? *Journal of Personality and Social Psychology, 94*, 245-264.

Eastwick, P. W., Finkel, E. J., & Eagly, A. H. (2011). When and why do ideal partner preferences affect the process of initiating and maintaining romantic relationships? *Journal of Personality and Social Psychology, 101*, 1012-1032.

Eastwick, P. W., Finkel, E. J., & Joel, S. (2023). Mate evaluation theory. *Psychological Review, 130*, 211-241.

Eastwick, P. W., Finkel, E. J., & Simpson, J. A. (2019). Best practices for testing the predictive validity of ideal partner preference-matching. *Personality and Social Psychology Bulletin, 45*, 167-181.

Eastwick, P. W., Joel, S., Carswell, K. L., Molden, D. C., Finkel, E. J., & Blozis, S. A. (2023). Predicting romantic interest during early relationship development: A preregistered investigation using machine learning. *European Journal of Personality, 37*, 276-312.

Eastwick, P. W., Luchies, L. B., Finkel, E. J., & Hunt, L. L. (2014). The predictive validity of ideal partner preferences: A review and meta-analysis. *Psychological Bulletin, 140*, 623-665.

Eastwick, P. W., & Neff, L. A. (2012). Do ideal partner preferences predict divorce? A tale of two metrics. *Social Psychological and Personality Science, 3*, 667-674.

Eastwick, P. W., Neff, L. A., Luchies, L. B., Finkel, E. J., & Hunt, L. L. (2014). Is a meta-analysis a foundation or just another brick? Comment on Meltzer, McNulty, Jackson, & Karney (2014). *Journal of Personality and Social Psychology, 106*, 429-434.

Eastwick, P. W., & Smith, L. K. (2018). Sex-differentiated effects of physical attractiveness on romantic desire: A highly powered, preregistered study in a photograph evaluation context. *Comprehensive Results in Social Psychology, 3*, 1-27.

Eastwick, P. W., Smith, L. K., & Ledgerwood, A. (2019). How do people translate their experiences into abstract attribute preferences?. *Journal of Experimental Social Psychology, 85*, 103837.

Eastwick, P. W., Sparks, J., & Finkel, E. (2024, July 2). The Preference-Matching Project. Retrieved from [osf.io/b29vu/?view\\_only=35a15592f8b04cdfb9ab32f45c73f3c6](https://osf.io/b29vu/?view_only=35a15592f8b04cdfb9ab32f45c73f3c6)

Finkel, E. J., Rustebult, C. E., Kumashiro, M., & Hannon, P. A. (2002). Dealing with betrayal in close relationships: Does commitment promote forgiveness? *Journal of Personality and Social Psychology, 82*, 956-974.

Finsaas, M. C., Baranger, D. A. A., Goldstein, B. L., Viz, e C., Lynam, D., & Olino, T. M. (2021). InteractionPoweR shiny app: Power analysis for interactions in linear regression. Retrieved 1-10-24.

Fletcher, G. J. O., Kerr, P. S., Li, N. P., & Valentine, K. A. (2014). Predicting romantic interest and decisions in the very early stages of mate selection standards, accuracy, and sex differences. *Personality and Social Psychology Bulletin, 40*, 540-550.

Fletcher, G. J., Overall, N. C., & Campbell, L. (2020). Reconsidering “best practices” for testing the ideal standards model: A response to Eastwick, Finkel, and Simpson (2019). *Personality and Social Psychology Bulletin, 46*, 1581-1595.

Fletcher, G. J. O., Simpson, J. A., Campbell, L., & Overall, N. C. (2019). *The Science of Intimate Relationships* (Second ed.). Malden, MA: Wiley-Blackwell.

Fletcher, G. J. O., Simpson, J. A., & Thomas, G. (2000). Ideals, perceptions, and evaluations in early relationship development. *Journal of Personality and Social Psychology, 79*, 933-940.

