

# Bounds on stomatal size can explain scaling with stomatal density in forest plants

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## Summary

- A prevailing hypothesis posits that achieving higher maximum rates of leaf carbon gain and water loss is constrained by geometry and/or selection to limit the allocation of epidermal area to stomata ( $f_s$ ). Under this 'stomatal-area minimization hypothesis', higher  $g_{s,max}$  is associated with greater numbers of smaller stomata because this trait combination increases  $g_{s,max}$  with minimal increase in  $f_s$ , leading to relative conservation of  $f_s$  semi-independent of  $g_{s,max}$  due to coordination in stomatal size, density, and pore depth. An alternative hypothesis is that the evolution of higher  $g_{s,max}$  can be enabled by a greater epidermal area allocated to stomata, leading to positive covariation between  $f_s$  and  $g_{s,max}$ ; we call this the 'stomatal-area adaptation hypothesis'. Under this hypothesis, the interspecific scaling between  $g_{s,max}$ , stomatal density, and stomatal size is a by-product of selection on a moving optimal  $g_{s,max}$ .
- We integrated biophysical and evolutionary quantitative genetic modeling with phylogenetic comparative analyses of a global data set of stomatal density and size from 2408 vascular forest species. The models present specific assumptions of both hypotheses and deduce predictions that can be evaluated with our empirical analyses of forest plants.
- There are three main results. First, neither the stomatal-area minimization nor adaptation hypothesis is sufficient to be supported. Second, estimates of interspecific scaling from common regression methods cannot reliably distinguish between hypotheses when stomatal size is bounded. Third, we reconcile both hypotheses with the data by including an additional assumption that stomatal size is bounded by a wide range and under selection; we refer to this synthetic hypothesis as the 'stomatal adaptation + bounded size' hypothesis.
- This study advances our understanding of scaling between stomatal size and density by mathematically describing specific assumptions of competing hypotheses, demonstrating that existing hypotheses are inconsistent with observations, and reconciling these hypotheses with phylogenetic comparative analyses by postulating a synthetic model of selection on  $g_{s,max}$ ,  $f_s$ , and stomatal size.

## Introduction

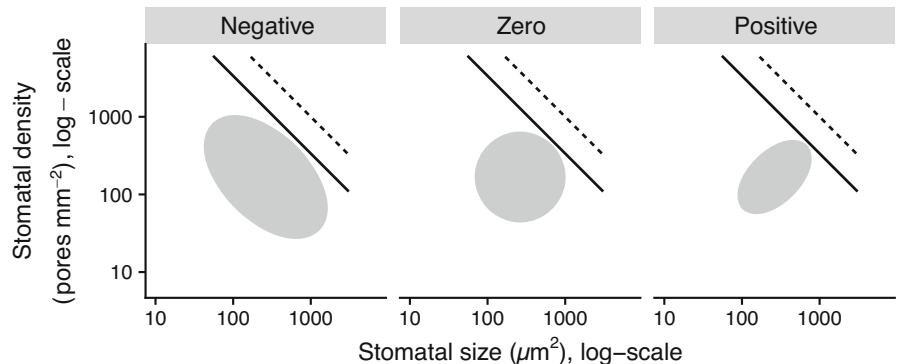
Stomatal pores are critical determinants of the function of plants and the composition of the atmosphere (Berry *et al.*, 2010). The stomatal conductance to diffusion of water vapor and CO<sub>2</sub> ( $g_s$ ) influences a broad spectrum of ecological processes at leaf,

community, and ecosystem scales, including photosynthesis, net primary production, and water-use efficiency (Cramer *et al.*, 2001; Haworth *et al.*, 2011). Stomata can regulate  $g_s$  either through evolutionary or plastic shifts in stomatal size or density (Jordan *et al.*, 2015) or through short-term stomatal aperture changes (Hetherington & Woodward, 2003). The  $g_s$ , and its typical operational value measured under field conditions ( $g_{s,op}$ ), can thus vary from near zero with stomata fully closed to  $g_{s,max}$

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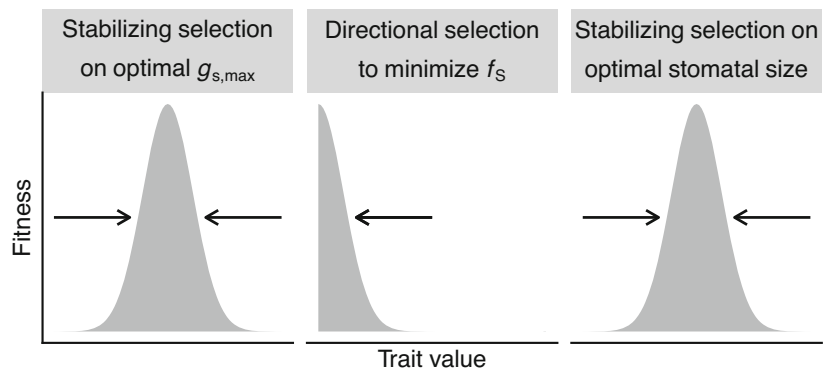
(a)

### Many relationships between stomatal size and density are geometrically possible



(b)

### Assumptions underlying competing hypotheses for inverse size–density scaling



**Fig. 1** Assumptions underlying competing hypotheses for stomatal size–density scaling make different predictions about the trait covariance structure. Maximum stomatal conductance ( $g_{s,max}$ ) and the fraction of epidermal area allocated to stomata ( $f_s$ ) are determined by stomatal density and size. On a log scale, they are the sum of log-stomatal density ( $d_s$ ) and log-stomatal size ( $a_s$ ) times a scaling exponent ( $\beta$ ), 0.5 for  $g_{s,max}$  and 1.0 for  $f_s$  (see the [Materials and Methods](#) section). (a) Many scaling relationships between stomatal size and density are possible as long as  $f_s$  does not exceed 1 (dashed line) or more realistically a value less than 1 to allow space between stomata (solid line,  $f_s = 0.34$ , the maximum value in our data set). The gray ellipses represent different possible scaling relationships with the same mean trait values in our data set ( $\bar{A}_s = 263 \mu\text{m}^2$ ,  $\bar{D}_s = 168 \text{ pores mm}^{-2}$ ). These are 95% quantile of covariance ellipses for a bivariate normal with trait correlations of  $-0.5$ ,  $0$ , and  $0.5$  and trait variances of  $0.75$ ,  $0.55$ , and  $0.45$  for 'negative', 'zero', and 'positive' relationships, respectively. (b) We consider nested hypotheses (see the [Materials and Methods](#) section) that all assume stabilizing selection on optimal  $g_{s,max}$  (left panel) and then add complexity by assuming selection to minimize  $f_s$  (middle panel) and stabilizing selection on stomatal size (right panel).

with stomata fully open. The  $g_{s,max}$  is a fundamental anatomical constraint, and across species measured under controlled conditions,  $g_{s,op}$  and  $g_{s,max}$  are correlated (Wilson, 1975; Franks *et al.*, 2013; Haworth *et al.*, 2013). For example, in woody angiosperms,  $g_{s,op}$  is consistently *c.* 25% of  $g_{s,max}$  (Murray *et al.*, 2020). Because of its importance in controlling leaf water and CO<sub>2</sub> fluxes, stomatal anatomy can provide critical information in global vegetation and crop models toward the current grand challenge of understanding how crops and forest trees are optimized for carbon gain vs water use (Ordoñez *et al.*, 2009; Yuan & Chen, 2009; Díaz *et al.*, 2016; Freschet *et al.*, 2017). The  $g_{s,max}$  varies substantially among extant species (Sack & Buckley, 2016; Murray *et al.*, 2020; Liu *et al.*, 2022) and changed over geological time in response to global climate change (Royer, 2001; Franks *et al.*, 2009; McElwain & Steinthorsdottir, 2017). Trade-offs between carbon gain and costs of stomata, such as water loss, likely contribute to much of the variation in  $g_{s,max}$  among extant species and over time, yet the relative importance of different costs among vascular land plants is not well-understood.

