

Systematic Reanalysis of KMTNet Microlensing Events, Paper II: Two New Planets in Giant-Source Events

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ABSTRACT

In this work, we continue to apply the updated KMTNet tender-love care (TLC) photometric pipeline to historical microlensing events. We apply the pipeline to a subsample of events from the KMTNet database, which we refer to as the giant source sample. Leveraging the improved photometric data, we conduct a systematic search for anomalies within this sample. The search successfully uncovers four new planet-like anomalies and recovers two previously known planetary signals. After detailed analysis, two of the newly discovered anomalies are confirmed as clear planets: KMT-2019-BLG-0578 and KMT-2021-BLG-0736. Their planet-to-host mass ratios are $q \sim 4 \times 10^{-3}$ and $q \sim 1 \times 10^{-4}$, respectively. Another event, OGLE-2018-BLG-0421 (KMT-2018-BLG-0831), remains ambiguous. Both a stellar companion and a giant planet in the lens system could potentially explain the observed anomaly. The anomaly signal of the last event, MOA-2022-BLG-038 (KMT-2022-BLG-2342), is attributed to an extra source star. Within this sample, our procedure doubles the number of confirmed planets, demonstrating a significant enhancement in the survey sensitivity.

1. INTRODUCTION

Gravitational microlensing is the effect observed in the light from a distant source star being lensed by a foreground stellar or planetary object. This phenomenon allows us to detect extrasolar planetary systems by observing the characteristic distortions in the light curves of background stars (Mao & Paczyński 1991; Gould & Loeb 1992). Microlensing surveys, including the Optical Gravitational Lens Experiment (OGLE, Szymański et al. 2011; Udalski et al. 2015), Microlensing Observations in Astrophysics (MOA, Bond et al. 2001; Sumi et al. 2003), Wise (Shvartzvald et al. 2016), and the Korea Microlensing Telescope Network (KMTNet, Kim et al. 2016), along with the followup observations powered by them, have yielded over 200 exoplanet discoveries, proving that microlensing is a powerful tool for planet detection.

KMTNet, with its three 1.6 m telescopes and 4 deg² field-of-view cameras at the Cerro Tololo Inter-American Observatory (CTIO) in Chile (KMTC), the South African Astronomical Observatory (SAAO) in South Africa (KMTS), and the Siding Spring Observatory (SSO) in Australia (KMTA), has been one of the most productive microlensing surveys since its commissioning in 2015. However, the current KMTNet survey has not reached its full potential in terms of detection sensitivity. This limitation is primarily due to the accuracy of photometric data.

Because of the intrinsic low probability (about $10^{-6} \sim 10^{-5}$ events per monitored star per year (Sumi et al. 2013; Mróz et al. 2019)), any microlensing survey needs to observe a huge number of stars ($10^8 \sim 10^9$) to ensure a sufficient number of detections. Additionally, planetary signals are perturbations on the main light curves. The short time scale of such signals requires dense observation over time. To meet these combined requirements, microlensing surveys must handle massive datasets with dense time sampling. Therefore, such surveys often establish a multi-level photometric pipeline, typically consisting of an automated, general-purpose pipeline and a less automated, customized pipeline.

In KMTNet, the first levels are the DIAPL (Wozniak 2000) pipeline and the online pySIS (Albrow et al. 2009) pipeline. The DIAPL pipeline aims to efficiently produce the light curve of all stars within the observed field to find microlensing events. The online pySIS pipeline is then performed on those discovered events and aims to better characterize the events and detect possible anomalous signals. Both of the above pipelines operate in real-time. However, their inability to fully account for the systematic errors in KMTNet images (e.g., highly irregular point spread functions, large seeing variations) limits the accuracy of the resulting data.

The second level is the Tender-Love-Care (TLC) pySIS pipeline, which aims to produce better data for detailed analysis and publication of the anomalies. Following recent updates by Yang et al. (2024), this pipeline is now more automatic and produces significantly more stable and accurate photometric data for KMTNet images.

Yang et al. (2024) (hereafter Paper I) applied the updated TLC pySIS pipeline to some historical microlensing events from KMTNet and searched the new photometry for anomalous signals. This approach, using the improved data, successfully identified a new anomalous signal in event MOA-2019-BLG-421. Although the signal was subtle and highly degenerate, it demonstrated that this procedure can indeed enhance the detection efficiency of anomalous signals.

In this work, we continue the idea of Paper I to systematically apply the updated TLC pySIS pipeline to historic KMTNet events and search for new signals. As a step further than Paper I, we have two goals. The first is to find clear new planets, as opposed to just planet candidates as in Paper I. The second is to evaluate the improvements of this procedure on planet detection efficiency. The results will determine whether it is worthwhile to apply this procedure to all historic (and future) KMTNet events. To achieve these two goals and limit the workload, we decided to start with a small subsample which we estimated should have a size of \sim hundreds of events.

A straightforward choice would be the high-magnification (HM) sample (e.g., Gould et al. 2010), as HM events are inherently more sensitive to planets. However, planetary

signals in HM events are relatively subtle, say $\Delta A/A \sim 1\%$, where A is the flux magnification due to microlensing. Achieving the necessary photometric precision to capture such signals often requires overcoming various sources of systematic noise, such as the detector’s non-linearity and photon-response non-uniformity. Therefore, anomaly signals in HM events, even when identified, often require independent follow-up observations for confirmation. Furthermore, HM planet candidates often suffer strong degeneracies. Follow-up observations can help break those degeneracies (e.g., Zhang et al. 2023). Therefore, HM samples are better constructed through follow-up programs (Gould et al. 2010; see also Zang et al. 2021a; Yang et al. 2022).

In this work, we choose a giant-source microlensing event sample. While not as sensitive as HM events, giant-source events still have greater sensitivity to planets than the median event. This is because a giant source has a significant finite-source effect, which increases its probability of encountering caustics. Moreover, planetary caustics are typically larger than central caustics (Chung et al. 2005; Han 2006), resulting in the majority of planetary signals arising in the low-magnification region, where they produce larger $\Delta A/A$ (e.g., “Hollywood” events like Yee et al. 2021). Additionally, giant sources are intrinsically brighter, enabling more accurate photometry.

We construct the giant source microlensing event sample using KMTNet survey data from the 2016 to 2022 seasons¹. The subsample contains 352 events. We apply the updated TLC pySIS pipeline to all events in this sample to yield accurate light curves. We subsequently developed an anomaly search algorithm to identify any anomalous signals within these events, successfully uncovering four new planet-like anomalies. The structure of this paper is as follows. Section 2 presents some new updates of the photometric pipeline. Section 3 describes the sample selection. After applying the pipeline, we describe the anomaly search in Section 4. In Section 5, we present the modeling of the newly discovered signals. Finally, Section 6 summarizes our findings based on this sample. A detailed analysis of the sample’s sensitivity and statistical properties will be presented in future work.

2. PHOTOMETRY PIPELINE

Building upon the work presented in Paper I (see Section 2 in Yang et al. 2024, for the overall procedures and details of the photometry pipeline), we have further refined the photometric pipeline to enhance its self-consistency and facilitate its application to a larger sample.

First, for the photometric algorithm, we now consider (linear) inter-pixel correlations. In Paper I, we develop a “blur” algorithm to handle images with very small full-width half-

maximum (FWHM). However, this process introduced undesirable pixel-to-pixel correlations, potentially leading to underestimates of the photometric errors. To address this issue, we now consider the linear correlation between a pixel and its left and upper neighbors. In the non-correlated linear optimizations, we minimize

$$\chi_{\text{non-corr.}}^2 = \sum_{i,j} \frac{\epsilon_{i,j}^2}{\sigma_{i,j}^2}, \quad \epsilon_{i,j} = \mathcal{D}_{i,j} - (b + f\mathcal{P}_{i,j}), \quad (1)$$

to obtain the difference flux f and background b . Where $\mathcal{D}_{i,j}$, $\mathcal{P}_{i,j}$ and $\sigma_{i,j}$ are the (i, j) -th pixel value of the difference image, the pixelated PSF model, and the uncertainty image, respectively. This process assumes independent and Gaussian-distributed $\epsilon_{i,j}$ values. Now, we consider a correlated error,

$$\epsilon_{i,j} = c_1 \epsilon_{i-1,j} + c_2 \epsilon_{i,j-1} + \nu_{i,j}, \quad (2)$$

where c_1 and c_2 are the constant correlation coefficients, and $\nu_{i,j}$ is the de-correlated error which should be independent and Gaussian-distributed. Then, the value to be minimized becomes

$$\chi_{\text{corr.}}^2 = \sum_{i,j} \frac{\nu_{i,j}^2}{\sigma_{i,j}^2}. \quad (3)$$

This minimization corresponds to a linear fit applied to the de-correlated difference image ($\mathcal{D}'_{i,j}$) and the de-correlated PSF model ($\mathcal{P}'_{i,j}$)

$$\mathcal{D}'_{i,j} = b + f\mathcal{P}'_{i,j} + \nu_{i,j}, \quad (4)$$

where

$$\mathcal{D}'_{i,j} = \mathcal{D}_{i,j} - c_1 \mathcal{D}_{i-1,j} - c_2 \mathcal{D}_{i,j-1} \quad (5)$$

and

$$\mathcal{P}'_{i,j} = \mathcal{P}_{i,j} - c_1 \mathcal{P}_{i-1,j} - c_2 \mathcal{P}_{i,j-1}. \quad (6)$$

The constant correlation coefficients (c_1 and c_2) are determined by a separate linear fit to $\epsilon_{i,j}$. In practice, we begin by assuming $c_1 = c_2 = 0$, calculate (f, b) , and then iteratively use these results to estimate (c_1, c_2) . To prevent unphysical values, we limit the correlation coefficients to be positive and less than 0.4. This constraint is consistent with the “blur” process.

Second, we updated the definition of the photometry goodness indicator to be

$$\sigma_{\text{res}} = \frac{1}{N_{\text{phot}}} \sum_{i,j} \frac{|\epsilon_{i,j}|}{\sigma_{i,j}^2 + \sigma_{\text{RON}}^2} \quad (7)$$

(Notation consistent with Paper I, Section 2.1). This new definition incorporates the noise term for a more general representation. Previously, the threshold for “bad” data points varied across events. The revised indicator allows for a universal threshold of $\log \sigma_{\text{res}} < 0.5$ which works for most events.

¹ excluding 2020, which was impacted by the COVID pandemic

Lastly, we construct a list of the CCD defects for all KMT-Net cameras so that those pixels can be automatically masked during image subtractions.

Following these improvements, the pipeline (hereafter referred to as “auto-TLC”) can now be efficiently and automatically applied to a large sample, minimizing the need for manual inspection.

3. SAMPLE SELECTION

Giant stars, with their larger radii, offer advantages in microlensing planet detection. The finite source size ρ is defined by

$$\rho \equiv \frac{\theta_*}{\theta_E}, \quad (8)$$

where θ_* and θ_E are the angular radius of the source star and the Einstein ring, respectively. The typical Einstein radius for Galactic microlensing events ranges from ~ 0.1 to 1 mas. The source sizes of main-sequence stars and red giants in the Galactic bulge are typically $\sim 0.4 - 0.8$ μ as and $\sim 3 - 15$ μ as, corresponding to ρ values of $\sim 10^{-3}$ and $\sim 10^{-2}$, respectively. When the planet-host separation is large or planet mass is small, the size of the planetary caustic can be less than $10^{-2}\theta_E$ (Han 2006). In these cases, the probability of a caustic crossing is dominated by ρ rather than the caustic size itself. Therefore, the giant sources are relatively more sensitive to these planets.

Before outlining the specific criteria for our giant source sample, we list a few key expectations. First, the sample should mostly consist of giant sources, so that the advantages of giant sources (bright, sensitive to planets) can be exploited. A small fraction of non-giant source events (e.g., a giant star acting as a blending source) is acceptable, provided the sample remains unbiased in terms of planet detections. Second, the sample size should be of order a few hundred events. This ensures that the photometry and anomaly search can be completed within a reasonable timescale (\sim months) given current computational resources and manpower. As we described in Section 1, the purpose of this starting subsample is, on the one hand, to discover clear previously missed planets, and on the other hand, to systematically confirm whether further research on a larger sample is worthwhile.

We describe the detailed sample selection criteria below. The selection uses the information from the KMTNet website². We focused on events discovered between 2016 and 2022 seasons for which seasons the EventFinder (Kim et al. 2018a) has been completed. Events from the 2020 season are excluded. The 2020 event search is incomplete because both CTIO and SAAO were shut down during the COVID-19 pandemic. For the remaining events, the criteria are as below:

- (1). $I_{\text{cat}} - A_I < 16$,
- (2). $I_s - A_I < 16$,
- (3). $I_{\text{baseline}} < 17$,
- (4). Located in fields with combined cadence ≤ 15 min.