Fletcher, G. J. O., Simpson, J. A., Thomas, G., & Giles, L. (1999). Ideals in intimate relationships. *Journal of Personality & Social Psychology, 76*, 72-89.

Gelfand, M. J., Raver, J. L., Nishii, L., Leslie, L. M., Lun, J., Lim, B. C., ... & Yamaguchi, S. (2011). Differences between tight and loose cultures: A 33-nation study. *Science, 332*, 1100-1104.

Gerlach, T. M., Arslan, R. C., Schultze, T., Reinhard, S. K., & Penke, L. (2019). Predictive validity and adjustment of ideal partner preferences across the transition into romantic relationships. *Journal of Personality and Social Psychology, 116*, 313-330.

Giner-Sorolla, R. (2018, January 24). Powering your interaction [Blog post]. Retrieved from <https://approachingblog.wordpress.com/2018/01/24/powering-your-interaction-2/>

Goldberg, L. R. (1993). The structure of phenotypic personality traits. *American Psychologist, 48*, 26-34.

Gosling, S. D., Rentfrow, P. J., & Swann Jr, W. B. (2003). A very brief measure of the Big-Five personality domains. *Journal of Research in Personality, 37*, 504-528.

Hammond, M. D., & Overall, N. C. (2014). Endorsing benevolent sexism magnifies willingness to dissolve relationships when facing partner-ideal discrepancies. *Personal Relationships, 21*, 272-287.

Hauser, D. J., Moss, A. J., Rosenzweig, C., Jaffe, S. N., Robinson, J., & Litman, L. (2023). Evaluating CloudResearch's Approved Group as a solution for problematic data quality on MTurk. *Behavior Research Methods*, 55(8), 3953-3964.

Hill, R. (1945). Campus values in mate-selection. *Journal of Home Economics*, 37, 554-558.

Hitsch, G. J., Hotacsu, A., & Ariely, D. (2010). What makes you click? Mate preferences in online dating. *Quantitative Marketing and Economics*, 8, 393-427.

Hopwood, C. J., & Donnellan, M. B. (2010). How should the internal structure of personality inventories be evaluated?. *Personality and Social Psychology Review*, 14, 332-346.

Hromatko, I., Bajoghli, H., Rebernjak, B., Joshaghani, N., & Tadinac, M. (2015). Relationship satisfaction as a function of mate value. *Evolutionary Behavioral Sciences*, 9(4), 242-256.

Joel, S., & MacDonald, G. (2021). We're not that choosy: Emerging evidence of a progression bias in romantic relationships. *Personality and Social Psychology Review*, 25, 317-343.

Kenny, D. A. (2019). *Interpersonal perception: The foundation of social relationships*. Guilford Publications.

Kito, M., Yuki, M., & Thomson, R. (2017). Relational mobility and close relationships: A socioecological approach to explain cross-cultural differences. *Personal Relationships*, 24, 114-130.

Knee, C., Nanayakkara, A., Vietor, N. A., Neighbors, C., & Patrick, H. (2001). Implicit theories of relationships: Who cares if romantic partners are less than ideal? *Personality & Social Psychology Bulletin*, 27, 808-819.

Kurzban, R., & Weeden, J. (2007). Do advertised preferences predict the behavior of speed daters? *Personal Relationships*, 14, 623-632.

Lackenbauer, S. D., & Campbell, L. (2012). Measuring up: The unique emotional and regulatory outcomes of different perceived partner-ideal discrepancies in romantic relationships. *Journal of Personality and Social Psychology*, 103, 472-488.

Lakens, D. (2017). Equivalence tests: A practical primer for t tests, correlations, and meta-analyses. *Social Psychological and Personality Science*, 8, 355-362.

Lam, B. C. P., Cross, S. E., Wu, T.-F., Yeh, K.-H., Wang, Y.-C., & Su, J. C. (2016). What do you want in a marriage? Examining marriage ideals in Taiwan and the United States. *Personality and Social Psychology Bulletin*, 42, 703-722.