The  $g_{s,max}$  is often modeled as a function (Eqn 2) of underlying anatomical traits stomatal density ( $D_s$ , number of pores per unit epidermal area) and size ( $A_s$ , area of guard cells surrounding each pore). The relationship between stomatal anatomy and conductance is derived from Fick's law of diffusion, with 'end-

correction' albeit with simplifying assumptions about guard cell geometry and homogeneity of the leaf lamina. Anatomical traits are widely used to study the adaptation and competition of plants because they correlate with  $g_{s,op}$  (Brown & Escombe, 1901; Parlange & Waggoner, 1970; Franks & Farquhar, 2001; Vaten & Bergmann, 2012; McElwain *et al.*, 2016; Conesa *et al.*, 2019; Murray *et al.*, 2020). Two long-standing observations about stomatal anatomy are that (1)  $A_s$  and  $D_s$  negatively covary and (2) higher  $g_{s,max}$  is associated with greater density of smaller stomata among species. Both patterns have been observed among extant species and in the fossil record. Similar patterns often occur within species (e.g. Franks & Beerling, 2009a,b, and references therein). Biologists have long observed an inverse relationship between  $A_s$  and  $D_s$  across diverse plant species, first recognized in 1865 (Weiss, 1865), and well-established more recently (Franks & Beerling, 2009a; de Boer *et al.*, 2016; Sack & Buckley, 2016; Haworth *et al.*, 2023; Liu *et al.*, 2023). More recent surveys also observe that species with the greatest  $g_{s,max}$  are on the high- $D_s$ , low- $A_s$  end of the spectrum (Franks & Beerling, 2009a; de Boer *et al.*, 2016). One might think that stomatal size and density negatively covary because of the geometric constraint that they cannot occupy more than the entire leaf surface, but this is not mathematically required (Fig. 1a). Evolutionary processes that affect stomata may result in predictable among-species scaling

exponents between size, density, and other traits (de Boer *et al.*, 2016).

Here, we consider two potential evolutionary hypotheses that are not mutually exclusive, but rather two endpoints along a continuum of explanations. The first hypothesis posits that epidermal space allocated for stomata is strongly constrained by selection maintaining other structures such as trichomes (Baird *et al.*, 2024) or bundle sheath extensions (Baresch *et al.*, 2019), or limiting pathogen infection (Muir, 2020), or maximizing internal leaf space for mesophyll cells (Lundgren *et al.*, 2019). Hence, there is likely a trade-off between allocating epidermal area to stomata, denoted  $f_s$  (Eqn 1), instead of other functions. The costs of allocating more stomatal area may increase as  $f_s$  increases because stomata crowd out other functions and even compromise the function of other stomata through stomatal interference (Lehmann & Or, 2015) and inadequate spacing to allow proper opening (Dow *et al.*, 2014). We refer to this as the ‘stomatal-area minimization hypothesis’ because it assumes that  $g_{s,max}$  is constrained by selection to minimize  $f_s$ . This hypothesis is not new. For example, Franks *et al.* (2009) proposed that competition for epidermal space would predict that  $f_s$  should be relatively constant and  $D_s$  and  $A_s$  do not positively covary. Our goal for this hypothesis was to clarify its assumptions of this hypothesis and more completely derive its predictions.

One consequence of stomatal-area minimization is that a response to selection for sufficiently high  $g_{s,max}$  can only be achieved through combinations of small stomata at high density (Franks & Beerling, 2009b). Because of the allometry between stomatal size and pore depth (de Boer *et al.*, 2016; Sack & Buckley, 2016), combinations of larger stomata at lower density with the same  $g_{s,max}$  require more epidermal space than combinations of smaller stomata at higher density. This hypothesis could therefore explain why high  $g_{s,max}$  leaves are associated with small stomata at high density. It can also explain why there is negative covariation between stomatal size and density. Leaves with the highest  $g_{s,max}$  values cannot obtain certain combinations of larger size and lower density without increasing  $f_s$  beyond its geometric limit. However, it is not clear under this hypothesis why there appears to be a deficit of leaves with small size at low density (Fig. 1).

There are two potential problems with the stomatal-area minimization hypothesis. The first is that actual  $f_s$  is typically far below its geometric limit of 100% (33.6%) is the maximum  $f_s$  in our data set of extant species (Cornelissen *et al.* 2003; solid line in Fig. 1a) and most leaves are well below this (the median  $f_s$  is 4.4% in our data; see the ‘Stomatal trait data from global forests’ in the Materials and Methods section, for a description of the data). Compared with other geological periods, the  $f_s$  among extant leaves is relatively high because extant species evolved in a low- $CO_2$  world, current anthropogenic emissions notwithstanding, where carbon limitation is likely. It seems unlikely that space allocated to stomata would trade off with other epidermal functions when  $f_s$  values are this low. We are not questioning whether stomata have no cost (we discuss these costs later), but rather the assumption that epidermal space imposes a major limitation on stomatal trait variation. At prevailing values of  $f_s$ , most species

could increase  $g_{s,max}$  through higher  $A_s$ ,  $D_s$ , or both; increasing  $g_{s,max}$  does not require that  $A_s$  and  $D_s$  evolve in the opposite direction to minimize  $f_s$  (Franks *et al.*, 2009). For example, consider two leaves with stomatal densities 250 and 200 pores  $mm^{-2}$  and stomatal areas 150 and 187.5  $\mu m^2$ . We use the following equations to calculate  $f_s$  and  $g_{s,max}$  from  $A_s$  and  $D_s$ :

$$f_s = D_s A_s \quad \text{Eqn 1}$$

$$g_{s,max} = b m D_s A_s^{0.5} \quad \text{Eqn 2}$$

where  $b$  and  $m$  are biophysical and morphological constants, respectively (Sack & Buckley, 2016) (see the Materials and Methods section for equations to calculate these constants). The two leaves have an identical  $f_s$  of 0.0375, but  $g_{s,max}$  at 25°C is 11% greater for the leaf with smaller stomata (1.47 vs 1.32  $mol\ m^{-2}\ s^{-1}$ ). Suppose the environment for the species with larger stomata (lower  $g_{s,max}$ ) changed such that the optimal  $g_{s,max}$  becomes 1.47  $mol\ m^{-2}\ s^{-1}$ , how would this species respond to selection? It could respond by increasing  $D_s$  to 224 pores  $mm^{-2}$ , which would change  $f_s$  to 0.042, or it could respond by increasing  $A_s$  to 234  $\mu m^2$  and  $f_s$  to 0.047. To minimize  $f_s$ , selection for higher  $g_{s,max}$  should favor higher  $D_s$  on the assumption that differences in  $f_s$  impose a fitness cost. We are not aware of empirical estimates of the selection gradient on  $f_s$ , and this will be an important area for future research.

A second potential problem with the stomatal-area minimization hypothesis is that directional selection for decreased  $f_s$  would eliminate variation in stomatal size. In the absence of any trade-offs on epidermal space allocated to stomata, the response to selection for higher  $g_{s,max}$  depends on genetic (co)variance in stomatal size and density (Lande, 1979). When  $f_s$  is limited because of trade-offs, this manifests as a greater pleiotropic fitness cost of achieving high  $g_{s,max}$  by increasing  $A_s$  and a lower pleiotropic fitness cost of achieving high  $g_{s,max}$  by increasing  $D_s$ . A logical extension of this hypothesis is that even stabilizing selection on  $g_{s,max}$  should favor greater  $D_s$  and lower  $A_s$  to minimize  $f_s$ . Unless genetic variance in  $D_s$  is near zero or there is strong positive genetic covariance between  $A_s$  and  $D_s$ , selection will favor smaller  $A_s$ . Such directional selection eventually eliminates variance, which contrasts with the wide variance in stomatal size among species. Responses to selection for smaller stomata to reduce  $f_s$  may be limited by factors such as genome size (Beaulieu *et al.*, 2008; Šimová & Herben, 2012; Roddy *et al.*, 2020) or inefficient diffusion through small pores (Hodgson *et al.*, 2010).