The first criterion is designed to select giant stars (Kim et al. 2021), down to approximately 1.5 magnitudes below the red clump center. The second criterion ensures the source indeed corresponds to the cataloged giant star. The third and fourth criteria do not hold specific physical meaning but only limit the observed signal-to-noise ratio and sample size.

As previously mentioned, giant source events are sensitive to wide-orbit planets. However, the standard KMTNet TLC pipeline is typically only applied to the event season itself, as most events are shorter than one year. In this work, to allow searching for distant signals far from the peak, images from multiple years are included to produce the photometric data. Because the probability of finding a $s > 15$ signal drops to almost nothing, we include only images within $t_0 \pm 15t_E$ for each event. If a portion of a season overlaps with the above time range, the entire season is used.

Following these criteria, a total of 352 events were selected. We then employ the auto-TLC pipeline for a systematic photometric re-reduction of these events. The majority of images of the KMTNet survey were taken in the I band, and about 9% were taken in the V band. The systematic re-reduction includes all I -band images and KMTC V -band images from the selected seasons for color measurements.

The auto-TLC pipeline encountered failures in 12 events. These events were all located near field boundaries, where the pipeline failed in automatic reference image selection. The failure rate ($12/352 \approx 3.4\%$) is consistent with the frequency of stars near the boundary of a field $300/9000 \approx 3.3\%$. Because the planet occurrence rate should not depend on the distance to the image boundary, we do not try to recover these events. Consequently, our subsequent anomaly signal search will focus on the remaining 340 events.

4. ANOMALY SEARCH

4.1. The Anomaly Search Algorithm

Zang et al. (2021b) highlighted the need for a search algorithm to comprehensively find the anomalies, particularly for a survey like KMTNet. Following this idea, here we develop a new algorithm to automatically detect candidate signals.

The first step for searching anomalies in a light curve, is to remove “normal” signals. In our case, the “normal” signal is the standard point-source point-lens (PSPL) microlensing light curve (Paczynski 1986). Three parameters are required to describe the magnification as a function of time, $A(t)$. They are u_0 , the impact parameter of the lens-source relative

² <https://kmtnet.kasi.re.kr/~ulens/>

trajectory (in units of the Einstein radius), t_E , the microlensing timescale, and t_0 , the closest time of the lens and source. The magnification is then

$$A(t) = \frac{u(t)^2 + 2}{u(t)\sqrt{u(t)^2 + 4}}, \quad u(t) = \sqrt{u_0^2 + \left(\frac{t - t_0}{t_E}\right)^2}. \quad (9)$$

Two flux parameters, f_S and f_B for each site and field are needed to represent the magnified and unmagnified fluxes. For each event, we fit the entire light curve with such a PSPL model and extract the residuals. Subsequent anomaly searches are mainly based on the residuals.

The existing KMTNet AnomalyFinder (Zang et al. 2021b, 2022a) employs a robust strategy for searching for signals in the online data, which are often noisy and contain many outliers. This approach involves fitting a series of PSPL models to the residuals. However, for the high-quality auto-TLC data used in this work, the photometric algorithms effectively identify and exclude most outliers, significantly reducing the false positives. Therefore, to improve the efficiency, we developed a new anomaly search algorithm that does not require a model for the anomaly.

Despite the improved stability of the auto-TLC data compared to online data, false positives can still occur due to reasons like poor seeing conditions, high sky background, defects in the image, or intrinsic variability in the source or the nearby blended stars. Therefore, we design our algorithms accordingly to reduce false positives caused by these reasons.

The flow chart outlining our anomaly search procedure is presented in Fig. 1. Here we first briefly list all the main steps and then follow this with a detailed explanation of each step.

- (a). Fit the light curve with microlensing models and extract the residuals.
- (b). Rescale the error bars of the data in FWHM bins for each site/field.
- (c). Define a time window for searching for anomalies.
- (d). In this time window, calculate the cumulative χ^2 and compare it to a threshold.
- (e). Check if the signal is significant compared to the baseline variability.
- (f). Check if the window contains data from multiple fields/sites.
- (g). Check if the signal is dominated by a small fraction of largest χ^2 points.
- (h). Check if the signal comes from smooth deviations (consistent over time) rather than scattered points.

- (i). If a signal passes all these criteria, it is labeled as an anomaly candidate.

When fitting the light curve in (a), we first use the static point-source point-lens (PSPL) model. We fit the light curve and extract the residuals ΔF (observed flux – model) and their corresponding uncertainties σ_F in flux space. Initially, any deviations from the static PSPL model are considered as potential anomaly candidates.

The purpose of step (b) is to reduce the false positives caused by seeing variations. Although the auto-TLC data are much more stable than the online data, correlations between flux and seeing still exist. Therefore, we rescale the error bars of the data points in a series of FWHM bins for each site/field. To allow each FWHM bin to have sufficient points, the bins are taken from the minimum to the maximum value with a bin width of $0.2''$. We calculate χ^2/N_{pt} (where N_{pt} is the number of data points) in these seeing bins, and use it to determine the scaling factor k ,

$$\sigma'_{F,i} = k\sigma_{F,i}, \quad (10)$$

that adjusts the errors to achieve

$$\chi^2 = \sum_i \left(\frac{\Delta F_i}{\sigma'_{F,i}} \right)^2 = 1, \quad (11)$$

where ΔF_i , $\sigma_{F,i}$, and $\sigma'_{F,i}$ are the residual, and the uncertainties of each measurement before and after the rescaling. Because the light curve covers at least 2 yr for these events, the χ^2 here is dominated by the unmagnified baseline data. Planetary signals, even if they exist, are only a small fraction of the whole light curve. Therefore, the false negatives caused by this step are negligible, but many false positives are reduced. Nevertheless, we limit k to be $1.0 \leq k \leq 3.0$ to prevent making false negatives. This step also reduces the influence of intrinsic source variability on false positives.

In (c), to find signals as short as ~ 0.2 d, we sample 16 logarithmic uniform values in 0.02 d – 2 d and 20 logarithmic uniform values in 2 d – 2000 d as the duration of the time window. For each window duration, we set the time window to start from the first data point and shift it with a step size of $1/10$ of the window length until the data ends. On average, this approach divides the data for each year into $\sim 600,000$ windows.

Then in (d), we first skip windows containing ≤ 3 data points because they are unreliable for the anomaly detection. On average, 75% of windows are skipped due to this criterion. For the windows with sufficient data points, N_{window} , we calculate the cumulative χ^2 in the window using

$$\chi^2_{\text{window}} = \sum_{i \in \text{window}} \left(\frac{\Delta F_i}{\sigma_{F,i}} \right)^2. \quad (12)$$

where the subscript i represents the index of the data point. Assuming the residuals follow an independent Gaussian distribution (no signal present), the χ^2_{window} should follow the χ^2 distribution. We select the threshold $\chi^2_{\text{thres}}(N_{\text{window}})$ as a function of N_{window} , such that random variables from a χ^2 distribution have a probability less than 6.33×10^{-5} (corresponding to a $> 4\sigma$ probability for a Gaussian distribution) of exceeding the threshold. Larger χ^2_{window} are considered anomalous and indicative of potential anomaly signals. Additionally, a minimum threshold of $\chi^2_{\text{thres,min}}(N_{\text{window}}) = N_{\text{window}} + 80$ is set to ensure the signal is significant enough compared to unrecognized non-Gaussian systematic errors. Similar minimum thresholds are widely used in previous microlensing planet detections (e.g., Suzuki et al. 2016; Poleski et al. 2021; Zang et al. 2021b), where the exact value varies regarding the features of the dataset. Only windows with $\chi^2_{\text{window}} > \chi^2_{\text{thres}}$ proceed to the next steps.

In (e), we check the source variability. Because giant sources often have intrinsic variability, we ensure reported anomalies are not dominated by these variations. The baseline variability is defined as the root-mean-square of $\Delta F/\sqrt{F}$, calculated as

$$\text{RMS} = \sqrt{\frac{1}{N_{\text{data}}} \sum_i^{N_{\text{data}}} \frac{\Delta F_i^2}{F_i}}, \quad (13)$$

where N_{data} is the total number of data points and F_i is the original measured flux. The reason why we adopt RMS of $\Delta F/\sqrt{F}$ instead of ΔF is that the magnified points tend to have larger ΔF s. We therefore normalize ΔF by \sqrt{F} to account for Poisson errors during microlensing events. Similarly, the RMS for a window is calculated as

$$\text{RMS}_{\text{window}} = \sqrt{\frac{1}{N_{\text{window}}} \sum_{i \in \text{window}} \frac{\Delta F_i^2}{F_i}}. \quad (14)$$

Assuming the residuals $\Delta F/\sigma_F$ follow independent Gaussian distributions, then $(\sqrt{N_{\text{window}}} \text{RMS}_{\text{window}}/\text{RMS})$ should follow a χ distribution. We adopt a threshold corresponding to $< 4.55\%$ probability ($> 2\sigma$ probability for a Gaussian distribution). If the variation in a window exceeds this threshold, the window proceeds to the next step.

In (f), we simply check if the window contains data from more than one site or field so that the signals can be cross-verified. If true, it is passed to the next step.

Then, in (g), we check if the signal is dominated by a few outliers. In many cases, the light curve can occasionally contain outliers. Real planetary signals in giant source events typically have a timescale of $t_{\text{anom}} \sim t_* = \rho t_E$ and thus cannot be extremely short or consist of only a few data points. Therefore, here we exclude the 5% (rounded down to an integer value, with a minimum of 1) highest χ^2 points and re-evaluate the remaining points against the criterion in step (d).

Note that because the number of points changes, the threshold also changes. This step also helps exclude cases where a very large window contains a very short anomaly, improving the efficiency for human reviewers.

Finally, in (h), we check if the “signal” is consistent over time, i.e., not just caused by randomly scattered points. Random signals are not smooth over time and are not considered candidate signals. We fit the residuals in the time window with a polynomial function and extract the corresponding χ^2_{poly} . Based on testing results, we empirically choose the polynomial order to be $\lfloor N_{\text{window}}/6 \rfloor$, with minimum order 2 and maximum order 8. We then check if χ^2_{poly} satisfies

$$\frac{\chi^2_{\text{poly}} - N_{\text{window}}}{\chi^2_{\text{window}} - N_{\text{window}}} < 0.3, \quad (15)$$

i.e., if the polynomial model can explain $\geq 70\%$ signal, $\Delta\chi^2_{\text{window}} = \chi^2_{\text{window}} - N_{\text{window}}$. If a signal is caused by randomly scattered points, then χ^2_{poly} will be comparable to χ^2_{window} , because a smooth polynomial function cannot reproduce the randomly scattered “signals”. On the other hand, if the “signal” is smooth over time, one can find χ^2_{poly} significantly smaller than χ^2_{window} . Signals that pass all the above criteria are classified as anomaly candidates and await further inspection and/or modelling.

Finally, for a light curve containing possible signals, the above search often returns many overlapping windows, which can be troublesome for human reviewers. Therefore, we group these reported windows as follows: If a smaller window is fully included in a larger window, we keep the one with larger $\Delta\chi^2_{\text{window}}$ and discard the other. The process starts from the smallest window, and continues until all windows are not fully included in each other. The final windows are reported to the reviewer.

Now, the anomaly search algorithm has been built. It is named `easyAnomalyFinder` or `eAF` for short. Because it avoids detailed model fitting, the typical time cost per event is only $\sim 1 - 5$ seconds on a single CPU core, which will eventually speed up the sample sensitivity calculation. Relevant results will be presented in future work. We note that the `eAF` algorithms can potentially have a broader usage, such as searching for various types of time-domain signals. We managed to make the code easier to customize. The reader can find the public code on [GitHub³](https://github.com/hongjingy/easyAnomalyFinder).

We then apply `eAF` to the auto-TLC data of the 340 giant source events. We review the reported signals, if they are false positives, we clean the data accordingly and search for the signals again. If the signals are likely true, they are modeled, and the residuals from the new model are fed back

³ <https://github.com/hongjingy/easyAnomalyFinder>

into eAF until no signals are reported. The search results are described in the next section.

4.2. Apply *easyAnomalyFinder* to Giant Source events

We apply eAF to the 340 events with new photometric data. Here we present the results. We first classify the events into two categories: events with no signal (or a signal caused by systematic errors or intrinsic variation) and events with candidate signals. A total of 59 events have candidate signals, while 281 have no signal.