Larsen, K. S., & Long, E. (1988). Attitudes toward sex-roles: Traditional or egalitarian? *Sex Roles*, 19, 1-12.

Lawless, H. T., & Heymann, H. (2010). *Sensory Evaluation of Food: Principles and Practices*: Springer Science & Business Media.

Ledgerwood, A. (2019). New developments in research methods. In E. J. Finkel & R. F. Baumeister (Eds.), *Advanced Social Psychology* (Second ed., pp. 39-61): Oxford University Press.

Ledgerwood, A., Eastwick, P. W., & Smith, L. K. (2018). Toward an integrative framework for studying human evaluation: Attitudes towards objects and attributes. *Personality and Social Psychology Review*, 22, 378-398.

Lewis, K. (2016). Preferences in the early stages of mate choice. *Social Forces*, 95, 283-320.

Li, N. P., & Meltzer, A. L. (2015). The validity of sex-differentiated mate preferences: Reconciling the seemingly conflicting evidence. *Evolutionary Behavioral Sciences*, 9, 89-106.

Li, N. P., Yong, J. C., Tov, W., Sng, O., Fletcher, G. J. O., Valentine, K. A., . . . Balliet, D. (2013). Mate preferences do predict attraction and choices in the early stages of mate selection. *Journal of Personality and Social Psychology, 105*, 757-776.

Locke, K. D., Barni, D., Morio, H., MacDonald, G., Mastor, K. A., de Jesús Vargas-Flores, J., ... & Ortiz, F. A. (2020). Culture moderates the normative and distinctive impact of parents and similarity on young adults' partner preferences. *Cross-Cultural Research, 54*, 435-461.

Lorenz, M., Aisch, G., & Kokkelink, D. (2012). Datawrapper: Create charts and maps [software]. Available from <https://www.datawrapper.de>

Mafra, A. L., Fisher, M. L., & Lopes, F. d. A. (2021). Does mate preference represent mate choice? A cross-cultural investigation. *Evolutionary Behavioral Sciences, 15*, 64-81.

Meltzer, A. L., McNulty, J. K., Jackson, G. L., & Karney, B. R. (2014). Sex differences in the implications of partner physical attractiveness for the trajectory of marital satisfaction. *Journal of Personality and Social Psychology, 106*, 418-428.

Moshontz, H., Campbell, L., Ebersole, C. R., IJzerman, H., Urry, H. L., Forscher, P. S., ... & Chartier, C. R. (2018). The Psychological Science Accelerator: Advancing psychology through a distributed collaborative network. *Advances in Methods and Practices in Psychological Science, 1*(4), 501-515.

Murray, S. L., Griffin, D. W., Derrick, J. L., Harris, B., Aloni, M., & Leder, S. (2011). Tempting fate or inviting happiness? Unrealistic idealization prevents the decline of marital satisfaction. *Psychological Science, 22*, 619-626.

Murray, S. L., Holmes, J. G., & Griffin, D. W. (1996). The self-fulfilling nature of positive illusions in romantic relationships: Love is not blind, but prescient. *Journal of Personality and Social Psychology, 71*, 1155-1180.

Neff, L. A., & Karney, B. R. (2003). The dynamic structure of relationship perceptions: Differential importance as a strategy of relationship maintenance. *Personality and Social Psychology Bulletin, 29*, 1433-1446.

Oishi, S., & Schimmack, U. (2010). Residential mobility, well-being, and mortality. *Journal of Personality and Social Psychology, 98*, 980-994.

Olejnik, S., Li, J., Supattathum, S., & Huberty, C. J. (1997). Multiple testing and statistical power with modified Bonferroni procedures. *Journal of Educational and Behavioral Statistics, 22*, 389-406.