An alternative to the ‘stomatal-area minimization’ hypothesis is what we call the ‘stomatal-area adaptation’ hypothesis, that is  $f_s$  should evolve to optimize  $g_{s,max}$ . We assume that the optimal  $g_{s,max}$  is determined by trade-offs between carbon gain, water loss, protection from pathogens, and other factors. In one sense, it is trivially true that the covariance between  $g_{s,max}$  and  $f_s$  is almost certainly positive because both depend on stomatal density and size, which is borne out in our data set (Supporting Information Fig. S1). Mathematically, negative covariance between  $g_{s,max}$  and  $f_s$  is only possible if the correlation between  $D_s$  and  $A_s$  is very

close to  $-1$  (Notes S1), which is inconsistent with observations. Instead, what we mean is that selection for higher  $g_{s,max}$  is much stronger than selection to minimize  $f_s$ , meaning that the response to selection will not be biased in the direction of minimizing surface allocation. This hypothesis recognizes that no relationship between  $A_s$  and  $D_s$  is mathematically required. If the fraction of the leaf surface taken up by stomata is much less than unity, qualitatively different relationships between stomatal size and density across species are geometrically possible, including negative, zero, and positive covariances (ellipses in Fig. 1a). Empirically, inverse scaling is common but not universal (Dunbar-Co *et al.*, 2009). The hypothesis also assumes that stomatal size and density can evolve at least somewhat independently within and among species. Stomata are spaced out by epidermal cells to open and close properly (Dow *et al.*, 2014), but the development of higher  $D_s$  can occur through the increased differentiation of epidermal cells into stomata (i.e. achieving higher stomatal differentiation rate, or stomatal index; Salisbury, 1928; Sack & Buckley, 2016), without any effect on stomatal size. Hence,  $A_s$  and  $D_s$  should be able to evolve independently. Empirically, populations can have genetic variation in stomatal size that is independent of genetic variation in density (e.g. Dittberner *et al.* (2018)). Under the stomatal-area adaptation hypothesis, many combinations of  $A_s$  and  $D_s$  have similar fitness through their effect on  $g_{s,max}$  and the covariance between  $A_s$  and  $D_s$  among species emerges indirectly through their effect on  $g_{s,max}$ . This hypothesis is an oversimplification because not all  $A_s$  values are possible or adaptive. Genome size and physics limit stomata from becoming too small, as mentioned in the previous paragraph, and large stomata may be inefficient in terms of adjusting aperture because of their low surface area to volume ratio (Drake *et al.*, 2013; Raven, 2014).

Distinguishing between these hypotheses will help better understand changes in stomatal traits in the fossil record and variation among species and communities today. The stomatal-area minimization hypothesis indicates that competition for epidermal space is essential for understanding variation in stomatal traits. By contrast, the stomatal-area adaptation hypothesis indicates that selection on optimal  $g_{s,max}$  is fundamental to understanding variation in stomatal traits. These are not mutually exclusive hypotheses. Stomatal-area minimization may be important as  $f_s$  approaches an upper bound, but the stomatal-area adaptation hypothesis may better explain variation when  $f_s$  is near zero.

Although multiple approaches are needed to evaluate these hypotheses, here we focus on among-species scaling between stomatal size and density. Because stomatal-area minimization and adaptation hypotheses assume different constraints on stomatal anatomy, they may result in different scaling relationships. Note that both  $f_s$  (Eqn 1) and  $g_{s,max}$  (Eqn 2) show similar mathematical dependence on  $D_s$  and  $A_s$  that we can generalize as:

$$Z_S = Z_0 D_S A_S^\beta \quad \text{Eqn 3}$$

where a composite stomatal trait  $Z_S$  (i.e.  $f_s$  or  $g_{s,max}$ ) is proportional to the product of constituent stomatal traits ( $A_s$  and  $D_s$ ), with scaling exponent  $\beta$  multiplied by a scalar  $Z_0$ , which reflects

stomatal dimension proportionalities and physical diffusion factors (Sack & Buckley, 2016). For  $g_{s,max}$ ,  $Z_0 = bm$  and  $\beta = 0.5$  (Eqn 1); for  $f_s$ ,  $Z_0 = 1$  and  $\beta = 1$  (Eqn 2). The different scaling exponents arise because diffusion is proportional the linear dimension of the pore and the relationship between maximum pore area and depth is constrained by a relatively constant guard cell shape. Specifically, the length, width, and depth of guard cells are proportional, which is captured by the morphological constant described in more detail later. The different scaling exponents set up the potential for  $g_{s,max}$  to increase semi-independently of  $f_s$  when plants evolve smaller and shallower stomata while stomatal density increases.

Since all traits are log-normally distributed, we log-transformed Eqn 3:

$$z_S = z_0 + d_S + \beta a_S \quad \text{Eqn 4}$$

where lowercase variables indicate log-transformation of uppercase counterparts. Log-transformation also has the advantage of linearizing the equation, and traits measured on different scales can be directly compared in terms of proportional changes.

It is tempting to think that stomatal-area minimization and adaptation hypotheses could be tested by estimating the scaling exponent  $\beta$  among species. For the composite stomatal trait  $z_S$  to remain relatively constant among species that vary in  $d_S$  and  $a_S$ , the slope between  $d_S$  and  $a_S$  would need to be  $-\beta$ , which can be estimated using linear regression. For example, de Boer *et al.* (2016) estimated a scaling exponent close to unity, in that the standardized major axis (SMA) slope between stomatal density and size on a log-log scale was close to  $-1$ . They interpret this as evidence that species have diversified along an axis that increases  $g_{s,max}$  with minimum increase in  $f_s$ . There are two potential problems with this interpretation. First, since the seminal evolutionary quantitative genetic (EQG) theory of Lande (1979), biologists have recognized that the among-species covariance does not necessarily map directly onto within-species covariance or short-term responses to selection. A second challenge is estimating scaling exponents because different methods, most commonly ordinary least-squares (OLS) and SMA regression, yield different estimates that may not support the same hypothesis. There is broad consensus that the choice of estimation procedure should be guided primarily by biological rather than by statistical considerations.

To resolve these problems and advance our understanding of the selective forces that underlie inverse stomatal size-density scaling in vascular land plants, it requires not only data but also theory. Mathematical models generate quantitative predictions and inform statistical decisions. The hypotheses for size-density scaling are based on responses to selection within species, but many studies analyze patterns among species. It is not clear how among species patterns map onto within species processes. EQG is especially well-suited to address this gap in our understanding because it explicitly bridges micro- and macroevolutionary scales (Arnold *et al.*, 2001). Models also inform whether OLS, SMA, or neither can estimate desired parameter values under competing



hypotheses. Theory helps empiricists gain biological insight by selecting statistical methods appropriate for the biological questions. Further background information on EQG and regression methods is provided in Notes S2. The five primary objectives in this study:

- (1) provide specific mathematical assumptions about selection on stomatal density and size under the stomatal-area minimization and adaptation hypotheses;
- (2) deduce the mean and expected genetic (co)variance between stomatal density and size within species at equilibrium (mutation-selection balance) under both stomatal-area minimization and adaptation hypotheses, as well as a new extended hypothesis we call the ‘stomatal adaptation + bounded size’ hypothesis (see the [Materials and Methods](#) section);
- (3) deduce the short-term response to selection on  $g_{s,max}$  based on the equilibrium genetic (co)variance between stomatal density and size within species under competing hypotheses;
- (4) deduce the expected among-species (co)variance between stomatal density and size at stationarity using the Ornstein–Uhlenbeck (OU) model; and
- (5) compare expected parameter values under different hypotheses to observed values estimated using a phylogenetic comparative data set of 2408 vascular forest plant species.

## Materials and Methods

### Assumptions of competing hypotheses

>To ground our hypotheses in EQG theory, we assume different individual fitness functions for each hypothesis. Under the stomatal-area adaptation hypothesis, we assume that  $\log g_{s,max}$  is under stabilizing selection around an optimal  $\log g_{s,opt}$ :

$$W = W_{\max} \exp \left( - \frac{(\log g_{s,max} - \log g_{s,opt})^2}{2\omega} \right) \quad \text{Eqn 5}$$

where  $W$  is absolute fitness,  $W_{\max}$  is the maximum absolute fitness, and  $\omega$  is inversely proportional to the strength of selection. The  $g_{s,opt}$  is the optimal value of  $g_{s,max}$ . We do not explicitly model demography; hence, only the relative fitness determines the change in trait frequency. Therefore, we can ignore  $W_{\max}$  because it cancels out (i.e. we are assuming soft selection). Fitness is modeled as a Gaussian function because this is the simplest, most general, and mathematically tractable form of stabilizing selection and, therefore, widely used in EQG (Walsh & Lynch, 2018). As with many metric traits, the stomatal traits we consider are normally distributed on a log-transformed scale; therefore, we analyze log-transformed values throughout. The above assumptions about soft selection, the functional form of stabilizing selection, and log-transformation apply to all hypotheses.

One further, essential assumption that applies to all models is that  $g_{s,opt}$  is not static, but changes dynamically through time, independently in every species. The specific mathematical

assumptions are described later with additional detail in Notes S3. One key point to emphasize here is that we are not attempting to model the changes in the environment and/or other traits that cause  $g_{s,opt}$  to fluctuate. Rather, our treatment is phenomenological and additional analysis would be required to determine how  $g_{s,opt}$  responds to specific factors.