In the no-signal group, 235 do not report any signals or the “signal” is due to outliers that can be easily eliminated by cleaning the data. Seven events are not real microlensing events but false positives of the KMTNet AlertFinder (Kim et al. 2018b) or EventFinder (Kim et al. 2018a). Thirty-four events exhibited very significant intrinsic variability, making it unlikely any real signals can be identified within them. An example is shown in the upper left panel of Fig. 2. The other 5 events have significant systematic errors that could not be cleaned.

Among the 59 events with candidate signals, we cross-matched them with known planetary and binary catalogs and further categorized them as: known planetary events, known stellar binary events, and new anomalous events. All the known planets (2) and known binaries (16) in this sample are successfully recovered. The two known planets are OGLE-2018-BLG-0383/KMT-2018-BLG-0900 (Wang et al. 2022) and OGLE-2016-BLG-0007/KMT-2016-BLG-1991 (Zang et al. 2025). The remaining 41 events are considered new anomalous events that require further investigation.

Sixteen of these events display asymmetric signals like KMT-2021-BLG-0147 (shown in Fig. 2, upper right panel). These features could be caused by the microlensing parallax effect (Gould 1992, 2000, 2004). This effect is a result of Earth’s orbital motion, which provides acceleration to the source-lens-Earth relative motion. Here we include the parallax effect in the modelling of these events, obtaining new residuals. No further signals were found using eAF on these new residuals. Therefore, we classify these events as possible parallax events. Note that some of these signals could be due to the “xallarap” effect, i.e., the source star’s orbital motion around a companion (Griest & Hu 1992; Han & Gould 1997). In many cases, the parallax and xallarap effects are degenerate. However, this work focuses on planetary candidates, so we do not investigate these signals in detail for now.

In the remaining 25 events, 4 of them have “flat-top” features. An example is shown in the lower left panel of Fig. 2. These features are likely caused by the finite-source effect when $\rho > u_0$. However, the extended source could also cross the central caustic produced by a planet in the lens. Therefore, we update the models to finite-source point-lens (FSPL)

models and re-evaluate the residuals. One event (KMT-2022-BLG-0330) shows no clear residuals after FSPL fitting, while three still have signals. We then search for a finite-source binary-lens (FS2L) model for these three events (methods are the same as will be described in Section 5.1). Finally, we find all of them are clear stellar binary events with secondary-to-primary mass ratio $q > 0.1$.

We fit for single-source binary-lens (1S2L) models for the remaining 21 candidate events. We find 17 of them are clear stellar binary events ($\log q \geq -1.5$) like KMT-2017-BLG-1630 (see lower right panel of Fig. 2).

The other 4 events are considered new candidate planetary events, they are KMT-2018-BLG-0831, KMT-2019-BLG-0578, KMT-2021-BLG-0736, and KMT-2022-BLG-2342. The light curves and the reported signals for these four events are shown in Fig. 3.

Table 1 lists the observational information of these four events. KMT-2018-BLG-0831 was first discovered by the OGLE Early Warning System (EWS, Udalski et al. 1994; Udalski 2003) and named OGLE-2018-BLG-0421. It was later independently identified by the KMTNet post-season EventFinder system (Kim et al. 2018a). KMT-2019-BLG-0578 was only found by the KMTNet real-time AlertFinder (Kim et al. 2018b) on April 29, 2019. KMT-2021-BLG-0736 was first found by the KMTNet AlertFinder on May 10, 2021 and then independently discovered by the MOA collaboration on June 2, 2021 as MOA-2021-BLG-152. KMT-2022-BLG-2342 was first identified on February 24, 2022 by MOA collaboration and denoted as MOA-2022-BLG-038, and later independently found by the KMTNet post-season EventFinder system.

Although the anomalies in these events were discovered using KMTNet data only, not all the events were initially identified by KMTNet. Hereafter, we use the names from the survey who first discovered them for these events, i.e., OGLE-2018-BLG-0421, KMT-2019-BLG-0578, KMT-2021-BLG-0736, and MOA-2022-BLG-038. In the following Section 5, we further investigate these four events to determine whether they are clear planets.

5. ANALYSIS OF PLANET CANDIDATES

In this section, we present the modeling details of the four planet-like anomalies identified in Section 4.2. Their light curves are shown in Fig. 3.

The anomaly signals can be categorized into three groups based on their morphologies. The event in the first group is KMT-2019-BLG-0578. Its anomaly shows an “M” shape over the peak. This feature can only be produced by a second object in the lens system. Therefore, only binary-source single-lens (2L1S) models need to be explored for this event. The second group consists of OGLE-2018-BLG-0421. The anomaly signal shows a slight brightening followed by a dim-

ming compared to a PSPL model. Such subtle asymmetries can potentially arise from microlensing parallax, xallarap effects, or the presence of a secondary lens. However, the short duration (~ 3 days) of the anomaly makes the parallax explanation (due to Earth's orbital motion) unlikely. Additionally, the anomaly timescale is significantly shorter than the overall event timescale (15 days). If caused by the source star's orbital motion (xallarap effect), the remaining light curve should also show periodic features (e.g., [Han & Gould 1997](#)). We attempted to model the anomaly with both parallax and xallarap effects, but these models either provided poor fits to the light curve or resulted in unphysical parameters. Therefore, only 2L1S models are needed to be explored for this event. The last group is KMT-2021-BLG-0736 and MOA-2022-BLG-038, both of which have an additional peak separate from the main peak. These features can be explained by either a cusp approach in the 2L1S scenario ([Mao & Paczyński 1991](#); [Gould 1992](#)), or the presence of a fainter, secondary source passing closer to the lens (1L2S model, [Gaudi 1998](#)). Therefore, both 2L1S and 1L2S models should be investigated for these two events.

While the detailed analysis is tailored to each event, a general procedure is followed. In Section 5.1, we first describe the general modelling methods. Then, the results for each individual event are presented in Sections 5.2-5.5.

5.1. Preamble

5.1.1. Light Curve Modelling

As mentioned above, we only explore 2L1S and 1L2S models in detail for these candidate events.

Standard 2L1S models require seven parameters to describe the magnification as a function of time, $A(t)$. The first three parameters, (t_0, u_0, t_E) , are identical to those used in the PSPL model. The difference is that t_0 and u_0 are defined relative to the magnification center of the two lenses⁴. The three following parameters are (s, q, α) : s and q represent the projected separation (in units of the angular Einstein radius) and the mass ratio of the two lenses, and α is the direction of source-lens relative proper motion. The last parameter is ρ , the angular size of the source in units of the angular Einstein radius. In addition to the magnification, two flux parameters, $(f_{S,i}, f_{B,i})$ represent the source and blend flux are needed for each dataset i . The total flux as a function of time is then

$$f_i(t) = f_{S,i}A(t) + f_{B,i}. \quad (16)$$

We employ the `VBBinaryLensing` ([Bozza 2010](#); [Bozza et al. 2018](#)) package to calculate the magnification for the 2L1S models.

⁴ When $s \leq 1$, the magnification center is the mass center. When $s > 1$, the magnification center is located at $\frac{q}{1+q}(s - \frac{1}{s})$ from the primary star.

Finding all possible solutions for a light curve with such a large number of parameters, especially considering the highly non-linear behavior of (s, q, α) , is not a trivial task. Therefore, we adopt a grid search approach followed by Markov-chain Monte-Carlo (MCMC) sampling to explore the parameter space and recover all possible solutions. The grid search starts with an initial set of parameters. By default, we sample 49 equally spaced values in $-1.2 \leq \log s \leq 1.2$, 61 equally spaced values in $-6.0 \leq \log q \leq 0.0$, 10 equally spaced values in $-3.4 \leq \log \rho \leq -0.7$, and 16 equally spaced values in $0^\circ \leq \alpha < 360^\circ$ as the initial parameters. Then, for each set of initial parameters, we fix $(\log s, \log q, \log \rho)$ and allow all other parameters to vary. We use the MCMC sampler `emcee` ([Foreman-Mackey et al. 2013](#)) to search for local minima of the χ^2 function. After the grid search, one or more local minima might be identified. The corresponding parameter sets represent candidate models that can potentially describe the light curve. For each candidate model, we perform another round of MCMC sampling allowing all parameters to be free. This refines the parameter estimates and provides their uncertainties. We then compare the goodness of fit for the models and decide which one(s) to adopt.

Although for simple cases like KMT-2021-BLG-0736 and MOA-2022-BLG-038, 2L1S solutions can be found using analytical estimates (e.g., [Ryu et al. 2022](#)), we consistently follow the above grid search approach for all events to ensure that all possible 2L1S solutions are explored.

The light curve for a 1L2S event is essentially the superposition of two individual 1L1S light curves. Following [Hwang et al. \(2013\)](#), at least eight parameters are required to model this scenario. Among them, $(t_{0,1}, u_{0,1}, \rho_1)$ and $(t_{0,2}, u_{0,2}, \rho_2)$ are the approach time, impact parameter, and the source size of the two sources, respectively. The two sources' Einstein timescales t_E are assumed to be the same. Finally, a flux ratio parameter

$$q_F = \frac{f_{S,2}}{f_{S,1}} \quad (17)$$

is needed. The total flux during the event is then

$$f_i(t) = f_{S,i} [A_1(t) + q_F A_2(t)] + f_{B,i}, \quad f_{S,i} \equiv f_{S,1,i}, \quad (18)$$

where $A_1(t)$ and $A_2(t)$ represent the magnifications from the 1L1S models for the primary and secondary sources, respectively. The flux ratio could vary depending on the observational band. Therefore, an independent flux ratio parameter $q_{F,\lambda}$ needs to be included for each band λ .

The fitting process for 1L2S models is generally more straightforward due to the linear nature of the superposition. Therefore we initially set $(t_{0,1}, u_{0,1}, \rho_1, t_E)$ as the best-fit 1L1S parameters of the primary event. Then, we choose a reasonable initial value for $t_{0,2}$ around the location of the

secondary peak in the light curve. Proper initial values for $(u_{0,2}, \rho_2, q_F)$ can also be estimated regarding the magnification excess. Finally, we optimize all the parameters using MCMC sampling to minimize χ^2 and obtain the final parameters of the 1L2S model.

5.1.2. Source Properties

This section describes the method for determining the source star's de-reddened magnitude and color. This information can be used to measure the source angular radius θ_* , and ultimately θ_E and the lens-source relative proper motion μ_{rel}

$$\theta_E = \frac{\theta_*}{\rho}, \quad \mu_{\text{rel}} = \frac{\theta_E}{t_E}. \quad (19)$$

These parameters provide information on the physical properties of both the source and lens systems.

We begin by measuring the source's I and V magnitudes within a $3.3' \times 3.3'$ square region centered on the source. We measure the flux of all stars on both I - and V -band master reference image using pyDIA⁵, and then calibrate the magnitudes to the OGLE-III (Szymański et al. 2011) system and plot them on a Color-Magnitude Diagram (CMD). Additionally, the source magnitude I_S can be obtained from the light curve modelling process. For single-source events, the source color $(V - I)_S$ remains constant throughout the event. Therefore, it can be measured by performing a linear regression on the time series data of the I and V light curves. For binary-source events, the source color is measured by modelling the independent flux parameters of the light curve.

After that, we measure the centroid of Red Clump (RC) stars (Yoo et al. 2004; Nataf et al. 2013) on the CMD. This centroid is then used to calculate the offset between the source(s) and the RC stars,

$$\Delta[(V - I), I] = [(V - I), I]_S - [(V - I), I]_{\text{RC}}, \quad (20)$$

where $[(V - I), I]_{\text{RC}}$ is the RC centroid. By comparing this offset with the known de-reddened color and magnitude of the Galactic bulge RC stars $[(V - I), I]_{\text{RC},0}$ (Bensby et al. 2013; Nataf et al. 2013), we can determine the de-reddened color and magnitude of the source(s),

$$[(V - I), I]_{S,0} = [(V - I), I]_{\text{RC},0} + \Delta[(V - I), I]. \quad (21)$$

Finally, using $[(V - I), I]_{S,0}$, we can estimate the source star's angular size θ_* according to Adams et al. (2018). If the light curve modelling provides a measurement or constraint on ρ , we can convert it to a measurement or constraint on θ_E and μ_{rel} using Eq. 19.