Overall, N. C., Fletcher, G. J. O., & Simpson, J. A. (2006). Regulation processes in intimate relationships: The role of ideal standards. *Journal of Personality and Social Psychology, 91*, 662-685.

Perusse, D. (1993). Cultural and reproductive success in industrial societies: Testing the relationship at the proximate and ultimate levels. *Behavioral and Brain Sciences, 16*, 267-322.

Rights, J. D., & Sterba, S. K. (2019). Quantifying explained variance in multilevel models: An integrative framework for defining R-squared measures. *Psychological Methods, 24*, 309-338.

Roberts, S. C., Havlíček, J., & Schaal, B. (2020). Human olfactory communication: current challenges and future prospects. *Philosophical Transactions of the Royal Society B, 375*, 20190258.

Robles, T. F., Slatcher, R. B., Trombello, J. M., & McGinn, M. M. (2014). Marital quality and health: a meta-analytic review. *Psychological Bulletin, 140*, 140-187.

Rodriguez, L. M., Hadden, B. W., & Knee, C. R. (2015). Not all ideals are equal: Intrinsic and extrinsic ideals in relationships. *Personal Relationships*, 22, 138-152.

Rogers, K. H., Wood, D., & Furr, R. M. (2018). Assessment of similarity and self-other agreement in dyadic relationships: A guide to best practices. *Journal of Social and Personal Relationships*, 35, 112-134.

Rohrer, J. M., Hünermund, P., Arslan, R. C., & Elson, M. (2022). That'sa lot to PROCESS! Pitfalls of popular path models. *Advances in Methods and Practices in Psychological Science*, 5(2), 25152459221095827.

Rosseel, Y. (2012). Lavaan: An R package for structural equation modeling and more. Version 0.5-12 (BETA). *Journal of Statistical Software*, 48, 1-36.

Rossignac-Milon, M., & Higgins, E. T. (2018). Epistemic companions: Shared reality development in close relationships. *Current Opinion in Psychology*, 23, 66-71.

Rusbult, C. E., Martz, J. M., & Agnew, C. R. (1998). The investment model scale: Measuring commitment level, satisfaction level, quality of alternatives, and investment size. *Personal Relationships*, 5, 357-387.

Ruvolo, A. P., & Veroff, J. (1997). For better or for worse: Real-ideal discrepancies and the marital well-being of newlyweds. *Journal of Social and Personal Relationships*, 14, 223-242.

Selterman, D., & Gideon, S. (2022). Experiences of romantic attraction are similar across dating apps and offline dates in young adults. *Journal of Social Psychology Research*, 1, 145-163.

Shackelford, T. K., & Buss, D. M. (1997). Marital satisfaction in evolutionary psychological perspective. In R. J. Sternberg & M. Hojjat (Eds.), *Satisfaction in Close Relationships* (pp. 7-25). New York, NY: Guilford.

Shaw, M., Rights, J. D., Sterba, S. S., & Flake, J. K. (2023). r2mlm: An R package calculating R-squared measures for multilevel models. *Behavior Research Methods*, 55(4), 1942-1964.

Simpson, J. A., Fletcher, G. J. O., & Campbell, L. (2001). The structure and function of ideal standards in close relationships. In G. J. O. Fletcher & M. S. Clark (Eds.), *Blackwell handbook of social psychology: Interpersonal Processes* (pp. 86-106). Malden, MA: Blackwell Publishers.

Snijders, T. A., & Bosker, R. J. (2012). *An Introduction to Basic and Advanced Multilevel Modeling* (2nd ed.). London: Sage.

Sparks, J., Daly, C. A., Wilkey, B. M., Molden, D. C., Finkel, E. J., & Eastwick, P. W. (2020). Negligible evidence that people desire partners who uniquely fit their ideals. *Journal of Experimental Social Psychology*, 90, 103968.

Strauss, C., Morry, M. M., & Kito, M. (2012). Attachment styles and relationship quality: Actual, perceived, and ideal partner matching. *Personal Relationships*, 19, 14-36.