The additional assumption specific to the stomatal-area minimization hypothesis is that there is selection for decreased  $f_s$ , all else being equal:

$$W = W_{\max} \exp \left( - \frac{(\log f_s - \log f_{s,min})^2}{2\varphi_f \omega} - \frac{(\log g_{s,max} - \log g_{s,opt})^2}{2\omega} \right) \quad \text{Eqn 6}$$

Although Eqn 6 appears to imply stabilizing selection on  $\log f_s$ , it results in directional selection for lower  $f_s$  if  $f_{s,min}$  is sufficiently low relative to  $g_{s,opt}$  (see the [Results](#) section). Pure directional selection on  $f_s$ , in the absence of countervailing constraints on stomatal size, leads to selection for infinitesimal stomatal size and infinite stomatal density. The strength of selection on  $f_s$  relative to  $g_{s,opt}$  is determined by  $\varphi_f$ , where values above unity indicate weaker selection on  $f_s$  than  $g_{s,opt}$ . In our formulation, the stomatal-area adaptation hypothesis is a special case of the stomatal-area minimization hypothesis, where  $\varphi_f = \infty$ . Hence, there is a continuum between these hypotheses modulated by  $\varphi_f$ , where lower values indicate stronger selection on minimized  $f_s$  relative to optimizing  $g_{s,opt}$ .

As described in the [Results](#) section and Tables 1 and 2, both hypotheses make predictions that are inconsistent with data from vascular land plants. In particular, the lack of constraint on stomatal size leads to unrealistic interspecific variation in this trait. It is therefore necessary to add an additional assumption about physical bounds on stomatal size, which we do by assuming stabilizing selection around an optimal stomatal size  $a_{s,opt}$ :

$$W = W_{\max} \exp \left( - \frac{f_s}{2\varphi_f \omega} - \frac{(\log a_s - \log a_{s,opt})^2}{2\varphi_a \omega} - \frac{(\log g_{s,max} - \log g_{s,opt})^2}{2\omega} \right) \quad \text{Eqn 7}$$

We refer to this extended model as the ‘stomatal adaptation + bounded size’ hypothesis. In contrast to Eqn 6, pure directional selection of  $f_s$  does not lead infinitesimal stomatal size because of the stabilizing selection term in Eqn 7. It is not an alternative model because the first two models are nested within this more complex model. In this model, directional selection for lower  $f_s$  is limited by stabilizing selection on both stomatal size and  $g_{s,max}$ . While we do not explicitly model the fitness costs of extreme

	Assumptions		Microevolutionary equilibria within species		Response to selection on $g_{s,max}$
Hypothesis	Fitness function	Fluctuating fitness optima	Trait means	Genetic (co)variance	Trait response
$H_1$ : Stomatal-area adaptation	Eqn 5	$g_{s,max}$	Neutrally stable (Eqn S32)	Infinite variance (S3.3.1.5)	Density responds more negative (Figs S8–S10)
$H_2$ : Stomatal-area minimization	Eqn 6	$g_{s,max}$	Stable (Eqn S49)	Finite variance (S3.3.2.5)	Variable trait response and covariance (Figs S8–S10)
$H_3$ : Stomatal adaptation + bounded size	Eqn 7	$g_{s,max}$ and $A_S$	Stable (Eqn S61)	Finite variance (S3.3.3.5)	Variable trait response Covariance usually positive (Figs S8–S10)

**Table 2** Macroevolutionary predictions of competing hypotheses for the scaling of stomatal size and density and comparisons with estimates from phylogenetic comparative analyses of 2408 species of forest vascular plants.

Hypothesis	Macroevolutionary stationarity among species		Trait variance ratios at stationarity			Scaling exponent at stationarity	
	Trait means	Trait (co)variance	$V_{ds}^*/V_{as}^*$	$V_{ds}^*/V_{gmax}^*$	$V_{as}^*/V_{gmax}^*$	SMA	OLS
$H_1$ : Stomatal-area adaptation	Neutral plane (Eqn S37)	Neutral plane (Eqn S42)	4	$\infty$	$\infty$	0.5 (Eqn S43)	0.5 (Eqns S44, S45)
$H_2$ : Stomatal-area minimization	Stable (Eqn S51)	Stable (Eqn S52)	1	4	4	1 (Eqn S53)	1 (Eqns S54, S55)
$H_3$ : Stomatal adaptation + bounded size	Stable (Eqn S63)	Stable (Eqn S64)	Likely > 1	Likely > 1	Variable	Variable (Eqn S66)	Variable (Eqns S67, S68)
Estimates	See Table 3		1.26 [1.16–1.36]	1.27 [1.16–1.32]	1.01 [0.895–1.11]	0.523 [0.436–0.563] 2.42 [2.24–3.01]	1.12 [1.08–1.17]

The models differ in whether they predict the among-species means and (co)variance in stomatal density and size form a neutral plane or a stable point. The equation number in parentheses cross references the relevant part of the [Supporting Information](#). We also compared the ratio of trait variances at stationarity, where  $V_{ds}^*$ ,  $V_{as}^*$ , and  $V_{gmax}^*$  are the stationary variances in stomatal density, size, and  $g_{s,max}$  on log scales, respectively. These quantities can be deduced from the associated equations for the stationary variance. We compared these predictions to estimated values from vascular plants shown in the bottom row. The range in brackets indicated the 95 % bootstrap confidence interval. Finally, we predict the scaling exponent that would be estimated using ordinary least-squares (OLS) and standardized major axis regression methods. The equation numbers in parentheses cross references the relevant part of the [Supporting Information](#). For comparison, the estimates using each method on our data set are given in the bottom row. The OLS estimates differ depending on which variable is treated as the explanatory variable, stomatal density (upper) or stomatal size (lower). The range in brackets indicated the 95 % bootstrap confidence interval.

and analysis. From the stationary distribution, we calculate the expected interspecific scaling exponent one would estimate regressing density against size using OLS or SMA methods on a random sample of species.

For all hypotheses, two key assumptions are that the fluctuations in the adaptive optima are sufficiently small, and the genetic variance is sufficiently large that every species closely tracks the fluctuating optima. This enables us to approximate the stationary distribution of species trait means using the stationary distribution of the adaptive optima. In Notes [S3](#), we discuss the parameter space we analyzed in which this assumption would be met and confirmed the validity of the assumption using recursion simulations. An additional simplifying assumption is that the OU process is homogeneous among species and through time (i.e. there are no major discontinuities in the environment or between adaptive regimes *sensu*; Uyeda & Harmon [2014](#)).

Stomatal trait data from global forests

The stomatal data set of global forests represents a total of 2408 plant species from natural forests (Fig. [2](#)), including novel field data collected from Chinese forest communities and a compilation of published literature values. The data set includes representatives of all major vascular plant clades (angiosperms, gymnosperms, and pteridophytes) covering 201 families and 934 genera. For each species, we calculated  $g_{s,max}$  and  $f_s$ , where  $f_s$  is proportional to the stomatal pore area index, which defined as the product of  $D_s$  and stomatal length ( $L$ ) squared (Sack *et al.*, [2003](#)), because  $A_s = mL^2$  (Sack & Buckley, [2016](#)).

We calculated  $g_{s,max}$  (Eqn [2](#)) to water vapor at a reference leaf temperature ( $T_{leaf} = 25^\circ\text{C}$ ) following Sack & Buckley ([2016](#)). They defined biophysical and morphological constants as:

$$b = \frac{D_{wv}}{\nu}$$
Eqn 8

$$m = \frac{\pi c^2}{j^{0.5}(4hj + \pi c)}$$
Eqn 9

$b$  is the diffusion coefficient of water vapor in air ( $D_{wv}$ ) divided by the kinematic viscosity of dry air ( $\nu$ ).  $D_{wv} = 2.49 \times 10^{-5} \text{ m}^2 \text{ s}^{-1}$  and  $\nu = 2.24 \times 10^{-2} \text{ m}^3 \text{ mol}^{-1}$  at  $25^\circ$  (Monteith & Unsworth, [2013](#)). For kidney-shaped guard cells,  $c = b = j = 0.5$ ; for dumbbell-shaped guard cells in the Poaceae,  $c = b = 0.5$  and  $j = 0.125$ . We used the species average  $g_{s,max}$  and  $f_s$  for all analyses.