5.1.3. Lens Properties

For the events that are confirmed as planetary events, the light curve modelling only provides relative parameters (s, q) of the lens system, but not the physical distance and mass. Determining the physical properties of the lens system requires at least two of the following three parameters: microlensing parallax π_E , angular Einstein radius θ_E , and the brightness of the lens star (see also Zang et al. 2022b). For example, if both π_E and θ_E are measured, the mass of the lens system can be uniquely determined using (Gould 2000)

$$M_L = \frac{\theta_E}{\kappa\pi_E}, \quad \kappa = \frac{4G}{c^2 \text{AU}} \simeq 8.144 \text{ mas}/M_\odot, \quad (22)$$

where G is the gravitational constant and c is the speed of light.

However, it is uncommon to measure both effects in a single event. Moreover, directly measuring the lens star's brightness requires resolving the lens and source stars. This requires a wait of $\gtrsim 5 - 10$ yr for the current large telescopes. In cases where none or only one of these parameters is available, we can only infer the physical properties of the lens system using a Bayesian approach with a Galactic model prior.

The Galactic model contains information about the stellar density, mass function, and the velocity distribution of the Milky Way. We adopt these distributions from "Model C" in Yang et al. (2021). We simulate a large number of microlensing events based on the Galactic model. Each simulated event is then assigned a weight that considers both the microlensing event rate and the likelihood function obtained from the light curve modelling. Specifically, the weight for the i -th simulated event is

$$w_i = \Gamma_i \times \mathcal{L}_i(t_E) \mathcal{L}_i(\theta_E) \mathcal{L}_i(\pi_E), \quad (23)$$

where $\Gamma_i \propto \theta_{E,i} \mu_{\text{rel},i}$ is the microlensing event rate, with $\mu_{\text{rel},i}$ being the relative proper motion. $\mathcal{L}(t_E)$, $\mathcal{L}(\theta_E)$, and $\mathcal{L}(\pi_E)$ are the likelihood functions for the corresponding parameters derived from the light curve modelling. If a particular observable is not measured in the real event, the corresponding likelihood function is set to be uniform. In addition, Gould (2022) pointed out that events with confirmed planetary signals might exhibit a different μ_{rel} distribution compared to other events, potentially due to observational bias. To account for this, when ρ is unmeasured, an additional term of $\mu_{\text{rel},i}^{-1}$ is incorporated into the weights.

After properly weighting all the simulated events, we can obtain the posterior distributions of the physical properties of the lens system (host mass, planet mass, and planet-to-host separation).

5.2. OGLE-2018-BLG-0421/KMT-2018-BLG-0831

⁵ MichaelDAlbrow/pyDIA: Initial Release on Github, doi:10.5281/zenodo.268049

Here we describe the exploration of 2L1S models for OGLE-2018-BLG-0421. We note that this event was also discovered and observed by OGLE collaboration (Udalski et al. 2015). Therefore, data from OGLE are included in the following process. The images from the OGLE survey were taken in the I band and the data were reduced by the Wozniak (2000) difference image pipeline first and then the updated pySIS (Albrow et al. 2009; Yang et al. 2024).

A preliminary analysis revealed significant long-term deviations in the light curve data for $t > 8250$. This deviation could be intrinsic variation of the source or unknown systematic errors. As this deviation (timescale ~ 1 yr) is not correlated with the ~ 3 d anomaly, we exclude data points beyond $8150 \leq t \leq 8250$ to avoid contamination.

We first conduct a standard grid search as described in Section 5.1. However, the local minima are not well covered, especially in $\log \rho$ space. Therefore, we adjust the default grid search to allow ρ to vary in the optimization process. The adjusted grid search returns fifteen local minima within $\Delta\chi^2 < 100$, $(C_1, \dots, C_7, R_1, R_2, W_1, \dots, W_6)$ as Fig. 4 shows. These initial models can be grouped into three categories regarding their mass ratios. Models with $\log q > -1$ ($C_1, C_5, C_7, W_1, W_5, W_6$) indicate stellar mass companions of the primary lens, models with $-2 < \log q < -1$ (C_4, C_6, W_4) are likely brown dwarf companions, and $\log q < -2$ ($C_2, C_3, W_2, W_3, R_1, R_2$) indicate planetary mass companions. The subtle nature of the signal in the light curve results in strong degeneracy among stellar, brown dwarf, and planetary models.

We then perform further optimization around each of the fifteen initial local minima. Table 2 lists the optimized parameters of the remaining five solutions with $\Delta\chi^2 < 15$. Model R_2 emerges as the best fit but only with a $\Delta\chi^2$ of approximately (3.8, 4.0, 4.7, 14.1) compared to the other models (C_1, C_7, C_2, W_4). Fig. 5 shows the light curves and models around the anomaly region for these five solutions in the left panels. The corresponding caustic geometries and source-to-lens trajectories are shown in the right panels of Fig. 5. As shown in Table 2 and Fig. 5, all three possible scenarios, stellar companion model (C_1, C_7), brown dwarf companion model (W_4), and planetary companion models (C_2, R_2) can describe the anomalous light curve well.

Notably, all five degenerate solutions share consistent constraints on source brightness and ρ constraints. This implies that even future follow-up observations measuring the lens brightness and θ_E may not be sufficient to break the degeneracy. Therefore, we conclude that OGLE-2018-BLG-0421 is an ambiguous event and do not investigate it further.

5.3. KMT-2019-BLG-0578

The event KMT-2019-BLG-0578, according to previous discussions, also needs only 2L1S modelling. A standard

grid search recovers only two local minima, C and W for this event. Unlike the previous event OGLE-2018-BLG-0421, the presence of strong caustic-crossing features in KMT-2019-BLG-0578 limited the number of potential solutions. The two minima have similar values for $\log q = -2.4$ and $\alpha \sim 258^\circ$, but have opposite $\log s = \pm 0.1$. This again reflects the well-known “close-wide degeneracy” in microlensing (Griest & Safizadeh 1998). We then perform MCMC optimization around both local minima. The results are listed in Table 3.

The two models are almost identical and differ only by $\Delta\chi^2 \sim 0.02$. The left panels of Fig. 6. present the light curves along with the models. Because the two 2L1S models do not have visual differences, only the model C is presented. The right panels of Fig. 6 show the caustic structures and the lens-source motion trajectories of the two models. As expected, the central caustics of the two models are practically identical.

Despite the degeneracy, both models consistently suggest the presence of a planet within the lens system, with a mass ratio of $q \sim 4 \times 10^{-3}$. Therefore, the event KMT-2019-BLG-0578 is a clear planetary event. Because the models provide a measurement of ρ , we proceed to estimate θ_E using the method outlined in Section 5.1.2

We utilize the KMTA03 reference images to create the CMD and calibrate it to the OGLE-III catalog. The CMD is shown in Fig. 9. The source star of this event is marked in the figure, it is offset to the bulge main sequence population. However, the red clump’s CMD of this field seems elongated (the dashed blue line as a hint), which indicates that the field has a significant extinction variation. The source star could be a member of the low-extinction main-sequence population. Nevertheless, no matter which population the source is in, the determined source angular radius θ_* values is consistent within $\sim 1\sigma$. Therefore, we still use the overall red clump centroid to continue the analysis. Table 4 summarizes the measured values, including the RC centroid, the source color and magnitude, and their de-reddened values. The intrinsic RC centroid are calculated based on a linear interpretation of the values presented in Nataf et al. (2013). Finally, we estimate the source size to be $\theta_* = 0.361 \pm 0.040$ μ as (Adams et al. 2018), and consequently $\theta_E = 0.213 \pm 0.029$ mas and $\mu_{\text{rel}} = 8.8 \pm 1.3$ mas/yr.

This event has a measurement of θ_E , but the short timescale (< 10 d) prevents the constraints on the microlensing parallax. We estimate the physical properties following the Bayesian approach mentioned in Section 5.1.3. We simulate 10^7 microlensing events and assign weights based on Eq. 23. The median and $\pm 1\sigma$ confidence intervals of the resulting posterior distribution are presented in Table 5. Based on the analysis, the planet is likely a giant planet with a mass of $\sim 1.2M_J$ (M_J denotes Jupiter mass). It is orbiting its

$\sim 0.28M_{\odot}$ host at a projected distance of either ~ 1.2 AU or ~ 2.0 AU.

In conclusion, although KMT-2019-BLG-0578 has two degenerate models, both of them suggest a clear $q \approx 4 \times 10^{-3}$ planet in the lens system. The signal is not visible in the on-line KMTNet data. The discovery of this planet relies on the new photometric data reduced by the updated pipeline. Using Galactic models, we estimate the planet to be a $\sim 1.2M_{\mathrm{J}}$ planet orbiting its $\sim 0.28M_{\odot}$ M dwarf host at a projected distance of about either 1.2 or 2.0 AU.

5.4. KMT-2021-BLG-0736

Both 1L2S and 2L1S models can potentially describe the anomaly of KMT-2021-BLG-0736. Therefore, here we explore both of them.

Data from MOA are included in the modelling process. The images from the MOA survey were mainly taken in the MOA-*Red* band, which is approximately the sum of the standard Cousins *R* and *I* bands. The MOA data were reduced by the [Bond et al. \(2001\)](#) difference image pipeline.

First, for 2L1S models, a standard grid search returns three solutions, W_{inner} with $(\log s, \log q) \sim (0.22, -4.2)$, W_{outer} with $(\log s, \log q) \sim (0.19, -4.0)$, and W_{cross} with $(\log s, \log q) \sim (0.20, -4.6)$. We further optimized these solutions using a MCMC. The resulting parameters are summarized in Table 6. The light curves and models around the anomaly region are presented in Fig. 7.

As illustrated in Fig. 7, these three models represent the source approaching/crossing the planetary caustics from different sides. The degeneracy between W_{inner} and W_{outer} is known as the “inner-outer” degeneracy ([Gaudi & Gould 1997](#); [Zhang et al. 2022](#); [Ryu et al. 2022](#)). However, in this case, the W_{inner} model is significantly disfavored by $\Delta\chi^2 \sim 162$ and can be ruled out. This is because W_{inner} and W_{outer} predict contrasting deviations in the light curve before and after the planetary peak. W_{inner} predicts a lower flux compared to the 1L1S model before the peak and a higher flux afterward, while W_{outer} suggests the opposite trend. The difference is well resolved thanks to the complete coverage of the observational data. In addition, Model W_{cross} is a caustic-crossing solution, where the large source covers the entire planetary caustic. It predicts a shorter “dip” after the planetary peak, thus is disfavored by $\Delta\chi^2 \sim 59$. This model can also be ruled out. Consequently, W_{outer} emerges as the only remaining 2L1S model for KMT-2021-BLG-0736.

We then try 1L2S models for this event. The initial parameters for the first source were set to the best-fit values from the 1L1S model. The second source was initialized with $(t_{0,2}, u_{0,2}, \rho_2) = (9376, 0.01, 0.01)$ and $q_{F,I} = q_{F,\mathrm{Red}} = 0.001$. We then performed an MCMC search with all parameters free. The final optimized parameters are listed in Table 6, and the corresponding model is plotted in Fig. 7. While

the 1L2S model has two additional parameters compared to the 2L1S models, the best-fit solution yields a significantly poorer fit to the light curve. The figure also demonstrates the inability of the 1L2S model to capture the signal. Therefore, the 1L2S interpretation is excluded.

In conclusion, the 2L1S W_{outer} is the only model of the light curve, indicating the presence of a planet in the lens system with a mass ratio of $q \approx 1.06 \times 10^{-4}$.

We then follow the methodologies in Sections 5.1.2-5.1.3 to estimate the source and lens properties. The Color-Magnitude Diagram (CMD) of stars surrounding the source is constructed and displayed in Fig. 9. The source of KMT-2021-BLG-0736 is a red giant that is bluer than the Red Clump. The measured de-reddened source color and magnitudes are $(V - I, I)_{S,0} = (0.489, 14.450)$, leading to an estimated angular source size of $\theta_* = 3.69 \pm 0.35 \mu\mathrm{as}$. Table 7 provides the detailed measurements.

Although the light curve model does not directly measure ρ , it provides constraints on its value. Based on the 3σ upper limit of ρ , we find $\theta_E > 0.27$ mas and $\mu_{\mathrm{rel}} > 5.0$ mas/yr. The above constraint on θ_E is considered in the Bayesian analysis. We simulate 10^8 events using the Galactic model and weight them according to Eq. 23. For bright source events, *Gaia* observations could provide additional information. We checked the source proper motion reported by *Gaia* DR3 ([Gaia Collaboration et al. 2023](#)), however, the source star has a renormalized unit weight error (RUWE) of 1.63, which indicates a problematic astrometric solution. We therefore measure the proper motion distribution for the red clump stars within a radius of $2'$ from the source, which has a mean value of $(\mu_l, \mu_b) = (-6.075, 0.055)$ mas/yr and a covariance matrix of $C_{\mu_l, \mu_b} = ((12.053, -1.230), (-1.230, 11.114))$. The distribution is incorporated into the simulated events as an additional prior of the source star’s proper motion. The resulting posterior distributions of the physical parameters are presented in Table 8. The estimated mass of the planet is $21 M_{\oplus}$, and it orbits its $\sim 0.6M_{\odot}$ host star at a distance of around 4.0 AU.