Sugiyama, L. S. (2005). Physical attractiveness: An adaptationist perspective. In D. M. Buss (Ed.), *The Handbook of Evolutionary Psychology* (pp. 292-343). Hoboken, NJ: John Wiley & Sons, Inc.

Tancredy, C. M., & Fraley, R. C. (2006). The nature of adult twin relationships: An attachment-theoretical perspective. *Journal of Personality and Social Psychology*, 90, 78-93.

Thomson, R., Yuki, M., Talhelm, T., Schug, J., Kito, M., Ayanian, A. H., ... & Ferreira, C. M. (2018). Relational mobility predicts social behaviors in 39 countries and is tied to

historical farming and threat. *Proceedings of the National Academy of Sciences*, 115, 7521-7526.

Triandis, H. C., & Gelfand, M. J. (1998). Converging measurement of horizontal and vertical individualism and collectivism. *Journal of Personality and Social Psychology*, 74, 118-128.

Todd, P. M., Penke, L., Fasolo, B., & Lenton, A. P. (2007). Different cognitive processes underlie human mate choices and mate preferences. *Proceedings of the National Academy of Sciences*, 104, 15011-15016.

Valentine, K. A., Li, N. P., Meltzer, A. L., & Tsai, M.-H. (2020). Mate preferences for warmth-trustworthiness predict romantic attraction in the early stages of mate selection and satisfaction in ongoing relationships. *Personality and Social Psychology Bulletin*, 46, 298-311.

van Buuren, S., & Groothuis-Oudshoorn, K. (2011). "Mice: Multivariate imputation by chained equations in R." *Journal of Statistical Software*, 45, 1-67.

Vohs, K. D., Schmeichel, B. J., Lohmann, S., Gronau, Q. F., Finley, A. J., Ainsworth, S. E., ... & Albarracín, D. (2021). A multisite preregistered paradigmatic test of the ego-depletion effect. *Psychological Science*, 32, 1566-1581.

Wood, D., & Brumbaugh, C. C. (2009). Using revealed mate preferences to evaluate market force and differential preference explanations for mate selection. *Journal of Personality and Social Psychology*, 96, 1226-1244.

Wood, D., & Furr, R. M. (2016). The correlates of similarity estimates are often misleadingly positive: The nature and scope of the problem, and some solutions. *Personality and Social Psychology Review*, 20, 79-99.

Wood, D., Lowman, G. H., Harms, P., & Roberts, B. W. (2019). Exploring the relative importance of normative and distinctive organizational preferences as predictors of work attitudes. *Journal of Applied Psychology*, 104, 270-292.

Wu, K., Chen, C., & Greenberger, E. (2018). A rosier reality: Incongruency in stated and revealed ingroup preferences among young asian american speed daters. *Social Psychology Quarterly*, 81, 340-360.

Yuki, M., & Schug, J. (2012). Relational mobility: A socioecological approach to personal relationships. In O. Gillath, G. Adams, & A. Kunkel (Eds.), *Decade of Behavior 2000-2010. Relationship Science: Integrating Evolutionary, Neuroscience, and Sociocultural Approaches* (p. 137-151). American Psychological Association.

Zentner, M. (2005). Ideal mate personality concepts and compatibility in close relationships: A longitudinal analysis. *Journal of Personality and Social Psychology*, 89, 242-256.

## Author Contributions

**Authorship order:** (1) Project coordinators (i.e., first four authors), (2) Participating labs, (3) Consultants (i.e., last author). Project coordinators are sorted by the size of their contribution of the project. Participating labs are sorted alphabetically by last name.