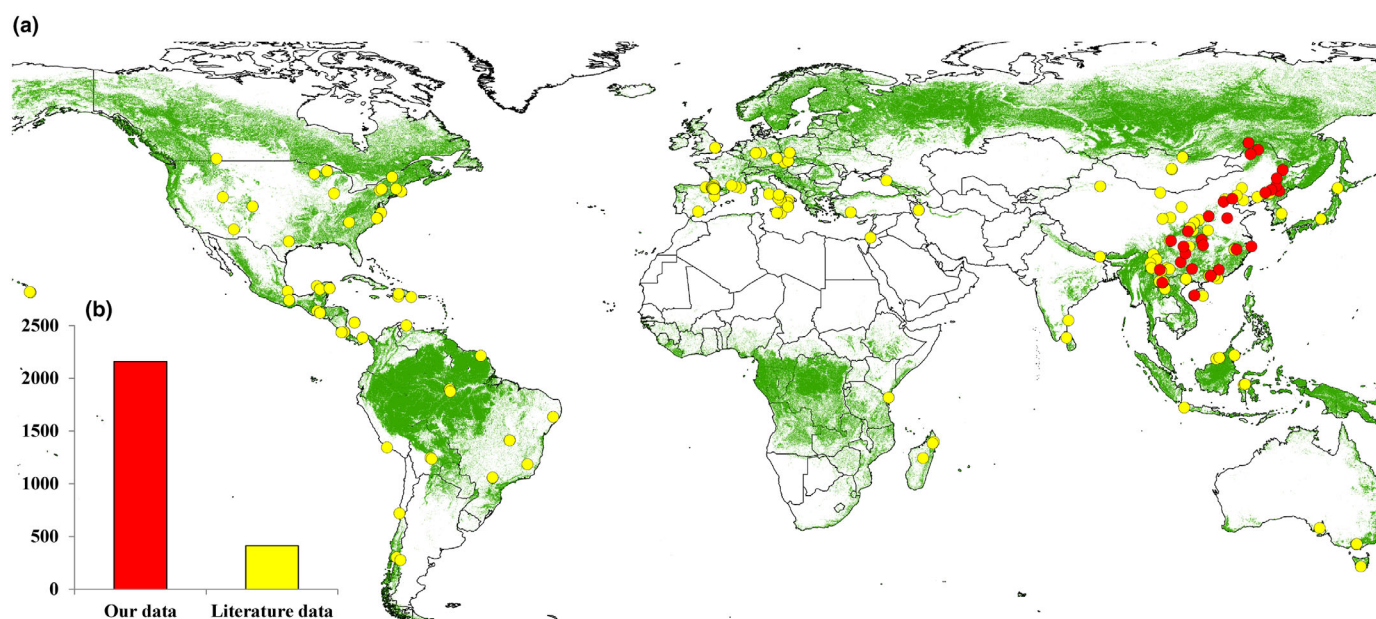
Phylogenetic comparative analyses

We implemented a multivariate phylogenetic comparative method to estimate scaling between stomatal density and size using the OU model. This provides an estimate of scaling as directly connected to our model predictions as possible. Our analysis revealed that neither OLS nor SMA regression, even when incorporating phylogenetic structure, would estimate an

**Table 3** Phylogenetic comparative estimates of among species stomatal trait means and variances at stationarity.

Macroevolutionary stationarity among species			
Trait mean	Trait (co)variance		
	Parameter estimates	Implied macroevolutionary parameters	
<b>Stomatal density</b>			
Angiosperms 174 [167–180] mm <sup>-2</sup>	$\begin{bmatrix} \hat{V}_{dS}^* & \\ \hat{V}_{dS,aS}^* & \hat{V}_{aS}^* \end{bmatrix} = \begin{bmatrix} 0.585 & \\ -0.242 & 0.465 \end{bmatrix}$	$\begin{bmatrix} \text{Var}(a_{S,\text{opt}}) & \\ \text{Cov}(a_{S,\text{opt}}, \log(g_{S,\text{opt}})) & \text{Var}(\log(g_{S,\text{opt}})) \end{bmatrix} = \begin{bmatrix} 0.465 & \\ -0.011 & 0.458 \end{bmatrix}$	
Gymnosperms 80.8 [62.3–105] mm <sup>-2</sup>	$\left( \begin{bmatrix} 0.547 & \\ -0.267 & 0.434 \end{bmatrix} - \begin{bmatrix} 0.633 & \\ -0.204 & 0.500 \end{bmatrix} \right)$	$\left( \begin{bmatrix} 0.434 & \\ -0.029 & 0.429 \end{bmatrix} - \begin{bmatrix} 0.500 & \\ 0.030 & 0.536 \end{bmatrix} \right)$	
Pteridophytes 73.1 [63.4–85.7] mm <sup>-2</sup>			
<b>Stomatal size</b>			
Angiosperms 250 [243–257] μm <sup>2</sup>			
Gymnosperms 677 [533–881] μm <sup>2</sup>			
Pteridophytes 762 [665–865] μm <sup>2</sup>			

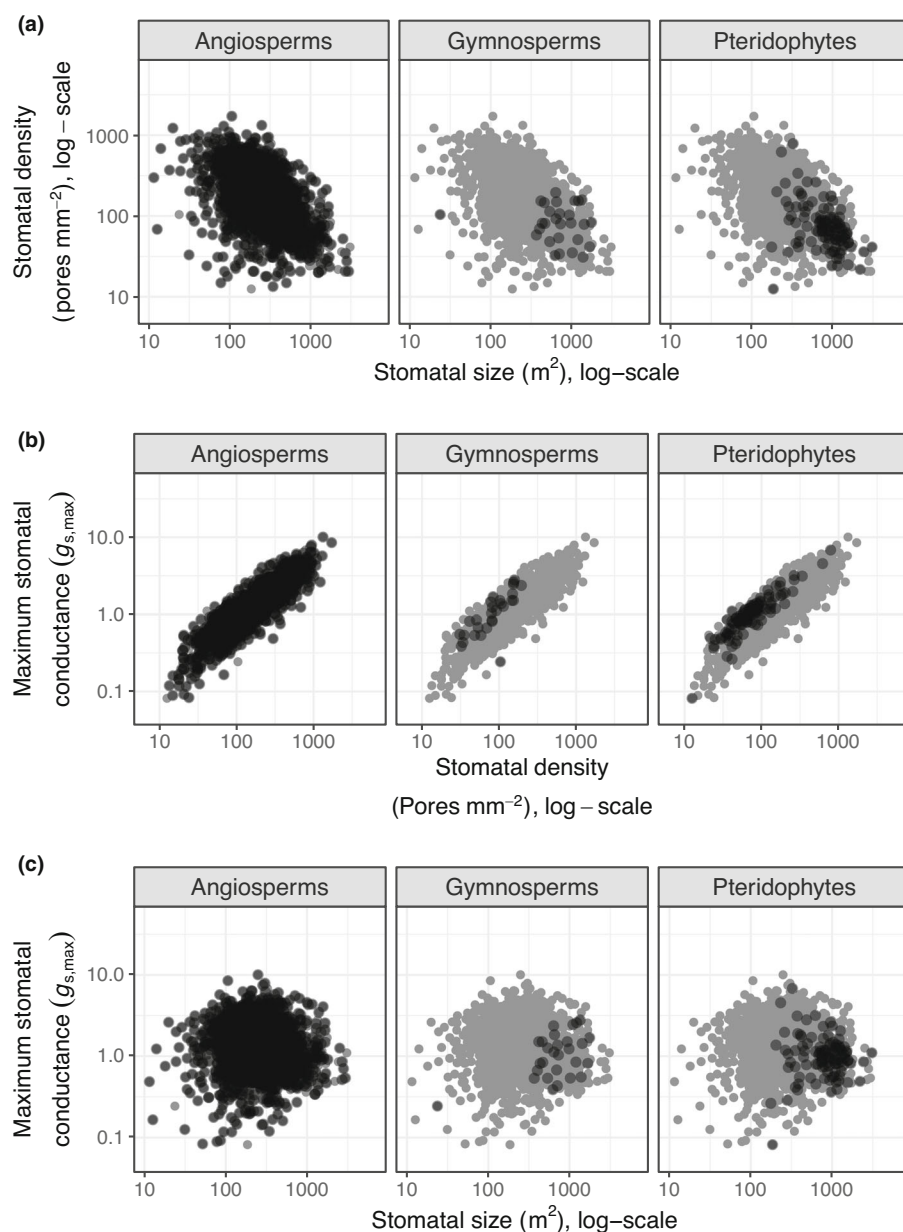
The trait means are estimated separately for Angiosperms, Gymnosperms, and Pteridophytes on a log scale, but transformed to measurement units here. The range in brackets is the 95% bootstrap confidence interval. Note that stomatal size refers to the area of the entire stomatal complex (guard cells and pore), not the pore area. We estimated a single stationary trait covariance matrix for all groups. The matrices in brackets are the 95% bootstrap confidence intervals for each element of the matrix.  $\hat{V}_{dS}^*$  and  $\hat{V}_{aS}^*$  are the stationary variances for stomatal density and size, respectively, on a log scale;  $\hat{V}_{dS,aS}^*$  is the stationary covariance on a log scale. The implied macroevolutionary parameters are the stationary variance in optimal stomatal size ( $\text{Var}(a_{S,\text{opt}})$ ), optimal  $g_{S,\text{max}}$  ( $\text{Var}(\log(g_{S,\text{opt}}))$ ), and their covariance ( $\text{Cov}(a_{S,\text{opt}}, \log(g_{S,\text{opt}}))$ ). See the Note S3 for explanation of implied macroevolutionary parameters.