This is another new clear planetary event from the systematic search using new photometric data.

5.5. MOA-2022-BLG-038

Event MOA-2022-BLG-038 also needs both 2L1S and 1L2S modelling. We note that this event is also located in MOA fields and was first discovered by MOA collaboration ([Bond et al. 2001](#); [Sumi et al. 2003](#)). Therefore, data from MOA are included in the following analysis.

We first check 2L1S models. The signal is similar to KMT-2021-BLG-0736, thus we also expect a group of “inner-outer-cross” solutions. However, after detailed modelling, we find the expected solutions merge into a single one within

the MCMC chains. The parameters of this unique solution are summarized in Table 9. The reason is that the anomaly signal lasts relatively long and has relatively high magnification, which favors the scenario where the source entirely crosses a planetary caustic rather than merely approaching a cusp (see the right panel of Fig. 8). Such a large ρ smooths out the distinction between the “inner”, “outer” and “cross” models, leading to the single solution. The light curves around the anomaly, along with the model, residuals, and the corresponding caustic geometry are shown in Fig. 8. However, the presence of systematic residuals suggests that the 2L1S model might not be the best model for the event.

We then investigate the 1L2S scenario. The initial parameters for the second source are set to $(t_{0,2}, u_{0,2}, \rho_2) = (9831.5, 0.01, 0.01)$, $q_{F,I} = q_{F,Red} = 0.001$. The remaining parameters related to the first source are initialized with the values obtained from the 1L1S model. MCMC optimization is then performed with all parameters free. The resulting optimized model parameters and their uncertainties are listed in Table 9. Fig. 8 also presents the model and residuals for this scenario.

When comparing the 1L2S model to the 2L1S model in Fig. 8, a clear improvement in the fit can be observed. The χ^2 values also support this observation, with the 2L1S model being significantly disfavored by $\Delta\chi^2 > 400$. Based on this comparison, we exclude the 2L1S interpretation for this event.

Despite favoring the 1L2S model, we perform an additional self-consistency check. The event is located in the highest cadence fields of KMTNet, and the anomaly lasts about 5 days. Therefore, KMTNet provides many V –band observations during the anomaly. We include V –band data from KMTNet CTIO to measure the color information. The last row in Table 9 lists the results obtained using all I –band data and KMTC V –band data. The flux ratios $q_{F,I}$ and $q_{F,V}$ allow us to calculate the magnitude and color difference between the two sources as following,

$$\Delta I_S = I_{S,2} - I_{S,1} = -2.5 \log q_{F,I}, \quad (24)$$

$$\Delta(V - I)_S = (V - I)_{S,2} - (V - I)_{S,1} = -2.5 \log \frac{q_{F,V}}{q_{F,I}}. \quad (25)$$

The positions of both sources are plotted on the CMD in Fig. 9. The first source is a red giant and the second source is a typical bulge main-sequence star. Therefore, the 1L2S model is reasonable, there is no evidence to contradict it.

In summary, the analysis of the anomaly signal in event MOA-2022-BLG-038 suggests the presence of a companion star to the the lensed source. The binary source system consists of a red giant and a main-sequence star. There is no indication that the lens system cannot be treated as a point

lens. The newly discovered candidate anomaly is not a planetary signal.

6. CONCLUSION AND DISCUSSION

In this work, we update the photometric pipeline based on Yang et al. (2024) and form an automatic “auto-TLC” pipeline. We define a giant source event sample based on the KMTNet database, including a total of 352 events. We then apply the auto-TLC pipeline to these events to produce high-quality photometric data. In this sample, the “auto-TLC” photometry is successfully produced for 340 events. We then develop an anomaly search algorithm, eAF, and use it to identify potential planetary signals within the light curves of these events. We recovered 2 previously known planetary signals and 16 previously known stellar binary signals in the sample, and find 4 new planet-like anomalies. The events with new candidate planetary anomalies are OGLE-2018-BLG-0421 (KMT-2018-BLG-0831), KMT-2019-BLG-0578, KMT-2021-BLG-0736, and MOA-2022-BLG-2342 (KMT-2022-BLG-2342).

Subsequent detailed modelling of the detected anomalies revealed that the nature of the anomalies in KMT-2019-BLG-0578 and KMT-2021-BLG-0736 are clear planetary signals. The planet in event KMT-2019-BLG-0578 has a mass-ratio of $q \sim 4 \times 10^{-3}$. It is likely to be a ~ 1.2 Jovian mass planet orbiting the M dwarf host at a distance of ~ 1.2 or ~ 2.0 AU, depending on which one of the degenerate solutions is correct. The M/K dwarf lens star in KMT-2021-BLG-0736 hosts a $q \sim 1 \times 10^{-4}$ (or $21M_{\oplus}$) Neptune-mass planet. The (projected) orbital distance is ~ 4.0 AU. The other two events OGLE-2018-BLG-0421 and MOA-2022-BLG-2342 remains ambiguous or suggests non-planetary interpretations.

By systematically re-analyzing the 340 giant-source events with improved photometry, we successfully identified two new planetary systems. This approach has led to the discovery of previously missed planetary signals, effectively doubling the number of confirmed planets within the analyzed sample. Figure 10 shows all the four planets in this sample in $(\log q, \log s)$ space.

To rigorously answer the question of how much planet-detection efficiency is improved, a comprehensive evaluation of sensitivity is needed. Fortunately, the eAF anomaly search algorithm developed here offers the potential to systematically calculate the sensitivity of the sample to planet detections. The detailed sensitivity calculations as well as the corresponding statistical analysis will be presented in a future paper.

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Software: NumPy (Harris et al. 2020), SciPy (Virtanen et al. 2020), astropy (Astropy Collaboration et al. 2013, 2018, 2022), pyDIA, PyAstronomy (Czesla et al. 2019), VBBinaryLensing (Bozza 2010; Bozza et al. 2018)

Table 1. Observational information of the new candidate planetary events

KMTNet name	KB180831	KB190578*	KB210736*	KB222342
Other name	OB180421*	—	MB21152	MB22038*
R.A. (J2000)	17:51:58.73	17:57:56.32	17:53:17.42	17:58:04.53
Dec. (J2000)	−31:51:58:10	−27:47:41.50	−29:42:44.71	−28:17:16.30
<i>l</i>	−1.806°	2.364°	0.194°	1.952°
<i>b</i>	−2.684°	−1.748°	−1.833°	−2.021°
KMTNet field	01,41	03,43	02,42	02,03,42,43

NOTE. KB180831 is the abbreviation of KMT-2018-BLG-0831, OB180421 is the abbreviation of OGLE-2018-BLG-0421, and MB21152 is the abbreviation of MOA-2021-BLG-152, and so on. The official name (based on first discovery) is marked by “*”.

Table 2. The 2L1S parameters for OGLE-2018-BLG-0421

Model	<i>C</i> ₁	<i>C</i> ₂	<i>C</i> ₇	<i>R</i>₂	<i>W</i> ₄
$\chi^2/\text{dof.}$	2105.03/2103	2105.99/2103	2105.03/2103	2101.13/2103	2115.22/2103
<i>t</i> ₀ (HJD'−8223)	0.8585 ± 0.0031	8223.8812 ± 0.0011	8223.8607 ± 0.0037	8223.8925 ± 0.0021	8223.8799 ± 0.0011
<i>u</i> ₀	0.1003 ± 0.0005	0.0997 ± 0.0005	0.1001 ± 0.0005	0.0945 ± 0.0006	0.0968 ± 0.0010
<i>t</i> _E (d)	15.710 ± 0.051	15.730 ± 0.040	15.728 ± 0.050	16.004 ± 0.043	16.257 ± 0.155
<i>ρ</i>	0.0898 ± 0.0008	0.0893 ± 0.0011	0.0897 ± 0.0009	0.0810 ± 0.0016	0.0876 ± 0.0013
<i>α</i> (°)	229.1 ± 1.9	228.9 ± 1.5	48.3 ± 1.9	−10.64 ± 0.59	105.6 ± 1.5
log <i>s</i>	−0.773 ± 0.012	−0.244 ± 0.048	−0.771 ± 0.015	−0.0252 ± 0.0027	0.540 ± 0.061
log <i>q</i>	−0.34 ± 0.11	−2.40 ± 0.13	0.44 ± 0.16	−3.108 ± 0.036	−1.43 ± 0.18
<i>s</i>	0.1686 ± 0.0047	0.574 ± 0.058	0.170 ± 0.006	0.9437 ± 0.0058	3.50 ± 0.50
<i>q</i>	0.47 ± 0.10	0.0042 ± 0.0023	3.0 ± 1.3	0.00078 ± 0.00006	0.041 ± 0.017
<i>f</i> _{S,KMTC01}	8.667 ± 0.039	8.666 ± 0.033	8.654 ± 0.039	8.388 ± 0.036	8.515 ± 0.045
<i>I</i> _S	15.8744 ± 0.0052	15.8745 ± 0.0045	15.8760 ± 0.0052	15.9099 ± 0.0049	15.8936 ± 0.0061

* HJD'=HJD-2450000. Model parameters and their 1σ uncertainties are presented. For unmeasured parameters, the 3σ limit are provided. The magnitudes have been calibrated to OGLE-III. The best solution is marked in boldface.

Table 3. The 2L1S parameters for KMT-2019-BLG-0578

Model	<i>C</i>	<i>W</i>
$\chi^2/\text{dof.}$	12674.00/12673	12673.98/12673
t_0 (HJD')	8599.3102 ± 0.0034	8599.3096 ± 0.0033
u_0	0.0137 ± 0.0010	0.0137 ± 0.0011
t_E (d)	8.84 ± 0.48	8.83 ± 0.50
ρ (10^{-3})	1.67 ± 0.12	1.69 ± 0.13
α ($^\circ$)	257.4 ± 1.4	257.8 ± 1.4
$\log s$	-0.105 ± 0.011	0.113 ± 0.012
$\log q$	-2.408 ± 0.063	-2.401 ± 0.072
s	0.786 ± 0.019	1.298 ± 0.036
q (10^{-4})	39.5 ± 5.8	40.3 ± 6.7
$f_{S,\text{KMTC01}}$	0.0364 ± 0.0026	0.0365 ± 0.0027
I_S	21.820 ± 0.074	21.816 ± 0.070

* HJD'=HJD-2450000. Model parameters and their 1σ uncertainty are presented. The magnitudes have been calibrated to OGLE-III.