**Conceptualization:** Paul W. Eastwick, Jehan Sparks, and Eli J. Finkel.

**Data curation:** Paul W. Eastwick, Jehan Sparks, and Eva M. Meza.

**Formal analysis:** Paul W. Eastwick and Jehan Sparks.

**Funding acquisition:** Paul W. Eastwick, Eli J. Finkel, Matúš Adamkovič, Aderonke A. Akintola, Maciej Behnke, Theofilos Gkinopoulos, Dmitry Grigoryev, Gul Gunaydin, Mirosław Kocur, Marta Kowal, Jackson G. Lu, Marc Y. Lucas, Wojciech Małecki, Michał Misiak, S. Craig Roberts, Ivan Ropovik, Robert M. Ross, Cristina E. Salvador, Emre Selcuk, Agnieszka Sorokowska, Piotr Sorokowski, and Yuki Yamada.

**Investigation:** Paul W. Eastwick, Eva M. Meza, Matúš Adamkovič, Peter Adu, Ting Ai, Aderonke A. Akintola, Laith Al-Shawaf, Denisa Apriliaawati, Patrícia Arriaga, Benjamin Aubert-Teillaud, Gabriel Baník, Krystian Barzykowski, Carlota Batres, Katherine J. Baucom, Elizabeth Z. Beaulieu, Maciej Behnke, Natalie Butcher, Deborah Y. Charles, Jane M. Chen, Jeong Eun Cheon, Phakkanun Chittham, Patrycja Chwiłkowska, Chin Wen Cong, Lee T. Copping, Nadia S. Corral-Frias, Vera Ćubela Adorić, Mikaela Dizon, Hongfei Du, Michael I. Ehnmowo, Daniela A. Escribano, Natalia M. Espinosa, Francisca Expósito, Gilad Feldman, Raquel Freitag, Martha Frias Armenta, Albina Gallyamova, Omri Gillath, Biljana Gjoneska, Theofilos Gkinopoulos, Franca Grafe, Dmitry Grigoryev, Agata Groycka-Bernard, Gul Gunaydin, Ruby Ilustrisimo, Pavol Kačmár, Young-Hoon Kim, Mirosław Kocur, Marta Kowal, Maatangi Krishna, Paul Danielle Labor, Jackson G. Lu, Marc Y. Lucas, Wojciech Małecki, Klara Malinakova, Sofia Meißner, Zdeněk Meier, Michał Misiak, Amy Muise, Lukas Novak, Jiaqing O, Asil A. Özdoğru, Haeyoung Gideon Park, Mariola Paruzel, Zoran Pavlović, Marcell Püske, Gianni Ribeiro, S. Craig Roberts, Jan P. Röer, Ivan Ropovik, Robert M. Ross, Ezgi Sakman, Cristina E. Salvador, Emre Selcuk, Shayna Skakoon-Sparling, Agnieszka Sorokowska, Piotr Sorokowski, Ognen Spasovski, Sarah C. E. Stanton, Suzanne L. K. Stewart, Viren Swami, Barnabas Szaszi, Kaito Takashima, Peter Tavel, Julian Tejada, Eric Tu, Jarno Tuominen, David Vaidis, Zahir Vally, Leigh Ann Vaughn, Laura Villanueva-Moya, Dian Wisnuwardhani, Yuki Yamada, Fumiya Yonemitsu, Radka Žídková, and Kristýna Žívná.

**Methodology:** Paul W. Eastwick, Jehan Sparks, Eli J. Finkel, and S. Craig Roberts.

**Project administration:** Paul W. Eastwick, Eva M. Meza, and Nicholas A. Coles.

**Resources:** Paul W. Eastwick, Eva M. Meza, Patrícia Arriaga, Lee T. Copping, Asil A. Özdoğru, and Cristina E. Salvador.

**Software:** Paul W. Eastwick and Jehan Sparks.

**Supervision:** Paul W. Eastwick, Eva M. Meza, Emily Impett, Lukas Novak, Sarah C. E. Stanton, Peter Tavel, and Nicholas A. Coles.