**Fig. 2** Geographic distribution of sampling sites (a) and the number of plant species (b) in this study. Green areas indicate forested biomes.

interpretable scaling exponent when stomatal size is bounded (see the Results section). Therefore, we estimated the among-species covariance directly and report the expected OLS and SMA scaling exponent estimates if one regressed  $d_S$  on  $a_S$  (Eqns S16 and S17 in the Supporting Information). The choice of regressing  $d_S$  on  $a_S$  or vice versa is arbitrary in this case and is unrelated to a cause-and-effect relationship.

We used the Plant List (<http://www.theplantlist.org>) to confirm species names; then, we assembled a synthetic phylogeny using S.PhyloMaker (Qian & Jin, 2015). We estimated the parameters of the multivariate OU using the Ornstein–Uhlenbeck for Comparative Hypotheses (OUCH) model in the R package MVSLOUCH v.2.7.6 (Bartoszek *et al.*, 2012, 2024) and the phylogeny described above. To test for shifts in OU





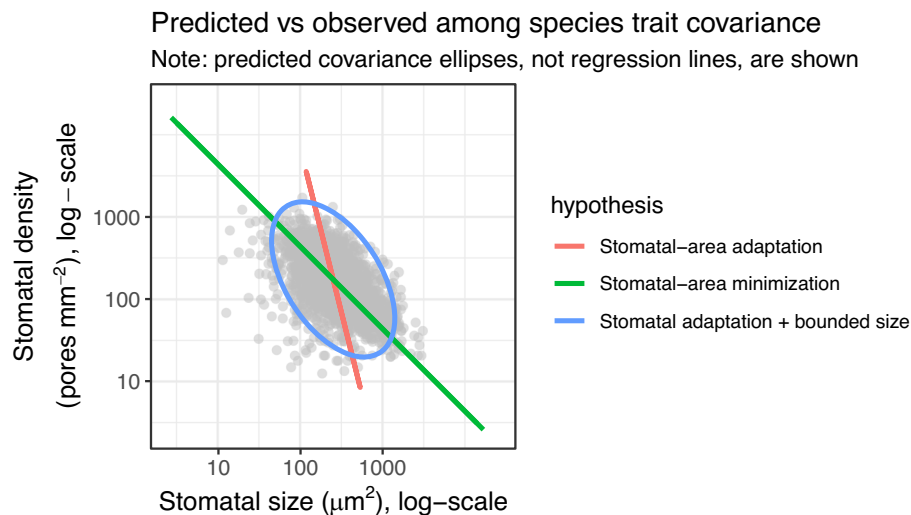
**Fig. 3** Inverse stomatal size–density scaling across vascular land plants and its relationship with anatomical maximum stomatal conductance ( $g_{s,max}$ ). In all panels, dark points represent species mean trait values from the focal group; gray background points are from all groups for comparison. The panels show the relationship between (a) stomatal size and density, (b) stomatal density and  $g_{s,max}$ , and (c) stomatal size and  $g_{s,max}$  for the same global data set of forest plants. All panels are on a log–log scale.

parameters between major groups, we estimated separate mean trait values for Angiosperms, Pteridophytes, and Gymnosperms. From the estimated multivariate OU parameters, we calculated the bivariate normal stationary distribution. Using the stationary distribution is appropriate here because the phylogenetic half-life (see the [Results](#) section) is much shorter than the root of the phylogeny, meaning there is sufficient time for stationarity to have been reached. Our covariance estimates account for phylogenetic nonindependence because the OU parameters were estimated using the phylogeny. We estimated 95% confidence intervals for all parameters using 1000 parametric bootstrap samples generated by simulating from the best-fit model and refitting. One gymnosperm species, *Torreya fargesii* Franch. (Taxaceae), had substantially lower stomatal size than would be predicted from its density (Fig. 3a). The results of the paper did not change if this outlier was excluded. Therefore, we

excluded this species from statistical analyses but show it in the figure for completeness. All data were analyzed in R v.4.4.2 (R Core Team, 2025).

## Results

The expression of stomatal-area adaptation and minimization hypotheses in the mathematics of EQG provides specific quantitative predictions about the central tendencies and (co)variances of stomatal density and size, both within and among species. Table 1 summarizes the model predictions within species; Table 2 summarizes the model predictions among species and compares them to estimates from phylogenetic comparative analysis of vascular forest plants. Complete derivations are available in Notes S3. In this section, we do not exhaustively analyze all parameter space but rather focus on whether or under what



**Fig. 4** The 'stomatal adaptation + bounded size' hypothesis, with tuning, predicts among species covariation in stomatal size and density, whereas simpler hypotheses do not. The gray points represent species mean trait values, as shown in Fig. 3. Colored polygons are 99% quantile ellipses predicted by each hypothesis, albeit tuned based on empirical constraints. The red and green ellipses appear as straight lines because these hypotheses incorrectly predict zero variance orthogonal to the main axis of variation. By contrast, the last hypothesis predicts significant variance orthogonal to the main axis of variation that is not due simply to measurement error, but real biological variation among species. For all hypotheses, the predicted ellipses are tuned to match estimated stationary trait means and variance in anatomical maximum stomatal conductance ( $g_{s,max}$ ). The plot is on a log-log scale.

conditions the hypotheses make realistic predictions. After that, we compare these predictions to parameter estimates from a phylogenetic comparison of vascular forest plants.

#### Neither stomatal-area adaptation nor minimization hypotheses are inconsistent with observations

The stomatal-area adaptation hypothesis assumes that selection only optimizes  $g_{s,max}$  and that the optimal  $g_{s,max}$  fluctuates following a univariate OU process. Note that there are many other assumptions in the EQG model, but these are the assumptions that differentiate this hypothesis from the other hypotheses. Under this hypothesis, stomatal density and size evolve as a by-product of selection on  $g_{s,max}$ . Problems with the stomatal-area adaptation hypothesis arise because there are no bounds on stomatal size. This results in infinite genetic variance in both stomatal density and size despite a well-defined genetic correlation between them (Table 1). Even if we assume finite genetic variance within species, the among-species distribution reaches a stable covariance structure, but it is not stationary. Hence, the hypothesis makes precise, testable predictions about SMA and OLS slopes (Table 2), but cannot be sufficient because it predicts infinite variances.

The stomatal-area minimization hypothesis adds directional selection for decreased  $f_s$  along with stabilizing selection on  $g_{s,max}$ . Therefore, this hypothesis constrains the within- and among-species variances to be finite (Tables 1, 2), but its predictions about interspecific scaling are inconsistent with the data. Specifically, this model predicts zero variance in  $f_s$  among species (Eqn S51 in the Supporting Information), which is inconsistent with observations (Fig. S1). Consequently, the hypothesis predicts much greater variance in  $a_s$  and  $d_s$  relative to  $g_{s,max}$  compared

with observations (Table 2). It also predicts a complete negative correlation between size and density among species (Fig. 4). These inconsistencies between predictions and observations arise because there are no bounds on how small or large stomata can be, an assumption that is not supported by our understanding of stomatal function in real leaves (see the Discussion section).