Table 4. Source properties and the derived microlensing parameters of KMT-2019-BLG-0578

Parameter	Value	Uncertainty
$(V - I)_{\text{RC}}$	3.484	0.021
I_{RC}	16.689	0.036
$(V - I)_{\text{RC},0}$	1.060	0.030
$I_{\text{RC},0}$	14.365	0.040
$(V - I)_S$	2.994	0.095
I_S	21.816	0.070
$(V - I)_{S,0}$	0.570	0.102
$I_{S,0}$	19.491	0.088
θ_* (μas)	0.361	0.040
θ_E (mas)	0.213	0.029
μ_{rel} (mas/yr)	8.9	1.3

Table 5. Physical properties from Bayesian analysis of planetary event KMT-2019-BLG-0578

Model	<i>C</i>	<i>W</i>
D_S (kpc)	$8.5^{+0.9}_{-0.7}$	$8.5^{+0.9}_{-0.7}$
D_L (kpc)	$7.3^{+0.7}_{-0.8}$	$7.3^{+0.7}_{-0.8}$
μ_{rel} (mas/yr)	$8.8^{+1.3}_{-1.1}$	$8.8^{+1.3}_{-1.1}$
M_{host} (M_\odot)	$0.28^{+0.28}_{-0.14}$	$0.28^{+0.28}_{-0.14}$
M_p (M_J)	$1.2^{+1.2}_{-0.6}$	$1.2^{+1.2}_{-0.6}$
a_\perp (AU)	$1.2^{+0.2}_{-0.2}$	$2.0^{+0.3}_{-0.3}$

Table 6. The 1L2S and 2L1S parameters for KMT-2021-BLG-0736

Model	1L2S	2L1S: W_{inner}	2L1S: W_{outer}	2L1S: W_{cross}
$\chi^2/\text{dof.}$	13621.90/13372	13536.13/13374	13374.06/13374	13433.27/13374
$t_{0,1}$ (HJD')	9385.284 ± 0.013	9385.198 ± 0.012	9385.196 ± 0.012	9385.222 ± 0.012
$t_{0,2}$ (HJD')	9375.844 ± 0.013
$u_{0,1}$	0.815 ± 0.013	0.871 ± 0.017	0.863 ± 0.019	0.866 ± 0.017
$u_{0,2}$	-0.0001 ± 0.0011
t_E (d)	20.92 ± 0.21	20.03 ± 0.26	20.12 ± 0.29	20.09 ± 0.26
ρ_1	< 0.27	< 0.012	< 0.013	0.0236 ± 0.0007
ρ_2 (10^{-3})	7.06 ± 0.51
$q_{F,I}$ (10^{-4})	2.71 ± 0.19
$q_{F,Red}$ (10^{-4})	3.01 ± 0.46
α ($^\circ$)	...	241.88 ± 0.17	241.94 ± 0.18	241.59 ± 0.16
$\log s$...	0.2164 ± 0.0034	0.1940 ± 0.0039	0.2016 ± 0.0034
$\log q$...	-4.120 ± 0.025	-3.972 ± 0.020	-4.613 ± 0.028
s	...	1.646 ± 0.013	1.563 ± 0.014	1.591 ± 0.012
q (10^{-4})	...	0.760 ± 0.044	1.068 ± 0.049	0.244 ± 0.016
$f_{S,1, \text{KMTC01}}$	4.73 ± 0.13	5.40 ± 0.20	5.33 ± 0.22	5.34 ± 0.20
$I_{S,1}$	16.419 ± 0.023	16.293 ± 0.023	16.313 ± 0.022	16.300 ± 0.025

* HJD'=HJD-2450000. Model parameters and their 1σ uncertainty are presented. For unmeasured parameters, their 3σ limits are provided. The magnitudes have been calibrated to OGLE-III. The final selected model is highlighted in boldface.

Table 7. Source properties and the derived microlensing parameters of KMT-2021-BLG-0736

Parameter	Value	Uncertainty
$(V - I)_{\text{RC}}$	2.657	0.008
I_{RC}	16.297	0.022
$(V - I)_{\text{RC},0}$	1.060	0.030
$I_{\text{RC},0}$	14.434	0.040
$(V - I)_S$	2.086	0.030
I_S	16.313	0.022
$(V - I)_{S,0}$	0.489	0.043
$I_{S,0}$	14.450	0.051
θ_* (mas)	3.69	0.35
θ_E (mas)	> 0.27	—
μ_{rel} (mas/yr)	> 5.0	—

Table 8. Physical properties from Bayesian analysis of planetary event KMT-2021-BLG-0736

Model	W_{outer}
D_{S} (kpc)	$8.9^{+0.7}_{-0.8}$
D_{L} (kpc)	$7.0^{+0.8}_{-1.5}$
μ_{rel} (mas/yr)	$6.9^{+1.8}_{-1.3}$
M_{host} (M_{\odot})	$0.60^{+0.37}_{-0.29}$
M_p (M_{\oplus})	21^{+13}_{-10}
a_{\perp} (AU)	$4.0^{+0.9}_{-0.8}$

Table 9. The 1L2S and 2L1S parameters for MOA-2022-BLG-038/KMT-2022-BLG-2342

Model	2L1S	1L2S	1L2S (+ KMTC V -band)
$\chi^2/\text{dof.}$	47772.65/47300	47336.15/47299	48937.30/48891
$t_{0,1}$ (HJD')	9671.674 ± 0.043	9671.816 ± 0.037	9671.830 ± 0.044
$t_{0,2}$ (HJD')	...	9831.5312 ± 0.0092	9831.5314 ± 0.0099
$u_{0,1}$	1.236 ± 0.022	1.0616 ± 0.0009	1.0615 ± 0.0009
$u_{0,2}$...	0.0087 ± 0.0005	0.0087 ± 0.0004
t_E (d)	45.78 ± 0.55	50.412 ± 0.056	50.406 ± 0.063
ρ_1	0.0528 ± 0.0020
ρ_2 (10^{-3})	...	< 9.9	< 9.9
$q_{F,I}$ (10^{-3})	...	1.240 ± 0.016	1.245 ± 0.014
$q_{F,Red}$ (10^{-3})	...	1.300 ± 0.100	1.314 ± 0.097
$q_{F,V}$ (10^{-3})	1.325 ± 0.214
α ($^{\circ}$)	339.76 ± 0.12
$\log s$	0.5993 ± 0.0049
$\log q$	-3.509 ± 0.014
s	3.975 ± 0.045
q (10^{-4})	3.10 ± 0.10
$f_{S,1,I,\text{KMTC01}}$	9.18 ± 0.38	6.561 ± 0.012	6.560 ± 0.012
$I_{S,1}$	15.638 ± 0.025	15.9906 ± 0.0024	15.9906 ± 0.0026
$f_{S,1,V,\text{KMTC01}}$	0.5081 ± 0.0010
$V_{S,1}$	19.2261 ± 0.0031

* HJD'=HJD-2450000. Model parameters and their 1σ uncertainty are presented. For unmeasured parameters, their 3σ limits are provided. No useful ρ_1 is measured in the 1L2S models. The magnitudes have been calibrated to OGLE-III. The final selected model is highlighted in boldface.

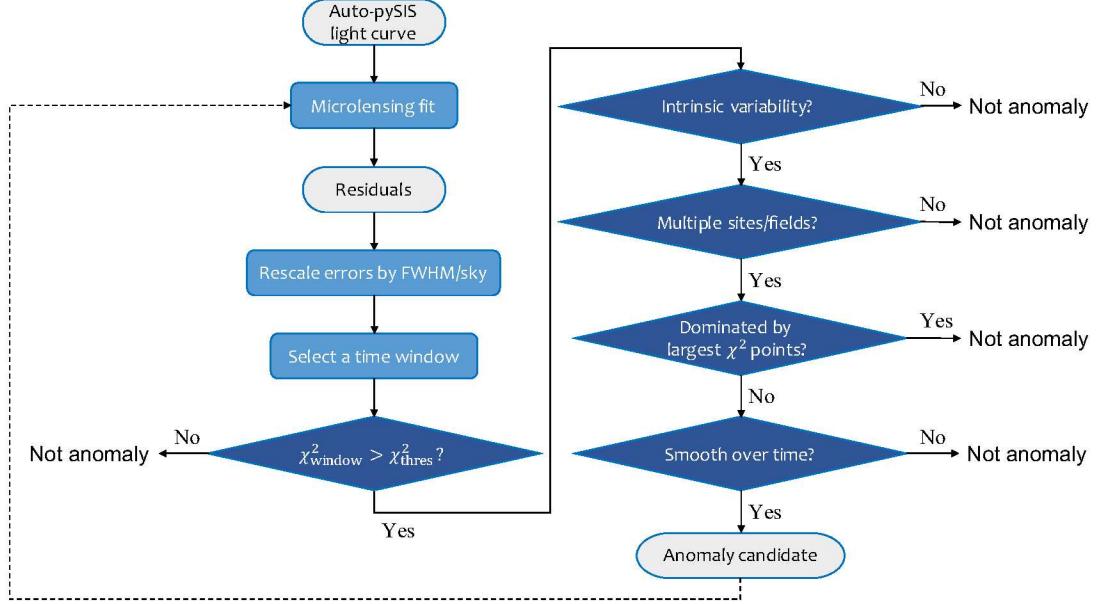


Figure 1. Flow chart of the anomaly search.

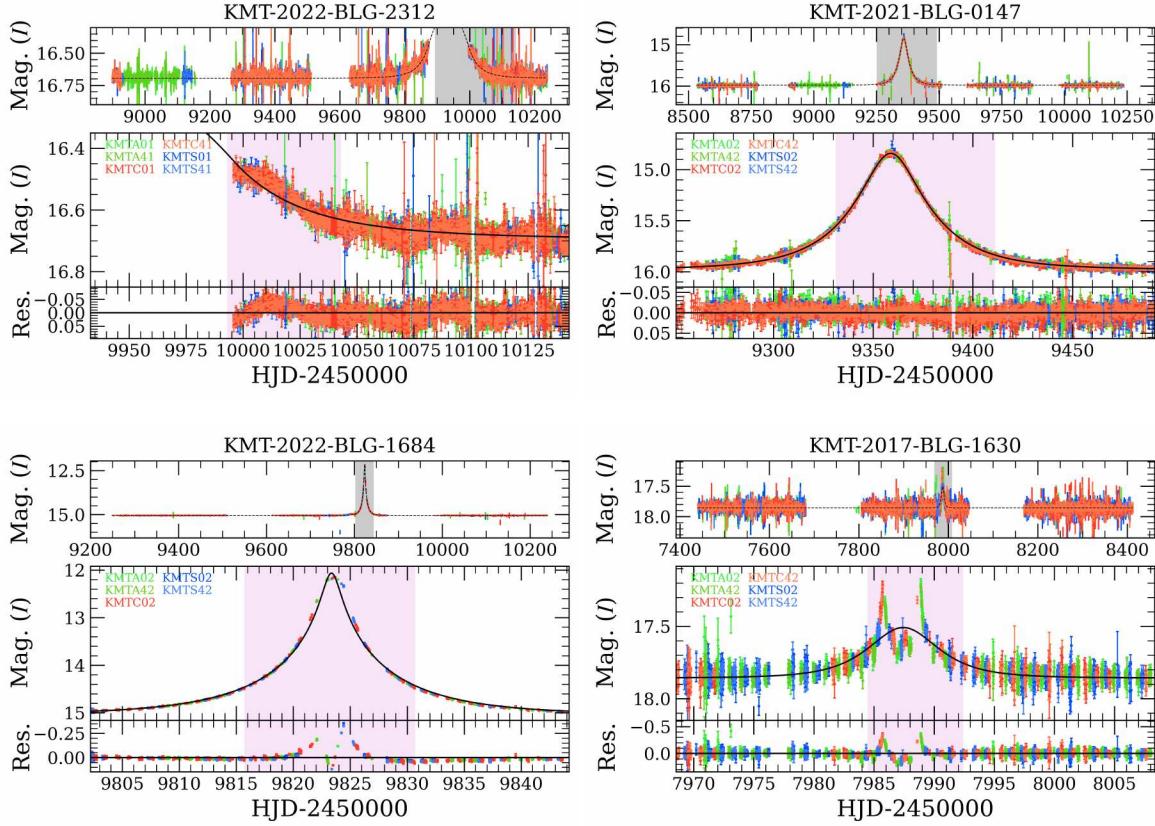


Figure 2. Light curves and reported signals of some example candidate anomalies. In each subfigure, the upper panel is the full light curve of the event and the lower is the zoom-in plot of the candidate signal region. The shadowed region marks the reported time window that has the candidate signal. The colors represent data from different sites and fields. The PSPL models are shown as black curves.