**Validation:** Paul W. Eastwick and Jehan Sparks.

**Visualization:** Paul W. Eastwick.

**Writing - original draft:** Paul W. Eastwick.

**Writing - review & editing:** Paul W. Eastwick, Jehan Sparks, Eli J. Finkel, Eva M. Meza, Matúš Adamkovič, Peter Adu, Ting Ai, Aderonke A. Akintola, Laith Al-Shawaf, Denisa Apriliawati, Patrícia Arriaga, Benjamin Aubert-Teillaud, Gabriel Baník, Krystian Barzykowski, Carlota Batres, Katherine J. Baucom, Maciej Behnke, Natalie Butcher, Deborah Y. Charles, Jeong Eun Cheon, Patrycja Chwiłkowska, Chin Wen Cong, Lee T. Copping, Nadia S. Corral-Frias, Vera Ćubela Adorić, Hongfei Du, Daniela A. Escribano, Raquel Freitag, Martha Frias Armenta, Omri Gillath, Biljana Gjoneska, Theofilos Gkinopoulos, Dmitry Grigoryev, Agata Groycka-Bernard, Gul Gunaydin, Emily Impett, Pavol Kačmár, Young-Hoon Kim, Mirosław Kocur, Marta Kowal, Maatangi Krishna, Jackson G. Lu, Marc Y. Lucas, Wojciech Małecki, Klara Malinakova, Zdeněk Meier, Michal Misiak, Amy Muise, Lukas Novak, Jiaqing O, Asil A. Özdogru, Haeyoung Gideon Park, Zoran Pavlović, Marcell Püski, Gianni Ribeiro, S. Craig Roberts, Jan P. Röer, Ivan Ropovik, Robert M. Ross, Cristina E. Salvador, Emre Selcuk, Shayna Skakoon-Sparling, Agnieszka Sorokowska, Piotr Sorokowski, Ognen Spasovski, Sarah C. E. Stanton, Suzanne L. K. Stewart, Viren Swami, Barnabas Szaszi, Peter Tavel, Eric Tu, Jarno Tuominen, David Vaidis, Leigh Ann Vaughn, Dian Wisnuwardhani, Yuki Yamada, Radka Žídková, Kristýna Živná, and Nicholas A. Coles.

## Funding Information

Benjamin Aubert-Teillaud was supported by ANRT CIFRE-WWF grant 2020/0738; Albina Gallyamova and Dmitry Grigoryev were supported by Basic Research Program at HSE University, RF; Mirosław Kocur, Marta Kowal, Wojciech Małecki, Michal Misiak, S. Craig Roberts, Agnieszka Sorokowska, and Piotr Sorokowski were supported by IDN Being Human Lab (University of Wroclaw); Nicholas A. Coles was supported by John Templeton Foundation (#62295), U.S. National Science Foundation (#2235066); Robert M. Ross was supported by John Templeton Foundation (Grant ID: 62631) and Australian Research Council (Grant ID: DP180102384); Yuki Yamada was supported by JSPS KAKENHI: JP20H04581, JP21H03784, and JP22K18263; Zoran Pavlović was supported by Ministry of Science, Technological Development and Innovation of the Republic of Serbia (grant no. 451-03-47/2023-01/200163); Paul W. Eastwick was supported by National Science Foundation Grant BCS-1941440, and a UC Davis Small Research Grant; Katherine J. Baucom was supported by NIH K23DK115820; Ivan Ropovik was supported by NPO EXCELES: Systemic Risk Institute (LX22NPO5101); Matúš Adamkovič was supported by PRIMUS/24/SSH/017 and APVV-22-0458; Gul Gunaydin and Emre Selcuk were supported by Sabanci University Integration Grant B.A.CG-20-02170; Maciej Behnke was supported by Statutory funds of the Department of Psychology and Cognitive Science, Adam Mickiewicz University,; Patrícia Arriaga was supported by UID/PSI/03125/2022; Pavol Kačmár was supported by VEGA 1/0853/21, APVV-19-0284.