#### The stomatal adaptation + bounded size hypothesis makes predictions consistent with observations

The inconsistencies between the previous two hypotheses and data forced us to modify them by making the additional, albeit realistic, assumption that stomatal dimensions are bounded by selection against extremely large or small pore size. We analyzed two versions of this hypothesis. The simpler version of this hypothesis assumed constant optimal stomatal size,  $a_{s,opt}$ . This simpler version predicts variance in  $d_s$ ,  $g_{s,max}$ , and  $f_s$ , but no variance in  $a_s$ , as noted previously in the Materials and Methods section. This is inconsistent with the wide variation in stomatal size and implies that  $a_{s,opt}$  must be a variable, not a constant. In other words, optimal stomatal size varies, presumably depending on the environment and other traits of a species. Although we did not explicitly model exogenous factors that would explain variation in  $a_{s,opt}$ , we review mechanistic hypotheses in the Discussion section. With the assumption that  $a_{s,opt}$  fluctuates according to an OU process, the model can predict realistic within-species patterns of trait means and covariances (Table 1), finite and realistic variances among species, and inverse size–density scaling (Table 2; Fig. 4). Importantly, the hypothesis is internally consistent in that the parameter values necessary to obtain realistic predictions are similar to those estimated using phylogenetic comparative methods (will be discussed later). The problem with this

hypothesis is that its predictions cannot be derived from first principles and must be tuned by parameter estimates from the data it is meant to explain. Therefore, the fact that the model predicts trait (co)variance that matches the observation (Fig. 4) does not necessarily indicate strong support for the hypothesis because it could be an artifact of tuning the parameters to match the data. To better test how well the hypothesis predicts trait (co)variance, we need an independent way to either predict from first principles or estimate the limits on optimal  $g_{s,max}$ ,  $f_s$ , and  $a_s$  in vascular land plants.

### Genetic covariance within species

Under the stomatal-area adaptation hypothesis, the genetic variance increased without bound to a stable covariance structure (Eqn S34) in which stomatal density and size were perfectly negatively correlated. The other hypotheses resulted in finite genetic (co)variance matrices (Table 1) that we solved for numerically over a range of parameter space. The equilibrium genetic variance increases as mutational variance increases and the strength of selection on  $g_{s,max}$  decreases (Figs S2, S3). Similarly, stronger selection to minimize  $f_s$  (Fig. S4) or optimize stomatal size (Fig. S5) reduces genetic variance. Under the stomatal-area minimization hypothesis, the genetic correlation between traits is always negative (Fig. S6), whereas under the stomatal adaptation + bounded size hypothesis, the correlation can be positive if the mutational correlation is sufficiently positive (Fig. S7).

### Microevolutionary responses to selection can be decoupled from macroevolutionary patterns

In EQG, the short-term (single generation) response to multivariate selection is determined by the additive genetic (co)variance between traits and the selection gradient (Lande, 1979). The additive genetic (co)variance at equilibrium is determined by the mutational variance and the curvature of the fitness surface. There is no mathematically necessary relationship between the genetic variance, the selection gradient, and the long-term macroevolutionary movement of adaptive optima. Indeed, our models predict these factors are generally decoupled from one another for stomatal size and density except under special cases and coincidental areas of parameter space. This is evident by the fact that the genetic (co)variance matrix does not appear in the equations for the equilibrium among-species (co)variance except the stomatal-area adaptation hypothesis (compare sections S3.3.1.6 to S3.3.2.6 and S3.3.3.6 in the Supporting Information).

The stomatal-area adaptation hypothesis is one special case in which the genetic variance, response to selection, and among-species covariance are aligned. It predicts a very specific negative genetic (co)variance structure (Table 1) that is insensitive to the mutational variance. It also predicts that the response to selection on  $g_{s,max}$  always results in opposing response directions for stomatal density and size (Figs S8–S10). Both the genetic (co)variance and response to selection are aligned with the among-species (co)variance (Table 1). For the other two hypotheses, there are few generalizations. For example, the stomatal-area minimization

hypothesis predicts a greater response in stomatal density relative to size, but the covariance of the response is sensitive to the genetic covariance (Fig. S8). Furthermore, size can respond more strongly if it has greater genetic variance (Fig. S9). Under the stomatal adaptation + bounded size hypothesis, stomatal size tends to respond more strongly to selection on  $g_{s,max}$  than density, but both generally respond in a coordinated manner under most parameter space (Figs S8, S9). As selection to minimize  $f_s$  strengthens, stomatal size responds more strongly than density because this allows leaves to increase  $g_{s,max}$  without increasing  $f_s$  as much (Fig. S10).

Except in special cases, these short-term responses to selection on  $g_{s,max}$  do not predict the interspecific (co)variance in stomatal density and size. This is because the mean trait values at *equilibrium* in a species do not necessarily depend on the specific short-term dynamics. We assumed sufficiently large population sizes and mutational variance relative to fluctuations in the adaptive optimum that all species could closely track their fitness optima. These assumptions will not always be met in nature, but this adaptive tracking model best fits observations for many traits in many species. With that assumption, we derived expected stationary distributions of stomatal trait values among species, and these equations do not include any microevolutionary parameters (Table 2). Simulations described in Notes S4 showed this approximation is valid over a wide range of variation in mutation (Fig. S11), mutational correlation (Fig. S12), and strengths of selection on  $g_{s,max}$  (Fig. S13),  $f_s$  (Fig. S14), and  $a_s$  (Fig. S15).

### Phylogenetic comparative estimates of stomatal trait variation among species

We estimated stomatal size–density scaling in 2408 forest plant species from new field-collected samples over 28 sites in China and global synthesis of data from the literature (Fig. 2). Stomatal density was lower and size was higher in gymnosperms and pteridophytes than in angiosperms (Table 3). Among all groups, stomatal density varies more than size, and there is strong negative covariance (Table 3; Fig. 3). As mentioned previously, the interspecific trait variance is inconsistent with either the stomatal-area adaptation or minimization hypotheses. The stomatal-area adaptation hypothesis predicts that the SMA and OLS estimates of the scaling exponent should be close to 0.5 and the variance of  $a_s$  and  $d_s$  should be much greater than that of  $g_{s,max}$  (Table 2). The OLS estimate was close to 0.5 only when size is used as the explanatory variable, but the other predictions were not supported. The stomatal-area minimization hypothesis predicts that SMA and OLS estimates of the scaling exponent should be close to unity and the variance of  $a_s$  and  $d_s$  should be similar and 4× greater than the variance in  $\log(g_{s,max})$  (Table 2). The SMA estimate was close to unity, similar to the results of de Boer *et al.* (2016), but the other predictions were not supported. The stomatal adaptation + bounded size hypothesis predicts among species trait (co)variances that are quantitatively consistent with the data as well as within species trait (co)variances that are qualitatively plausible. Based on the stationary trait distribution (Eqn S64), the phylogenetic

## Discussion

We identified two distinct hypotheses for inverse stomatal size–density scaling that we label ‘stomatal-area adaptation’ and ‘stomatal-area minimization’ hypotheses. We are not the first to describe these hypotheses, but the labels and associated fitness functions clarify key assumptions and testable predictions. Under the stomatal-area adaptation hypothesis, selection only optimizes  $g_{s,\max}$  (Eqn 5), meaning that  $f_S$  will increase when selection favors greater  $g_{s,\max}$ , and vice versa for selection to lower  $g_{s,\max}$ . Inverse size–density scaling emerges because there is an ellipse in density–size space for a given amount of variation in  $g_{s,\max}$ , either within or among species. However, the size of this ellipse is unstable. Variance in density and size grows unchecked, and their covariance becomes negative because of limited variance in  $g_{s,\text{opt}}$ . This model clearly shows that additional bounds on stomatal traits must be acting in nature. The stomatal-area minimization hypothesis adds a constraint on  $f_S$ , favoring changes in  $g_{s,\max}$  that minimize increases in  $f_S$  (Eqn 6). The  $g_{s,\max}$  and  $f_S$  can evolve semi-independently because of differential scaling and apparent

In comparative ecology, it is commonly assumed that when trait–trait relationships are governed by equations with the form



shown in Eqns 1 and 2, scaling relationships can be estimated using regression methods, typically SMA or OLS. Our evolutionary analysis reveals that this assumption holds when trait variance is unconstrained but breaks down if there are additional bounds on trait variance (Table 2). In this case, bounds and selection in stomatal size alter the covariance with stomatal density. With stabilizing selection on size, stomatal density is relatively less constrained in responding to selection, resulting in a weaker association between size and density than would otherwise occur. We visualized this by plotting the among-species covariance ellipses predicted at stationarity by different hypotheses but constrained by empirical estimates of the mean trait values and variance in  $g_{s,max}$  (Fig. 4). The simpler hypotheses predict highly constrained trait covariance along a single axis of variation, whereas the stomatal adaptation + bounded size hypothesis accommodates the variation orthogonal to the main axis of variation. Even if the simpler hypotheses were correct, empirical data sets would be variable because of measurement error. However, the biological variation in our very broad survey of forest plants is large relative to measurement error. Uncovering the evolutionary forces that gave rise to stomatal anatomical diversity among vascular forest plants cannot be determined from bivariate slopes alone. Instead, we recommend analyzing the entire (co)variance within and among species and estimating key micro- and macroevolutionary parameters that can inform the relative importance of selection  $g_{s,max}$ ,  $f_s$ , and stomatal size (Tables 1, 2).