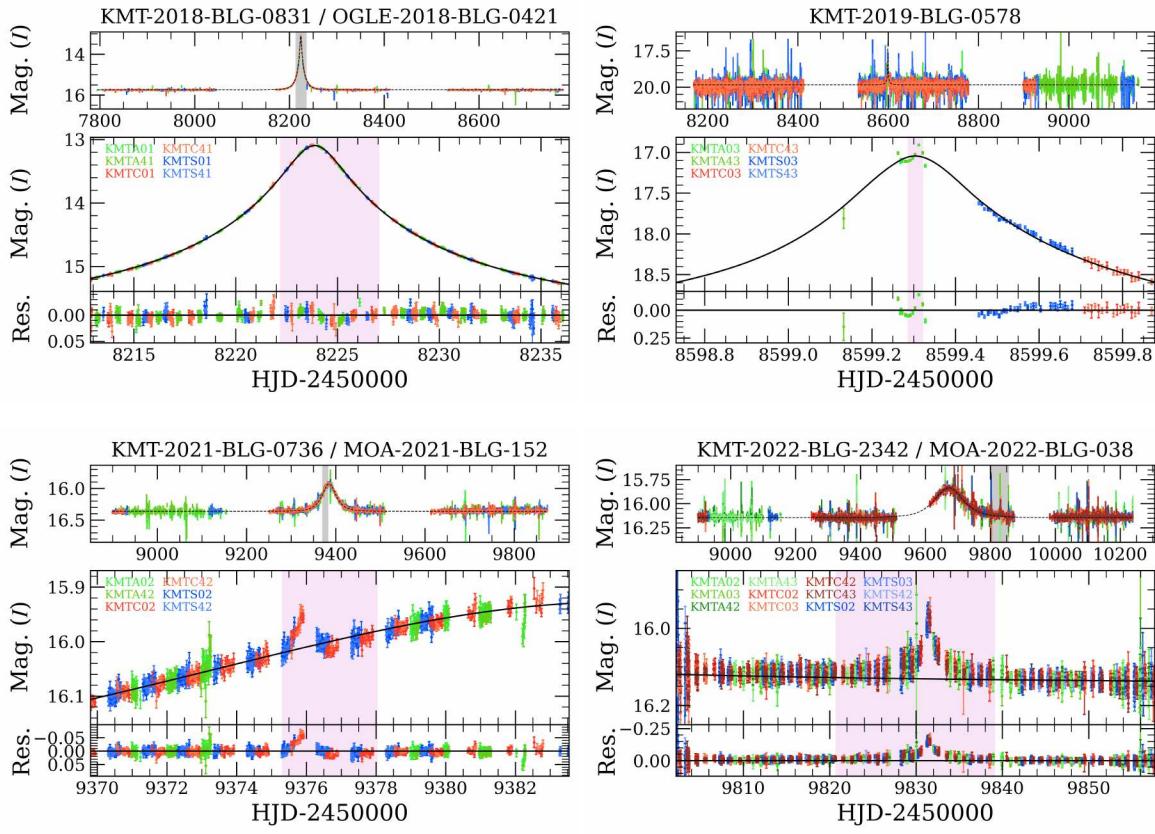


Figure 3. Same as Fig. 2 but for events with newly identified planet-like signals.

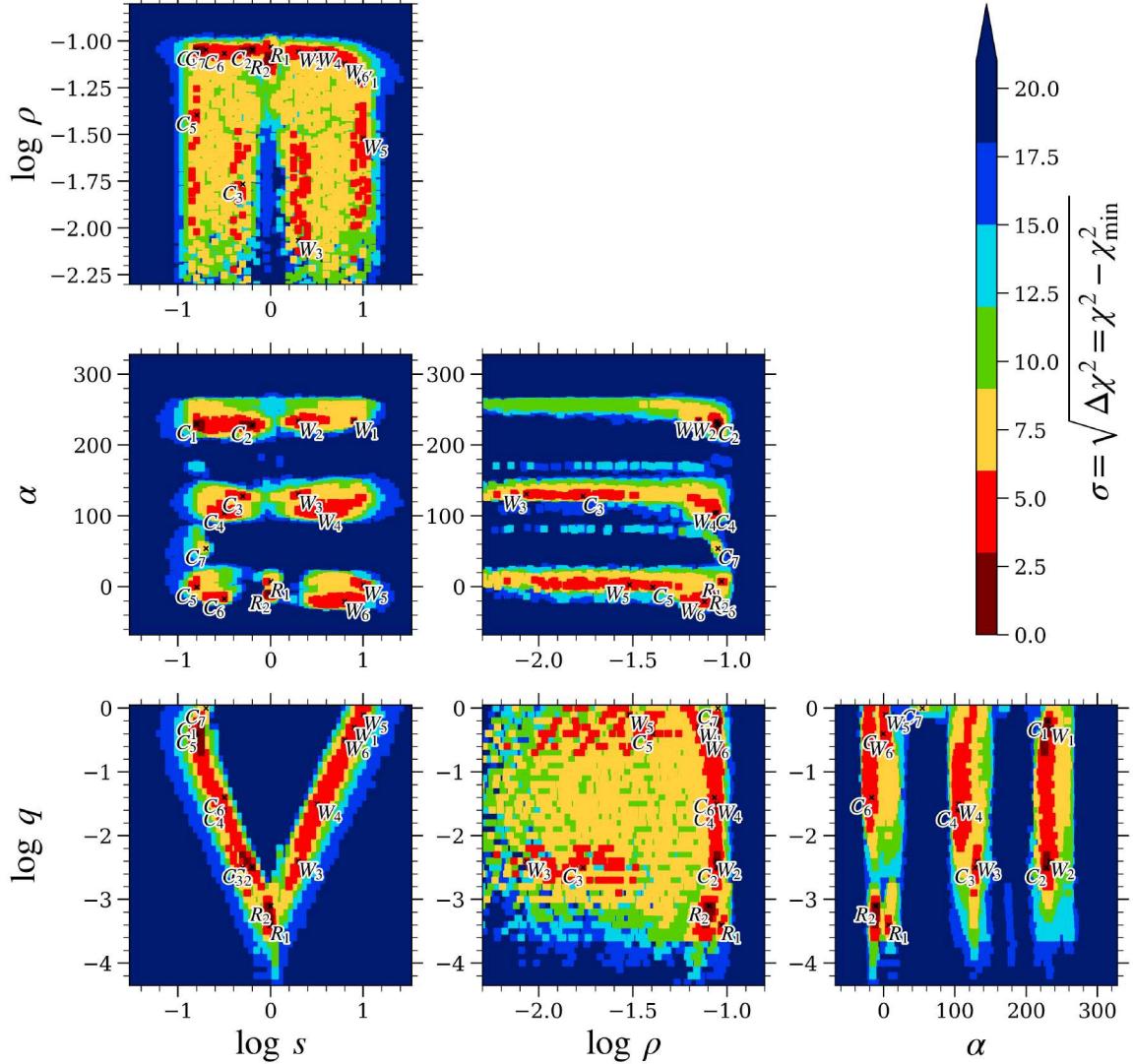


Figure 4. The 2L1S grid search χ^2 distribution for OGLE-2018-BLG-0421 in $(\log s, \log q, \alpha, \log \rho)$ space (α is units of degrees). The colors are coded by the $\Delta \chi^2$. The recognized local minima within 10σ are labeled with their names.

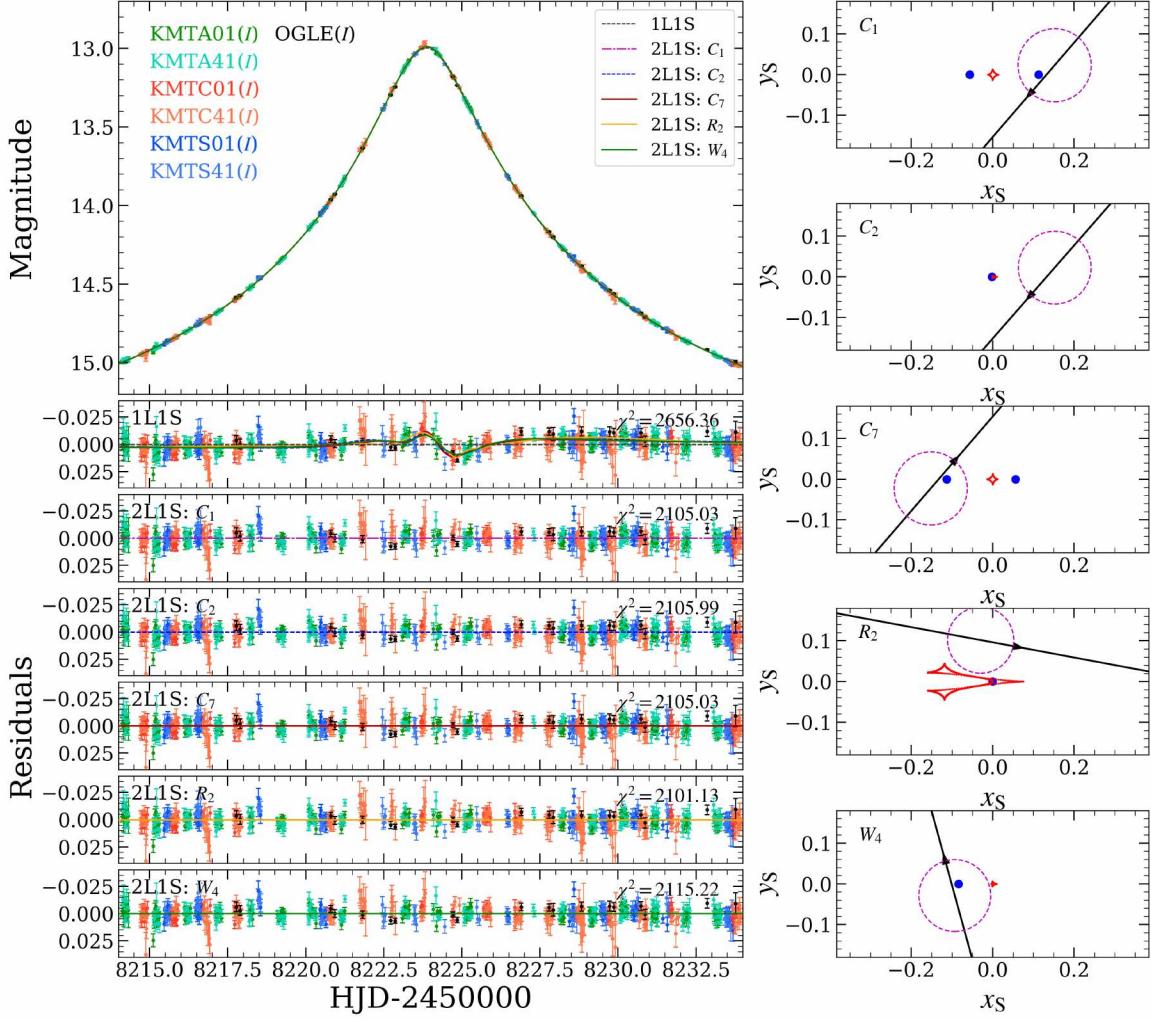


Figure 5. *Left:* The light curve, models, and residuals of OGLE-2018-BLG-0421 around the anomalous region. The residuals of each model are shown in separate panels with their χ^2 . *Right:* Caustics and trajectories of 2L1S models for OGLE-2018-BLG-0421. The red curves are the caustics and the black lines with arrows are the trajectories of the source-to-lens motion. The dashed circles mark the finite source sizes ρ .

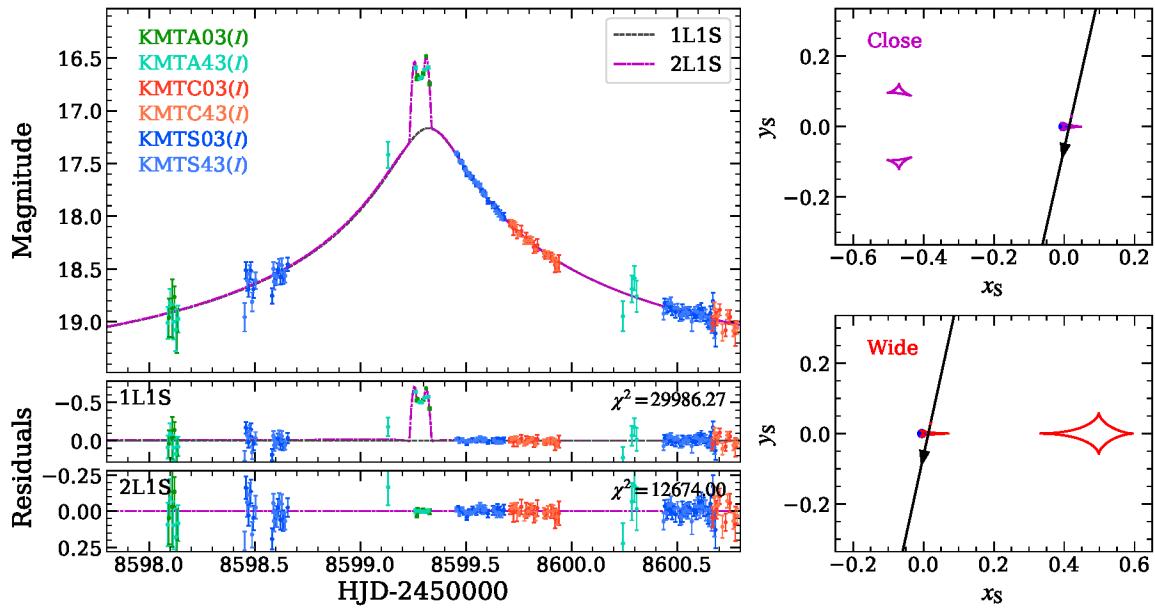


Figure 6. *Left:* The light curve, models, and residuals of KMT-2019-BLG-0578 around the anomalous region. The residuals of each model are shown in separate panels and their χ^2 are presented. The two 2L1S models are visually identical, thus only one (C) is presented. *Right:* Caustics and trajectories of 2L1S models for KMT-2019-BLG-0578. The curves are the caustics and the lines with arrows are the trajectories of the source-to-lens motion.