## Research Ethics Committee/ Institutional Review Board Information

The research reported in this article was approved by (or declared exempt by) the following protocols: University of California, Davis (1898056-1 “The Preference Matching Project”); Adam Mickiewicz University (No. 2/11/2022 “The Preference-Matching Project”); Anglia Ruskin University (ETH2223-3477 “The Preference-Matching Project”); Christ University (RCEC/88/07/22 “A Worldwide Test of the Predictive Validity of Ideal Partner Preference-Matching”); Chulalongkorn University (660026 “A Worldwide Test of the Predictive Validity of Ideal Partner Preference Matching”); Duke University (2023-0187 “The Preference-Matching Project”); Franklin & Marshall College (R#\_2vhJ3np8V4EspbV “A Worldwide Test of the Predictive Validity of Ideal Partner Preference-Matching”); HSE University (nPOTOK0/1 01/10.2022 “The Preference-Matching Project”); University Institute of Lisbon (86/2022 “Cross-cultural study on Preference Matching”); Institute of Social Sciences CSPS SAS (“The Preference-Matching Project”); Ithaca College (Protocol #365 “Preference Matching”); Kyushu University (2023-003 “A Worldwide Test of the Predictive Validity of Ideal Partner Preference Matching”); Macedonian Academy of Sciences and Arts (“A worldwide test of the predictive validity of pairing based on romantic preferences for ideal partners”); Redeemer’s University (RUN/REC/2022/11 “The Preference-Matching Worldwide Project: Ideal Partner Preference-Matching”); Sabancı University (FASS-2022-32 “Preference Matching Project”); SWPS University of Social Sciences and Humanities (7875342DoPSc “A Worldwide Test of the Predictive Validity of Ideal Partner Preference-Matching”); Teesside University (2022 Jul 10487 Copping “Mate Preference Matching Project”); The University of Edinburgh (337-2223/2 “The Relationships and Context Study”); The University of Queensland (2022/HE001268 “A Worldwide Test of the Predictive Validity of Ideal Partner Preference Matching”); Toronto Metropolitan University (REB 2023-025 “The Relationships and Context Study”); UIN Sunan Kalijaga (B-292/Un.02/L3/TL/01/2023 “A Worldwide Test of the Predictive Validity of Ideal Partner Preference Matching”); Universidade Federal de Sergipe (CAAE: 64709722.1.0000.5546 “Um Teste Mundial da Validade Preditiva da Correspondência da Preferência do Parceiro Ideal”); Université de Paris (1898056-1 “The Preference-Matching Project”); University of Belgrade (Protocol #2022-68 “Ideal Partner Preference-Matching Across Cultures”); University of Chester (DOPEC SST270622 “A Worldwide Test of the Predictive Validity of Ideal Partner Preference-Matching”); University of Colorado – Colorado Springs (2023-061 “The Relationships and Context Study”); University of Granada (2908/CEIH/2022 “Preference Matching Project”); University of Hong Kong (EA220448 “A Worldwide Test of the Predictive Validity of Ideal Partner Preference-Matching, Hong Kong data collection”); University of Ibadan (UI/SSHREC/2022/0025 “A Worldwide Test of the Predictive Validity of Ideal Partner Preference Matching”); University of Kansas (STUDY00148609 “Ideal Partner Preference-Matching Across Cultures”); University of the Philippines Diliman (CSSPERB-2023-007 “A Worldwide Test of the Predictive Validity of Ideal Partner Preference Matching”); University of Turku (“Relationships and Context Study”); University of Wrocław (2022/FBMOQ “The Preference Matching Project”); University of Zadar (2198-1-79-41/22-04 “A Worldwide Test of the Predictive Validity of Ideal Partner Preference Matching”); Witten/Herdecke University (declared exempt); Yonsei University (202208-HR-2923-02); and York University (e2022-233 “Preference Matching Study”).