Since our model demonstrates that interspecific scaling alone cannot uniquely identify which selective processes shape stomatal size and density, we considered the entire covariance structure in a phylogenetic comparative framework. Under the stomatal adaptation + bounded size hypothesis, fluctuating selection on  $g_{s,max}$  and  $a_s$ , along with selection to minimize  $f_s$ , determines a unique covariance among stomatal density and size (Eqn S64). These predictions enable a novel interpretation of the estimated covariance in stomatal density and size (Table 3). Specifically, the covariance pattern is consistent with independently fluctuating optimal  $g_{s,max}$  and  $a_s$  (the range of implied covariances overlaps zero, Table 3). This interpretation makes sense of three observations that would not otherwise be obvious. First, it explains why stomatal density varies more than stomatal size among species under most areas of parameter space. The hypothesis also explains two related observations –  $d_s$  and  $g_{s,max}$  positively covary (Fig. 3b), whereas  $a_s$  and  $g_{s,max}$  covary little within major groups (Fig. 3c). Based on Eqn 2, a naïve null model might predict similar variance in  $a_s$  and  $d_s$  and positive covariance between both  $a_s$  and  $d_s$  with  $g_{s,max}$ . However, when the adaptive optima for  $a_s$  and  $g_{s,max}$  fluctuate independently and species track those optima, our model predicts no covariance between  $a_s$  and  $g_{s,max}$  at stationarity. This results in greater variance in  $d_s$  and stronger positive covariance between  $d_s$  and  $g_{s,max}$  because  $d_s$  is less constrained to respond to selection on  $g_{s,max}$ . Within species,  $a_s$  often cannot respond as much to selection on  $g_{s,max}$  because if  $a_s$  deviates far from  $a_{s,opt}$ , fitness will decline;  $d_s$  is relatively free to respond since it only indirectly affects fitness through its effect on  $g_{s,max}$ . At a macroevolutionary scale, this results in greater variance (less constraint) in  $d_s$  and positive covariance between  $d_s$  and  $g_{s,max}$ . It

might seem counterintuitive that the optima for  $a_s$  and  $g_{s,max}$  fluctuate independently since small stomata are necessary to achieve the greatest observed  $g_{s,max}$  (Franks & Beerling, 2009a,b). The shift toward higher stomatal density and smaller size in Angiosperms (Table 3), whatever its ultimate cause (e.g. genome downsizing), contributes to this pattern. But our analysis shows that loosening the lower bounds on guard cell size simultaneously increases the variance in  $a_s$  and makes the covariance with  $d_s$  more negative (Eqn S64). Thus, changing constraints on stomatal size can open space for leaves with high  $g_{s,max}$  and small stomata even when they evolve independently.

A major implication of our models is that bounds on stomatal size and fluctuating selection on optimal stomatal size within these bounds play an important role in shaping the size–density scaling relationship. Minimum guard cell size is a function of genome size and its scaling relationships with nuclear and primordial cell size (Beaulieu *et al.*, 2008; Šimová & Herben, 2012; Roddy *et al.*, 2020). An alternative hypothesis is that extremely small stomata are deleterious because they inhibit gas exchange when collisions with guard cell walls dominate the diffusion (Hodgson *et al.*, 2010). This effect becomes important and reduces stomatal conductance when pore width approaches 0.5–1  $\mu\text{m}$  (Csiro, 1983). Hence, in plant species with large genomes, guard cell size may set the minimum stomatal pore length, whereas in species with small genomes, selection against extremely small pores may set the minimum stomatal size. On the other extreme, large guard cells are likely deleterious because the lower surface area to volume ratio of larger cells slows the rate of stomatal closure (Drake *et al.*, 2013; Raven, 2014), all else being equal. Plants can partially compensate for guard cell geometry by increasing the rate of transport per unit surface area (Lawson & Blatt, 2014), but this would likely involve costs such as constructing and maintaining transporter proteins. In summary, it is near certain that factors such as genome size, diffusion, and guard cell mechanics set bounds on stomatal size independent of the effect on  $g_{s,max}$ .

Within these overall bounds on guard cell size, stomatal size per se is likely under selection independent of its effect on  $g_{s,max}$ , but the adaptive significance of variation in stomatal size is not well understood. One hypothesis is that smaller guard cells close and open faster in response to environmental stimuli because of their greater surface area to volume ratio (Drake *et al.*, 2013), as discussed in the previous paragraph. Faster responses allow leaves to closely track variable light, humidity, and other factors, keeping stomatal conductance closer to its short-term optimum. All else being equal, faster response increases water-use efficiency for a given operational stomatal conductance (Drake *et al.*, 2013; Lawson & Viallet-Chabrand, 2019). Smaller stomata respond faster to environmental change within groups of closely related species (Drake *et al.*, 2013; Yoshiyama *et al.*, 2024), but size is less predictive of speed in broader taxonomic comparisons (Elliott-Kingston *et al.*, 2016; McAusland *et al.*, 2016; Haworth *et al.*, 2018). These observations suggest that the speed of stomatal response is modulated by factors other than guard cell size, such as the rate of membrane transport and mechanical advantage of adjacent epidermal cells (Lawson & Blatt, 2014). This hypothesis predicts that selection for faster stomatal response would

result in smaller stomata over short or moderate evolutionary timescales. Consistent with those predictions, smaller stomata are associated with greater water-use efficiency in more arid populations of *Arabidopsis* (Dittberner *et al.*, 2018) and *Populus trichocarpa* (Klein *et al.*, 2025). However, it is unclear what the cost of small stomata is and, hence, why selection would favor larger stomata, holding  $g_{s,max}$  constant. Faster response in small guard cells may be more energetically demanding (Raven, 2014) and therefore selected against in more predictable and/or less variable environments. However, this hypothesis has not been tested to our knowledge. Despite incomplete understanding of how selection operates on stomatal size, associations between stomatal size and environment within and among species indicate that this trait is under selection independent of its effects on  $g_{s,max}$ .

## Conclusion

We addressed the long-standing observation of inverse stomatal size–density scaling among plant species by considering evolutionary hypotheses based on optimizing maximum stomatal conductance ( $g_{s,max}$ ) and minimizing epidermal surface area allocated to stomata ( $f_s$ ). Considering only these factors predicts more variance in size and density relative to  $g_{s,max}$  than we observe in a global data set of forest plants. Selection and bounds on stomatal size reduce variation in both size and density relative to  $g_{s,max}$ . If selection on  $g_{s,max}$  and stomatal size fluctuates independently, this can explain both inverse size–density scaling and the strong positive covariance between  $g_{s,max}$  and  $f_s$ . The estimated scaling exponents using common statistical methods are sensitive to bounds on stomatal size and therefore future research testing these hypotheses should complement interspecific comparative studies with quantitative genetic and phenotypic selection approaches.

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## Competing interests

None declared.

## Author contributions

NH and GY designed field sampling. CDM, LS, NH, CL and HJB conceived the initial ideas. CL, NH, YL, JZ, ZZ, ML and

LX collected the data; CL wrote the first draft, and CDM contributed the final mathematical derivations, data analysis and wrote the final manuscript. LS, HJB, CL, NH, GY and XH revised the manuscript. All authors gave final approval for publication. CL and CDM contributed equally to this work.

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## Data availability

The raw data and R code supporting the findings of this study are openly available in Zenodo at doi: [10.5281/zenodo.17155106](https://doi.org/10.5281/zenodo.17155106).

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**Table S3** Range of parameter values for numerically solving within species genetic covariance under  $H_3$ .

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