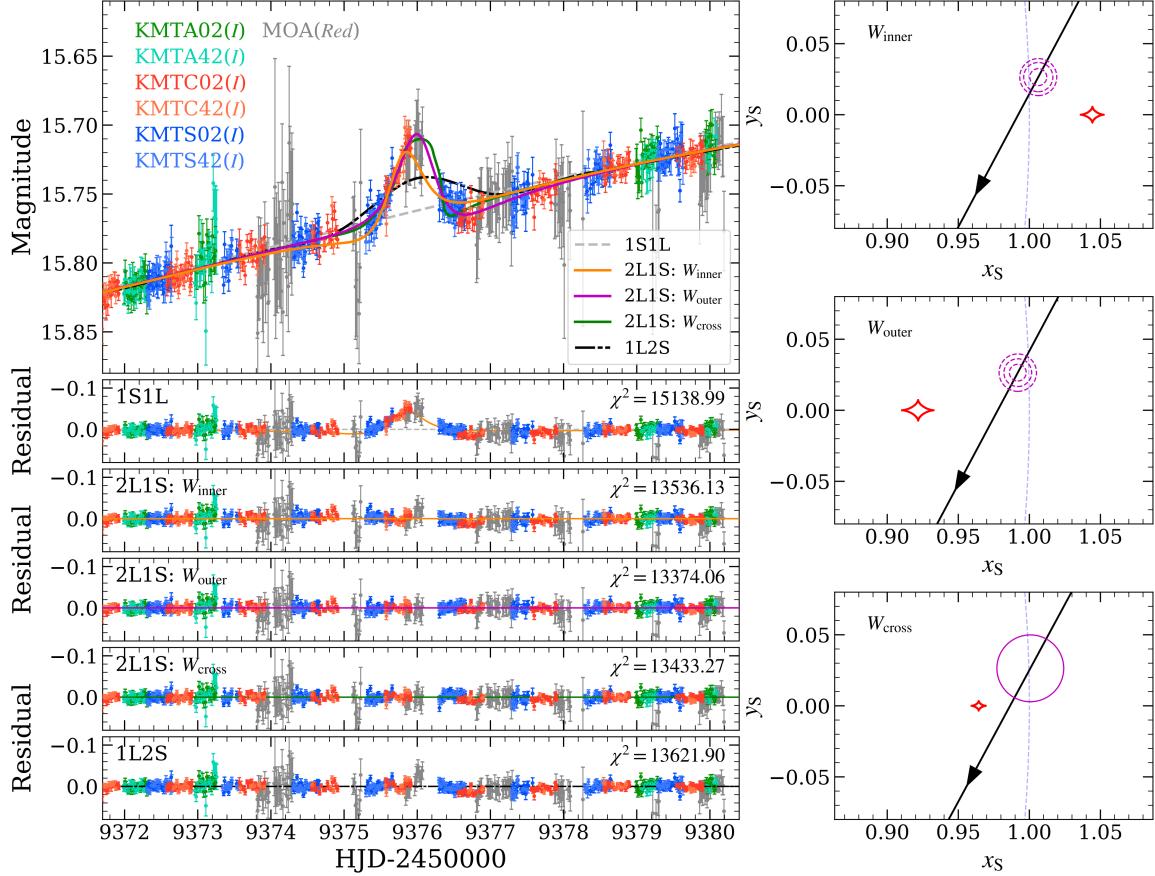


Figure 7. *Left:* The light curve, models, and residuals of KMT-2021-BLG-0736 around the anomalous region. The residuals of each model are shown in separate panels and their χ^2 are presented. *Right:* Caustics and trajectories of 2L1S models for KMT-2021-BLG-0736. The curves are the (planetary) caustics and the lines with arrows are the trajectories of the source-to-lens motion. The magenta dashed circles mark the $(1, 2, 3)\sigma$ upper limits of ρ and the solid circle (bottom right) marks the measured values of ρ . The blue (nearly vertical) lines represent the Einstein radius.

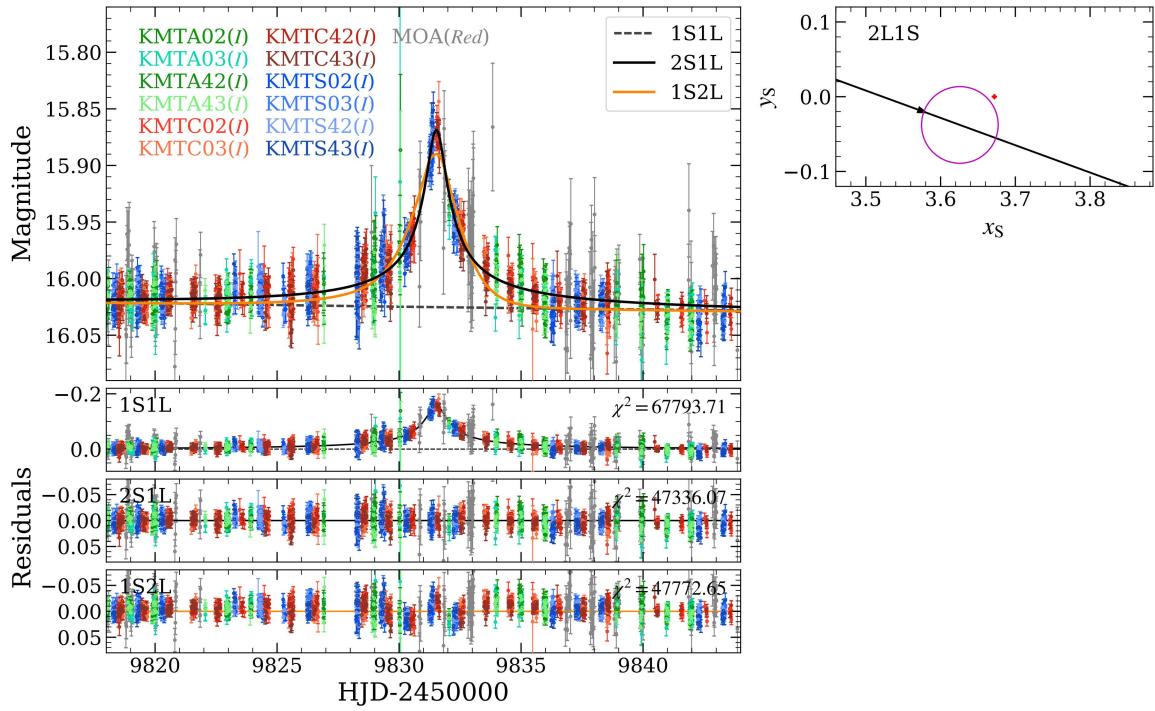


Figure 8. *Left:* The light curves, models, and residuals for MOA-2022-BLG-038 around the anomalous region. The residuals of each model are shown in separate panels and their χ^2 are presented. *Right:* Caustics and trajectories of 2L1S models for MOA-2022-BLG-038. The curves are the (planetary) caustics and the lines with arrows are the trajectories of the source-to-lens motion. The magenta circle marks the source size ρ .

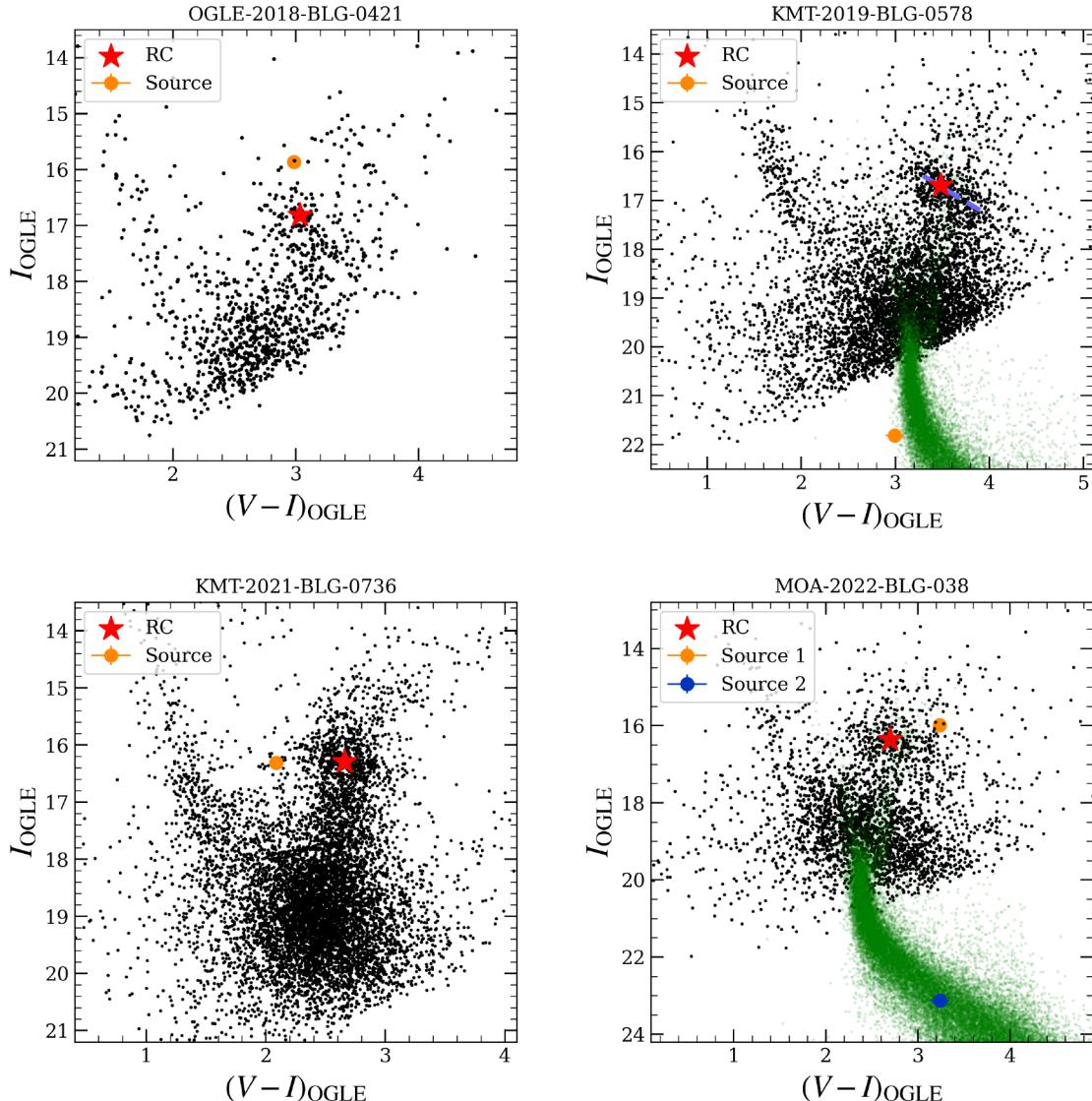


Figure 9. Color-Magnitude Diagrams (CMDs) for the four new candidate planetary events. The black dots are the field stars within $2' \times 2'$ of each lensed source. The green dots stars are from HST observations toward Baade's window (Holtzman et al. 1998) and have been offset to match the RC centroid. Colors and magnitudes are calibrated to OGLE-III (Szymański et al. 2011).

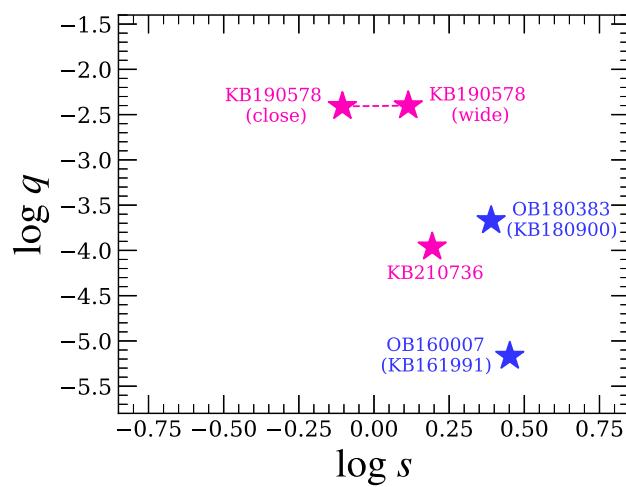


Figure 10. The $(\log s, \log q)$ distribution of the newly discovered (pink) and recovered (blue) clear planets in the sample.

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